Does Childhood Trauma Moderate Polygenic Risk for Depression? A Meta-analysis of 5,765 Subjects From the Psychiatric Genomics Consortium

Supplemental Information

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Farnush Farhadi Hassan Kiadeh 43 Hilary K Finucane 44, 45 Andreas J Forstner 35, 36, 46, 47 Josef Frank 48 Héléna A Gaspar 27 Michael Gill 49 Fernando S Goes 50 Scott D Gordon 51 Jakob Grove 7, 8, 9, 52 Lynsey S Hall 11, 53 Christine Søholm Hansen 9, 18 Thomas F Hansen 54, 55, 56 Stefan Herms 35, 36, 47 lan B Hickie 57 Per Hoffmann 35, 36, 47 Georg Homuth 58 Carsten Horn 59 Jouke-Jan Hottenga 10 David M Hougaard 9, 18 Marcus Ising 60 Rick Jansen 19, 19 Eric Jorgenson 61 James A Knowles 62 Isaac S Kohane 63, 64, 65 Julia Kraft 4 Warren W. Kretzschmar 66 Jesper Krogh 67 Zoltán Kutalik 68, 69 Yihan Li 66 Penelope A Lind 28 Donald J MacIntyre 70, 71 Dean F MacKinnon 50 Robert M Maier 2 Wolfgang Maier 72 Jonathan Marchini 73 Hamdi Mbarek 10 Patrick McGrath 74 Peter McGuffin 27

Sarah F Medland 28 Divya Mehta 2, 75 Christel M Middeldorp 10, 76, 77 Evelin Mihailov 78 Yuri Milaneschi 19, 19 Lili Milani 78 Francis M Mondimore 50 Grant W Montgomery 1 Sara Mostafavi 79, 80 Niamh Mullins 27 Matthias Nauck 81, 82 Bernard Ng 80 Michel G Nivard 10 Dale R Nyholt 83 Paul F O'Reilly 27 Hogni Oskarsson 84 Michael J Owen 85 Jodie N Painter 28 Carsten Bøcker Pedersen 9, 12, 13 Marianne Giørtz Pedersen 9, 12, 13 Roseann E. Peterson 17, 86 Erik Pettersson 22 Wouter J Peyrot 19 Giorgio Pistis 26 Danielle Posthuma 87, 88 Jorge A Quiroz 89 Per Qvist 7, 8, 9 John P Rice 90 Brien P. Riley 17 Margarita Rivera 27, 91 Saira Saeed Mirza 37 Robert Schoevers 92 Eva C Schulte 93, 94 Ling Shen 61 Jianxin Shi 95 Stanley I Shyn 96 Engilbert Sigurdsson 97

Grant C B Sinnamon 98 Johannes H Smit 19 Daniel J Smith 99 Hreinn Stefansson 100 Stacy Steinberg 100 Fabian Streit 48 Jana Strohmaier 48 Katherine E Tansey 101 Henning Teismann 102 Alexander Teumer 103 Wesley Thompson 9, 55, 104, 105 Pippa A Thomson 106 Thorgeir E Thorgeirsson 100 Matthew Traylor 107 Jens Treutlein 48 Vassily Trubetskoy 4 André G Uitterlinden 108 Daniel Umbricht 109 Sandra Van der Auwera 110 Albert M van Hemert 111 Alexander Viktorin 22 Peter M Visscher 1, 2 Yunpeng Wang 9, 55, 105 Bradley T. Webb 112 Shantel Marie Weinsheimer 9, 55 Jürgen Wellmann 102 Gonneke Willemsen 10 Stephanie H Witt 48 Yang Wu 1 Hualin S Xi 113 Jian Yang 2, 114 Futao Zhang 1 Volker Arolt 115 Bernhard T Baune 14 Klaus Berger 102 Dorret I Boomsma 10 Sven Cichon 35, 47, 116, 117 Udo Dannlowski 115 EJC de Geus 10, 118 J Raymond DePaulo 50 Enrico Domenici 119 Katharina Domschke 120 Tõnu Esko 5, 78 Hans J Grabe 110

