Web Material for

"Sensitivity Analyses to the Missing at Random Assumption Using Multiple Imputation with delta-adjustment: Application to a Tuberculosis/HIV Prevalence Survey with Incomplete HIV-Status Data"

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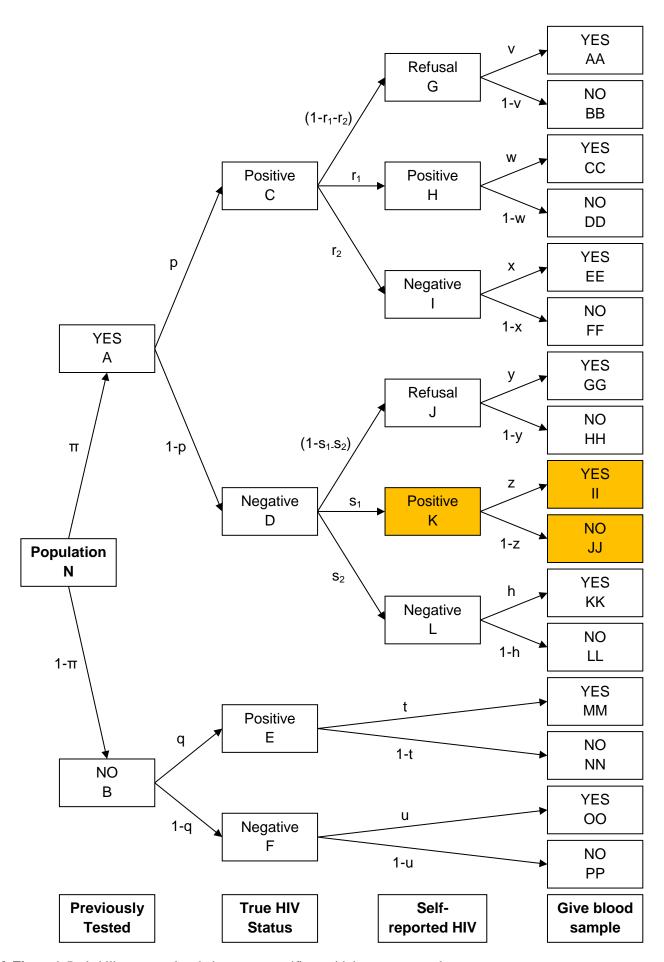
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WEB APPENDIX 1: SELECTING SENSITIVITY PARAMETER VALUES

Plausible values for the probabilities in web figure 1 can be informed by the observed data, the published literature, pilot studies and expert opinion.

- Suppose the total population is of size **N**.
- Let π denote the proportion of the population with a previous HIV test.
- Suppose that no individual fails to report a prior HIV test.
- Suppose that individuals with a prior HIV test are asked to disclose the result of their most recent test.
- If we assume that prior testing is independent of true HIV status, then p=q.
- If we assume that no individual who tested negative at their most recent test will self-report as HIV positive, then $s_1=0$ and consequently K=II=JJ=0.

A range of plausible delta values can be derived by varying the values of these probabilities and/or relaxing one or more of these assumptions.



Web Figure 1. Probability tree used to derive group-specific sensitivity parameter values.

Overall:

HIV prevalence among those who give a blood sample = $\frac{(AA + CC + EE + MM)}{(AA + CC + EE + GG + II + KK + MM + OO)}$

HIV prevalence among those who do not give a blood sample = $\frac{(BB + DD + FF + NN)}{(BB + DD + FF + HH + JJ + LL + NN + PP)}$

$$OR = \frac{(AA + CC + EE + MM)(HH + JJ + LL + PP)}{(GG + II + KK + OO)(BB + DD + FF + NN)} = exp(\delta)$$

Among those who self-report as HIV-negative:

HIV prevalence among those who give a blood sample = $\frac{EE}{(EE + KK)}$

HIV prevalence among those who do not give a blood sample = $\frac{FF}{(FF + LL)}$

$$OR = \frac{(\text{EE})(\text{LL})}{(\text{FF})(\text{KK})} = exp(\delta_1)$$

Among those who self-report as HIV-positive:

HIV prevalence among those who give a blood sample = $\frac{CC}{(CC + |||)}$

HIV prevalence among those who do not give a blood sample = $\frac{DD}{(DD + JJ)}$

$$OR = \frac{(CC)(JJ)}{(II)(DD)} = exp(\delta_2)$$

If we further assume that $s_1=0$, then II=JJ=0 and $OR = \exp(\delta_2) = 1.0$.

