Supplement to Usefulness of Released Cardiac Myosin Binding Protein-C as a Predictor of Cardiovascular Events

Carl W. Tong MD PhD^{a,b*}, Giuseppina F. Dusio PhD^{c#}, Suresh Govindan PhD^{d#}, Dustin W. Johnson MD^b, David T. Kidwell MD^b, Lisa M. De La Rosa MD^b, Paola C. Rosas MD PhD RPh^a, Yang Liu MD PhD^a, Elizabeth Ebert MD^b, M. Karen Newell-Rogers MD^a, Jeffrey B. Michel MD^b, Jerome P. Trzeciakowski PhD^a, Sakthivel Sadayappan PhD^{d,e*}

a Department of Medical Physiology, Texas A&M University Health Science Center College of Medicine, Temple, TX; $\frac{b}{c}$ Baylor Scott & White Health – Central Texas, Internal Medicine/Cardiology Division, Temple, TX; ^c University of Arkansas for Medical Sciences, Little Rock, AR; ^dDepartment of Cell and Molecular Physiology, Health Sciences Division, Loyola University Chicago, Maywood, IL; ^eDepartment of Internal Medicine, Heart, Lung and Vascular Institute, Cardiovascular Center, University of Cincinnati College of Medicine, Cincinnati, OH

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*# Equally contributed

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Address for correspondence: Sakthivel Sadayappan, PhD, MBA. Department of Internal Medicine, Heart, Lung and Vascular Institute, Cardiovascular Center, University of Cincinnati College of Medicine, 231 Albert Sabin Way, Cincinnati, OH 45267-0575, USA; Phone: +1 513-558-7498; FAX: +1 513-558-2884; Email: sadayasl@ucmail.uc.edu

Table S1: Diagnoses Encountered for Ordering Exercise Stress Echocardiography

Due to concerns for patients' safety, the study institution's protocol excluded patients with known low left ventricular ejection fraction (LVEF) from exercise stress study. Thus, narrow range of normal LVEFs will not provide bases for good correlation analysis. However, Sa (tissue Doppler of maximum myocardial contractile velocity about the mitral valve annulus during systole), which corresponds to $+dP/dt$ ax of contraction,¹ can be used as measure of contractility. The basal tissue Doppler Sa correlated with pre-stress cMyBP-C (Table S2). This correlation grew stronger when subjects who suffered the primary outcome were removed from the analysis (Table S2). The basal tissue Doppler e', which is peak myocardial relaxation velocity during early diastole, did not correlate with pre-stress cMyBP-C. The basal E/e' ratio, where E is the peak blood flow velocity Doppler during early diastole, did not correlate with pre-stress cMyBP-C. Thus, contractility positively correlated with pre-stress cMyBP-C levels but diastolic function (e' and E/e' ² did not correlate with pre-stress cMyBP-C levels.

Table S2: Correlation of Basal Echocardiography to Pre-Stress Serum Cardiac Myosin Binding Protein-C Levels

Variables	Pearson	Two-tail
	Correlation r	
Pre-Stress Sa vs. Pre-Stress cMyBP-C	0.161	$0.047*$
Pre-Stress Sa vs. Pre-Stress cMyBP-C in subjects who did not	0.276	$0.002*$
suffer a primary outcome event		
Pre-Stress e' vs. Pre-Stress cMyBP-C	0.026	0.753
Pre-Stress e' vs. Pre-Stress cMyBP-C in subjects who did not	0.115	0.197
suffer a primary outcome event		
Pre-Stress E/e' vs. Pre-Stress cMyBP-C	-0.087	0.289
Pre-Stress E/e' vs. Pre-Stress cMyBP-C in subjects who did	-0.169	0.057
not suffer a primary outcome event		

Cardiac Myosin Binding Protein-C (cMyBP-C). e': tissue Doppler of peak myocardial relaxation velocity about the mitral valve annulus during early diastole. E: Doppler of peak blood inflow velocity to the left ventricle across the mitral valve during early diastole. Sa: tissue Doppler of peak myocardial contraction velocity during systole. $*$ p<0.05 for existence of significant correlation.

There are 9 subjects in the study who did not have cardiovascular risk factors, exhibited normal echocardiographic findings, performed normal exercise stress echocardiography, and did not suffer primary outcome. Consequently, these 9 subjects can be used as true normal for the study. This true normal population appeared to differ from the rest of the study population by lower cMyBP-C levels. Like rest of the study population, exercise stress also increased serum cMyBP-C in normal subgroup. However, the apparent lower cMyBP-C levels are significant only by using unequal variances based analysis (See Table S3).

