Supplemental Materials: Is there a dose-response relation of dietary glycemic load to risk of type-2 diabetes? Metaanalysis of prospective cohort studies¹⁻⁵

Geoffrey Livesey¹ Richard Taylor² Helen Livesey¹ and Simin Liu³

¹ From Independent Nutrition Logic, Wymondham, Norfolk NR18 0QX, United Kingdom.

² From Merton College, Oxford, United Kingdom.

³ From the Center for Metabolic Disease Prevention and Departments of Epidemiology,

Medicine, and Obstetrics and Gynecology, University of California, Los Angeles, CA 90095-

1772, USA

⁴ The review was initiated by Independent Nutrition Logic Ltd and funded unconditionally by Beneo-Palatinit Gmbh, Mannheim, Germany, a producer of low glycemic carbohydrates. Acknowledged authors of reviewed publications provided additional data.

⁵ Address correspondences to G. Livesey. Independent Nutrition Logic Ltd, 21 Bellrope Lane, Wymondham, Norfolk NR18 0OX, United Kingdom. Email: glivesey@inlogic.co.uk

SUPPLEMENTAL MATERIALS

Search strategies

Excluded studies

Reasons for exclusion.

Included studies

Table S1. Extracted and calculated data for the included studies.

Table S2. Study identities, region, ethnicities, outcome ascertainment, population sample size, and number of cases accumulated.

 Table S3. Characteristics of the dietary instrument used.

Table S4. Characteristics of the study participants, duration of study, number o

 quantiles, and study baseline exclusions.

Table S5. Assessment of assumptions about accuracy of data used in the two-stage meta-analysis with covariates applied to all 24 studies.

Supplemental analyses

Table S6. Parameter estimates according the two-step meta-analysis approaches used aside a one-step approach, each on the full dataset of 24 studies.

Table S7. Influence of study factors on incremental RR values and β -coefficients for the covariates SEX, CORR, FUY and ETH (n=24 studies).

FIGURE S1. Meta-analysis of change in RR with increase in GL from lowest to highest quantile, by sex group.

FIGURE S2. Funnel plot of residuals for model 5 obtained by the two-step approach to meta-analysis.

FIGURE S3. Factors hypothesized to affecting the size of the relative risk for T2D per 100g glycemic load.

FIGURE S4. Sensitivity of covariates to study deletions.

FIGURE S5. Cumulative meta-regression analysis.

Protocol for study quality assessment

SUPPLEMENTAL MATERIALS REFERENCES

End of contents table

SUPPLEMENTAL MATERIALS

Search strategies

Strategies shown are updates on prior searches conducted in 1999, and March 2012.

*MEDLINE/PROQUEST**/ Royal Society of Medicine, UK; *1997 to week 3 of August 2012* Yield, 907 records. Hits, 18 records. Identified omissions, 1 record Patel et al (15).

- 1. MESH.EXPLODE("Cohort Studies")
- 2. MESH.EXPLODE("Prospective Studies")
- 3. MESH.EXPLODE("risk")
- 4. s1 or s2 or s3
- 5. MESH.EXPLODE("Diabetes Mellitus")
- 6. ti,ab(Glycemic index)
- 7. ti,ab(Glycemic load)
- 8. s6 or s7
- 9. yr(1997-2012)
- 10. s4 and s5 and s8 and s9
- 11. ab(glycemic index OR glycemic load) AND ab(Diabetes)
- 12. ab(risk OR association)
- 13. yr(>2010)
- 14. s11 and s12 and s13
- 15. s10 or s14

where 's' followed by a number abbreviates for search at line number; MESH, medical subject heading, ti, title; ab, abstract; and yr, year.

EMBASE/ PROQUEST (DialogTM, National Health Service, UK) / Royal Society of Medicine,

UK; 1997 to week 3 of August 2012

Yield, 1474 records. Hits, 18 records. Identified omissions, 1 record van Woudenberg et al (30).

- 1. EMB.EXACT("cohort analysis")
- 2. EMB.EXACT("prospective study")
- 3. EMB.EXACT("risk assessment")
- 4. EMB.EXACT("risk reduction")
- 5. EMB.EXACT("risk factor")

- 6. EMB.EXACT("follow up")
- 7. EMB.EXACT("hazard ratio")
- 8. EMB.EXACT("incidence")
- 9. s1 or s2 or s3 or s4 or s5 or s6 or s7 or s8
- 10. EMB.EXPLODE("Diabetes Mellitus")
- 11. EMB("glucose blood")
- 12. s10 or s11
- 13. TI,AB("carbohydrat*")
- 14. TI,AB("glycemic index")
- 15. TI, AB("glycemic load")
- 16. s13 or s14 or s15
- 17. s9 and s12 and s16
- 18. s17 and yr(1997-2012)
- 19. TI, AB("glycemic load")
- 20. TI,AB("glycemic index")
- 21. s19 or s20
- 22. TI, AB (diabetes)
- 23. TI,AB(risk)
- 24. TI, AB (association)
- 25. TI,AB(incidence)
- 26. TI,AB(cohort)
- 27. TI, AB(prospective)
- 28. s23 or s24 or s25 or s26 or 27
- 29. s21 and s22 and s28 and YR(>2010)
- 30. s18 or s29

where 's' followed by a number abbreviates for search at line number; EMB, the EMBASE

equivalent of MESH (both are medical subject headings) in Medline; TI, title, and AB, abstract.

'Glycemic' included the alternative spelling 'Glycaemic'.

WWW/CENTRAL (<u>http://www.thecochranelibrary.com/view/0/index.html)</u> / INLogic Ltd, UK. All dates to week 3 August 2012.

Yield, 66 records. Hits, 0 records. Identified omissions, 19 records, which is consistent with CENTRAL being focused on interventions.

All text (Glyc*emic load and Diabetes)

WWW/INLOGIC, UK 1997 to week 3 of August 2012.

Yield, zero additional records identified via the following online sources:

CDC, UK Centre for Reviews and Dissemination (<u>www.york.ac.uk/inst/crd</u>).

PROSPERO Register of Systematic Reviews

(www.york.ac.uk/inst/crd/projects/register.htm).

The Centers for Disease Control and Prevention (<u>www.cdc.gov</u>).

The National Institute of Health (<u>www.nih.gov</u>).

Google Scholar (<u>http://scholar.google.co.uk</u>).

Excluded studies

Articles identified by title and abstract but on examination of the full article did not meet the inclusion/exclusion criteria for the following reasons:

1) Was not an original study (2 reports):

- Pereira 2008 (1) was a commentary on the original study of Sahyoun et al (2) already included.
- Hu et al 2001(3) reviews data from the Nurses Health Study of Salmeron 1997 (4) amongst other lifestyle data.

2) Was not analyzed as a prospective cohort design (1 reports):

• Mohan et al 2009 (5) was a cross-sectional study.

3) Used an ineligible population (2 reports):

- Schulz et al 2006 (6) used a population that did not exclude T2D patients at baseline.
- Mayer-Davies et al 2006 (7) used a population that did not exclude T2D patients at baseline and summarizes their report noted immediately above (6).

4) Did not address the questions asked (2 reports):

- Fung et al 2002 (8) focused on whole grain and T2D in men of the Health Professionals Follow-up Study. Information on GL was not independent of that report earlier by Salmerón et al 1997 (9).
- Barclay et al 2007 (10) provided data on glycemic index only- no data on glycemic load was presented.

5) Dietary or other details were insufficient (1 report):

• Yu et al 2011 (11) provide limited information on glycemic load and T2D among 690 Honk Kong adults in a prospective cohort study with follow up of 9 to 14y, and report for their most adjusted model a non-significant effect of OR of 1.03 (CI 0.78-1.34) per 1 SD intake of GL unadjusted by the residual method for energy (equivalent to an OR of approx. 1.12 for the range of intakes of about 4SD, with potentially for higher value for energy adjusted GL intake. For this small study, a prior publication reported on validity of the FFQ used (12) but neither glycemic load nor any aspect of carbohydrate intake was addressed (CORR was unknown). 6) Reports of misidentified studies:

• Two publications (13, 14) incorrectly cited information about a mixed-sex population study (15). Information was available for the mixed-sex population only and did not report results for men or women separately. Correspondence with the first author of the original study (15), and with the first author in one citing the original study (13), indicates mistaken data extraction and/or misreporting. These errors are not perpetuated further in the present work.

Included studies

These are listed in Tables S1-4 below and in FIGURE 2 of the main article.

| Quantile | | RR | | Glycemic _ load (g/d | Reference food | Study e inta | energy ke | Cases | Non-case |
|----------------------|---------------|--------------|-------------------|-------------------------------------|--------------------------------|--------------------|--------------|-------|--------------------|
| | Median | L95CI | U95CI | reported, adjusted to energy) | (White bread or glucose) | Median or mean | units | n | n |
| Salmerón et al 1997 | (16) in wom | ien, RR bas | ed on rate | ratios. | | | | | |
| 1 | 1 | 3 | 3 | 111 | | | | 156 | ~12879 4 |
| 2 | 1.24 | 0.99 | 1.55 | 131 | | | | 189 | ~12846 4 |
| 3 | 1.22 | 0.97 | 1.54 | 144 | WB | 7424 ⁵ | kJ/d | 185 | ~12850 4 |
| 4 | 1.25 | 0.99 | 1.59 | 157 | | | | 179 | ~12856 4 |
| 5 | 1.47 | 1.16 | 1.86 | 178 | | | | 206 | ~12829 4 |
| Salmerón et al 1997 | (9) in men, 1 | RR based o | n odds rati | os | | | | | |
| 1 | 1 | | | 119 | | | | 120 | ~8432 6 |
| 2 | 1.07 | 0.82 | 1.41 | 144 | | | | 120 | ~8432 ⁶ |
| 3 | 1.04 | 0.78 | 1.39 | 160 | WB | 1995 ⁷ | kcal/d | 103 | ~8449 ⁶ |
| 4 | 1.13 | 0.83 | 1.54 | 177 | | | | 93 | ~8459 ⁶ |
| 5 | 1.25 | 0.90 | 1.73 | 203 | | | | 87 | ~8465 ⁶ |
| Meyer et al 2000 (17 | 7), RR based | on rate rati | ios. ⁸ | | | | | | |
| 1 | 1 | | | 94 | | | | 247 | ~6951 ⁹ |
| 2 | 0.96 | 0.79 | 1.15 | 110 | | | | 236 | ~6962 ⁹ |
| 3 | 0.86 | 0.71 | 1.05 | 120 | WB | 7531 ¹⁰ | kJ/d | 220 | ~6978 ⁹ |
| 4 | 0.92 | 0.75 | 1.12 | 129 | | | | 214 | ~6984 ⁹ |
| 5 | 0.95 | 0.78 | 1.16 | 145 | | | | 224 | ~6974 ⁹ |
| | | | | | | | | | continued. |