Among those who refuse to disclose their status:

HIV prevalence among those who give a blood sample = $\frac{AA}{(AA + GG)}$

HIV prevalence among those who do not give a blood sample = $\frac{BB}{(BB + HH)}$

$$OR = \frac{(AA)(HH)}{(GG)(BB)} = exp(\delta_3)$$

Among those with no prior test:

HIV prevalence among those who give a blood sample = $\frac{MM}{(MM + OO)}$

HIV prevalence among who do not give a blood sample = $\frac{NN}{(NN + PP)}$

$$OR = \frac{(MM)(PP)}{(OO)(NN)} = exp(\delta_4)$$

Among those with a previous test:

HIV prevalence among those who give a blood sample = $\frac{(AA + CC + EE)}{(AA + CC + EE + GG + II + KK)}$

HIV prevalence among those who do not give a blood sample = $\frac{(BB + DD + FF)}{(BB + DD + FF + HH + JJ + LL)}$

$$OR = \frac{(\text{AA} + \text{CC} + \text{EE})(\text{HH} + \text{JJ} + \text{LL})}{(\text{GG} + \text{II} + \text{KK})(\text{BB} + \text{DD} + \text{FF})}$$

Worked Example

Individuals participating in the ZAMSTAR TB/HIV prevalence survey were asked the following series of questions:

1. Do you know your status?

IF YES:

2. Are you willing to disclose your status?

IF YES:

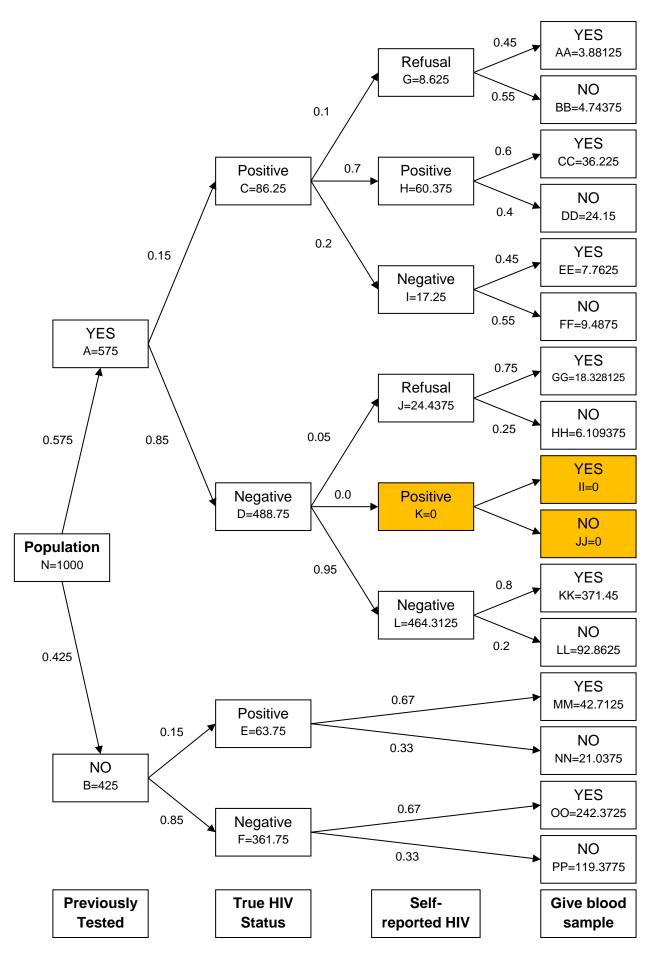
4. Are you HIV-positive or HIV-negative?

Plausible values for the probabilities in the probability tree shown in web figure 2 were informed by data collected in the ZAMSTAR trial itself (1), assumptions detailed in (2), and findings from a study investigating patterns in HIV refusal over time in the Karonga District of Malawi (3). We derive delta values for the entire adult population, but delta values could also be derived for males and females separately.

- Suppose we have a population of size **N=1000**.
- Suppose that the proportion of the population with a prior HIV test is $\pi = 0.575$ (ZAMSTAR trial data)
- Assume that the HIV prevalence among individuals without a prior HIV test is $\mathbf{q} = \mathbf{0.15}$ (ZAMSTAR trial data)
- Assume that no individual fails to report a past HIV test.
- Assume that prior testing is independent of true HIV status (2), p = q = 0.15.
- Among individuals without a prior test, assume that test acceptance is independent of true HIV status (2), $\mathbf{t} = \mathbf{u} = \mathbf{0.67}$ (ZAMSTAR trial data).
- Among those who are truly HIV-positive and have previously been tested, $\mathbf{r}_1 = \mathbf{0.7}$ self-report as HIV-positive (ZAMSTAR trial data), $\mathbf{r}_2 = \mathbf{0.2}$ report as HIV-negative, and $\mathbf{1} \mathbf{r}_1 \mathbf{r}_2 = \mathbf{0.1}$ refuse to disclose their status.
- Among those who are truly HIV-negative and have previously been tested, $s_1 = 0$ self-report as HIV-positive, $s_2 = 0.95$ report as HIV-negative (3), and $1 s_1 s_2 = 0.05$ refuse to disclose their status.
- Among those who are truly HIV-positive, have previously been tested and self-report as HIV-positive, $\mathbf{w} = \mathbf{0.6}$ provide a blood sample for HIV testing (ZAMSTAR trial data).
- Among those who are truly HIV-positive, have previously been tested and self-report as HIV-negative, $\mathbf{x} = \mathbf{0.45}$ provide a blood sample for HIV testing (3).
- Among those who are truly HIV-positive, have previously been tested but refuse to disclose their status, $\mathbf{v} = \mathbf{0.45}$ provide a blood sample for HIV testing (expert opinion based on ZAMSTAR trial data).