Steven P Hamilton 121 Caroline Hayward 122 Andrew C Heath 90 Kenneth S Kendler 17 Stefan Kloiber 60, 123, 124 Glyn Lewis 125 Qingqin S Li 126 Susanne Lucae 60 Pamela AF Madden 90 Patrik K Magnusson 22 Nicholas G Martin 51 Andrew M McIntosh 11, 34 Andres Metspalu 78, 127 Ole Mors 9, 128 Preben Bo Mortensen 8, 9, 12, 13 Bertram Müller-Myhsok 15, 16, 129 Merete Nordentoft 9, 130 Markus M Nöthen 35, 36 Michael C O'Donovan 85 Sara A Paciga 131

Nancy L Pedersen 22 Brenda WJH Penninx 19 Roy H Perlis 39, 132 David J Porteous 106 James B Potash 133 Martin Preisig 26 Marcella Rietschel 48 Catherine Schaefer 61 Thomas G Schulze 48, 94, 134, 135, 136 Jordan W Smoller 39, 40, 41 Kari Stefansson 100, 137 Henning Tiemeier 37, 138, 139 Rudolf Uher 140 Henry Völzke 103 Myrna M Weissman 74, 141 Thomas Werge 9, 55, 142 Cathryn M Lewis* 27, 143 Douglas F Levinson* 144 Gerome Breen* 27, 145 Anders D Børglum^{*} 7, 8, 9 Patrick F Sullivan* 22, 146, 147,

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Dichotomous Childhood Trauma Questionnaire (CTQ) score

The CTQ covers the five domains of sexual abuse (SA), physical abuse (PA), emotional abuse (EA), emotional neglect (EN), and physical neglect (PN). Each domain is assessed by five questions (scored 1 to 5) resulting in a domain score ranging from 5 to 25. Per domain, cutoffs were applied to define a narrow definition of childhood trauma separating no or mild trauma from moderate or severe trauma, based on cut-offs for moderate/severe of > 7 (SA), > 9 (PA), > 12 (EA), > 14 (EN), > 9 (PN) respectively. These cut-offs are based on the CTQ manual. From this, an overall dichotomous CTQ indicator was constructed to separate trauma in any of the five domains (1) from trauma in none of the domains (0).

Childhood trauma in DGN and QIMR

In the Depression Gene Network (DGN) cohort, sexual abuse was assessed with two questions: "Someone touched parts of your body in a sexual way, or had you touch parts of the person in a sexual way"; and "Someone had or attempted to have oral sex, anal sex, or sexual intercourse with you". Physical abuse in DGN was also assessed with two questions: "Someone outside your household physically attacked or assaulted you, threatened you with a weapon or held you captive"; and "Your mother, father or another adult household member hurt you on purpose (for example, beat, choked, kicked, cut or burned you)". The narrow definition was defined as at least one of four questions occurring frequently versus sometimes, rarely or never, and the broad definition as at least one of four questions occurring frequently or sometimes versus rarely or never. For data from the Queensland Institute of Medical Research (QIMR), two instruments were used to assess childhood trauma before the age of 18. Most QIMR individuals were assessed with an instrument covering sexual abuse: touching your sexual parts, you touching their sexual parts, or sexual intercourse (SA assessed with one question for family members and one question for non-family); and physical abuse: being punished by hitting (one question), hurting from punishment next day (one question), being physically injured on purpose (one question). The other QIMR individuals (on the QIMR 3 genotype-batch labeled as M7) were assessed with a questionnaire covering sexual abuse as the occurrence of: exposure to sexual organs, exposure to masturbation, being touched, attempt to have sex, and have sex (SA specified in 16 separate questions); and for physical abuse the occurrence of: being hit, kicked, choked, throttled or locked in by either father, father-figure, mother, or mother-figure (PA specified in 13 separate questions). For QIMR the narrow and broad definitions were defined as above, except for physical abuse from the second questionnaire (QIMR 3 M7) that didn't distinguish between occurring "frequently" and "sometimes" resulting in converging of the narrow and broad definitions. For the analyses, we applied the broad definition.

