

# Primary Analysis of SARA

The primary analysis for the two hypotheses mentioned in the “Inference criteria” section is as follows:

## 1 Hypothesis 1: 4PM push notification

### 1.1 Overview of hypothesis

$H_1$ : The 4PM push notification with inspirational quote will increase the full completion of survey and/or active task the same day as compared to no inspirational quote ( $p < 0.025$ )

$H_0$ : The 4PM push notification with inspirational quote will *not* increase the full completion of survey and/or active task the same day as compared to no inspirational quote

- Primary outcome: participants fully complete the survey and/or active tasks in the evening of the same day. This outcome is binary.
- Independent variable: push notification with an inspirational message vs no push notification with an inspirational message. This variable is binary.
- Covariates: whether the survey and/or active tasks were fully completed prior-day (binary); whether text or phone calls were made in the last 24 hours, i.e. after 4PM from the previous day to before 4PM the current day (binary); whether the app was opened in last the prior 72 hours outside of when survey and/or active task were completed, i.e. after 4PM from the 3-days ago to before 4 PM the current day (binary)

### 1.2 Notation

Denote the primary outcome for person  $i$  on day  $t$  ( $t = 1, \dots, 30$  and  $i = 1, \dots, n$  where  $n$  is the total sample size.) by  $Y_{it}$ . There are 3 binary covariates in both of the above analyses. Denote the 3 by 1 covariate vector for the  $i$ th person on the  $t$ th day by  $X_{it}$ . Denote the binary indicator of treatment for the  $i$ th person on the  $t$ th day by  $A_{it}$ .

### 1.3 Construction of the Test Statistic

We will use a multiplicative structural nested log linear model for a binary outcome. To conduct the hypothesis tests we estimate the marginal effect of treatment on the log linear scale. The marginalization is over time and the individual's data,  $H_{it}$ . Technically this marginal effect is given by:

$$\log \left( \frac{\sum_{t=1}^{30} E[Y_{it} = 1 | A_{it} = 1]}{\sum_{t=1}^{30} E[Y_{it} = 1 | A_{it} = 0]} \right) = \beta_0.$$

$\beta_0$  is a regression coefficient in a multiplicative structural nested log-linear mean model [1, 2].

We use an analysis method that permits us to include covariates,  $X_{it}$  to reduce the noise. In particular we use the working model:

$$e^{X_{it}^T \alpha_0} \approx E[e^{-A_{it} \beta_0} Y_{it} | X_{it}].$$

This model need not be correct for the estimator of  $\beta_0$  to be consistent. This model is used to reduce estimation variance.

We estimate the marginal treatment effect by solving

$$0 = \sum_{i=1}^n \sum_{t=1}^{30} e^{-A_{it} \hat{\beta}} \left( Y_{it} - e^{X_{it}^T \hat{\alpha} + A_{it} \hat{\beta}} \right) \begin{pmatrix} (A_{it} - 1/2) \\ -e^{X_{it}^T \hat{\alpha}} X_{it} \end{pmatrix} \quad (1)$$

to obtain  $\hat{\beta}$  and  $\hat{\alpha}$ .

The approximate variance-covariance matrix

$$\text{Var} \begin{pmatrix} \sqrt{n}(\hat{\beta} - \beta_0) \\ \sqrt{n}(\hat{\alpha} - \alpha_0) \end{pmatrix} = M_n^{-1} \Sigma_n (M_n^{-1})^T$$

where

$$M_n = (1/n) \sum_{i=1}^n \sum_{t=1}^{30} \begin{pmatrix} (A_{it} - 1/2) e^{-A_{it} \hat{\beta}} Y_{it} A_{it} & (A_{it} - 1/2) e^{X_{it}^T \hat{\alpha}} X_{it}^T \\ e^{-A_{it} \hat{\beta}} Y_{it} A_{it} e^{X_{it}^T \hat{\alpha}} X_{it} & e^{X_{it}^T \hat{\alpha}} X_{it} X_{it}^T \end{pmatrix}$$

and

$$\Sigma_n = (1/n) \sum_{i=1}^n \left( \sum_{t=1}^{30} \left( e^{-A_{it} \hat{\beta}} Y_{it} - e^{X_{it}^T \hat{\alpha}} \right) \begin{pmatrix} (A_{it} - 1/2) \\ e^{X_{it}^T \hat{\alpha}} X_{it} \end{pmatrix} \right) \left( \sum_{t=1}^{30} \left( e^{-A_{it} \hat{\beta}} Y_{it} - e^{X_{it}^T \hat{\alpha}} \right) \begin{pmatrix} (A_{it} - 1/2) \\ e^{X_{it}^T \hat{\alpha}} X_{it} \end{pmatrix} \right)^T$$

### 1.4 Test Statistic

For the test  $H_0 = \beta_0 \leq 0$  versus  $H_1 = \beta_0 > 0$ , we reject  $H_0$  in favor of  $H_1$  if  $T_n > z_{0.025}$  where  $z_{0.025}$  is the 97.5th percentile of the standard normal distribution and  $T_n = \frac{\sqrt{n} \hat{\beta}}{\sigma_n}$  and  $\sigma_n$  is the square root of the (1, 1) entry in  $M_n^{-1} \Sigma_n (M_n^{-1})^T$ .