- Among those who are truly HIV-negative, have previously been tested and self-report as HIV-negative, $\mathbf{h} = \mathbf{0.8}$ provide a blood sample for HIV testing (3).
- Among those who are truly HIV-negative, have previously been tested but refuse to disclose their status, y = 0.6 provide a blood sample for HIV testing (expert opinion based on ZAMSTAR trial data).

As mentioned previously, we can derive a range of plausible delta values by varying the values of these probabilities and/or relaxing one or more of the assumptions made.



Web Figure 2. Probability tree used to derive group-specific sensitivity parameter values in the worked example.

Overall:

HIV prevalence among those who give a blood sample = 90.58125/722.731875 = 0.125

HIV prevalence among those who do not give a blood sample = 59.41875/277.768125 = 0.214

$$OR = \exp(\delta) = 1.90$$

Among those who self-report as HIV-negative:

HIV prevalence among those who give a blood sample = 7.7625/379.2125 = 0.020

HIV prevalence among those who do not give a blood sample = 9.4875/102.350 = 0.092

$$OR = exp(\delta_1) = 5.06$$

Among those who self-report as HIV-positive:

HIV prevalence among those who give a blood sample = 1.0

HIV prevalence among those who do not give a blood sample = 1.0

$$OR = \exp(\delta_2) = 1.0$$

Among those who refuse to disclose their status:

HIV prevalence among those who give a blood sample = 3.88125/22.209375 = 0.175

HIV prevalence among those who do not give a blood sample = 4.74375/10.853125 = 0.439

$$OR = \exp(\delta_3) = 3.13$$

Among those with no prior test:

HIV prevalence among those who give a blood sample = 42.7125/285.085 = 0.150

HIV prevalence among who do not give a blood sample = 21.0375/140.415 = 0.150

$$OR = \exp(\delta_4) = 1.00$$

Among those with a previous test:

HIV prevalence among those who give a blood sample = 47.86875/437.646875 = 0.109

HIV prevalence among those who do not give a blood sample = 38.38125/137.353125 = 0.279

$$OR = 3.16$$

WEB APPENDIX 2: EXAMPLE R CODE FOR IMPLEMENTING THE IMPUTATION PROCEDURE

```
# Load the mice package(4)
library(mice)
# Declare a new imputation function, mice.impute.logreg.sens
# This is a minor modification of the standard mice.impute.logreg
function contained in (4)
# The argument delta must be specified by the user
# delta is the difference in the log-odds of Y=1 for those with missing
Y values compared to those with observed Y values
mice.impute.logreg.sens <- function(y, ry, x, delta,...) {
     # The method consists of the following steps:
     # 1. Fit a logistic regression model, and find (bhat, V(bhat))
     # 2. Add delta to the linear predictor values
     # 3. Draw beta from N(bhat, V(bhat))
     # 4. Compute predicted scores for units with missing data,
     # logit^{-1}(X beta)
     \# 5. Compare the score to a random (0,1) deviate, and impute.
     # 1. Fit a logistic regression model, and find (bhat, V(bhat))
     x \leftarrow cbind(1, as.matrix(x))
     expr <- expression(glm.fit(x[ry, ], y[ry],</pre>
                     family = binomial(link = logit))
     fit <- suppressWarnings(eval(expr))</pre>
     fit.sum <- summary.glm(fit)</pre>
     # Fitted coefficient values
     beta <- coef(fit)</pre>
     # 2. Add delta to the linear predictor values
     beta[1] <- beta[1] + delta</pre>
     # 3. Draw beta from N(bhat, V(bhat))
     rv <- t(chol(fit.sum$cov.unscaled))</pre>
     beta.star <- beta + rv %*% rnorm(ncol(rv))</pre>
     # 4. Compute predicted scores for units with missing data,
     # logit^{-1}(X beta)
     p <- 1/(1 + exp(-(x[!ry, ] %*% beta.star)))
     # 5. Compare the score to a random (0,1) deviate, and impute.
     vec <- (runif(nrow(p)) <= p)</pre>
     vec[vec] <- 1</pre>
     if (is.factor(y)) {
           vec \leftarrow factor(vec, c(0, 1), levels(y))
     }
  return (vec)
}
```

```
# Suppose we have a dataset with incomplete HIV status and fully
observed data on age, gender, region, TB status and self-reported HIV
status.
dta <- c("HIV", "age", "gender", "region", "TB", "selfrepHIV")
# Initialise arguments of the mice function
ini <- mice(data=dta, maxit=0)</pre>
# Default predictor matrix
pred <- ini$pred</pre>
# Default imputation function specification
meth <- ini$meth</pre>
# By default, all variables will be included in the imputation model for
HIV status
# We can remove variables from the imputation model for HIV as follows:
pred["HIV", c("selfrepHIV")] <- 0</pre>
# Specify the imputation function for HIV
meth["HIV"] <- "logreg.sens"</pre>
# Create imputed datasets
# data - incomplete dataset with missing values recorded as NA
# pred - predictor matrix
# meth - imputation function to be used
# m - Number of imputed datasets to be created
# maxit - Number of sampler iterations
imp <- mice(data=dta, pred=pred, meth=meth,</pre>
               m=25, maxit=1, delta=log(2.0))
# As per the mice package documentation, we can obtain an estimate of
the HIV prevalence as follows:
m < - imp$m
Q \leftarrow rep(NA, m)
U < - rep(NA, m)
# Estimate the HIV prevalence in each imputed dataset
for (i in 1:m) {
     Q[i] <- mean(complete(imp, i)$HIV)
     U[i] <- var(complete(imp, i)$HIV) / nrow(dta) # (standard error)^2</pre>
}
# Combine the imputation-specific estimates using Rubin's Rules
pool.scalar(Q, U, n=nrow(dta), method = "rubin")
```