Peyrot *et al.* Supplement **Supplement** (Supplement) and the supplement of the supplement \sim

Simulation study 1: impact of gene-environment correlation in tests for GxE-interaction

Tests of genotype by environment interaction are known to be scale dependent. In a linear regression model, where a continuous phenotype is regressed on a measured genetic variant (e.g. a candidate gene) and a measured exposure, non-normality of the phenotypic distribution can give rise to spurious interaction effects. We considered this issue given logistic regression of a binary phenotype by means of a small simulation study. We generated phenotypic data based on 12 binary symptoms, which were related to an underlying normally distributed depression liability by a Rasch model (1). The parameters of the Rasch model were chosen so that the distribution of the sum scores based on the 12 symptoms was highly skewed. We dichotomized the sum score of these 12 symptoms to arrive at the binary phenotype with a prevalence of .20. The underlying normally distributed depression liability was subject to main effects of genes (A; explaining 38.8% of the liability variance) and the main effects of a given exposure (explaining 11.1%). There was no interaction effect (AxE). We considered the type I error rate α of the interaction effect, where we regressed the binary phenotype on A, the dichotomized exposure variable (E; prevalence .10) and on the interaction AxE. We set the nominal α at .05. We varied the correlation between the exposure and the genetic variable. Based on 10,000 replications, we observed an inflated type I error rate of the interaction effect as a function of the correlation between the genetic variable and the exposure. However, this inflation was relatively small. The observed type I error rate was .046 (zero correlation), .056 (correlation .15) and .0752 (correlation .30). Note that .056 and .0752 both deviate significantly from the nominal value of .5 (p=.003 and p<.0001, respectively). So in this scenario, which is based on the NESDA and Radiant-UK data, we note that we expect some type I error rate inflation. However, we conclude that the type I error rate inflation in test of GxE in the present setup is small and does not render the test useless. Specifically, in the NESDA and Radiant-UK data the correlation between the genetic variable (polygenic risk score) and the exposure (childhood trauma) is likely to be very low (Table S5).

Simulation study 2

The aim of this simulation study is to aid interpretation of interaction analyses with polygenic risk score (PRS) by simulating different underlying genetic architectures.

Liability threshold model and the *impact of childhood trauma* (CT) on major depressive disorder *(MDD)*

Simulation is based on the liability-threshold model, which can be modeled as MDD underpinned by an unobserved liability, l_{MDD} , where individuals are affected when liability exceeds disease threshold, T_{MDD} . The liability is assumed to be normally distributed and scaled to a population mean of 0 and variance of 1 (which defines T_{MDD} given the prevalence of MDD K_{MDD}), and to result from independent normally distributed environmental (e_{MDD}) and genetic effects (g_{MDD}) with l_{MDD} = $g_{MDD}+e_{MDD}$, where $var(g_{MDD})/var(l_{MDD})=var(g_{MDD})=h_{l,MDD}^2$, the heritability of MDD on the liability scale. Here, we subdivide the environmental effects as $e_{MDD} = CT_{liability scale} +$ $e_{residual, MDD}$. We assume that $CT_{observed scale}$ is represented by a dichotomous measure that labels individuals as exposed (1) or unexposed (0) with an odd ratio for MDD of exposed of OR_{CT} . For a prevalence of MDD of $K_{MDD} = 0.15$, prevalence of CT of $K_{CT} = 0.25$ and $OR_{CT} = 3.2$, the $CT_{observed scale}$ can be transformed to $CT_{liability scale}$ as -0.16 (unexposed) and 0.47 (exposed), and explains 7.4% of variation on the liability scale (Appendix A). Assuming a heritability of MDD of $h_{l,MDD}^2 = 0.35$, the variance explained by the residual environmental effects $e_{residual, MDD}$ follows as 57.6% (assuming that $CT_{liability scale}$, $e_{residual, MDD}$, and g_{MDD} are all independent). For Model 1, we consider CT as part of the environmental effects on MDD, but we note that CT has been found to be heritable itself (2); the consequences of which will be discussed later. In Model 1, we will, further, assume that the genetic and residual environmental effects are equal in those exposed and those unexposed to CT, which can thus be thought of as a "pure additive" model on the liability scale of $CT_{liability scale}$, $e_{residual, MDD}$, and g_{MDD} (i.e. no GxE-interaction). After describing simulation of SNP data, we will discuss decreasing the correlation of SNP-effects between those exposed and those unexposed to CT (Model 2), increasing a genetic contribution to CT through introducing a heritability for CT (Model 3), increasing magnitude of SNP-effects on MDD in those exposed compared to those unexposed to CT (Model 4), and decreasing magnitude of residual environmental effects on MDD in those exposed compared to those unexposed to CT (Model 5).

