SI for 'From global scaling to the dynamics of individual cities'

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1. Dataset description

The dataset is freely available [\(1\)](#page-5-0) and the methodology is described in the Urban mobility report, 2012 of the Texas A&M Transportation Institute (TTI), College Station, Texas [\(2\)](#page-5-1).

This dataset has also been studied in [\(3\)](#page-5-2) and contains the total hours of delays, excess fuel consumption, and excess *CO*² emission due to congestion for 101 of the largest urban centers in the US. The data spans a 30-year period from 1982 to 2011. Other information such as the population size, number of commuters, the freeway's lane-miles, and the lane-miles of arterial streets, are also available at the same source.

A. Population size. The group of the 101 urban centers described in this dataset is very heterogeneous and contains cities with very different population (see Fig. S1). We see on this figure that indeed the population of the 101 cities varies from 10^5 to

Fig. S1. Histogram for the population of the 101 cities considered in the dataset used in this study.

very large numbers of the order $10⁷$.

B. Spatial distribution of cities. The spatial distribution of the cities in this dataset appears to be uniform as can be seen on the map shown in Fig. [S2.](#page-1-0) This points to the probable absence of spatial bias in the selection of these cities.

Fig. S2. Spatial distribution of the cities studied in the dataset [19].

2. Exponents

A. Type-1 cities. For this set of cities, the total annual delay behaves as

$$
\sim P^{\beta} \tag{S1}
$$

We report in the table [S1](#page-2-0) the list of values for the exponent *β* for cities in this set.

 $δτ$

Table S1. Table for the exponent *β* **for type-1 cities.**

We checked that the exponent is not correlated with for example the final value of the population (see Fig. [S3,](#page-2-1) left), but seems to display some non-negligible correlation with the average growth rate of a city (Fig. [S3,](#page-2-1) right): a linear fit gives a value of −42*.*8 and a *p*−value of order 1% (we have however to be careful with these results as the number of points is small 35). Certainly more work is needed here in order to study and understand these correlations.

Fig. S3. Exponent *β* for type-1 cities versus (left) the population in 2014 and (right) the average population growth rate.

B. Type-2 cities. By definition, for these cities, the total annual delay displays two regimes:

$$
\delta \tau \sim \begin{cases} P^{\beta_1} & \text{when } P < P^* \\ P^{\beta_2} & \text{when } P > P^* \end{cases}
$$
 [S2]

We report in the table $S2$ the values for the exponents β_1 and β_2 computed for cities in this set.

Table S2. Table for the exponents *β*¹ **and** *β*² **for type-2 cities.**

3. Correlation between $β_1$ and $β_2$

For type-2 cities we plot β_2 versus β_1 in the Fig. [S4.](#page-4-0) We observe in this figure that there are no significant correlations between

Fig. S4. Type-2 cities: exponent *β*² versus *β*1.

these exponents (the *p*−value for the regression is 0*.*27 *>* 0*.*05 and we cannot reject the hypothesis of no correlations).

4. Distribution of T^* , P^* , $(\delta \tau/P)^*$

For type-2 cities, we show here the distributions of the quantities defined at the change of slope: *T* ∗ is the time at which the slope happened, P^* is the corresponding population and $(\delta \tau / P)^*$ is the delay per capita when it happened.

Fig. S5. Empirical histograms for T^* , P^* (in unit of million inhabitants) and $(\delta \tau/P)^*$ (in unit of hours). In particular the histogram for $(\delta \tau/P)^*$ shows that the changes of slope in type-2 cities appears approximately at the same value of about 40 hours per year and per capita of congestion induced delay.

Bibliography

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- 1. <http://tti.tamu.edu/documents/ums/congestion-data/complete-data.xlsx>
2. <http://mobility.tamu.edu/ums/methodology>
3. Chang YS, Lee YJ, Choi SS (2017) Is there more traffic congestion in larger cities? -Scaling analysis o

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