 Table S1. Extracted and calculated data for the included studies ^{1,2}

... continued

Stevens et al 2002 (18), white participants, RR is based on a rate ratio. Other data used are in footnotes¹¹

| | (10), white pu | ano punto, i | | on a rate ratio | o and autu | used are mir | oomotes | | |
|---------------------|----------------|--------------|-------------|--------------------------|------------|--------------------|---------|------------------|------------------|
| 1 | 1 | | | | | | | nr ¹¹ | nr 11 |
| 2 | | | | | | | | nr | nr |
| 3 | | | | 146 ¹² | WB | 1625 ¹³ | kcal/d | nr | nr |
| 4 | | | | | | | | nr | nr |
| 5 | 1.10 | 0.90 | 1.39 | | | | | nr | nr |
| Stevens et al 2002 | (18), African | Americans, | RR is a rat | e ratios . ¹⁴ | | | | | |
| 1 | 1 | | | | | | | nr ¹⁴ | nr ¹⁴ |
| 2 | | | | | | | | nr | nr |
| 3 | | | | 154 ¹⁵ | WB | 1602 ¹⁶ | kcal/d | nr | nr |
| 4 | | | | | | | | nr | nr |
| 5 | 0.97 | 0.73 | 1.35 | | | | | nr | nr |
| Schulze et al 2004 | (19), RR is ba | used on rate | ratios. | | | | | | |
| 1 | 1 | | | 139 | | | | 184 | ~18066 17 |
| 2 | 1.31 | 1.05 | 1.64 | 159 | | | | 192 | ~18058 17 |
| 3 | 1.20 | 0.92 | 1.56 | 172 | WB | 1811 ¹⁸ | kcal/d | 141 | ~18109 17 |
| 4 | 1.14 | 0.84 | 1.55 | 187 | | | | 115 | ~18135 17 |
| 5 | 1.33 | 0.92 | 1.91 | 211 | | | | 109 | ~18141 17 |
| Hodge et al 2004 (2 | 20), RR is bas | ed on odds | ratios. 19 | | | | | | |
| 1 | 1 | | | 91.8 | | | | 82 | 7828 |
| 2 | 0.86 | 0.61 | 1.20 | 101.2 | C | 002020 | 1 7 / 1 | 70 | 7840 |
| 3 | 1.17 | 0.86 | 1.60 | 118.9 | G | 8830 | KJ/d | 111 | 7799 |
| 4 | 0.92 | 0.65 | 1.30 | 155.7 | | | | 102 | 7809 |
| | | | | | | | | | |

....continued

Zhang et al 2006 (21), RR is based on rate ratios.

| 1 | 1 | | _ | 137 | | | | 210 | 2045 ²¹ |
|-------------------|------------------|--------------|--------------------------|-------------------|-----------------|--------------------|--------|-----|---------------------|
| 2 | 0.99 | 0.76 | 1.28 | 157 | | | | 174 | ~2404 21 |
| 3 | 0.89 | 0.65 | 1.22 | 171 | WB | 1813 ²² | kcal/d | 128 | 2621 ²¹ |
| 4 | 1.21 | 0.84 | 1.74 | 186 | | | | 145 | ~2660 21 |
| 5 | 1.61 | 1.02 | 2.53 | 212 | | | | 139 | 2576 ²¹ |
| Villegas et al 20 | 007 (22), RR is | based on ra | te ratios. ²³ | | | | | | |
| 1 | 1 | | _ | 164 | | | | 221 | ~12624 24 |
| 2 | 1.06 | 0.88 | 1.27 | 181 | | | | 256 | ~12589 24 |
| 3 | 0.97 | 0.81 | 1.17 | 190 | G | 1683 ²⁵ | kcal/d | 253 | ~12592 24 |
| 4 | 1.23 | 1.03 | 1.46 | 200 | | | | 349 | ~12496 24 |
| 5 | 1.34 | 1.13 | 1.58 | 235 | | | | 526 | ~12319 24 |
| Krishnan et al 2 | 007 (23), RR is | s based on r | ate ratios. | | | | | | |
| 1 | 1 | | | 82 | | | | 463 | ~7553 ²⁶ |
| 2 | 1.00 | 0.85 | 1.17 | 99 | | | | 368 | ~7648 ²⁶ |
| 3 | 1.09 | 0.92 | 1.31 | 109 | G | 1715 ²⁷ | kcal/d | 369 | ~7647 ²⁶ |
| 4 | 1.10 | 0.91 | 1.33 | 120 | | | | 362 | ~7654 ²⁶ |
| 5 | 1.22 | 0.98 | 1.51 | 142 | | | | 376 | ~7640 ²⁶ |
| Mosdol et al 200 | 07 (24), RR is l | based on rat | te ratios. | | | | | | |
| 1 | 1 | | _ | 121 ²⁸ | | | | 119 | 1721 ²⁹ |
| 2 | 1.05 | 0.76 | 1.44 | 145 | G ³⁰ | 2095 ³¹ | kcal/d | 117 | 1755 ²⁹ |
| 3 | 0.8 | 0.51 | 1.26 | 169 | | | | 93 | 1793 ²⁹ |
| | | | | | | | | | |

continued...

Patel et al 2007 (15), data is available for a mixed sex population only, RR is based on rate ratios.

| I dioi o | ui 2007 (13) |), aata 15 a i | | | · population o | my, na is ou | | u (105. | 22 | 22 |
|-----------|----------------|----------------|-------------|-------------|-------------------|---------------|-------------------------|----------------|--------------------|---------------------|
| | 1 | 1 | _ | — | 93 ³² | | | | nr ³³ | nr ³³ |
| | 2 | | _ | _ | | | | | nr | nr |
| | 3 | | | | 129 ³² | WB | 1494 ³⁴ | kcal/d | nr | nr |
| | 4 | | | | | | | | nr | nr |
| | 5 | 1.15 | 1.06 | 1.25 | 163^{32} | | | | nr | nr |
| Sahyoun e | et al 2008 (2) | , RR is bas | ed on odds | ratios. 35 | | | | | | |
| | 1 | 1 | _ | _ | 95 | | | | 17 | 362 ³⁶ |
| | 2 | 1.50 | 0.70 | 3.00 | 117 | | | | 22 | 359 ³⁶ |
| | 3 | 1.00 | 0.50 | 2.20 | 127 | G | 1835 ³⁵ | kcal/d | 18 | 360 ³⁶ |
| | 4 | 1.50 | 0.70 | 3.20 | 138 | | | | 20 | 361 ³⁶ |
| | 5 | 1.30 | 0.60 | 2.70 | 162 | | | | 22 | 357 ³⁶ |
| Halton et | al 2008 (25) | , RR is bas | sed on rate | ratios. | | | | | | |
| | 1 | 1 | | | 62 ³⁷ | | | | ~279 ³⁸ | ~8227 39 |
| | 3 | 1.23 | 1.00 | 1.49 | 79 ³⁷ | | | | ~348 | ~8158 ³⁹ |
| | 5 | 1.56 | 1.24 | 1.97 | 89 ³⁷ | G | 1560 ⁴⁰ | kcal/d | ~436 | ~8070 ³⁹ |
| | 7 | 1.88 | 1.45 | 2.45 | 99 ³⁷ | | | | ~525 | ~7981 ³⁹ |
| | 10 | 2.47 | 1.75 | 3.47 | 122 ³⁷ | | | | ~690 | ~7816 ³⁹ |
| Hopping e | et al 2010 (26 | 5), Europea | n Americai | n (Caucasia | n) men, RR is | based on rate | e ratios. ⁴¹ | | | |
| 11 0 | 1 | 1 | | ` <u> </u> | 84 ⁴¹ | | | | 257 | 2766 ⁴² |
| | 2 | 1.08 | 0.89 | 1.31 | 120 | | | | 236 | 2788 ⁴² |
| | 3 | 1.09 | 0.87 | 1.36 | 150 | G | 9045 | kJ/d | 202 | 2821 ⁴² |
| | 4 | 1.31 | 1.01 | 1.68 | 186 | - | | | 207 | 2816 ⁴² |
| | 5 | 1.54 | 1.12 | 2.10 | 256 | | | | 178 | 2845 ⁴² |
| | - | 1.6 . | | | | | | | 1.0 | |

... continued

Hopping et al 2010 (26), European American (Caucasian) women, RR is based on rate ratios.

| 1 | 1 | | | 71 | | | | 141 | 2787 ⁴² |
|--------------------|------------------|------------|--------------|----------------|--------------|-------|------|-----|--------------------|
| 2 | 1.34 | 1.04 | 1.73 | 100 | | | | 158 | 2771 ⁴² |
| 3 | 1.48 | 1.10 | 1.99 | 125 | G | 7144 | kJ/d | 152 | 2777 ⁴² |
| 4 | 1.47 | 1.03 | 2.08 | 155 | | | | 131 | 2798 ⁴² |
| 5 | 2.13 | 1.37 | 3.31 | 211 | | | | 133 | 2795 ⁴² |
| Hopping et al 2010 |) (26), Japanes | e America | n men, RR is | s based on rat | e ratios. | | | | |
| 1 | 1 | | — | 103 | | | | 369 | 2945 ⁴² |
| 2 | 1.06 | 0.92 | 1.23 | 141 | | | | 527 | 2788 ⁴² |
| 3 | 1.08 | 0.92 | 1.26 | 173 | G | 9052 | kJ/d | 574 | 2740 ⁴² |
| 4 | 1.09 | 0.91 | 1.29 | 213 | | | | 647 | 2668 ⁴² |
| 5 | 1.05 | 0.85 | 1.31 | 281 | | | | 560 | 2754 ⁴² |
| Hopping et al 2010 |) (26), Japanes | e America | n women, R | R is based on | rate ratios. | | | | |
| 1 | 1 | | — | 86 | | | | 284 | 3450 ⁴² |
| 2 | 1.17 | 0.99 | 1.38 | 117 | | | | 475 | 3260 ⁴² |
| 3 | 1.24 | 1.02 | 1.50 | 144 | G | 7150 | kJ/d | 542 | 3192 ⁴² |
| 4 | 1.23 | 0.98 | 1.54 | 175 | | | | 569 | 3166 ⁴² |
| 5 | 1.18 | 0.88 | 1.58 | 235 | | | | 504 | 3230 ⁴² |
| Hopping et al 2010 |) (26), Native] | Hawaiian n | nen, RR is b | ased on rate r | atios. | | | | |
| 1 | 1 | | — | 101 | | | | 119 | 795 ⁴² |
| 2 | 0.89 | 0.67 | 1.17 | 147 | | | | 110 | 804 ⁴² |
| 3 | 0.98 | 0.73 | 1.32 | 193 | G | 10628 | kJ/d | 122 | 792 ⁴² |
| 4 | 0.93 | 0.68 | 1.27 | 247 | | | | 154 | 760 ⁴² |
| 5 | 1.10 | 0.76 | 1.61 | 335 | | | | 293 | 620 ⁴² |
| | | | | | | | | | |

continued...

