THE LANCET **Planetary Health**

Supplementary appendix

This appendix formed part of the original submission and has been peer reviewed. We post it as supplied by the authors.

Supplement to: Janko M M, Irish S R, Reich B J, et al. The links between agriculture, Anopheles mosquitoes, and malaria risk in children younger than 5 years in the Democratic Republic of the Congo: a population-based, cross-sectional, spatial study. *Lancet Planet Health* 2018; **2:** e74–82.

Supplemental Material

STROBE Statement—Checklist of items that should be included in reports of *cross-sectional studies*

zones."

(*b*) Describe any methods used to examine subgroups and interactions

The mothers for all eligible participants assented to their children being included in the study.

(c) Consider use of a flow diagram

We did not categorize continuous variables.

We include the following statements in the manuscript:

The sponsors of the study had no role in study design, data collection, data analysis, data interpretation, or writing of the report. The corresponding author had full access to all the data in the study and had final responsibility for the decision to submit for publication. Parental consent for children's participation in the 2013-2014 Demographic and Health Surveys (DHS) was obtained by the DHS Program. The 2013-2014 DRC DHS was reviewed and approved by the Institutional Review Board (IRB) at ICF International—the implementing agency of the DHS—and the University of Kinshasa IRB (Comité d'Ethique de l'Ecole de Santé Publique de l'Université de Kinshasa). This study was also approved by the IRB at the University of North Carolina at Chapel Hill.

The authors acknowledge support from the National Institutes of Health (grant 5R01AI107949 to Steven R. Meshnick), the National Science Foundation (grant BCS-1339949 to Michael Emch), the Gates Foundation (grant OPP1161913 to Brian J. Reich). Mark Janko received support from the Royster Society of Fellows at UNC-CH. Mark Janko and Marc Peterson were supported by the Population Research Infrastructure Program awarded to the Carolina Population Center (P2C HD050924) by the *Eunice Kennedy Shriver* **National Institute of Child Health and Development. Seth Irish is funded by the US President's Malaria Initiative.**

*Give information separately for exposed and unexposed groups.

Note: An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at http://www.plosmedicine.org/, Annals of Internal Medicine at http://www.annals.org/, and Epidemiology at http://www.epidem.com/). Information on the STROBE Initiative is available at www.strobe-statement.org.

Model Specifications for Models on Probability of Malaria Infection:

Our outcome of interest is each individual's PCR-diagnosed malaria status, a binary indicator taking the value 1 if an individual is infected with malaria, and 0 otherwise. Typically, binary data are handled using logistic regression. However, spatial models for point-referenced data become computationally intensive very quickly as the number of spatial locations increases. This computational burden is further increased due to the lack of conjugacy between the prior distributions for model parameters and the data likelihood in logistic regression. As such, we adopt a probit specification in which we introduce latent variables that are assumed to follow a normal distribution with unit variance. Such a specification also has a scientific rationale. For example, we can think of these latent variables as a propensity to become infected with malaria, with values above 0 indicating increased propensity to become infected with malaria, and vice versa. To see this connection, observe that we can represent the probability of malaria infection, given covariates, as coming from a linear model. For example, let y_i^* be the binary indicator for whether or not individual i (i in $1 \dots n$) has malaria. Then the probability of malaria infection is given by:

$$
Pr(y_i^* = 1 | X) = Pr(x_i^T \beta + \epsilon_i > 0)
$$

= Pr(x_i^T \beta > -\epsilon_i)
= Pr(\epsilon_i < x_i^T \beta)
= \Phi(x_i^T \beta)

Where $\Phi(\cdot)$ is the CDF of a standard normal distribution.

Analysis of DHS data consisted of fitting three hierarchical probit regression models, differing only in the correlation structures specified for the random effects. Below, we outline the MCMC procedure for drawing posterior samples from the full conditional distributions for all model parameters for each model. We begin with the hierarchical probit model in which the random effects are assumed to be independent across space, and we then introduce spatial correlation in these random effects, beginning with the intercept, and then extending this to model a spatially-varying slope as well via a separable model.

The basic setup for all these models is as follows. Let:

$$
Y = X\beta + Z\theta + \epsilon
$$

Where Y is an $n \times 1$ vector latent normal responses, X is an $n \times p$ row vector of covariates (including an intercept) for individual, β is a $p \times 1$ column vector regression coefficients linking the covariates to the response, Z is an $n \times$ q random effects design matrix, where q is the number of DHS clusters in the dataset, 331 in this analysis. θ is a $q \times$ 1 random intercept that varies across DHS clusters. Finally, ϵ is a white noise error assumed to follow a standard normal distribution.