Simulation of SNP data and genetic effects

We simulated individuals in a population one-by-one until a total of 9,000 cases and 9,000 controls were obtained, from which 10,000 were used as discovery and 8,000 as target set. Therefore, we

Peyrot *et al.* Supplement **Supplement** (Supplement) and the supplement of the supplement \sim

first simulated the SNPs following the method of Golan et al (3), and subsequently modeled CT and MDD. Briefly, the properties of 10,000 SNPs in full linkage equilibrium were first defined by drawing their minor allele frequencies (MAF) from the uniform distribution from 0.05 to 0.5, and a proportion of 30% of these SNPs were set to have an effect on MDD with effects drawn from a normal distribution with variance $h_{l,MDD}^2/3,000$ while the effects of the other SNPs were set at 0. With these SNP effects, an individual i was simulated by first drawing its allele count $(x_{ij}; 0,1 \text{ or } 2)$ with probabilities of $(1 - MAF_j)^2$, $2(1 - MAF_j)MAF_j$, and MAF_j^2 respectively for all SNP j, and, second, defining its genetic effects as $g(i)_{MDD} = \sum_i effect_i(x_{ii} - 2MAF_i)/(2(1 - MAF_i)MAF_i)$. Childhood trauma status of individual i was assigned with probability K_{CT} , and transformed to the liability scale $CT(i)_{liability scale}$ as described in Appendix A. The residual environmental effect $e(i)_{residual, MDD}$ was drawn from a normal distribution with variance $1-h_{l,MDD}^2-var(CT_{liability\, scale})$, so that the liability of individual i followed as $l(i)=g(i)_{MDD}+1$ $CT(i)_{liability scale} + e(i)_{residual, MDD}$. Individual i was deemed affected with MDD when $l(i)$ > T_{MDD} and non affected otherwise, where disease threshold T_{MDD} was defined such that K_{MDD} = $P(z > T_{MDD} | z \sim N(0,1))$. This procedure was repeated until a total of 9,000 cases and 9,000 controls were obtained. Subsequently, a genome-wide association study (GWAS) was conducted with PLINK on 5,000 cases and 5,000 controls (4), the results of which were used to prepare polygenic risk scores in the target set of the other 4,000 cases and 4,000 controls. For every parameterization, the simulation was repeated 10 times.

Simulation - Model 1

For the base assumption of the genetic architecture we assumed a prevalence of MDD of K_{MDD} = 0.15, a heritability of MDD of $h_{l,MDD}^2 = 0.35$, a prevalence of CT of $K_{CT} = 0.25$, no impact of SNPs in CT ($h_{l,CT}^2 = 0$), and odds ratio for MDD in those exposed to childhood trauma of $OR = 3.2$, and pure additivity on the liability scale (identical genetic and residual environmental effects in those exposed and those unexposed to childhood trauma).