Hopping et al 2010 (26), Native Hawaiian women, RR is based on rate ratios.

| | 1 | 1 | — | — | 84 | | | | 110 | 1078 ⁴² |
|-----------|---------------|---------------------|---------------------|---------------------|--------------------|---------------|-----------------------|--------|------------------|----------------------|
| | 2 | 0.97 | 0.73 | 1.28 | 126 | | | | 111 | 1077 ⁴² |
| | 3 | 1.13 | 0.84 | 1.51 | 163 | G | 8625 | kJ/d | 145 | 1044 ⁴² |
| | 4 | 1.32 | 0.97 | 1.81 | 212 | | | | 204 | 984 ⁴² |
| | 5 | 1.44 | 0.98 | 2.12 | 329 | | | | 373 | 815 ⁴² |
| Sluijs et | al 2010 (27 |), RR is a ra | te ratio, ba | sed on othe | r data in footn | otes 43-45 | | | | |
| | 1 | 1 | — | | | | | | nr ⁴³ | nr ⁴³ |
| | 2 | — | — | | | | | | nr | nr |
| | 3 | — | — | | 118 | G | 2053 | kcal/d | nr | nr |
| | 4 | — | | | | | | | nr | nr |
| | 5 | ~1.83 ⁴⁴ | ~1.30 ⁴⁴ | ~2.53 ⁴⁴ | ~141 ⁴⁵ | | | | nr | nr |
| Simila e | t al 2011 (2 | 8), RR is bas | sed on rate | ratios | | | | | | |
| | 1 | 1 | | | 144 | | | | 280 | ~ 4909 ⁴⁶ |
| | 2 | 0.95 | 0.79 | 1.14 | 162 | | | | 241 | ~4948 ⁴⁶ |
| | 3 | 0.88 | 0.71 | 1.09 | 175 | G | 10800 47 | kJ/d | 203 | ~4986 ⁴⁶ |
| | 4 | 0.88 | 0.69 | 1.11 | 188 | | | | 195 | ~4994 ⁴⁶ |
| | 5 | 0.88 | 0.65 | 1.17 | 208 | | | | 179 | ~5010 46 |
| Sakurai e | et al 2012 (2 | 29), RR is ba | sed on rate | ratios. Pub | lished GL has | units of g/10 | 000kcal ⁴⁸ | | | |
| | 1 | 1 | | | 62.7 | | | | 23 | 377 |
| | 2 | 1.16 | 0.66 | 2.06 | 78.0 | | | | 26 | 375 |
| | 3 | 1.56 | 0.89 | 2.71 | 87.2 | G | 2198 ⁴⁹ | kcal/d | 34 | 364 |
| | 4 | 1.07 | 0.57 | 1.99 | 97.1 | | | | 23 | 377 |
| | 5 | 1.24 | 0.65 | 2.24 | 114.4 | | | | 27 | 369 |
| | | | | | | | | | | |

... continued

| Van Woudenbergh | et al 2011 (30 | U), RR 18 b | based on rate | e ratio | | | | | |
|-------------------|----------------|---------------|---------------|-------------------|-----------------|--------------------|--------|------|---------------------|
| 1 | 1 | | _ | 107 | | | | 173 | ~1282 50 |
| 2 | 0.91 | 0.71 | 1.16 | 126 | G^{51} | 1981 ⁵² | kcal/d | 149 | ~1306 50 |
| 3 | 1 | 0.74 | 1.36 | 146 | | | | 134 | ~1321 50 |
| Mekary et al 2011 | (31) (32), RR | R is based or | n rate ratio | | | | | | |
| 1 | 1 | | _ | 58 | | | | 1239 | 14173 ⁵³ |
| 2 | 1.02 | 0.94 | 1.11 | 80 ⁵⁴ | | | | 1283 | ~12820 53 |
| 3 | 1.13 | 1.03 | 1.23 | 99 | G ⁵⁵ | 1743 ⁵⁶ | kcal/d | 1390 | 14450 ⁵³ |
| 4 | 1.22 | 1.10 | 1.35 | 118 ⁵⁴ | | | | 1466 | ~12637 53 |
| 5 | 1.32 | 1.16 | 1.51 | 153 | | | | 1572 | 14491 ⁵³ |
| | | | | | | | | | |

Footnotes:

¹ Values in normal font without superscripts are data published the citation tabulated.

Values in italics were supplied on correspondence with authors of the citation-see corresponding footnotes.

Values in normal font with superscripts are calculated and regard as exact as a published value unless preceded by a tilda (~) when the values are approximate. The approximations were made to enable the meta-analytical procedures where small errors are of little consequence to the assessment of dose response—see corresponding footnotes.

 2 Other extracted data and author supplied information are given in subsequent footnotes.

³ All such in this column in rows for Q1, authors of the original reports provide 95CI values for relative risks from Q_1

to Q_n defining the relative risk at Q1 as one with zero degrees of freedom, hence no 95CI values are given for Q_1 .

⁴ Calculated: Number of participants (65173) divided by the number of quantiles (5), less the number of cases tabulated (16).

⁵ Calculated: Mean of quintile values (7253+7636+7594+7531+7106)÷5 (16).

⁶ Calculated: Number of participants (42759) divided by the number of quantiles (5), then less the number of cases tabulated (9)

⁷ Calculated: Mean of quintile values (1960+2010+2016+2016+1971)÷5 from reference (9).

⁸ Author response confirmed further information was not available or not readily accessible (17).

⁹ Calculated. Number of participants (35988) divided by the number of quantiles (5), then less the number of cases tabulated (17).

¹⁰ Calculated: Mean of ten energy intake values (6879+6879+7297+7945+8577+8368+7075+7046+7226+8021)÷10 (kJ/d) (17).

- ¹¹ Other extracted data for European Americans: incremental RR per 1sd of energy adjusted GL (mean and 95%CI) 1.13 (1.0 to 1.276) meant that case and control data were not needed to obtain rates of change in RR with GL in the first step of two-step analysis. 1SD of energy adjusted GL was calculated at 62g for the mean energy intake shown and is the combined SD values obtained on pooling means and SDs for quantiles of energy adjusted GL in Tables 1 and 2 of the original publication (18).
- ¹² Calculated: The range of GL from quantile 1 to quantile 5 was obtained assuming a normal distribution calculated from study mean and SD for energy adjusted GL intakes in Tables 1 and 2 of the original publication. The study average of glycemic load was derived from the mean of two sets of ten quintiles values (18), thus (144+130+136+148+172+122+141+150+159+160) ÷10. A value for 1SD of energy adjusted GL was calculated at 62g by combining the SD values for each quantile, and accounting for the SD between quantiles. This complex arrangement was used because information on GL intakes by quantile was available not for GL quantiles directly but was available for fiber and glycemic index quantiles, while correspondence with authors was not able to provide answers.

¹³ Calculated: Mean of ten energy intake values (1796+1531+1528+1562+1708+1566+1647+1658+1673+1581)÷10 (18).

- ¹⁴ Hazard ratio for slope (mean and 95%CI) 0.999 (0.966-1.002) for African-Americans (18) was extracted, which meant that case and control data were not needed to obtain rates of change in RR with GL in the first step of two-step analysis.
- ¹⁵ Calculated: Study average of glycemic load was derived from the mean of two sets of 5 quintiles values $(165+135+141+151+177+136+156+164+161+151) \div 10$ (18).
- ¹⁶ Calculated: Mean of ten energy intake values (1606+1654+1674+1587+1483+1780+1456+1485+1551+1740)÷10 (18).
- ¹⁷ Calculated: Total number of participants (91249) divided by the number of quantiles (5), then less the number of cases tabulated.
- ¹⁸ Calculated: Using glycemic load (g/d) and glycemic index to calculate carbohydrate intake (g/d), followed by use of carbohydrate intake per unit energy intake (kcal/100kcal energy) to calculate energy intake (19).
- ¹⁹ Data provided by correspondence with the first author of the original report (20), who kindly re-analyzed their data with GL adjusted for energy intake by the residual method.
- ²⁰ Calculated: Mean of four energy intake values (8803+8038+8559+9919); 4.
- ²¹ Calculated: Total number of participants in the quantile less the number of cases shown. Participant numbers were 2255, 2749, and 2718 in the 1st, 3rd, and 5th quantiles and interpolated for the 2nd and 4th quantiles with adjustments to ensure the correct total number of participants (21).
- ²² Calculated: Mean of six energy intake values reported (1822+1833+1792+1790+1856+1783)÷6 (21).
- ²³ Values for GL were obtained by correspondence with the first author of the original report (22) and were: $Q_1 = 164.4$, $Q_2 = 180.5$, $Q_3 = 190.0$, $Q_4 = 200.2$ and $Q_5 = 234.7$ g GL/d.
- ²⁴ Calculated: Total number of participants (64227) divided by the number of quantiles (5), then less the number of cases

tabulated (22).

