

Deprivation-specific life tables using multivariable flexible modelling - trends from 2001 to 2011, Portugal

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Supplementary Material S1

Estimation of confidence intervals for model-based predicted mortality rate ratios

Considering a GLM model with Poisson error for the number of deaths by age, deprivation quintile and period:

$$\begin{aligned} \log(d_{age,i,j}) = & \beta_0 + f(age) + \sum_{k=2}^5 \beta_{dep_k} \cdot dep_k + g_1(age * dep_i) + \beta_{period_j} \cdot period_j + \\ & + \sum_{k=2}^5 \beta_{dep_period_k} \cdot period_j * dep_k + g_2(age * period_j) + \log(pyr_{s_{age,i,j}}) \end{aligned}$$

Reference categories:

Deprivation: first quintile (least deprived)

Period: 2000-02

Considering linear interaction between age and deprivation:

$$g_1(age * dep_i) = \sum_{i=2}^5 \beta_{age_dep_i} \cdot age * dep_i$$

Considering interaction of period with splines for age with 1 internal knot:

$$g_2(age * period_j) = \beta_{period_age_1} \cdot (period_j * age)_S + \beta_{period_age_2} \cdot v(period_j * age),$$

where $(period_j * age)_S$ represents the standardised $period * age$ variable and $v(period_j * age)$ represents the orthogonalised spline basis.

Mortality rate:

$$\begin{aligned} \log(R_{age,i,j}) = & \beta_0 + f(age) + \sum_{k=2}^5 \beta_{dep_k} \cdot dep_k + g_1(age * dep_i) + \beta_{period_j} \cdot period_j + \\ & + \sum_{k=2}^5 \beta_{dep_period_k} \cdot period_j * dep_k + g_2(age * period_j) \end{aligned}$$

where

$$\log(R_{age,i,j}) = \log(d_{age,i,j}) - \log(pyr_{age,i,j}) = \log\left(\frac{d_{age,i,j}}{pyr_{age,i,j}}\right)$$

Mortality rate ratio between two different deprivation groups, same period:

$$\begin{aligned} \log\left(\frac{R_{age,i=a,j}}{R_{age,i=b,j}}\right) &= \\ &= \beta_0 + f(age) + \beta_{d_{i=a}} + \beta_{age_dep=a} \cdot age + \beta_{period_j} \cdot period_j + \beta_{dep_period=a} \cdot period_j + \\ &+ g_2(age * period_j) - \\ &[\beta_0 + f(age) + \beta_{d_{i=b}} + \beta_{age_dep=b} \cdot age + \beta_{period_j} \cdot period_j + \beta_{dep_period=b} \cdot period_j + \\ &+ g_2(age * period_j)] = \\ &= \beta_{d_{i=a}} + \beta_{age_dep=a} \cdot age + \beta_{dep_period=a} \cdot period_j - (\beta_{d_{i=b}} + \beta_{age_dep=b} \cdot age + \beta_{dep_period=b} \cdot period_j) \\ &= (\beta_{d_{i=a}} - \beta_{d_{i=b}}) + (\beta_{age_dep=a} - \beta_{age_dep=b}) \cdot age + (\beta_{dep_period=a} - \beta_{dep_period=b}) \cdot period_j \end{aligned}$$

Calculate Confidence Interval assuming normality of $\log(RR)$.

The variance of the $\log(RR)$ can be estimated using the delta method:

$$VAR[\log(RR)(\beta)] \simeq \left[\frac{\partial \log(RR)}{\partial \beta}\right]_{\beta=\hat{\beta}} \times VAR[\hat{\beta}] \times \left[\frac{\partial \log(RR)}{\partial \beta}\right]_{\beta=\hat{\beta}}^T$$

$$\left[\frac{\partial \log(RR)}{\partial \beta}\right]_{\beta=\hat{\beta}} = \begin{bmatrix} 1 & -1 & age & -age & period & -period \end{bmatrix}$$

$$VAR[\hat{\beta}] = \begin{bmatrix} VAR[\beta_{d_{i=a}}] & COV[\beta_{d_{i=a}}, \beta_{d_{i=b}}] & \cdot & \cdot & \cdot & \cdot \\ \cdot & VAR[\beta_{d_{i=b}}] & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & COV[\beta_{dep_period=b}, \beta_{dep_period=a}] & VAR[\beta_{dep_period=b}] \end{bmatrix}$$