The hierarchical model can be written as follows:

$$
Y|\beta, \theta, \sigma^2 \sim N(X\beta + Z\theta, I_n)
$$

$$
\beta|\sigma_{\beta}^2 \sim N(0, \sigma_{\beta}^2 I_p)
$$

$$
\theta|\sigma^2 \sim N(0, \sigma^2 I_q)
$$

$$
\sigma^2 \sim IG(a, b)
$$

MCMC procedure for multilevel probit with independently varying intercept:

for j in 1:n.posterior.samples $\{$

Step 1: Draw from full conditional distribution of latent normal random variable Y , as follows:

Let y_i^* denote the binary indicator observed, taking the value 1 if the respondent has malaria, and 0 otherwise. Updating the latent variable y_i proceeds from sampling from a truncated normal distribution:

$$
f(y_i|rest) \sim \begin{cases} N(x_i^T \beta + z_i^T \theta, 1, upper = 0), if y_i^* = 0\\ N(x_i^T \beta + z_i^T \theta, 1, lower = 0), if y_i^* = 1 \end{cases}
$$

Where $upper = 0$ indicates sampling from a truncated standard normal distribution truncated above by 0, while $lower = 0$ indicates sampling from a standard normal truncated below by 0.

Step 2: Draw from full conditional distribution of β :

Define $\gamma = Y - Z\theta$

$$
\beta\vert rest \sim N\left(\left((X^TX + \frac{I_p}{\sigma_\beta^2})^{-1} \right) X^T \gamma, \left(X^T X + \frac{I_p}{\sigma_\beta^2} \right)^{-1} \right)
$$

Step 3: Draw from full conditional distribution of θ :

Define $\mu = Y - X\beta$

$$
\theta \vert rest \sim N \left(\left(Z^T Z + \frac{I_q}{\sigma^2} \right)^{-1} Z^T \mu, \left(Z^T Z + \frac{I_q}{\sigma^2} \right)^{-1} \right)
$$

Step 4: Draw from full conditional distribution of σ^2 :

$$
\sigma^2 | rest \sim IG(a^*, b^*)
$$

Where $a^* = a + \frac{q}{2}$ and $b^* = \frac{\theta^T \theta}{2}$ $\frac{b}{2} + b$.

$$
\big\}
$$

MCMC procedure for hierarchical spatial probit with spatially varying intercept:

The spatial model has the same general form as the probit specification above, but with additional parameters introduced into incorporate spatial structure. The hierarchical model thus has the following form:

$$
Y|\beta, \theta, \sigma^2 \sim N(X\beta + Z\theta, I_n)
$$

$$
\beta|\sigma_{\beta}^2 \sim N(0, \sigma_{\beta}^2 I_p)
$$

$$
\theta|\sigma^2, \phi \sim N(0, \sigma^2 \Sigma(\phi))
$$

$$
\sigma^2 \sim IG(a, b)
$$

$$
\phi \sim U(\phi_a, \phi_b)
$$

Where everything is as before, except the variance for the random effects θ , where we introduce spatial structure through $\Sigma(\phi)$, which is a $q \times q$ matrix of pairwise distances between DHS clusters whose correlation decays according to an exponential correlation function with parameter ϕ . The prior for ϕ is chosen such that unmeasured confounding is spatially correlated from between 100 meters and 225 kilometers, roughly 10% of the breadth of DRC. Samples from the posterior distributions for all model parameters can be obtained by using the following MCMC steps:

for 1 in j:n.posterior.samples{

Step 1: Draw from full conditional distribution of latent normal random variable Y . Same as before. Step 2: Draw from full conditional distribution of latent normal random variable β . Same as before. Step 3: Sample from full conditional distribution of θ

Define $\mu = Y - X\beta$

$$
\theta \vert rest \sim N((Z^TZ + \sigma^2 \Sigma(\phi)^{-1})^{-1}Z^T\mu, (Z^TZ + \sigma^2 \Sigma(\phi)^{-1})^{-1})
$$

Step 4: Sample from full conditional distribution of σ^2

$$
\sigma^2 | rest \sim IG(a^*, b^*)
$$

Where $a^* = a + \frac{q}{2}$ and $b^* = \frac{\theta^T \Sigma(\phi)^{-1} \theta}{2}$ $\frac{\varphi_j - b}{2} + b.$

Step 5: Sample from full conditional distribution of ϕ :

First transform ϕ to have support on the real line using:

$$
\phi^* = \log((\phi - \phi_a)/(\phi_b - \phi))
$$

Then draw a proposal ϕ_p^* from:

 $N(\phi^*, v(\phi^*)),$

where $v(\phi^*)$ is a tuning variance. Then, back transform to obtain proposal draw, ϕ_p , using:

$$
\phi_p = (\phi_b \exp(\phi_p^*) + \phi_a)/(1 + \exp(\phi_p^*))
$$

Calculate log acceptance ratio:

$$
LAR = \frac{1}{2} \left(\log(\det(\Sigma(\phi)) - \log(\det(\Sigma(\phi))) + \frac{1}{2\sigma^2} \theta^T \left(\Sigma(\phi)^{-1} - \Sigma(\phi_p)^{-1} \right) \theta + \phi_p^* - \phi^* + \right)
$$