- ²⁵ Calculated: Mean of energy intakes by quintile $(1773.2 + 1643.9 + 1609.5 + 1602.6 + 1784.1) \div 5$ (22).
- ²⁶ Calculated: Total number of participants (40078) divided by the number of quantiles (5), then less the number of cases tabulated (23).
- ²⁷ Calculated: Mean for study energy intakes reported for quantiles (1966+1429+1882+1582+1697+1638+1946+1516+1779)÷9 (23).
- ²⁸ Calculated: GL for the mixed population is calculated from the reported GL values for men (127, 152 & 176 g/d for Q₁ to Q₃) and women (108, 129 & 152 g/d for Q₁ to Q₃) and the fraction of the population that were men (0.71) (24).
- ²⁹ Calculated: Number of persons per quantile reported in the original report (24) less the number of cases tabulated.
- ³⁰ Based on very low reported central-quantile GI values of 56 and 54.5 for men and women (24), a glucose reference standard was assumed. This appears corroborated by a value of 86 for the same community at a time when white bread was usually a standard (33). Two corresponding authors were not available to report differently.
- ³¹ Calculated: Based on the reported fat and carbohydrate intakes (24), calorie conversion factors of 9 and 3.75 kcal/g for fat and carbohydrate as monosaccharide respectively, and 14.8% energy as protein average across sexes and tertiles for this population (34).
- ³² Calculated: Based on reported values of GL (g/d) (15) of 145 sd 32 for men, and 114 sd 23 in women, a normal distribution and the fraction of men in the population of 0.46 being applied to all quantiles.
- ³³ Case and control data were not needed when obtaining the rate of change in RR with GL in the first step of two-step analysis because the rate estimate is based on only one quantile versus referent. Case and control data were only needed when there was multiple data within the study when the case and control data help account for non-independence of observations from the same study (27).
- ³⁴ Calculated from values for each quantile in men and women separately and the fraction of the population that were men, $(0.46x(1723+1732+1726+1727+1690)\div 5) + (1-0.46)x(1288+1336+1326+1291+1268)\div 5$
- ³⁵ By correspondence, the first author of the original report (2) indicates that GL was adjusted for energy intake in men and women separately, with means of 2016.7 kcal/d in men and 1608.4 kcal/d in women, with a combined sex mean of 1835 kcal/d. Correspondence confirms GL values were based on the glucose standard, and that all non-European American participants were African-American.
- ³⁶ Calculated: Number of persons per quantile (379, 381, 378, 381, 379) less the number of cases per quantile tabulated (2).
- ³⁷ By analysis, assuming a normal distribution, a mean GL from the original report (25) and a range of 60 given between lowest and highest deciles by Lui & Chou (13).
- ³⁸ Calculated from the total number of cases distributed according to the relative risks in each quantile.
- ³⁹Calculated: Total number of participants (85059) divided by the number of quantiles (10), then less the number of cases

tabulated.

- ⁴⁰ Calculated: Mean of nine reported energy values (1553+1559+1559+1550+1555+1551+1565+1552+1591)÷9 (25).
- ⁴¹ Authors explained by correspondence that the published and author provided values of GL for this study (shown above) had not been energy adjusted. Prior to meta-analysis, a factor of 1.62 was applied to approximate this adjustment, which is the ratio of energy adjusted variance in GL for the similar whole multiethnic cohort in Howarth et al (35) to the variance in GL in the multiethnic cohort in the Hopping et al study (26) after adjustments for differences in energy intakes reported.
- ⁴² Calculated: Number participants less the number of cases, by quantile, data supplied by authors. Values agrees to 1 in 3000 with values calculated as the total number of participants divided by the number of quantiles, then less the number of cases by quantile for the published data (26).
- ⁴³ Case and non-case data was not used because the authors supplied rate information: RR was reported to increase by 1.27 (95%CI: 1.11,1.44) per 1SD rise in reported GL (g/2053kcal) of 21.2 g (27). This information was re-expressed per 100g GL in 2000kcal. Operationally this was via lnRR per 1SD rise in glycemic load
- ⁴⁴ Data not used in the two-step analysis, but approximated for the meta-analysis of rise in lnRR from the lowest to highest quantile (Fig S1 in the Supplemental Materials online). Data was calculated from information in footnotes 43 & 45.
- ⁴⁵The median glycemic load for quantile 5 was approximated using the reported glycemic load of 117.9g and its SD 21.2 g (27). Using these values a normal distribution was simulated for 100000 observations, divided into quintiles, and the median for the fifth quintile obtained. A normal distribution was indicated by the authors reporting an SD value for glycemic load among other data showing interquartile ranges when the normality assumption was not justified.
- ⁴⁶ Calculated: Total number of participants (25943) divided by the number of quantiles (5), then less the number of cases tabulated.
- ⁴⁷ Calculated as the mean of six values expressed in MJ (10.8 + 11 + 10.7 + 10.8 + 11 + 10.5)÷6
- ⁴⁸ Values for GL were reported in g per 1000 kcal (29).
- ⁴⁹ Calculated as the mean of five values $(2394 + 2299 + 2183 + 2104 + 2011) \div 5$
- ⁵⁰ Calculated: Total number of participants (4366) divided by the number of quantiles (3), then less the number of cases tabulated.
- ⁵¹ Correspondence with the first author of the original study confirms.
- ⁵² Calculated as the mean of three quntile values (1967 + 2005 + 1971)÷3
- ⁵³ Calculated approximately: Total number of participants less the number of participants in Q1, Q3 and Q4, this remainder divided between Q2 and Q4, each less the published number of cases in Q2 and Q4 respectively.
- ⁵⁴ Values at Q2 and Q4 were not published. We used mid-range values for these quantiles.
- ⁵⁵ Based on very low reported GI values and published correspondence comparing values in this and the prior study of Halton et al (25), a glucose reference standard was evident, as in the prior study from this group at 20y follow-up.

⁵⁶ Reported in published correspondence (32).

| | First authorian and (citat | or, date tion) | Region | Ethnicity | Ascertainment ² of outcome | Number of quantiles | Years of follow-up | Population sample (n) | No. Cases (n) |
|----|----------------------------|-------------------|---------------|-----------|---------------------------------------|---------------------------|--------------------|-----------------------------|---------------------|
| 1 | Salmerón | 1997 (f) (16) | USA | EA | Clinical report | 5 | 6 | 65173 | 915 |
| 2 | Salmerón | 1997 (m)(9) | USA | 95% EA | Clinical report | 5 | 6 | 42759 | 523 |
| 3 | Meyer | 2000 (17) | USA | EA | Self report | 5 | 6 | 35988 | 1141 |
| 4 | Stevens | 2002 (18) | USA | EA | Clinical report | 5 | 9 | 9529 | 971 |
| 5 | Stevens | 2002 (18) | USA | AA | Clinical report | 5 | 9 | 2722 | 478 |
| 6 | Schulze | 2004 (19) | USA | EA | Clinical report | 5 | 8 | 91249 | 741 |
| 7 | Hodge | 2004 (20) | Australia | EAu | Self report | 4 | 4 | 31641 | 365 |
| 8 | Zhang | 2006 (21) | USA | EA | Self report GDM | 5 | 8 | 13110 | 796 |
| 9 | Villegas | 2007 (22) | China | CH | Mixed reports ³ | 5 | 4.6 | 64227 | 1605 |
| 10 | Krishnan | 2007 (23) | USA | AA | Self report | 5 | 8 | 40078 | 1938 |
| 11 | Patel | 2007 (15) | USA | mixed | Self report | 5 | 9 | 124907 | ~2700 |
| 12 | Mosdol | 2007 (24) | Europe | Eu | Clinical report | 3 | 13 | 5598 | 329 |
| 13 | Sahyoun | 2008 (2) | USA | 67% EA | Clinical report | 5 | 4 | 1898 | 99 |
| 14 | Halton | 2008 (25) | USA | EA | Clinical report | 10 | 20 | 85059 | 4670 |
| 15 | Hopping | 2010 (26) | Hawaii- men | EA | Clinical report | 5 | 14 | 15116 | 1080 |
| 16 | Hopping | 2010 (26) | Hawaii- women | EA | Clinical report | 5 | 14 | 14643 | 715 |
| 17 | Hopping | 2010 (26) | Hawaii- men | JA | Clinical report | 5 | 14 | 16572 | 2677 |
| 18 | Hopping | 2010 (26) | Hawaii- women | JA | Clinical report | 5 | 14 | 18672 | 2364 |
| 19 | Hopping | 2010 (26) | Hawaii- men | NH | Clinical report | 5 | 14 | 4568 | 798 |
| 20 | Hopping | 2010 (26) | Hawaii- women | NH | Clinical report | 5 | 14 | 5941 | 943 |
| 21 | Sluijs | 2010 (27) | Europe | Eu | Clinical report | 5 | 10.1 | 37846 | 915 |
| 22 | Simila | 2011 (28) | Europe | Eu | Clinical report | 5 | 12 | 25943 | 1098 |
| 23 | Sakurai | 2011 (29) | Japan | Jp | Clinical report | 5 | 6 | 1995 | 133 |
| 24 | Van Wou | denbergh | Europe | Eu | Clinical report | 3 | 12.4 | 4366 | 456 |

 Table S2 Study identities, region, ethnicities, outcome ascertainment, population sample size, and number of cases accumulated ¹

| | 2011 (30) | | | | | | | |
|----|------------------|-----|----|-----------------|---|----|-------|------|
| 25 | Mekary 2011 (31) | USA | EA | Clinical report | 5 | 26 | 81827 | 6950 |

¹ Abbreviation: USA, United States of America; EA, European-American; AA, African-American; EAu European-Australian; CH, Chinese; mix, mixed ethnicities; JA, Japanese-American; NH, Native Hawiian; Eu, European; Jp, Japanese; T2D, Type 2 diabetes; GDM, gestational diabetes.

² Medical reports include hospital or medical doctor's records or biochemical tests.

³ Of 1608 self-reported cases, 896 were confirmed by medical record.

| First auth and (citat | or, date ion) | Instrument used for dietary assessment | Number of food items in the instrument | Instrument correlation with food records ² | Whether correlation was deattenuated | Validity of instrument for cohorts analyzed | No. of assessments made with instrument(s) |
|--------------------------|------------------|---|---|--|---|--|---|
| 1 Salmerón | 1997 (f) (16) | FFQ | 134 | 0.64 | yes | yes | 1 |
| 2 Salmerón | 1997 (m) (9) | FFQ | 131 | 0.73 | yes | yes | 1 |
| 3 Meyer | 2000 (17) | FFQ | 127 | 0.45 | yes | yes | 1 |
| 4 Stevens | 2002 (18) | FFQ | 66 | 0.45 | yes | no | 1 |
| 5 Stevens | 2002 (18) | FFQ | 66 | 0.45 | yes | no | 1 |
| 6 Schulze | 2004 (19) | FFQ | 133 | 0.64 | yes | yes | 2 |
| 7 Hodge | 2004 (20) | FFQ | 121 | $0.41 (0.56)^{3}$ | no (~yes) ³ | no (~yes) ³ | 1 |
| 8 Zhang | 2006 (21) | FFQ | 133 | 0.64 | yes | yes | 2 |
| 9 Villegas | 2007 (22) | FFQ | 77 | 0.66 (0.71) ⁴ | no (yes) 4 | yes | 2 |
| 10 Krishnan | 2007 (23) | FFQ | 68 | 0.43 | yes | yes | 1 |
| 11 Patel | 2007 (15) | FFQ | 68 | 0.62^{5} | yes | yes | 1 |
| 12 Mosdol | 2007 (24) | FFQ | 127 | 0.50 | yes | yes | 1 |
| 13 Sahyoun | 2008 (2) | FFQ | 108 | 0.65 | yes | yes | 1 |
| 14 Halton | 2008 (25) | FFQ | 61,116,134 ⁶ | 0.45,0.61,0.64 7 | yes | yes | 6 |
| 15 Hopping | 2010 (26) mEA | FFQ | 125 | 0.68 | yes | yes | 1 |
| 16 Hopping | 2010 (26) fEA | FFQ | 125 | 0.80 | yes | yes | 1 |
| 17 Hopping | 2010 (26) mJA | FFQ | 125 | 0.56 | yes | yes | 1 |
| 18 Hopping | 2010 (26) fJA | FFQ | 125 | 0.54 | yes | yes | 1 |
| 19 Hopping | 2010 (26) mNH | FFQ | 125 | 0.62 8 | yes | no ⁸ | 1 |
| 20 Hopping | 2010 (26) fNH | FFQ | 125 | 0.67 ⁸ | yes | no ⁸ | 1 |