WEB APPENDIX 3: PARAMETRIC CAUSAL MEDIATION ANALYSIS

Parametric mediation analysis involves fitting two parametric regression models to the data: a regression of the outcome on the exposure, mediator and other confounders and a regression of the mediator on exposure and other confounders. Using slightly different notation to Valeri and VanderWeele (5), suppose we have a binary outcome Y, a categorical exposure A with two or more levels, a binary mediator M and vector of confounders C. The model for the outcome is given by

$$logit\{P[Y=1|A=a,M=m,C=c]\} = \theta_0 + \theta_A a + \theta_M m + \theta_{AM} a m + \theta_C' c$$

and the model for the mediator is given by

$$logit\{P[M=1|A=a, C=c]\} = \beta_0 + \beta_A a + \beta_{C} c$$

The natural direct effect (NDE), natural indirect effect (NIE) and the total effect (TE) are identified assuming that there is no unobserved confounding of the outcome-exposure, outcome-mediator or mediator-exposure relationships and that every confounder of the outcome-mediator relationship is unaffected by the exposure (5).

If these assumptions are satisfied, the average natural direct effect (NDE) of setting the exposure to level a compared to level a^* , conditional on C = C, is given by

$$ORNDE \cong \frac{exp(\theta_A a)\{1 + exp(\theta_M m + \theta_{AM} a + \beta_0 + \beta_A a^* + \beta'_C c)\}}{exp(\theta_A a^*)\{1 + exp(\theta_M m + \theta_{AM} a^* + \beta_0 + \beta_A a^* + \beta'_C c)\}}$$

In the absence of exposure-mediator interaction, $\theta_{AM}=0$, this simplifies to

$$ORNDE \cong exp(\theta_A(a-a^*))$$

and does not depend on the values of the confounding variables.

The average natural indirect effect (NIE) of setting the exposure to level a compared to level a^* , conditional on $\mathbf{C} = \mathbf{c}$, is given by

$$OR^{NIE} \cong \frac{\{1 + \exp(\beta_0 + \beta_A a^* + \beta'_C c)\}\{1 + \exp(\theta_M m + \theta_{AM} a + \beta_0 + \beta_A a + \beta'_C c)\}}{\{1 + \exp(\beta_0 + \beta_A a + \beta'_C c)\}\{1 + \exp(\theta_M m + \theta_{AM} a + \beta_0 + \beta_A a^* + \beta'_C c)\}}$$

Note that in this expression all four components of the numerator and denominator include terms for the confounding variables: thus estimates of the average natural indirect effect are conditional on the values taken by the confounding variables. Note also that sensitivity of this quantity to departures from MAR is primarily attributable to sensitivity of β_A .

The average natural total effect (TE) of setting the exposure to level a compared to level a^* , conditional on C = c, is then given by

$$ORTE = OR^{NDE}OR^{NIE}$$

Standard errors for these three quantities can be obtained by bootstrapping or through application of the multivariate delta method (5).

Interpretation of the natural direct, natural indirect and total effect odds ratios

Suppose that we wish to estimate the causal effects, as mediated by HIV status, of setting educational attainment to College/university rather than to Primary.

The average natural direct effect is the odds ratio of active pulmonary TB, conditional on $\mathbf{C} = \mathbf{c}$, for individuals with College/university educational attainment and HIV status set to what it would have been had their educational attainment been set to Primary compared to individuals with Primary educational attainment and their natural HIV status at this exposure level.

The average natural indirect effect is the odds ratio of active pulmonary TB, conditional on $\mathbf{C} = \mathbf{c}$, for individuals with College/university educational attainment and their natural HIV status at this exposure level compared to individuals with College/university educational attainment and HIV status set to what it would have been had their educational attainment been set to Primary.

The total effect is the odds ratio of active pulmonary TB, conditional on $\mathbf{C} = \mathbf{c}$, for individuals with College/university educational attainment compared to individuals with Primary educational attainment. It is the product of the average natural direct effect and the average natural indirect effect on the odds ratio scale.