Simulation - Model 2

A clear case of GxE interaction would be when the individual SNP-effects on MDD in those exposed would differ from the effects in those unexposed, i.e. when

 $r_q = cor(effect_{SNP} i | cr=1, effect_{SNP} i | cr=0}) = 0$ for the 3,000 effective SNPs. To model this scenario, we further assumed that the effects are on the same 3,000 SNPs and the variance explained is constant, that is $var(effect_{SNP} |_{l \subset T=1}) = var(effect_{SNP} |_{l \subset T=0}) = 0.35.$

Simulation - Model 3

For the Models 1, 2, 4 and 5 we have assumed that CT is purely environmental, but heritability of childhood trauma has been estimated at around 0.5 (2). Therefore, an impact of SNPs effects on CT is considered here. For this, we assume that CT is a "disease trait" itself with underlying liability as described above for MDD (not suggesting that children are to blame for the trauma they experience, rather we hypothesize that heritability arises from transmitted alleles that affect personality characteristics in parents). Nevertheless, we drew SNP-effects for CT from a random normal distribution with variance $h_{l,CT}^2 = 0.5$ and environmental effects from a normal distribution with variance $1-h_{l,CT}^2$ to construct a liability of CT l_{CT} , and individuals were deemed exposed to CT when $l_{CT}(i) > T_{CT}$ with the threshold defined such that $K_{CT} = P(z > T_{CT} | z \sim N(0,1))$. The effects were assigned to the same 3,000 SNPs impacting MDD, but drawn from an independent normal distribution. Given the CT status thus simulated. MDD was derived as described above.

Simulation - Model 4

Another way to think about GxE interaction is that environmental stress might potentiate genetic effects. This was modeled by setting a proportion of genetic effects on MDD in those exposed to those unexposed to CT as $var(effect_{SNPj|CT=1})/var(effect_{SNPj|CT=0}) = 3$ while keeping $cor\left(effects_{NP\ j\mid CT=1}, effect_{SNP\ j\mid CT=0}\right) = 1$. The variances of SNP-effects where chosen in such way that the variance of genetic effects in the full population were fixed at 0.35, while the residual environmental effects had the same variance in those exposed and those unexposed to CT (Appendix B).

Simulation - Model 5

A hypothetical scenario could be that environmental risk factors for MDD (such as socioeconomic status and life-stress in adulthood) cluster in those exposed to CT; the link between these environmental risk factors would be captured in estimates of the OR of CT, but could in addition result in less residual environmental variation in those exposed compared to those unexposed to childhood trauma. We modeled this as $var(e_{residualMDD|CT=1})/var(e_{residualMDD|CT=0}) = 1/$ 3 while assuming constant genetic effects in those exposed and those unexposed to CT, $\frac{effect_{SNP}}{effect_{SNP}}$ $\frac{f cert_{SNP}}{fert_{SNP}}$ $\frac{f cert_{SNP}}{fert_{SNP}}$ (Appendix C).

Appendix A. Transformation of OR to liability scale

To transform the OR from CT on MDD to the liability scale the approach of Witte et al was applied (5). Therefore, the OR (set at 3.2) was first transformed to the RR (2.6) and consequently to the risk