Table S3. Characteristics of the dietary instrument used 1

| 21 Sluijs | 2010 (27) | FFQ | 178 | 0.75 | yes | yes | 1 |
|--------------|-------------------|-----|--------------------------|--------------------------|-----------------------|-----|---|
| 22 Simila | 2011 (28) | DHQ | 276 | 0.55 (0.71) ⁹ | no (yes) ⁹ | yes | 1 |
| 23 Sakurai | 2011 (29) | DHQ | 147 | 0.62 | yes | yes | 1 |
| 24 van Woude | enbergh 2011 (30) | FFQ | 170 | 0.79 | yes | yes | 1 |
| 25 Mekary 20 | 11 (31) | FFQ | 61,116,134 ¹⁰ | 0.45,0.61,0.64 11 | yes | yes | 7 |

¹ Abbreviations: FFQ, food frequency questionnaire; DHQ, diet history questionnaire.

² Correlations were for carbohydrate intake, and are reproduced either from the citation or from its referenced validation study. Values are after adjustment for energy intake (unless specified differently) and de-attenuation (unless also accompanied by bracketed values, when values in brackets indicated approximate deattenduated values obtained as described in the main article. The correlation shown is for validation of one application of the instrument. To aid comparability between studies, correlations obtained by repeated measures were not used.

³ As discussed in the citation (20), a discrepancy appears between the published validation of the instrument, which was on a population external to the population sampled for the cohort study, and the reproducibility of the instrument in a sample of the cohort studied. Within the study the FFQ showed only "fair" to "moderate" agreement—interpretable from tables of kappa as 0.21-0.4 and 0.41 to 0.60 respectively, for which the mid-range of 0.41 was used as a crude estimate. Adjustments to approximate an energy-adjusted deattenuated value suggest a value of approx. 0.56 compared with the questionnaires validation, which gave 0.78 for in a different population.

⁴ Crude value as reported in the validation publication, in which the authors claim an energy adjustment did not change the result appreciably. Value in parenthesis is after approximate adjustment at present for de-attenuation.

⁵ A value for the mixed sex population was the average of values for men (0.73) and women (0.51).

⁶ Mean number of foods for the three FFQ used 116=(61x4/20 + 116x2/20 + 134x14/20) weighted by years of use (4, 2, 20) over the 20 year follow-up.

⁷ Mean correlation for the three FFQ used 0.60=(0.45x4/20 + 0.61x2/20 + 0.64x14/20) weighted by years of use (4, 2, 20) over the 20 year follow-up. Note, for comparison with other studies this corresponds to a single representative FFQ validation weighted by the years of use as opposed to a higher correlation obtainable by repeated measures.

⁸ An average was used for men and another average for women, obtained from among the population of non-native Hawaiians. (26)

⁹Energy adjusted deattenuated value (0.71) from validation paper.

¹⁰ Mean number of foods for the three FFQ used, 119 = (61x4/26 + 116x2/26 + 134x20/26) weighted by years of use (4, 2, 20) over the 26-year follow-up.

¹¹ Mean correlation for the three FFQ used, 0.61 = (0.45x4/26 + 0.61x2/26 + 0.64x20/26) weighted by years of use (4, 2, 20) over the 26-year follow-up. Note that, for comparison with other studies, this corresponds to a single representative FFQ validation weighted by the years of use as opposed to a higher correlation such as obtainable by repeated measures.

| | | | * * | • <i>í</i> | 14 | • / | | / U | | |
|----|------------|---------------|------------|------------|--------------------|--------|------------------------|--------------------------------|---------------|-------------|
| | | | Sample | Mean BMI | Mean age of sample | Mean | Range of GL intake | Reasons for | Newcastle | |
| | | | population | of sample | population | energy | O_1 to O_{max} | excluding | Ottawa | Conflict of |
| | First auth | or, date | as male | population | at baseline | intake | (g per | participants | quality scale | interest |
| | and (citat | ion) | (fraction) | (kg/m^2) | (y) | (kcal) | 2000kcal) ² | at baseline 1 | 3 | declared |
| 1 | Salmerón | 1997(f) (16) | 0 | 25 | 53 | 1774 | 88 - 140 | dm.ca.cvd.iei.mis | 8 | nr |
| 2 | Salmerón | 1997(m)(9) | 1 | 25 | 58 | 1995 | 83 - 142 | dm ca cyd iei mis | 8 | nr |
| 2 | Moyor | 2000 (17) | 0 | 25 | 50 62 | 1800 | 73 113 | dm jej mis | 6 | nr |
| 5 | Stavana | 2000 (17) | 0.46 | 27 | 54 | 1600 | 73 - 113 62 180 | dm iai mia ina ath | 0 | 111 |
| 4 | Stevens | 2002 EA (18) | 0.46 | 27 | 54 | 1625 | 62 - 189 | dm,iei,mis,ipc,eth | 8 | nr |
| 5 | Stevens | 2002 AA (18) | 0.37 | 29 | 53 | 1602 | 63 - 206 | dm,1e1, m1s,1pc,eth | 8 | nr |
| 6 | Schulze | 2004 (19) | 0 | 25 | 36 | 1811 | 107 - 163 | dm,ca,cvd,iei,mis | 8 | nr |
| 7 | Hodge | 2004 (20) | 0.5 | 26 | 55 | 2110 | 87 - 148 | dm,chd,preg,iei,mis | 6 | none |
| 8 | Zhang | 2006 (21) | 0 | 23 | 32 | 1813 | 106 - 164 | dm,mg,cvd,ca,iei,mis | 7 | nr |
| 9 | Villegas | 2007 (22) | 0 | <30 4 | 51 | 1683 | 195 - 279 | dm,cvd,cam | 8 | nr |
| 10 | Krishnan | 2007 (23) | 0 | <31 5 | 38 | 1715 | 96 - 166 | dm,ca,iei,igl,mis ⁶ | 6 | none |
| 11 | Patel | 2007 (15) | 0.46 | 26 | 63 | 1494 | 88 - 154 | dm,1yd,ca,iei,mis | 7 | none |
| 12 | Mosdol | 2007 (24) | 0.71 | 25 | 49 | 2095 | 116 - 161 | dm, em, mis, iei | 7 | none |
| 13 | Sahyoun | 2008 (2) | 0.45 | 27 | 75 | 1835 | 104 - 177 | dm,iei,mis | 8 | none |
| 14 | Halton | 2008 (25) | 0 | 24 | 46 | 1560 | 79 - 156 | dm,ca,cvd,iei,mis | 8 | none |
| 15 | Hopping | 2010 (26) mEA | 1 | 26 | 57 | 2162 | 101 - 199 | dm,oe,mis,sr | 8 | none |
| 16 | Hopping | 2010 (26) fEA | 0 | 26 | 58 | 1707 | 108 - 208 | dm,oe,mis,sr | 8 | none |
| 17 | Hopping | 2010 (26) mJA | 1 | 25 | 59 | 2163 | 120 - 222 | dm,oe,mis,sr | 8 | none |
| 18 | Hopping | 2010 (26) fJA | 0 | 24 | 59 | 1709 | 126 - 234 | dm,oe,mis,sr | 8 | none |
| 19 | Hopping | 2010 (26) mNH | 1 | 28 | 56 | 2540 | 107 - 221 | dm,oe,mis,sr | 8 | none |
| 20 | Hopping | 2010 (26) fNH | 0 | 27 | 56 | 2061 | 111 - 257 | dm,oe,mis,sr | 8 | none |

Table S4. Characteristics of the study participants, duration of study, number of quantiles, and study baseline exclusions.¹

| 21 Sluijs 2010 (27) | 0.26 | 26 | 51 | 2053 | 89 - 141 | dm,iei,mis | 8 | none |
|---------------------------------|------|----|----|------|-----------|--------------------|---|------|
| 22 Simila 2011 (28) | 1 | 26 | 57 | 2629 | 110 - 158 | dm,ns | 8 | none |
| 23 Sakurai 2011 (29) | 1 | 23 | 46 | 2000 | 125 - 229 | dm, mis,iei | 7 | none |
| 24 van Woudenbergh 2011 (30) | 0.4 | 26 | 67 | 1981 | 108 - 147 | dm,mis,hcrp,ini | 8 | nr |
| 25 Mekary 2011 (31) | 0 | 26 | 46 | 1743 | 66-176 | dm,cvd,ca,mis,iei, | 7 | none |

Abbreviations: BMI, body mass index (kg/m²); f, female; m, male; nr, not reported; dm, diabetes mellitus; ca, cancer; cvd, cardiovascular disease; chd, coronary heart disease; iei, implausible energy intakes; mis, missing or inadequately compete information; ipc, inadequate number of participants within a field centre; eth, ethnicity; preg, pregnancy; igl, implausible glycemic load; mg, multiple gestations; 1yd, one year deaths to minimize undiagnosed disease at baseline; oe, other ethnicities; ns non-smokers; hcrp, high C-reactive protein; ini, implausible nutrient intakes.

² Calculated values, energy adjusted for glycemic load.

³ The Newcastle-Ottawa observational study quality scale ranges from 0 to 9 representing a minimum to maximum quality (37).

⁴ An approximate estimate made using the percentage persons in categories of BMI was $\sim 26 \text{ kg/m}^2$.

⁵ An approximate estimate made using the percentage persons in categories values of BMI $\sim 26 \text{ kg/m}^2$.

⁶ Other exclusions: pregnancy, age less than 30y.