 $2log((1 + exp(\phi^*)/(1 + exp(\phi^*))))$

Update ϕ according to the following:

if
$$
(\log(U(0,1) < \text{LAR})
$$

\n $\phi = \phi_p$,
\nelse $\phi = \phi$

}

MCMC procedure for multilevel spatial probit with spatially varying intercept and slope:

As with the model for the spatial intercept, the model for the spatial intercept and slope process differs only in how the random effects are specified. The hierarchical model can be written as follows:

$$
Y|\beta, \theta, \sigma^2 \sim N(X\beta + Z\theta, I_n)
$$

$$
\beta|\sigma_{\beta}^2 \sim N(0, \sigma_{\beta}^2 I_p)
$$

$$
\theta|H, \phi \sim N(0, \Sigma(\phi) \otimes H)
$$

$$
H \sim IW(d+1, I_d)
$$

Where instead of a single spatial variance parameter, we represent the spatial variance-covariance matrix for the intercept and slope processes using the 2×2 matrix *H* (i.e. d=2 in the above specification). Note here too that the random effects design matrix Z is now $n \times 2q$, with the additional q columns containing the agricultural exposure around each DHS cluster. We specify an Inverse Wishart distribution with 3 degrees of freedom and a 2×2 identity scale matrix as the prior. Samples from the posterior distributions for all model parameters can be obtained by using the following MCMC steps:

for j in 1:n.posterior.samples{

Step 1: Draw from full conditional distribution of latent normal random variable Y . Same as before.

Step 2: Draw from full conditional distribution of latent normal random variable β . Same as before.

Step 3: Draw from full conditional distribution of θ :

$$
f(\theta | rest) \sim N((Z^T Z + \Sigma(\phi)^{-1} \otimes H^{-1})^{-1} Z^T \mu, (Z^T Z + \Sigma(\phi)^{-1} \otimes H^{-1})^{-1})
$$

Step 4: Draw from full conditional distribution of ϕ :

First transform ϕ to have support on the real line using:

$$
\phi^* = \log((\phi - \phi_a)/(\phi_b - \phi))
$$

Then draw a proposal ϕ_p^* from:

$$
N(\phi^*, v(\phi^*)),
$$

where $v(\phi^*)$ is a tuning variance. Then, back transform to obtain proposal draw, ϕ_p , using:

$$
\phi_p = (\phi_b \exp(\phi_p^*) + \phi_a)/(1 + \exp(\phi_p^*))
$$

Calculate log acceptance ratio:

$$
LAR = \frac{1}{2} (\log(\det(\Sigma(\phi) \otimes H) - \log(\det(\Sigma(\phi) \otimes H)) +
$$

$$
\frac{1}{2\sigma^2} \theta^T (\Sigma(\phi)^{-1} \otimes H^{-1} - \Sigma(\phi_p)^{-1} \otimes H^{-1}) \theta +
$$

$$
\phi_p^* - \phi^* +
$$

 $2log((1 + exp(\phi^*)/(1 + exp(\phi^*))))$

Update ϕ according to the following:

if
$$
(\log(U(0,1) < \text{LAR})
$$

\n $\phi = \phi_p$,
\nelse $\phi = \phi$

Step 5: Draw from full conditional distribution of H :

$$
H|rest \sim IW(q+3,\theta^T\Sigma(\phi)^{-1}\theta + I_2)
$$

}

All models were run for 120,000 iterations, with the first 20,000 discarded as burn-in and the Markov chain thinned such that inference about model parameters is based on 10,000 posterior samples. Model convergence was assessed by inspecting traceplots of model parameters, and final inferences are based on the best fitting model.

Model fitting results

Spatial models were initially compared by randomly withholding a third of the spatial locations and predicting those data out-of-sample. Performance was identical across both models, as can be seen above, and all models were re-fit to the full data, with final inferences presented in the manuscript being based off of the model incorporating a random intercept, as it had the lowest DIC.

While the non-spatial model exhibited the best fit to the data, we show results for the spatial processes from both models here, as these can be suggestive of potential areas of future concern. Supplementary Figure 1 below shows the spatial intercept surface, together with corresponding uncertainty.

Considerable variability in the spatial random intercept process persists after accounting for other risk factors, with areas of the DRC exhibiting both strong increased and decreased risk of infection, particularly in northern regions. Notably, however, these estimates are accompanied by considerable imprecision, preventing definitive conclusions about areas of increased or decreased residual risk.

Supplementary Figure 2 below shows the spatial intercept and slope surfaces for the model incorporating both a spatially varying intercept and a spatially varying slope for the effect of agriculture on malaria risk.

Incorporating the spatial random slope leads to slight attenuation in the intercept process, although the spatial pattern broadly remains. Further, there is slight evidence of possible attenuation of the effect of agriculture in two places in

DRC, one in the central-northern region, which is largely forest, and the other in central DRC in what is largely Savannah. This latter area also shows pockets of increased risk. In both cases, however, inferences on the intercept and slope processes are accompanied by considerable imprecision