	True Normal $n=9$	All Others $n=149$	P for unequal variances	P for equal variances
Pre-Stress cMyBP-C (ng/ml)	205 ± 40	$708 \pm 113*$	< 0.001	0.278
Post-Stress cMyBP-C (ng/ml)	254 ± 41 #	997 ± 155 *#	< 0.001	0.241
Absolute Change in $cMyBP-C$ (ng/ml)	49 ± 6	$289 \pm 50*$	< 0.001	0.270
% Change in cMyBP-C	$29 + 5$	$53 + 7$	0.007	0.410

Table S3: Comparison of True Normal to Rest of Study Population

*denotes significant difference p<0.05 from true normal (chose p for unequal variance if standard deviations were very different). # denotes significant difference $p<0.05$ from its corresponding pre-stress values. Values are Mean \pm SE.

Power analyses were done to assess the ability of serum cMyBP-C in differentiating a population that is more likely to suffer cardiovascular disease event from a normal population. Versions of SPSS and Stata used for the analyses did not have the capability to perform power analyses on this study's particular type of data. Thus, power analyses were done manually.

For estimating the sample size needed for a PreC or PostC as dichotomous covariates, we used the formula developed by Schoenfeld $(1983)^3$:

$$
n = \frac{(z_{1-\alpha/2} + z_{1-\beta})^2}{PR(1-R)(\log B)^2}
$$

Where, as above, α = significance level (.05); $(1 - \beta)$ = desired power; *B* = Hazard ratio to be detected (8.1 for PreC; 4.8 for PostC); P = proportion of subjects exhibiting the primary outcome; R = proportion of subjects for whom PreC = 1 (0.77) or PostC = 1 (0.57) - i.e. number of subjects with [cMyBP-C] above the respective threshold concentrations. For multiple Cox Regresssion, the estimated *n*'s were adjusted based on the Pseudo- R^2 between PreC or PostC and the remaining covariates. We then used the known n=158, R, α = 0.05, and hazard ratio to calculate power = (1 – β). The calculated power is > 0.999 for both PreC and PostC.

Because cMyBP-C follows a log-normal distribution, we used log10([cMyBP-C]) in our Cox regression and power analysis. We use the following formula (Hsieh and Lavori $(2000)^4$ for the estimated sample size *n* for a continuous covariate *x* when *x* is normally distributed:

$$
n = \frac{\left(z_{1-\alpha/2} + z_{1-\beta}\right)^2}{P\sigma^2 (\log B)^2}
$$

Where α = significance level (.05); $(1 - \beta)$ = desired power; *B* = Hazard ratio to be detected (13.8) for pre-stress or 14.3 for post-stress); $P =$ proportion of subjects exhibiting critical events; $\sigma =$ variance of x (0.719 for pre-stress log10[cMyBP-C] or 0.727 for post-stress log10([cMyBP-C]). For multiple Cox Regresssion, the estimated *n*'s were adjusted based on the multiple correlation coefficient, ρ, between log[cMyBP-C] and the remaining covariates:

$$
n^* = \frac{n}{1 - \rho^2}
$$

We then used the known n=158, R, $\alpha = 0.05$, and hazard ratios to calculate power = $(1 - \beta)$. The calculated power is >0.999 for both pre-stress and post-stress cMyBP-C levels.

Although the power analyses suggest that circulating cMyBP-C has >99% chance of finding population differences, we caution about exuberance of its potential use. The positive findings of the current study population of patients who were referred for stress echocardiography likely caused the very high calculated power. A more generalize population may provide a different yield. Power provides a measure on the ability of a method to differentiate between populations; however, the ability of serum cMyBP-C to predict outcome on individual bases will vary.

References for Supplement

1. Seo JS, Kim DH, Kim WJ, Song JM, Kang DH, Song JK. Peak systolic velocity of mitral annular longitudinal movement measured by pulsed tissue Doppler imaging as an index of global left ventricular contractility. *Am J Physiol Heart Circ Physiol* 2010;298:H1608-1615.

2. Nagueh SF, Smiseth OA, Appleton CP, Byrd BF, 3rd, Dokainish H, Edvardsen T, Flachskampf FA, Gillebert TC, Klein AL, Lancellotti P, Marino P, Oh JK, Popescu BA, Waggoner AD. Recommendations for the Evaluation of Left Ventricular Diastolic Function by Echocardiography: An Update from the American Society of Echocardiography and the European Association of Cardiovascular Imaging. *J Am Soc Echocardiogr* 2016;29:277-314.

3. Schoenfeld DA. Sample-size formula for the proportional-hazards regression model. *Biometrics* 1983;39:499-503.

4. Hsieh FY, Lavori PW. Sample-size calculations for the Cox proportional hazards regression model with nonbinary covariates. *Control Clin Trials* 2000;21:552-560.