Table S5. Assessment of assumptions about accuracy of data used in the two-stage meta-analysis with covariates applied to all24 studies (model 5 as reported in Tables 2 & 3 of the main article). Observations and comments apply to the currentapplication and dataset only.

| | | | | Approach | Overall outcome for the fully adjusted RR for | | |
|---|--|---|--|---|---|--------------------|-------------------------|
| | Potential weakness | | A | or new | 12D (cf Table 2 | Comment 1 | C |
| | in data | Studies affected | Assumption made | assumption | main article) | Comment 1 | Comment 2 |
| 1 | Assumas all data | A 11 | Hypothetically | No now | 1 45 (1 21 1 61) | | |
| 1 | colloted in Toble S1 | | nopo | assumption | 1.45 (1.51, 1.01) | | |
| | are accurate | | none | assumption | | | |
| 2 | Case numbers in each quantile were not available, and so were approximated when needed. | Halton et al (25) | Approximation of these values based on both the total case numbers and value for RR in each quantile has negligible effect on the overall outcome | Case numbers approximated for all studies, not just the one study affected. | 1.45 (1.31, 1.61) | Agrees with line 1 | Assumption justified |
| 3 | Person-years in each quantile were | Salmeron et al f (16) Schulze et al (19) | Approximation of these values | Person-years by | 1.45 (1.31, 1.61) | Agrees with line 1 | Assumption |
| | not available for all | Stevens et al EA (18) | based on the | reported for the | | | Justinou |
| | studies, and so | Stevens et al AA (18) | number of | unaffected studies | | | |
| | were approximated | Sahyoun et al (2) | participants per | were used, only | | | |
| | consistently for all | Hodge et al (20) | quantile and the | approximating | | | |
| | studies based on the | Patel et al (15) | number of years | these values when | | | |
| | number of | Zhang et al (21) | has negligible | not available. | | | |
| | participants per | Hopping et al (both | effect on the | (May introduce | | | |

| | quantile and the | sexes and all | overall outcome | hias between | | | |
|---|---------------------|---|---------------------|--------------------|------------------|--------------|------------|
| | number of follow | othnicition) (26) | | those that remark | | | |
| | | $\frac{\text{cumicules}}{\text{Holton at al}} (20)$ | | | | | |
| | up years. In some | Halton et al (25) | | person-years and | | | |
| | studies also the no | Sluijs et al (27) | | those that don't) | | | |
| | of participants in | Simila et al (28) | | | | | |
| | each quantile was | | | | | | |
| | assumed to equal | | | | | | |
| | the total number of | | | | | | |
| | participants in the | | | | | | |
| | study divided by | | | | | | |
| | the number of | | | | | | |
| | quantiles. | | | | | | |
| 4 | Study average | Mosdol et al (24) | Errors in | Energy assumed | 1.45 (1.31,1.61) | Agrees with | Assumption |
| | energy intake was | Schulze et al (19) | calculations are | 10% too low for | | line 1 | justified |
| | calculated from | | $\leq 10\%$ so have | Mosdol et al (24) | | | |
| | reported diet | | negligible effects | Energy assumed | 1.45 (1.31,1.61) | Agrees with | |
| | compositions and | | the outcome | 10% too high for | | line 1 | |
| | food energy | | | Mosdol et al (24) | | | |
| | conversion factors | | | Energy assumed | 1.45 (1.31,1.61) | Agrees with | |
| | | | | 10% too low for | | line 1 | |
| | | | | Schultz et al (19) | | | |
| | | | | Energy assumed | 1.46 (1.31,1.61) | Assumption | |
| | | | | 10% too high for | | used is | |
| | | | | Schultz et al | | conservative | |
| 5 | Glycemic load | Stevens et al (18) EA | Errors in | GL assumed 10% | 1.45 (1.31.1.61) | Agrees with | Assumption |
| | values were not | Stevens et al (18) AA | calculations are | too high for | | line 1 | iustified |
| | reported directly | Patel et al (15) | <10% so have | Stevens et al (18) | | | J |
| | but were | Halton et al (25) | negligible effects | EA. | | | |
| | approximated | | the outcome | GL assumed 10% | 1.45 (1.31.1.61) | Agrees with | |
| | assuming a normal | | | too low for | | line 1 | |
| | distribution from | | | Stevens et al (18) | | | |
| | values for mean | | | EA | | | |
| | and SD of GL in | | | GL assumed 10% | 1 45 (1 31 1 61) | Agrees with | |
| | the nonulation | | | too high for | 1.15 (1.51,1.01) | line 1 | |
| | population | | | Stevens et al (18) | | | |
| | | | | | | | |
| | | | | nn. | | 1 | |

| | | | | GL assumed 10% too low for Stevens et al (18) AA. GL assumed 10% too high for Patel et al (15) GL assumed 10% too low for Patel et al (15) GL assumed 10% too high for Halton et al (26) GL 10% too low for Halton et al | 1.45 (1.31,1.61) 1.45 (1.31,1.60) 1.45 (1.31,1.62) 1.45 (1.31,1.61) 1.45 (1.31,1.61) | Agrees with line 1 Agrees with line 1 Agrees with line 1 Agrees with line 1 Agrees with line 1 | |
|---|---|--------------------------------------|---|--|--|--|---|
| 6 | Standard for GI as bread or as glucose was clearly rational assumption, but unconfirmed via correspondence. | Mosdol et al (24) | Standard used was glucose | Standard used was bread | 1.45 (1.31-1.61) | Agrees with line 1 | Assumption leads to no appreciable error overall |
| 7 | Analytical values for all glycemic load values are imprecise for foods and diets. | Potentially, all studies reviewed | Errors are random leading to random errors among studies, and if excessive would both underestimate the outcome for RR and elevate I ² | Effect of additional random error in GL causing deviation in RR per unit GL was examined (error of mean 0 and SD 10% of the study range of GL was applied) | 1.42 (1.28, 1.58) | RR was significantly lower (t= - 9.5 for 10 reps) indicating random error among GL intakes leads to a conservative estimate | Assumption reasonable and conservative |
| | | | | | | | |

| | | I ² was elevated | Confirms | |
|--|--|-----------------------------|-------------------|-----------|
| | | from 2% to | that this | |
| | | 13% | random error | |
| | | (average of 10 | elevates I^2 , | |
| | | reps) | so that the I^2 | |
| | | 1 / | of 2% | |
| | | | indicates | |
| | | | random error | |
| | | | in analytical | |
| | | | values for | |
| | | | GL appear of | |
| | | | limited | |
| | | | concern | continued |

| 8 | Glycemic load | Hopping et al for both | The adjustment is | Error from this | I^2 decreased to a | I^2 behaved as | Assumption |
|---|------------------|--------------------------|--------------------|----------------------------|---------------------------|---------------------------|-----------------|
| | values were | sexes and all three | imputable from | assumption would | minimum and | expected. | made (Column |
| | reported without | ethnicities (26). This | the ratio of | contribute towards | increased again | | 4) was |
| | adjustment for | although the | variance for the | elevation of I^2 . | as the | With the | reasonable, and |
| | energy intake | corresponding | multiethnic study | | adjustment was | imputation I ² | on balance |
| | | validation study for the | of Hopping et al | | raised from | was already | conservative. |
| | | dietary instrument | (26) and the | | below to above | low at 2% | |
| | | used energy-adjusted | related | | that indicated by | implying the | |
| | | carbohydrate intakes | multiethnic study | | the imputation | assumption | |
| | | | of Howarth et al | | from the study | was | |
| | | | (35), who report | | of Howarth et al | reasonable. | |
| | | | GL with energy | | (35). | | |
| | | | adjustments . | | | The imputed | |
| | | | | | | information | |
| | | | | | | indicated a | |
| | | | | | | conservative | |
| | | | | | | estimate for | |
| | | | | | | RR was | |
| | | | | | | reached at | |
| | | | | | | line 1 | |
| | | | | The alternative | At the minimum | | |
| | | | | assumption was to | $(I^2 = 0.14\%)$ | | |
| | | | | apply an energy | RR was: | | |
| | | | | adjustment than | 1.49 (1.33,1.67) | | |
| | | | | minimized I ² , | | | |
| | | | | which is made | | | |
| | | | | viable because of | | | |
| | | | | the multiple | | | |
| | | | | observations by | | | |
| | | | | sex and ethnicity. | 2 | | |
| 9 | Values for CORR | Potentially, all studies | Values for CORR | Effect of | I ² was raised | There was | Assumption |
| | are imprecise | reviewed | are accurate other | additional random | trom 2 to 8% | no evidence | made (column |
| | | | than for random | error in CORR | | ot significant | 4) was |
| | | | error | was simulated to | RR per unit GL | error in RR | reasonable |
| | | | | examine deviation | was marginally | per unit GL | |
| | | | | in RR and the β - | higher, but not | due to this | |

| | | | | coefficient for CORR. (Error of mean 0 and SD 10% of the study range of CORR was applied.) | significantly so (t=0.66 for 10 repeats) to: 1.47 (1.30, 1.65) The β - coefficient for CORR fell from 2.05 to 1.89 (P=0.05, t-test, 10 repeats) | level of random error in CORR | |
|----|---|--|--|--|--|-------------------------------------|--|
| 10 | Adjustments were made towards deattenuated values for CORR when non- deattenuated values were | Villegas et al (22) reported CORR of 0.66, which became 0.71 following our adjustment towards a deattenuated value | An energy adjusted and deattenuated value for CORR is preferred. | No adjustment is made, and the non-deattenuated value 0.66 would have been preferred. | 1.47 (1.30,1.67) | Little different from line 1 | Assumption (Column 4) was reasonable, and on balance conservative. |
| | reported, or a deattenuated CORR within a study was estimable only approximately. | Simila et al (28) reported CORR of 0.55. However, the validation study (36) reports a value of 0.71 after adjustment for energy and deattenuation (Table S3 footnote 9). | An energy adjusted and deattenuated value for CORR is preferred. | The non-adjusted value 0.55 (28) should have been used rather than the deattenuated energy adjusted value reported in the validation study | 1.48 (1.31,1.67) | Little different from line 1 | Assumption (Column 4) was reasonable, and on balance conservative. |
| | | Hodge et al (20) describe agreement categorically between repeated FFQs, consistent with a CORR unadjusted for energy and deattenuation of 0.41, | An energy adjusted and deattenuated value for CORR is preferred and that estimated is reasonable | No adjustment is made, and the non-adjusted value of 0.41 would have been preferred | 1.46 (1.31,1.61) | Little different from line 1 | Assumption (Column 4) was reasonable, and on balance conservative. |