Identification and Interpretation of Controlled Direct Effects

The Valeri and VanderWeele framework can also be used to estimate controlled direct effects (CDE). Assuming that there is no unobserved confounding of the outcome-exposure or outcome-mediator relationships, the controlled direct effect of setting the exposure to level a rather than level a^* with the mediator held fixed at level m is given by

$$ORCDE = \exp\{(\theta_A + \theta_{AM}m)(a - a^*)\}.$$

While the CDE is equal to the NDE in the absence of exposure-mediator interaction, it is important to recognise that this quantity is identified even when the exposure affects one or more confounders of the outcome-mediator relationship and carries a different causal interpretation. In our setting, the controlled direct effect is the odds ratio of active pulmonary TB for individuals with College/University educational attainment compared to individuals with Primary educational attainment where the HIV status of all individuals is set to level m.

Educational Attainment as a Counterfactual Cause

There is some disagreement in the literature as to whether it is appropriate to consider educational attainment and other social determinants of health such as race and gender as well-defined causes under the potential outcomes framework (6, 7). This is because interventions based on such exposures could plausibly violate key assumptions of counterfactual consistency and no interference in some settings (6). Naimi et al. (7) have argued that the controlled direct effect (CDE) can represent a suitable target of inference in the presence of a modifiable mediator of the outcome-exposure relationship even when the exposure is not well-defined. For example, suppose we are willing to accept HIV status but not educational attainment as a counterfactual cause. These authors would interpret estimates of the CDE of educational attainment on active pulmonary TB obtained in the current study as the magnitude of educational disparity in the prevalence of active pulmonary TB that would remain were the investigator to intervene to equalise the distribution of HIV status across all categories of educational attainment (7).

WEB APPENDIX 4: SUPPLEMENTARY TABLES AND FIGURES

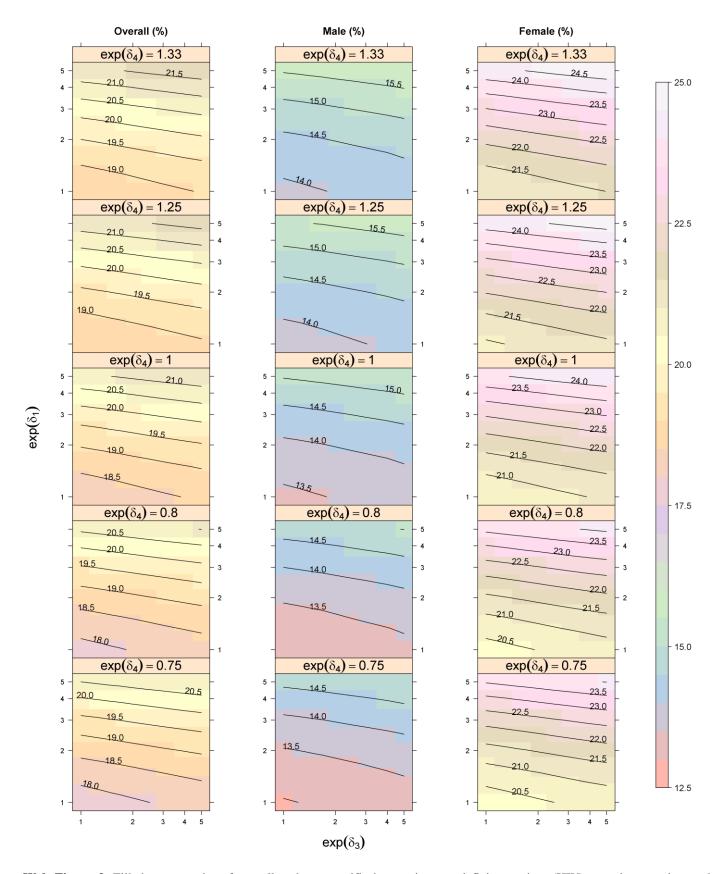
Web Table 1. Distribution of Self-Reported Human Immunodeficiency Virus Status Among Zambian Adults^a by Age, Sex and Region, Zambia-South Africa TB and AIDS Reduction Study, 2006-2010.

