Supplementary Information

Temporal and demographic variation in partial migration of the North Atlantic right whale

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Supplement S1

Multistate capture-recapture models that include an unobservable state have been used to estimate survival and temporary emigration probabilities. However, some parameters in these models may not be identifiable when there is only one sampling occasion during each primary period [20, 22]. Even when constraints are imposed to reduce model complexity, parameter estimates can still have poor precision or be biased due to finite sample sizes (Table S1). A robust sampling design [21, 24] with multiple secondary sampling occasions during each primary period can improve parameter estimability (Table S2).

The robust design, however, typically requires a closure assumption (i.e., all individuals in an observable state are present and available for capture during all sampling occasions within a primary period) that is often difficult to achieve in practice. For example, right whales appear to have staggered arrival to and departure from their southeastern U.S. wintering grounds [47]. Kendall [25] presented an "emigration only" model that relaxes this closure assumption by requiring all observable individuals to be present and available for capture on the first sampling occasion but permits them to become unavailable (*e.g*., leave the study area) before sampling ends (on subsequent sampling occasions). All sampling occasions after the first are pooled together, resulting in two sampling occasions per primary period for analysis.

In addition to allowing emigration before sampling ends, we extended Kendall's model to allow individuals to arrive after the first sampling occasion and to have multiple arrivals and departures within a primary period. In our model, we assume that for a primary period with *T* total sampling occasions (*e.g*., survey days), no new individuals arrive after sampling occasion *ta*, where $t_a \leq T$. Data are pooled for analysis across occasions 1 through t_a and $t_a + 1$ through *T*, resulting in two sampling occasions per primary period with the same assumptions as Kendall's emigration only model: one occasion where all individuals are available for capture and another where some individuals may have become unavailable. Performance of this model (Table S3) was comparable to that of the robust design model requiring full closure (Table S2). Note that this closure violation reduces the effective capture probabilities, and the level of reduction is determined by the underlying arrival and departure probabilities.

Table S1. Results of multistate model with one sampling occasion per primary period from 1,000 simulated datasets. Datasets were based on a Jolly-Seber model with an initial population size of 100 individuals, 8 primary periods each with a constant recruitment probability of 0.3, and capture probability ($p = 0.7$) and survival probability ($S = 0.9$) constant across time. Transition probability from observable to unobservable state $(\psi_{t,A,X})$ and transition probability from unobservable to observable state $(\psi_{t, X, A})$ varied across time. Model metrics include the mean, coefficient of variation (CV), root mean square error normalized by the true parameter value

(rRMSE), and coverage (from 95% confidence limits) for parameter estimates across all simulations. Coverage values < 0.85 in bold; transition probability from unobservable to observable state during the first time interval $(\psi_{2,X,A})$ is intrinsically unidentifiable.

Table S2. Results of multistate model with multiple sampling occasions per closed primary period from 1,000 simulated datasets. Datasets were based on a Jolly-Seber model with an initial population size of 100 individuals, 8 primary periods each with 3 secondary sampling occasions and a constant recruitment probability of 0.3, and capture probability ($p = 0.33$ for each secondary occasion) and survival probability $(S = 0.9)$ constant across time. Transition probability from observable to unobservable state $(\psi_{tA,X})$ and transition probability from unobservable to observable state $(\psi_{t,X,A})$ varied across time. Model metrics are same as in Table S1. Coverage values < 0.85 in bold; transition probability from unobservable to observable state during the first time interval $(\psi_{2,X,A})$ is intrinsically unidentifiable.

Table S3. Results of multistate model with multiple sampling occasions per open primary period from 1,000 simulated datasets. Datasets were based on a Jolly-Seber model with an initial population size of 100 individuals, 8 primary periods each with 6 secondary sampling occasions and a constant recruitment probability of 0.3, and capture probability ($p = 0.7$ for each secondary occasion) and survival probability $(S = 0.9)$ constant across time. Transition probability from observable to unobservable state $(\psi_{t,A,X})$ and transition probability from unobservable to observable state $(\psi_{t,X,A})$ varied across time. Additional parameters (not estimated by the model) for each open primary period were constant across primary periods and as follows: arrival (and re-entry) probabilities prior to each secondary sampling occasion ($a_1 = 0.1$, $a_2 = 0.2$, $a_3 = 0.3$, a_4 $= 0.4$, $as = 0$, $as = 0$) and departure probabilities for each secondary occasion interval ($d_{1,2} = 0.1$, $d_{2,3} = 0.2$, $d_{3,4} = 0.3$, $d_{4,5} = 0.7$, $d_{5,6} = 0.9$). No arrivals occur in secondary occasions 5 or 6; data were pooled for analysis across occasions 1 through 4 into *p1*, and 5 through 6 into *p2*. Model metrics are same as in Table S1. Coverage values < 0.85 in bold; transition probability from unobservable to observable state during the first time interval $(\psi_{2,X,A})$ is intrinsically unidentifiable.

 Table S4. Number of individual North Atlantic right whales identified in the southeastern U.S. seasonal management area by Early Warning System aerial surveys by demographic group, 1994-2015.

Figure S1. Number of individual North Atlantic right whales identified in the southeastern U.S. seasonal management area by Early Warning System aerial surveys, 1994-2015. Excludes first-year calves.

Figure S2. Gulf of Maine annual *C. finmarchicus* anomaly for spring (January – June) in the western Gulf of Maine (west of 68°W; open squares, dashed line) and for fall (July – December) in the eastern Gulf of Maine (east of 68°W; closed triangles, solid line). Data from NOAA EcoMon plankton surveys.

Supplement S3

Table S5. Model selection results for capture (*p*), survival (*S*), and state transition (ψ) probabilities. Models with the lowest AICc (Akaike's Information Criterion corrected for small sample sizes) values were deemed the best model in each step. Models for *p* include interaction effects of state and survey effort, additive effects of state and effort, effect of survey effort, effect of state, and no effects (.). Models for *S* include no effects and an effect of age-class (juveniles different than adults). Models for ψ include eight different models for demographic effects that interact with eleven time effects. Demographic models (see text for more details) include: separate intercepts and time effects for each demographic group (1); non-Markovian transitions (2); no temporary emigration (3); no differences between juveniles and adults (4); separate intercepts, but time effects additive across all groups (5); separate intercepts but no differences across sex, and time effects additive within reproductive state (6); separate intercepts, and time effects additive within reproductive state (7); and separate intercepts but no differences between juveniles and adults transitioning from the breeding state, and time effects additive within reproductive state and age-class (8). Time effects include: no variation across time (.); full time variation (time); summer Gulf of Maine sea surface temperature anomaly (GoM_SST); North Atlantic Oscillation index from the concurrent winter (NAO); North Atlantic Oscillation index from two winters prior (NAO_lag2); Gulf of Maine *Calanus finmarchicus* annual index (CAL); Gulf of Maine *C. finmarchicus* annual index averaged over the preceding two years (CAL_2avg); western Gulf of Maine *C. finmarchicus* spring index (WCAL); western Gulf of Maine *C. finmarchicus* spring index averaged over the preceding two years (WCAL_2avg); eastern Gulf of Maine *C. finmarchicus* fall index (ECAL); and eastern Gulf of Maine *C. finmarchicus* fall index averaged over the preceding two years (ECAL_2avg).

Step 2: *S*

			Parameters Deviance AICc		AAICc
	1 time	state*effort	280	4932.2 5607.1	0.0
ageclass 1.time		state*effort	281	4933.4 5611.2	4 I