on MDD in exposed $(CT = 1$ with MDD proportion 0.28) and unexposed $(CT = 0$ with MDD in proportion 0.11) assuming a population prevalence of $K_{MDD} = 0.15$ and $K_{CT} = 0.25$. The liability disease threshold for MDD in the full population was found as $T_{MDD, full\ population} = \Phi^{-1}(1 K_{MDD}$) = $\Phi^{-1}(1 - 0.15) = 1.0364$. First assuming a liability variance of 1 in both exposed and unexposed, the threshold in exposed was found as $T_{MDD|CT=1} = \Phi^{-1}(1 - 0.28) = 0.589$ and in unexposed as $T_{MDD|CT=1} = \Phi^{-1}(1 - 0.11) = 1.241$. In line with Witte et al, the mean liability in exposed was found at $\mu_{l|CT=1} = T_{MDD, full\ population} - T_{MDD|CT=1}$ and in unexposed at $\mu_{l|CT=0}$ = $T_{MDD, full\ population} - T_{MDD|CT=0}$, allowing to merge exposed and unexposed while ensuring the disease risks of 0.28 and 0.11 respectively. However, because the variance in both exposed and unexposed was assumed to equal 1, the merged sample had a variance larger than 1 introduced by the variance of CT and a mean slightly different from zero. To ease modeling of genetic effects, we rescaled to mean of zero and variance one, also correcting the disease threshold in this manner. With this, a model was derived transposing CT status of exposed and unexposed to the liability scale, while the overall variance of liability was set at 1, and mean at 0, as usual.

Appendix B. Modeling increased magnitude of SNP-effects in CT=1 compared to CT=0

When aiming to model increased variance of SNP effects in those exposed compared to those unexposed to CT, arbitrary choices have to be made about the residual environmental effects in exposed and unexposed, and the variance of liability, genetic effects and environmental effects in the overall population. We choose to fix the full population variance of liability at 1, variance of genetic effects at $h_{l,MDD}^2 = 0.35$, and variance of environmental effects at $1 - h_{l,MDD}^2 = 0.65$ (the latter including both the variance of $CT_{liability}$ as well as residual environmental effects). To obtain e.g. a variance of genetic effects in exposed three times the variance of genetic effects in unexposed $\left(\text{var}(\text{effect}_{SNP(i|CT=1)}/\text{var}(\text{effect}_{SNP(i|CT=0})=3)\right)$, the variance of genetic effects followed as $var(effect_{SNP\ i\,|\,CT=1}) = 0.56$ and $var(effect_{SNP\ i\,|\,CT=0}) = 0.28$ thereby ensuring that the variance of genetic effect in the full population equals $var(effect_{SNP i})$ =

 $0.25 \mu_{effect_{SNP}j|CT=1}^2 + 0.75 \mu_{effect_{SNP}j|CT=0}^2 - (0.2 \mu_{effect_{SNP}j|CT=1} + 0.8 \mu_{effect_{SNP}j|CT=0})^2 =$ $0.25(0.56 + 0^2) + 0.75(0.28 + 0^2) - 0 = 035$. We choose to fix the residual variance in both exposed and unexposed first at $var(e_{residual|CT=1}) = var(e_{residual|CT=1}) = 0.65$, and the overall variance of liability was thus larger in exposed than in unexposed. As a result, the sums in Appendix A were slightly adjusted as the variance and mean of the merged sample differed slightly to the above, and therefore correction to obtain variance of 1 and mean of zero in the full population also differed.

Appendix C. Decreased environmental variation in individuals exposed to CT

When aiming to model a smaller variance of residual environmental effects in those exposed compared to those unexposed to CT, several model choices have again to be made. We chose to fix the full population variance of liability at 1, variance of genetic effects at $h_{l,MDD}^2 = 0.35$ equal in exposed and unexposed, and variance of environmental effects at $1 - h_{l,MDD}^2 = 0.35$ (the latter including both the variance of $CT_{liability}$ as well as residual environmental effects).