Men (N=11484)						Women (N=22314)						
Age (years)	Reported HIV- negative result	Reported HIV- positive result	Refused to Disclose Test Result	Never Tested	Frequency	Reported HIV- negative result	Reported HIV- positive result	Refused to Disclose Test Result	Never tested	Frequency		
Rural, Low Annual Risk of TB Infection (N=9796)												
18-24	^b 38.7	0.4	2.8	58.2	1,270	61.4	2.4	3.0	33.2	2,053		
25-29	42.8	3.9	4.1	49.2	465	66.0	8.5	3.7	21.8	1,060		
30-34	42.8	7.6	5.0	44.6	397	54.6	13.7	4.7	26.9	780		
35-39	44.1	10.5	4.0	41.5	354	49.0	15.4	5.1	30.4	565		
40-49	36.9	12.5	2.6	48.0	425	38.8	12.2	3.5	45.5	745		
50+	25.4	5.2	1.5	67.8	653	19.0	5.2	1.6	74.1	1,029		
All age	37.6	5.0	3.1	54.4	3,564	50.5	7.7	3.4	38.5	6,232		
Urban, Low Annual Risk of TB Infection (N=7339)												
18-24	41.7	0.6	2.1	55.6	981	54.6	3.4	4.7	37.3	1,655		
25-29	44.4	5.3	5.6	44.7	360	58.5	8.8	5.6	27.0	907		
30-34	38.5	8.6	3.3	49.7	338	49.0	15.7	6.3	29.0	635		
35-39	32.2	17.0	4.8	46.1	230	36.8	20.7	6.4	36.1	421		
40-49	29.5	13.3	4.4	52.7	315	35.4	16.6	4.6	43.4	523		
50+	27.0	6.4	1.3	65.3	392	22.3	6.5	1.7	69.4	582		
All age	37.2	6.1	3.1	53.6	2,616	46.9	9.5	4.9	38.7	4,723		
			Urban (n	ot Lusaka	a), High Annua	l Risk of TB	Infection (N	(= 6565)				
18-24	43.7	0.3	1.5	54.5	890	66.2	2.6	2.1	29.1	1,478		
25-29	49.7	3.5	2.2	44.7	318	70.0	8.5	4.1	17.5	813		
30-34	44.6	8.9	1.9	44.6	258	59.6	14.6	4.4	21.4	542		
35-39	47.8	10.0	2.0	40.3	201	51.0	16.0	4.1	28.9	363		
40-49	35.4	11.4	4.1	49.1	271	39.7	18.7	2.7	38.9	486		
50+	30.1	3.8	1.1	64.9	365	27.6	5.7	1.4	65.3	580		
All age	41.9	4.4	1.9	51.8	2,303	56.5	8.7	2.9	31.9	4,262		
	Lusaka, High Annual Risk of TB Infection (N=10098)											
18-24	32.8	0.3	1.4	65.5	1,153	61.7	2.6	1.9	33.8	2,618		
25-29	40.9	2.2	2.2	54.7	508	65.2	8.8	4.0	22.0	1,370		
30-34	34.3	6.1	2.8	56.9	362	59.7	13.3	2.0	24.9	959		
35-39	34.3	13.1	2.2	50.4	268	49.2	22.4	2.1	26.3	585		
40-49	34.6	11.4	2.5	51.4	280	41.5	16.2	1.5	40.8	779		
50+	28.4	4.4	1.9	65.3	430	23.7	6.7	0.8	68.8	786		
All age	34.0	4.1	1.9	59.9	3,001	54.6	8.8	2.2	34.4	7,097		

Abbreviations: HIV, human immunodeficiency virus; TB, tuberculosis.

^a Participants responded to a 2010 survey on the prevalence of tuberculosis and human immunodeficiency virus and had an evaluable TB sputum sample

^bRow percentages



Web Figure 3. Filled contour plot of overall and sex-specific human immunodeficiency virus (HIV) prevalence estimates by degree of subgroup-specific departure from the missing at random (MAR) assumption in the HIV test result variable, Zambia-South Africa TB and AIDS Reduction Study, 2006-2010.

 δ_1 , δ_3 and δ_4 capture the degree of departure from MAR for individuals who self-reported as HIV-negative, individuals who refused to disclose their status and individuals who had never been tested, respectively. δ_2 captures the degree of departure from MAR for individuals who self-reported as HIV-positive and is fixed at zero.

Web Table 2. Estimated Odds Ratios from a Logistic Regression of Human Immunodeficiency Virus Test Result Against Educational Attainment, Age, Sex and Region Fitted Among Zambian Adults^a, Zambia-South Africa TB and AIDS Reduction Study, 2006-2010.