The 95% confidence interval for the mortality rate ratio is then given by:

$$\exp\left[\log\left(\frac{R_{age,i=a,j}}{R_{age,i=b,j}}\right) \pm 1.96 \times \sqrt{VAR[\log(RR)]}\right]$$

Considering that the deprivation group $i = b$ is the reference group, the expressions above simplify to:

$$\begin{aligned} \log\left(\frac{R_{age,i=a,j}}{R_{age,i=b,j}}\right) &= \\ &= \beta_{d_i=a} + \beta_{age_dep=a} \cdot age + \beta_{dep_period=a} \cdot period_j \end{aligned}$$

$$\left[\frac{\partial \log(RR)}{\partial \beta}\right]_{\beta=\hat{\beta}} = \begin{bmatrix} 1 & age & period \end{bmatrix}$$

$$VAR[\hat{\beta}] = \begin{bmatrix} VAR[\beta_{d_i=a}] & COV[\beta_{d_i=a}, \beta_{age_dep=a}] & \cdot \\ \cdot & VAR[\beta_{age_dep=a}] & \cdot \\ \cdot & COV[\beta_{age_dep=a}, \beta_{dep_period=a}] & VAR[\beta_{dep_period=a}] \end{bmatrix}$$

Mortality rate ratio between two different periods ($period = 1/period = 0$), same deprivation group:

$$\begin{aligned} \log\left(\frac{R_{age,i,j=1}}{R_{age,i,j=0}}\right) &= \\ &= \beta_0 + f(age) + \beta_{d_i} + \beta_{age_dep=i} + \beta_{period_1} + \beta_{dep_period=i1} + g_2(age * period_1) - \\ &- (\beta_0 + f(age) + \beta_{d_i} + \beta_{age_dep=i} + g_2(age * period_0)) = \\ &= \beta_{period_1} + \beta_{dep_period=i1} + g_2(age * period_1) - g_2(age * period_0) \end{aligned}$$

Assuming only one knot for the interaction age*period as stated above:

$$\begin{aligned} \log\left(\frac{R_{age,i,j=1}}{R_{age,i,j=0}}\right) &= \beta_{period_1} + \beta_{dep_period=i1} + \beta_{period_age_1} \cdot ((period_1 * age)_S - (period_0 * age)_S) + \\ &+ \beta_{period_age_2} \cdot (v(period_1 * age) - v(period_0 * age)) \end{aligned}$$

Again, the variance of the $\log(RR)$ can be estimated using the delta method:

$$VAR[\log(RR)(\beta)] \simeq \left[\frac{\partial \log(RR)}{\partial \beta}\right]_{\beta=\hat{\beta}} \times VAR[\hat{\beta}] \times \left[\frac{\partial \log(RR)}{\partial \beta}\right]_{\beta=\hat{\beta}}^T$$

$$\left[\frac{\partial \log(RR)}{\partial \beta}\right]_{\beta=\hat{\beta}} = \begin{bmatrix} 1 & 1 & (period_1 * age)_S - (period_0 * age)_S & v(period_1 * age) - v(period_0 * age) \end{bmatrix}$$

$$VAR[\hat{\beta}] = \begin{bmatrix} VAR[\beta_{period_1}] & COV[\beta_{period_1}, \beta_{dep_period=i1}] & \cdot & \cdot \\ \cdot & VAR[\beta_{dep_period=i1}] & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & COV[\beta_{period_age_2}, \beta_{period_age_1}] & VAR[\beta_{period_age_2}] \end{bmatrix}$$

The 95% confidence interval for the mortality rate ratio is then given by:

$$\exp \left[\log \left(\frac{R_{age,i,j=1}}{R_{age,i,j=0}} \right) \pm 1.96 \times \sqrt{VAR[\log(RR)]} \right]$$