Figure S1. Distribution of the 5-domain continuous childhood trauma measure

Table S1. Demographic information for contributing cohorts of major depressive disorder cases and unaffected controls

CT=childhood trauma

Table S2. Correlation of childhood trauma domains (N=3850)

| | EA | PА | SA | ΕN | PN | SUM | |
|--|-------|-------|-------|-------|-------|------------|--|
| Childhood Trauma Questionnaire subscales (continuous measures) | | | | | | | |
| Emotional Abuse (EA) | 1 | 0.596 | 0.387 | 0.609 | 0.481 | 0.803 | |
| Physical Abuse (PA) | 0.596 | 1 | 0.387 | 0.413 | 0.410 | 0.681 | |
| Sexual Abuse (SA) | 0.387 | 0.387 | 1 | 0.246 | 0.285 | 0.539 | |
| Emotional Neglect (EN) | 0.609 | 0.413 | 0.246 | 1 | 0.632 | 0.805 | |
| Physical Neglect (PN) | 0.481 | 0.410 | 0.285 | 0.632 | 1 | 0.728 | |
| Sum score (SUM) | 0.803 | 0.681 | 0.539 | 0.805 | 0.728 | 1 | |
| Dichotomous indicator of sexual or physical abuse | | | | | | | |
| SA/PA (dichotomous) | 0.367 | 0.542 | 0.754 | 0.203 | 0.201 | 0.497 | |

The Pearson correlation coefficients (all p-value<2e-16) are displayed between the five domains of the Childhood Trauma Questionnaire (CTQ) by applying the residuals of linear regression of the domains on sex and cohort (COFAMS, NESDA, Radiant-UK, SHIP). It can be seen that sexual abuse is slightly less correlated than the other domains, and that there seems no clear distinction between the abuse and neglect domains. In addition, the Spearman's rho correlation coefficient is displayed of the CTQ domains with the dichotomous indicator of sexual abuse and/or physical abuse (SA/PA) that was available for two additional cohorts.

| | COFAMS | DGN | NESDA | QIMR 3 | QIMR 6 | | QIMR C RAD. UK | SHIP-0 | SHIP-T |
|---------------|---------------|--------------------------|--------------|---------|-----------|--------------------------|--------------------------|--------------------------|-----------|
| COFAMS | 771,120 | $\overline{}$ | | | | | $\overline{}$ | | |
| DGN | 741,245 | 1,051,603 | | | | | $\overline{}$ | | |
| NESDA | 675,669 | 851,244 | 924.741 | | | $\overline{}$ | $\overline{}$ | | |
| QIMR 3 | 626,026 | 775,291 | 702.250 | 821.960 | | $\overline{}$ | $\overline{}$ | | |
| QIMR 6 | 716,604 | 930,576 | 822,954 | 803.446 | 1,000,453 | $\overline{}$ | $\overline{}$ | $\overline{}$ | |
| QIMR C | 711,902 | 746,328 | 683.496 | 635.209 | 724,195 | 772,404 | $\overline{}$ | | |
| RAD. UK | 729,795 | 954,007 | 840.621 | 811.506 | 983,793 | 736,767 | 1,028,612 | | |
| SHIP-0 | 706,975 | 905,732 | 907.329 | 737.015 | 871,372 | 713,690 | 890,930 | 992,050 | |
| SHIP-T | 762,091 | 1,037,269 | 903,725 | 809.699 | 981,370 | 765,093 | 1,008,254 | 967,781 | 1,131,800 |
| | | | | | | | | | |