	Complete Case Analysis		Best Case Analysis ^b		Worst Case Analysis ^c		Multiple Imputation under MAR ^d	
	OR	95% CI	OR	95% CI	OR	95% CI	OR	95% CI
Educational Attainment								
None	0.81	0.68, 0.97	0.80	0.68, 0.95	0.98	0.88, 1.09	0.80	0.69, 0.94
Primary	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
Lower secondary	1.02	0.93, 1.11	0.95	0.87, 1.03	1.09	1.03, 1.16	1.03	0.95, 1.11
Upper secondary	0.75	0.68, 0.83	0.67	0.61, 0.74	1.17	1.10, 1.24	0.77	0.70, 0.85
College/university	0.64	0.55, 0.75	0.51	0.44, 0.59	1.34	1.23, 1.46	0.66	0.57, 0.76
Sex								
Male	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
Female	4.42	3.47, 5.63	4.30	3.38, 5.47	1.21	1.12, 1.31	4.33	3.42, 5.48
Age, years								
18-24	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
25-29	5.49	4.12, 7.32	5.18	3.90, 6.89	1.40	1.25, 1.58	5.55	4.25, 7.24
30-34	9.94	7.53, 13.12	8.85	6.73, 11.63	1.82	1.61, 2.07	9.79	7.51, 12.77
35-39	14.37	10.82, 19.07	11.67	8.85, 15.38	2.40	2.09, 2.75	14.08	10.84, 18.30
40-49	14.41	10.95, 18.97	11.64	8.90, 15.22	2.36	2.07, 2.68	14.90	11.47, 19.35
50+	4.20	3.13, 5.64	3.79	2.83, 5.07	1.40	1.25, 1.57	4.34	3.26, 5.78
Age:Sex								
18-24:Female	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
25-29:Female	0.44	0.32, 0.60	0.42	0.31, 0.58	1.14	0.99, 1.32	0.44	0.33, 0.59
30-34:Female	0.34	0.25, 0.46	0.32	0.24, 0.43	1.14	0.98, 1.32	0.35	0.27, 0.47
35-39:Female	0.27	0.20, 0.38	0.27	0.20, 0.37	0.96	0.81, 1.14	0.30	0.23, 0.40
40-49:Female	0.21	0.15, 0.28	0.22	0.17, 0.30	0.77	0.66, 0.90	0.21	0.17, 0.29
50+:Female	0.24	0.17, 0.34	0.25	0.18, 0.34	0.84	0.73, 0.97	0.24	0.19, 0.34
Region and TB risk								
Rural, low ARTI	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
Urban, low ARTI	2.07	1.86, 2.30	1.85	1.68, 2.05	1.43	1.34, 1.52	1.83	1.67, 2.01
Urban (not Lusaka), high ARTI	1.56	1.40, 1.74	1.54	1.38, 1.71	1.14	1.07, 1.22	1.42	1.28, 1.57
Lusaka, high ARTI	1.42	1.29, 1.57	1.66	1.51, 1.82	0.76	0.72, 0.81	1.29	1.18, 1.40

Abbreviations: ARTI: annual risk of tuberculosis infection; CI, confidence interval; HIV, human immunodeficiency virus; MAR, missing at random; OR, odds ratio; TB, tuberculosis.

^a Participants responded to a 2010 survey on the prevalence of tuberculosis and human immunodeficiency virus, had an evaluable TB sputum sample and agreed to be tested for HIV (n=23093)

^b All missing HIV test result values were imputed as positive

^c All missing HIV test result values were imputed as negative

^d Imputation model included age, region, active pulmonary TB, household wealth index, educational attainment, current TB treatment, past TB treatment, marital status, diabetes status, smoking status, alcohol consumption, hunger in past 3 months, household crowding, circumcision status (males only), current cough, persistent cough for more than 2 weeks, current chest pain, current fever, current night sweats, current shortness of breath, unintentional weight loss in past month, and self-reported HIV status

Web Table 3. Estimated Odds Ratios from a Logistic Regression of Active Pulmonary Tuberculosis Against Human Immunodeficiency Virus Test Result, Educational Attainment, Age, Sex and Region Fitted Among Zambian Adults^a, Zambia-South Africa TB and AIDS Reduction Study, 2006-2010.

	Complete Case Analysis		Best Case Analysis ^b		Worst Case Analysis ^c		Multiple Imputation under MAR ^d	
	OR	95% CI	OR	95% CI	OR	95% CI	OR	95% CI
Education								
None	1.88	0.97, 3.64	1.43	0.77, 2.66	1.40	0.75, 2.59	1.45	0.78, 2.70
Primary	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
Lower secondary	0.71	0.45, 1.11	0.78	0.54, 1.12	0.76	0.52, 1.09	0.77	0.54, 1.12
Upper secondary	0.69	0.41, 1.08	0.71	0.48, 1.04	0.65	0.44, 0.95	0.73	0.50, 1.07
College/university	0.33	0.12, 0.93	0.29	0.13, 0.64	0.24	0.11, 0.53	0.30	0.14, 0.66
Sex								
Male	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
Female	0.89	0.42, 1.88	0.98	0.52, 1.86	1.04	0.55, 1.97	0.85	0.45, 1.63
Age, years								
18-24	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
25-29	1.72	0.69, 4.31	3.28	1.64, 6.55	3.47	1.74, 6.92	2.72	1.35, 5.48
30-34	3.83	1.73, 8.46	4.45	2.28, 8.69	4.86	2.50, 9.45	3.40	1.72, 6.72
35-39	1.35	0.47,3.83	2.08	0.91, 4.78	2.28	1.00, 5.21	1.47	0.63, 3.43
40-49	0.99	0.33, 2.98	2.13	0.97, 4.68	2.32	1.06, 5.09	1.50	0.67, 3.34
50+	1.39	0.54, 3.60	1.54	0.69, 3.45	1.55	0.70, 3.47	1.35	0.60, 3.03
Age:Sex								
18-24:Female	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
25-29:Female	0.61	0.21,1.84	0.69	0.16, 0.92	0.38	0.16, 0.90	0.40	0.17, 0.96
30-34:Female	0.18	0.06, 0.54	0.18	0.07,0.46	0.17	0.07, 0.44	0.19	0.08, 0.50
35-39:Female	0.76	0.22, 2.78	0.59	0.21,1.68	0.57	0.20, 1.62	0.65	0.23, 1.85
40-49:Female	0.67	0.17, 2.58	0.52	0.19, 1.42	0.50	0.19, 1.36	0.62	0.23, 1.67
50+:Female	0.12	0.02, 0.69	0.13	0.03, 0.56	0.13	0.03, 0.53	0.15	0.04, 0.64
Region and TB risk								
Rural, low ARTI	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
Urban, low ARTI	1.64	0.93, 2.87	1.69	1.10, 2.59	1.72	1.13, 2.64	1.55	1.01, 2.38
Urban (not Lusaka),	1.10	0.58, 2.09	1.32	0.83, 2.11	1.36	1.23, 2.17	1.28	0.80, 2.04
high ARTI								
Lusaka, high ARTI	2.05	1.25, 3.36	1.61	1.08, 2.41	1.83	1.75, 2.73	1.64	1.10, 2.44
HIV test result								
Negative	1.00	Referent	1.00	Referent	1.00	Referent	1.00	Referent
Positive	4.06	2.80, 5.89	3.02	2.19, 4.17	2.37	1.75, 3.20	4.69	3.32, 6.63