| | Assumption | and which became 0.57 following our adjustment towards a deattenuated and energy adjusted value (see Table S3 footnote 3). | variates (cf Table 3 i | n the main article) | | Comments |
|-----|---|--|------------------------|---------------------|-------------------|---|
| | investigated | | (| , | | |
| | | SEX | CORR | FUY (per 10 y) | ETH | |
| 1 * | Model 5 | 0.22 (0.02,0.46) | 2.05 (0.6,4.7) | 0.00 (-0.19.0.20) | 0.22(0.05,0.41) | |
| 2-6 | Highest value | 0.24 (0.03,0.50) | 2.21 (0.7,5.0) | 0.04 (-0.16,0.24) | 0.22(0.05,0.41) | |
| 26 | T (1 | | 1.07 (0.6.4.2) | 0.02(0.01.0.16) | | Alternative assumptions lead to |
| 2-6 | Lowest value | 0.19 (0.01, 0.41) | 1.87 (0.6, 4.3) | -0.03(-0.21,0.16) | 0.20 (0.04,0.39) | relatively small differences |
| | Added random error to GL (see above at 7) | 0.22 (0.01,0.47) | 1.78 (0.5, 4.3) | -0.01(-0.18,0.16) | 0.20 (0.06,0.38) | immediately left) |
| 8 | Alternative energy adjustment for obs. from Hopping et al (26) | 0.28 (0.04,0.56) | 2.42 (0.8,5.5) | 0.01(-0.13,0.30) | 0.23(0.05,0.44) | Adjustments based on imputations from Howarth et al (35) appear conservative compared with a best fitting solution. |
| 9 | Added random error to CORR (see above at 9) | 0.22 (0.06, 0.62) | 1.89 (0.50, 4.59) | -0.01(-0.14, 0.59) | 0.21 (0.03, 0.42) | Random error in CORR leads to an underestimation of the β coefficient for CORR |
| 10 | Adjustments to CORR. Highest values | 0.26 (0.04,0.53) | 2.29 (0.52,0.61) | -0.02(-0.22,0.29) | 0.21(0.04,0.41) | Alternative assumptions lead to relatively small differences compared with our preferred |
| 10 | Adjustments to CORR. Lowest values | 0.21(0.01,0.45) | 2.02 (0.052, 6.14) | -0.07(-0.28,0.15) | 0.18(0.01,0.39) | assumptions leading to coefficients at line 1 immediately left |

Abbreviations: CORR, the energy adjusted and deattenuated dietary instrument correlation for carbohydrate; FUY, the duration of follow-up in years; ETH, European American ethnicity versus all other ethnicities examined combined; f, female; I², percentage of total variance due to among-

studies variance. SEX, the proportion of study participants that are male (reported as RR for females > RR for males); RR, relative risk for T2D of 100g GL increment in 2000kcal diets.

* The number and all such below in this column refer to the corresponding assumptions and simulations noted in rows for corresponding numbers in this column above.

Table S6. Parameter estimates according the two-step meta-analysis approaches used aside a one-step approach, each on the full dataset of 24 studies

| Parameter | | Steps | Method ¹ | RR or ΔRR^2 | 95%CI ² | Р |
|-----------|--|----------------------|--|---------------------|-------------------------------|------------------|
| Adjuste | d RR | Two-step One-step | GLST then metareg Pooled GLST ³ | 1.45 1.46 | (1.31, 1.61) (1.31, 1.61) | <0.001 <0.001 |
| SEX | (ΔRR per 100g GL; F>M) | Two-step One-step | GLST then metareg Pooled GLST | 0.22 0.22 | (0.02,0.46) (0.03,0.45) | 0.031 0.024 |
| CORR | (ΔRR per 100g GL; over CORR to 1) | Two-step One-step | GLST then metareg Pooled GLST | 2.05 1.98 | (0.6, 4.7) (0.5, 4.8) | <0.001 0.001 |
| FUY | (ΔRR per 100g GL; over 10y) | Two-step One-step | GLST then metareg Pooled GLST | 0.00 0.01 | (-0.19,0.20) (-0.18,0.19) | 0.96 0.94 |
| ETH | (ΔRR per 100g GL; EA>Other) | Two-step One-step | GLST then metareg Pooled GLST | 0.22 0.22 | (0.05,0.41) (0.03,0.43) | 0.011 0.018 |
| | | | | | \mathbf{R}^2 \mathbf{I}^2 | P for I^{24} |
| Statistic | S | Two-step One-step | GLST then metareg Pooled GLST | | 97 2 100 0 | 0.43 0.78 |

^{*I.*} All models here assume relationships are linear.

^{2.} Analyses conducted in log form and displayed here in unlogged form.

^{3.} One step analysis of lnRR in pooled GLST meta-regression was versus 5 determinants: increment in GL dose (GL from Q_1 to $Q_{>1}$); and increments in four dose-covariate interactions (dose-x-SEX, dose-x-CORR, dose-x-FUY, dose-x-ETH, where SEX, CORR, FUY and ETH were centered).

⁴. The P-values: Two-step, Q-test for among-studies variance or model deviance from linearity; One step, chi² test for goodness-of-fit.

| | | | | | Hypothesiz | zed covaria | ate |
|---|------------------------|---------|--------|-----------|--------------|--------------------------|---------------------|
| _ | Units or subjective | Z-score | | | | | |
| Characteristic or influence factor ² | score | 3 | IRR | SEX | CORR | FUY | ETH |
| | | | Perc | entage c | hange in IF | R or β co | efficient |
| | | | due to | o additio | n of the po | tentially in | nfluențial |
| Dietary factor | | | stu | dy chara | cteristic as | s a 5 th cova | ariate ¹ |
| Study mean glycemic load | (g/2000 kcal) | -0.11 | < 0.1 | 4 | 2 | _ 4 | -2* |
| Study mean energy intake | (kcal/d) | -0.03 | 0.4 | -15* | 6 | - | -1 |
| Glucose or bread reference | (Glucose=1, bread=0) | 0.38 | -0.5 | 1 | -5 | - | 5* |
| No. of dietary assessments | (n) | 0.08 | -0.8 | 1 | -1 | - | 0 |
| No. foods in the dietary instruments | (n) | -1.40 | -3.7 | -10 | 19* | - | 9 |
| Dietary instrument (CORR) ⁵ | (Fractional) | | (| (Included | d in model | 5) | |
| " applicability within population | (Yes=1, doubtful=0) | 0.45 | -3.0 | 5 | -10 | - | 8 |
| " used energy-adjusted intakes | (Yes=1, no=0) | 1.75 | -4.0 | -23 | 4 | - | -9 |
| Fibre excluded as confounder | (Yes=1, no=0) | 0.08 | 0.1 | 3 | 1 | - | -1 |
| Population factor | | | | | | | |
| T2D excluded at baseline 6 | (Yes=1, no=0) | 0.00 | 0 | 0 | 0 | - | 0 |
| Population sample analysed (n) | (n) | 0.47 | -3.1 | 2 | -5 | - | 4 |
| European American versus other (ETH) ⁵ | (ETH=1, others = 0) | | | (Include | d in model | 5) | |
| Gender (SEX) ⁵ | (Male=1, Female=0) | | (| Include | d in model | 5) | |
| Body mass index | (kg/m^2) | -0.40 | -1.0 | 2 | -5 | - | -1 |
| Age at baseline | (y) | -1.44 | 9.7 | -23 | 11 | - | -9 |
| Progress factor | | | | | | | |
| Person-years | (ny) | 0.51 | -2.3 | 3 | -4 | - | 4 |
| Total incidents (cases) in study | (n) | 0.86 | -2.4 | 1 | -1 | - | 12 |
| Follow-up years (FUY) ⁵ | (years) | | (| (Included | d in model | 5) | |
| Outcome ascertainment (method of | · · · · | | | | | | |
| diagnosis) | (Clinical =1, self =0) | -0.37 | -0.1 | 2 | 1 | - | 3 |

Table S7. Influence of study factors on incremental RR values and β -coefficients for the covariates SEX, CORR, FUY and ETH (n=24 studies)¹.

| Study quality factors | | | | | | | |
|--|----------------------|-------|------|----|----|----|----|
| Selection criteria: | | | | | | | |
| Truly or somewhat representative of | | | | | | | |
| population | (Yes=1, no=0) | -0.45 | 3.5 | 0 | 1 | - | -5 |
| Cohorts selected from the same population | | | | | | | |
| 0 | (Yes=1, no=0) | 0.00 | 0.0 | 0 | 0 | - | 0 |
| Ascertainment of exposure (CORR >0.5) | (Yes=1, no=0) | _7 | _7 | _7 | _7 | _7 | _7 |
| Outcome of interest not present at start of | | | | | | | |
| study ⁶ | (Yes=1, no=0) | 0.00 | 0.0 | 0 | 0 | - | 0 |
| Total score for selection criteria | (Calc. range 0 to 4) | -0.23 | -3.0 | 0 | 6 | - | -5 |
| Outcome criteria: | | | | | | | |
| Secure assessment (clinical or self-report) | (Yes=1, no=0) | -0.37 | -0.1 | 2 | 1 | - | 3 |
| Follow-up period sufficient long (>4 y) 6 | (Yes=1, no=0) | 0.00 | 0 | 0 | 0 | - | 0 |
| Few subjects lost or lost explained ⁶ | (Yes=1, no=0) | 0.00 | 0 | 0 | 0 | - | 0 |
| Total score for outcome criteria | (Calc. range 0 to 3) | -0.37 | -0.1 | 2 | 1 | - | 3 |
| Comparability criteria: | | | | | | | |
| Study control for non-nutrient risk factors | (Yes=1, no=0) | 1.01 | 3.5 | 1 | 5 | - | -2 |
| Study control for nutrient risk factors | (Yes=1, no=0) | -0.04 | -0.9 | 2 | -1 | - | 0 |
| Total score for comparability criteria | (Calc. range 0 to 2) | 0.24 | 2.7 | 0 | 6 | - | -6 |
| Sum for all quality criteria | (Calc. range 0 to 9) | -0.33 | 1.1 | 1 | 3 | - | 2 |
| | | | | | | | |

Abbreviations: IRR, incremental relative risk for energy-adjusted GL increment of 100g after adjustment for centred covariates; SEX, of adult males or females ; CORR, dietary instrument correlation with food intake record for carbohydrate; FUY, number of followup years; EA, European American ethnicity versus other ethnicities; Calc., calculated as sum of the immediate above. * Indicates factors that correlate or potentially correlate (monovariate P<0.05) with the covariate as found in Table 1 in the main article.