Table S3. Number of overlapping SNPs between cohorts for GRM-based analyses

Table S4. Impact of CTQ subdomain continuous measures on MDD

| | Mean (SD) | | |
|--------------------------|-------------|------------|-----------------|
| Subset | Cases | Controls | OR (p-value) |
| Emotional Abuse | | | |
| Male & Female | 9.3(4.8) | 6.2(2.3) | 2.40 (1.1e-06) |
| Male | 8.5(4.2) | 6.0(2.0) | 2.01 (7.1e-05) |
| Female | 9.6(5.0) | 6.3(2.5) | 2.46 (2.1e-07) |
| Physical Abuse | | | |
| Male & Female | 6.3(2.8) | 5.6(1.6) | 1.51 (4.6e-05) |
| Male | 6.3(2.6) | 5.7(1.6) | $1.41(1.1e-04)$ |
| Female | 6.2(2.9) | 5.5(1.5) | 1.51 (8.8e-05) |
| Sexual Abuse | | | |
| Male & Female | 6.3(3.4) | 5.2(1.3) | $1.64(1.6e-03)$ |
| Male | 5.8(2.3) | 5.1(0.9) | 1.25 (3.4e-03) |
| Female | 6.5(3.8) | 5.3(1.7) | 1.95 (2.9e-03) |
| Emotional Neglect | | | |
| Male & Female | 12.6(5.4) | 8.9(4.0) | 2.08 (8.4e-06) |
| Male | 12.6(5.2) | 9.2(4.1) | 1.87 (2.8e-04) |
| Female | 12.5 (5.4) | 8.6(3.9) | 2.14 (4.7e-06) |
| Physical Neglect | | | |
| Male & Female | 7.8(3.0) | 6.8(2.4) | 1.75 (8.4e-05) |
| Male | 7.9(2.9) | 7.0(2.5) | 1.54 (2.9e-04) |
| Female | 7.8(3.1) | 6.6(2.3) | 1.79 (9.3e-04) |
| Overall CTQ score | | | |
| Male & Female | 42.4 (15.1) | 32.7(8.4) | 2.62 (1.4e-05) |
| Male | 41.3 (13.4) | 33.0(8.2) | $2.18(1.1e-04)$ |
| Female | 42.8 (15.8) | 32.3 (8.6) | 2.74 (3.6e-05) |

 $CTQ =$ Childhood Trauma Questionnaire; MDD = major depressive disorder; OR = odds ratio; SD = standard

deviation

Table S5. Impact of polygenic risk score (based on MDD discovery p<1) on childhood trauma (i.e. gene-environment correlation)

The impact of the polygenic risk scores (PRS) (based on major depressive disorder [MDD] discovery results $p<1$) on childhood trauma (CT) is displayed in all individuals, MDD cases only and controls only for the continuous Childhood Trauma Questionnaire (CTQ) measure covering five domains (applied in main Table 2) and the dichotomous measure covering sexual and/or physical abuse (applied in main Table 3). However, the potential bias of gene-environment correlation in gene-environment interaction analyses depends on the correlation in the full population. Therefore, cases were randomly sampled such that cases/controls=0.15/0.85 to mimic results in the full population. Sampling was repeated 100 times, and conducted for those cohorts with more than 100 controls only. The Pearson correlation was estimated for the continuous CTQ measure, and the Spearman correlation for the dichotomous CT measure, and analyses were corrected for sex and three principal components.

Table S6. Interaction-analyses for male and female separetely with the PRS based on MDD-PRS including all SNPs (discovery $p<1$ in the sample of $N=112,268$)

Table S7. Interaction-analyses for the separate CT domains with the MDD-PRS including all SNPs (discovery p<1)

Sum = sumscore of all five CT domains; EA = Emotional abuse; PA = Physical Abuse ; SA = Sexual Abuse ; EN =

Emotional Neglect ; PN = Physical Neglect

Table S8. Comparing different discovery samples for MDD

Table S9. Polygenic risk scores analyses with simulated data

Simulated data of 10,000 SNPs were based on five models, all assuming heritability of MDD of 0.35, prevalence of MDD of 0.15, prevalence of CT of 0.25 and an odds ratio (OR) of CT on MDD of 3.2 (see Supplemental Methods). Model 1: SNP-effects are the same in exposed and unexposed; Model 2: correlation of 0 between SNP-effects in exposed and unexposed; Model 3: SNP-effects on MDD are the same in exposed and unexposed, heritability of CT of 0.5 (for Models 1,2,4, and 5, heritability of CT was set at 0); Models 4: same direction of SNP-effects in exposed and unexposed (correlation of 1), but 3 times larger variance of effects in exposed than unexposed; Model 5: SNPeffects the same in exposed and unexposed, but three times smaller environmental variance in exposed. Simulation was repeated ten times, the means of which are displayed with the standard error (SE) between brackets.

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