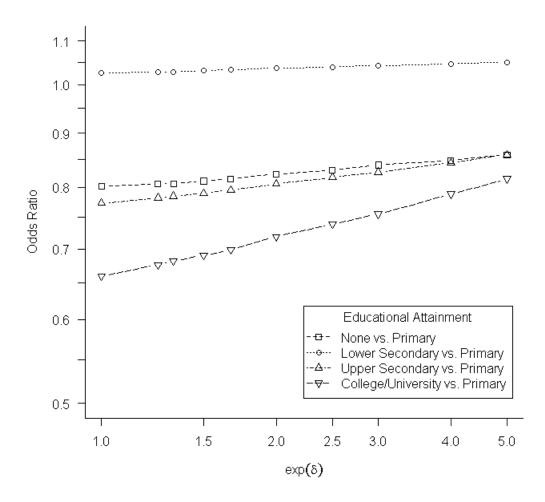
Abbreviations: ARTI: annual risk of tuberculosis infection; CI, confidence interval; HIV, human immunodeficiency virus; MAR, missing at random; OR, odds ratio; TB, tuberculosis.

^a Participants responded to a 2010 survey on the prevalence of tuberculosis and human immunodeficiency virus, had an evaluable TB sputum sample and agreed to be tested for HIV (n=23093)

^b All missing HIV test result values were imputed as positive

^c All missing HIV test result values were imputed as negative

^d Imputation model included age, region, active pulmonary TB, household wealth index, educational attainment, current TB treatment, past TB treatment, marital status, diabetes status, smoking status, alcohol consumption, hunger in past 3 months, household crowding, circumcision status (males only), current cough, persistent cough for more than 2 weeks, current chest pain, current fever, current night sweats, current shortness of breath, unintentional weight loss in past month, and self-reported HIV status



Web Figure 4. Association between positive human immunodeficiency virus (HIV) test result and individual educational attainment adjusted for age, sex and region by degree of departure ($\delta = \delta_1 = \delta_2 = \delta_3 = \delta_4$) from the missing at random assumption in the HIV test result variable, Zambia-South Africa TB and AIDS Reduction Study, 2006–2010.

REFERENCES

- 1. Alyes H, Muyoyeta M, Du Toit E, et al. Effect of household and community interventions on the burden of tuberculosis in Southern Africa: the ZAMSTAR community-randomised trial. *Lancet*. 2013; **382**(9899):1183-1194.
- 2. Reniers G, Eaton J. Refusal bias in HIV prevalence estimates from nationally representative seroprevalence surveys. *AIDS*. 2009: **23**(5):621-629.
- 3. Floyd S, Molesworth A, Dube A, et al. Underestimation of HIV prevalence in surveys when some people already know their status, and ways to reduce the bias. *AIDS*. 2013; **27**(2):223-242.
- 4. van Buuren S, Groothuis-Oudshoorn K. MICE: Multivariate imputation by chained equations in R. *J Stat Softw.* 2011; **45**(3):1-67.
- 5. Valeri L, VanderWeele TJ. Mediation analysis allowing for exposure-mediator interactions and causal interpretation: Theoretical assumptions and implementation with SAS and SPSS macros. *Psychol Methods*. 2013; **18**(2):137-150.
- 6. VanderWeele TJ, Robinson WR. On the causal interpretation of race in regressions adjusting for confounding and mediating variables. Epidemiology. 2014; **24**:473-84.
- 7. Naimi AI, Kaufman JS. Counterfactual Theory in Social Epidemiology: Reconciling Analysis and Action for the Social Determinants of Health. *Curr Epidemiol Rep.* 2015; **2**(1):52-60.