## 2 Hypothesis 2: After-survey-completion reward

### 2.1 Overview of hypothesis

$H_1$ : Among individuals who complete the survey, offering a post-survey-completion meme will increase the full completion of survey and/or action task the next day as compared to not offering a meme after survey completion ( $p < 0.025$ )

$H_0$ : Among individuals who complete the survey, offering a post-survey-completion meme will *not* increase the full completion of survey and/or action task the next day as compared to not offering a meme after survey completion

- Primary outcome: whether participants fully complete the survey and/or active tasks the following day. This outcome is binary
- Independent variable: offering a meme vs. not offering a meme after survey completion. This variable is binary.
- Covariates: whether the survey and/or active tasks were fully completed the prior day (binary); whether text or phone calls were made in the last 30 hours (binary); whether the app was opened in the prior last 80 hours outside of when survey and/or active task were completed, i.e. after 6PM from the 3-days ago to before data collection time on the current day (binary)

### 2.2 Notation

Denote the primary outcome for person  $i$  on day  $t$  ( $t = 1, \dots, 29$  and  $i = 1, \dots, n$  where  $n$  is the total sample size.) by  $Y_{i(t+1)}$ . There are 3 binary covariates in both of the above analyses. Denote the 3 by 1 covariate vector for the  $i$ th person on the  $t$ th day by  $X_{it}$ . Denote the binary indicator of treatment for the  $i$ th person on the  $t$ th day by  $A_{it}$ . In this second analysis, person  $i$  is only available for treatment on day  $t$  if this person fully completed the survey. Let  $I_{it} = 1$  if person  $i$  completed the survey on day  $t$  and set  $I_{it} = 0$  otherwise. Note that in the first analysis  $I_{it} = 1$  for all  $i, t$ .

### 2.3 Construction of the Test Statistic

We will use a multiplicative structural nested log linear model for a binary outcome. To conduct the hypothesis tests we estimate the marginal effect of treatment on the log linear scale. The marginalization is over time and the individual's data,  $H_{it}$ . Technically this marginal effect is given by:

$$\log \left( \frac{\sum_{t=1}^{29} E[Y_{i(t+1)} = 1 | A_{it} = 1, I_{it} = 1]}{\sum_{t=1}^{29} E[Y_{i(t+1)} = 1 | A_{it} = 0, I_{it} = 1]} \right) = \beta_0.$$

$\beta_0$  is a regression coefficient in a multiplicative structural nested log-linear mean model [1, 2].

We use an analysis method that permits us to include covariates,  $X_{it}$  to reduce the noise. In particular we use the working model:

$$e^{X_{it}^T \alpha_0} \approx E[e^{-A_{it} \beta_0} Y_{i(t+1)} | X_{it}, I_{it} = 1].$$

This model need not be correct for the estimator of  $\beta_0$  to be consistent. This model is used to reduce estimation variance.

We estimate the marginal treatment effect by solving

$$0 = \sum_{i=1}^n \sum_{t=1}^{29} I_{it} e^{-A_{it} \hat{\beta}} \left( Y_{i(t+1)} - e^{X_{it}^T \hat{\alpha} + A_{it} \hat{\beta}} \right) \begin{pmatrix} (A_{it} - 1/2) \\ -e^{X_{it}^T \hat{\alpha}} X_{it} \end{pmatrix} \quad (2)$$

to obtain  $\hat{\beta}$  and  $\hat{\alpha}$ .

The approximate variance-covariance matrix

$$\text{Var} \begin{pmatrix} \sqrt{n}(\hat{\beta} - \beta_0) \\ \sqrt{n}(\hat{\alpha} - \alpha_0) \end{pmatrix} = M_n^{-1} \Sigma_n (M_n^{-1})^T$$

where

$$M_n = (1/n) \sum_{i=1}^n \sum_{t=1}^{29} I_{it} \begin{pmatrix} (A_{it} - 1/2) e^{-A_{it} \hat{\beta}} Y_{i(t+1)} A_{it} & (A_{it} - 1/2) e^{X_{it}^T \hat{\alpha}} X_{it}^T \\ e^{-A_{it} \hat{\beta}} Y_{i(t+1)} A_{it} e^{X_{it}^T \hat{\alpha}} X_{it} & e^{X_{it}^T \hat{\alpha}} X_{it} X_{it}^T \end{pmatrix}$$

and

$$\Sigma_n = (1/n) \sum_{i=1}^n \left( \sum_{t=1}^{29} I_{it} \left( e^{-A_{it} \hat{\beta}} Y_{i(t+1)} - e^{X_{it}^T \hat{\alpha}} \right) \begin{pmatrix} (A_{it} - 1/2) \\ e^{X_{it}^T \hat{\alpha}} X_{it} \end{pmatrix} \right) \left( \sum_{t=1}^{29} I_{it} \left( e^{-A_{it} \hat{\beta}} Y_{i(t+1)} - e^{X_{it}^T \hat{\alpha}} \right) \begin{pmatrix} (A_{it} - 1/2) \\ e^{X_{it}^T \hat{\alpha}} X_{it} \end{pmatrix} \right)^T$$

## 2.4 Test Statistic

For the test  $H_0 = \beta_0 \leq 0$  versus  $H_1 = \beta_0 > 0$ , we reject  $H_0$  in favor of  $H_1$  if  $T_n > z_{0.025}$  where  $z_{0.025}$  is the 97.5th percentile of the standard normal distribution and  $T_n = \frac{\sqrt{n} \hat{\beta}}{\sigma_n}$  and  $\sigma_n$  is the square root of the (1, 1) entry in  $M_n^{-1} \Sigma_n (M_n^{-1})^T$ .

## References

- [1] James M Robins. Correcting for non-compliance in randomized trials using structural nested mean models. *Communications in Statistics-Theory and methods*, 23(8):2379–2412, 1994.
- [2] James M Robins. Causal inference from complex longitudinal data. *Latent variable modeling and applications to causality*, pages 69–117, 1997.