- ¹ Influence values are shown as positive when the influence is to increase the absolute value of a coefficient and negative when lowering the absolute value. Thus the influence of each jth coefficient by each vth factor in the studies was expressed as $100^{*}(\Delta\beta)/\beta$ were β is the coefficient in meta-regression model 5 and $\Delta\beta$ is the change in β due to addition to the model of the vth among-study variable.
- 2 Influence factors were centered before assessing influence on model 5 coefficients.
- ³ Z-score for the β -coefficient accompanying this factor alongside the hypothesized covariates SEX, CORR, FUY and ETH.
- ⁴ Over all studies the β -coefficient for FUY was non-significant and near zero.
- ⁵ Factor included in the model.
- ⁶ These factors scored equally for all studies, so that zero effects on the β coefficients shown have zero degrees of freedom.
- ⁷ The full range of CORR is accounted in fully adjusted model 5.

FIGURE S1 (right). Meta-analysis of incremented RR with increase in GL from lowest to highest quantile, by sex group. Data points (■) for each study vary in size larger points for studies of greatest weight in the metaanalysis. The associated horizontal lines indicate the corresponding 95% confidence interval for a study-arrow heads indicating truncations. Diamond show the combined study means (upper and lower tips) and the corresponding 95% confidence intervals (left and right tips). Note that the scale for ΔRR is logarithmic with values showing the means for studies and untransformed combined means for subtotals and overall total. Abbreviations: Correlation, 100 x the correlation for carbohydrate intake (dietary instrument versus objective measure); RR, relative risk; LCI and UCI, lower and upper 95% confidence intervals; %WT, percentage weight based on random effects; P, level of statistical significance (z-test for RR, Q-test for I^2); I^2 , heterogeneity = $100 \cdot \tau^2 / (\tau^2 + se^2)$ where τ^2 is the among-studies variance; HaltMeka 08-11, data from Halton et al 2008 (25) and Mekary et al 2011 (31) are from the same study differing in the duration of follow-up (20 and 26y) and so were combined prior to the meta-analysis.

| First author, | No. | | | | | |
|-----------------------------------|-----------|------------------------------|--------|------|-------|------|
| year and ethnicity | subjects | Correlation | RR | LCI | UCI | %WT |
| Famala | | | | | | |
| Krishnan 2007 (23) | 40078 | 43 | 1 22 | 0.08 | 1 5 1 | 6.03 |
| Mover 2000 (17) | 35088 | 45 | 0.05 | 0.30 | 1.01 | 6.47 |
| Hopping 2010 (26) JA | 18672 | 54 | 1 18 | 0.70 | 1.10 | 4 42 |
| HalfMoka (08-(11 (25 31) | 83443 | 60.5 | → 1 70 | 0.00 | 3.63 | 1 15 |
| Zhang 2006 (21) | 13110 | 64 | 1.61 | 1.02 | 2.53 | 2 42 |
| Schulze 2004 (19) | 91249 | 64 | 1.33 | 0.92 | 1.91 | 3 34 |
| Salmeron 1997 (16) | 65173 | 64 | 1.47 | 1.16 | 1.86 | 5.56 |
| Hopping 2010 (26) NH | 5941 | 67 | 1 44 | 0.98 | 2 12 | 3 09 |
| Villegas 2007 (22) | 64227 | 71 | 1.34 | 1.13 | 1.58 | 7.30 |
| Hopping 2010 (26) EA | 14643 | 80 | → 2.13 | 1.37 | 3.31 | 2.54 |
| Subtotal $(l^2 = 50.3\%)$ | P = 0.03 | 4) | 1.32 | 1.50 | 1.51 | |
| | 0.00 | | | | | |
| Male | | | | | | |
| Hopping 2010 (26) JA | 16572 | 56 | 1.05 | 0.85 | 1.31 | 6.03 |
| Sakurai 2011 (29) | 1995 | 62 | 1.24 | 0.65 | 2.24 | 1.46 |
| Hopping 2010 (26) NH | 4568 | 62 | 1.10 | 0.76 | 1.61 | 3.21 |
| Hopping 2010 (26) EA | 15116 | 68 | 1.54 | 1.12 | 2.10 | 4.06 |
| Simila 2011 (28) | 25943 | | 0.88 | 0.65 | 1.17 | 4.40 |
| Salmeron 1997 (9) | 42759 | 73 | 1.25 | 0.90 | 1.73 | 3.87 |
| Subtotal (l ² = 32.6%, | P = 0.19 | 1) | 1.14 | 0.97 | 1.34 | |
| Mixed sexes | | | | | | |
| Stevens 2002 (18) AA | 2722 | 45 < | 0.97 | 0.73 | 1.35 | 4.17 |
| Stevens 2002 (18) EA | 9529 | 45 • | 1.10 | 0.90 | 1.39 | 6.00 |
| Mosdol 2007 (24) | 5598 | 49.5 | 0.80 | 0.51 | 1.26 | 2.44 |
| Hodge 2004 (20) | 31641 | 56 | 0.92 | 0.65 | 1.30 | 3.58 |
| Patel 2007 (15) | 124907 | 62 | 1.15 | 1.06 | 1.25 | 9.60 |
| Sahyoun 2008 (2) | 1898 | 65 < | 1.30 | 0.60 | 2.70 | 1.03 |
| Sluijs 2010 (27) | 37846 | 75 • | 1.83 | 1.27 | 2.53 | 3.61 |
| van Woudenbergh 2011 (30 |)) 4366 | 79 🔶 🔸 | 1.00 | 0.74 | 1.36 | 4.22 |
| Subtotal (l² = 46.3%, | P = 0.07 | 1) | 1.11 | 0.97 | 1.28 | |
| Overall (l² = 48.0%, F | e = 0.005 |) | 1.20 | 1.11 | 1.30 | 100 |
| Random effects analysis | | | | | | |
| | | | 20 | | | |
| | | U./5 1 1.5 Worse Better | 3.0 | | | |
| Relative | risk of | diabetes Q ₁ to Q | bv | sex | | |
| | | | ax, ~J | | | |

FIGURE S2 (right). Funnel plot of residuals (B) for model 5 obtained by the two-step approach to meta-analysis (cf Tables 2 and 3 in the main article). Points (•) represent individual studies. Triangular sides represent 95% CIs which ideally bound 95% of studies. Trim-and-fill analysis indicated no hypothetical points were required to eliminate bias if it were present.



FIGURE S3 (right). Factors hypothesized to affecting the size of the relative risk for T2D per 100g glycemic load. Panels A, the proportion of males in the sampled population (SEX). Panel; B, the validity of the dietary instrument (CORR). Panel; C, the number of followup y (FUY). Panel; D, European American participation (EA) versus all other ethnicities. Study means (O) with larger bubbles have greater weight. Lines are combined trends or means and their 95% CIs for 24 studies summarized in Table 3 of the min article by the fully adjusted model 5.



FIGURE S4 (right). Sensitivity of covariates to study deletions. With the 24 studies no more than 2-3 studies were expected to fall outside the box shown. To ± 1 study this was evident for each β -coefficient. Another feature, except for FUY, each box height was small relative to the z-score for the corresponding β coefficient (indicated by positive and negative values on the y-scale).



FIGURE S5 (below). **Cumulative meta-regression analysis**. Data show relative risk for type-2 diabetes associated with an increase in energy-adjusted glycemic load of 100g/2000kcal and the adjustment coefficients of covariance as they occur with each additional study (n=24) or update (n=25). *Note the number for updates exceeds number of studies by one as information from Mekary et al 2011 (31) for 26y follow-up updates that from Halton et al 2008 (25) at 20y follow-up in the present analysis. Points (**■**) show values reached after study inclusion based on the multivariate model 5 (Tables 2 and 3 in the main article). Horizontals are 95%CIs. Arrow heads truncate. Shaded areas represent the size of relation finally reached. Natural log scales are used and labels are unlogged. P-vales are shown right of points (z-test for RR and covariates). Units are as described in Tables 2 & 3 in the main article. Studies are identified by first author, date and other marks. *Abbreviations*: RR, relative risk; F or f, females; M or m, males; CORR, correlation for dietary instrument validity; FUY, follow-up y; AA, African-American; EA, European-American; ETH, European-American ethnicities versus other ethnicities combined; JA, Japanese-American; NH, Native Hawaiian.



Newcastle-Ottawa score of study quality (NOS) as used in the present study

While generally accepted that individual study quality should be assessed and reported when conducting systematic reviews, no method has been validated for non-randomized studies such as prospective cohort studies. The value of study quantity assessment remains for the present primarily in providing a measure to which a study has been conducted and reported according to generally recognized practices for studies deemed of high quality. Individual quality items and groups of quality items are generally recognized as potential determinants of a successful study and may correlate with study outcomes, but this should not be expected automatically and there is increasing recognition that study quality score should not be used as if a determinant of a study outcome.

The following reproduces the protocol as encountered (36) with insert in bold italics to adapt it to the present study.

Note: A study can be awarded a maximum of one star (*point*) for each numbered item within the Selection and Outcome categories. A maximum of two stars (*points*) can be given for Comparability

Selection for healthy persons representative of a community aiming for national (and eventually global) representation.

1) Representativeness of the exposed cohort

a) truly representative of the average *__adult mixed gender or male or female __* in the community ? *

b) somewhat representative of the average <u>adult mixed gender or male</u> or female_ in the community ?* For example not full age range of the community for which type-2 diabetes is incident.

c) selected group of users eg nurses, volunteers

d) no description of the derivation of the cohort

2) Selection of the non exposed cohort

a) drawn from the same community as the exposed cohort ? *

b) drawn from a different source

c) no description of the derivation of the non exposed cohort

3) Ascertainment of exposure

a) secure record (e.g. surgical records) ?* Dietary instrument used and

reported to be validated

b) structured interview ?*

c) written self report

d) no description

4) <u>Demonstration that outcome of interest</u> (*type-2 diabetes*) was not present at start of study

a) yes ?*

b) no

Comparability

1) Comparability of cohorts on the basis of the design or analysis

a) study controls for <u>exposure to known non-nutrient risk factors</u> <u>age, BMI, smoking, physical activity.</u> *

b) study controls for any additional factor ? *Exposure to suspected macronutritional risk factors, at least two from intakes of dietary fiber* (or cereal fiber) intake, energy intake, fat intake, and alcohol intake.*

Outcome

- 1) Assessment of outcome *
 - a) independent blind assessment ?
 - b) record linkage ? *Clinical report* *
 - c) self report
 - d) no description

2) Was follow-up long enough for outcomes to occur.

a) yes? Select yes if four or more years of follow-up (low to allow duration of follow up to be assessed as a covariate) *
b) no

3) Adequacy of follow up of cohorts

a) complete follow up - all subjects accounted for ? *

b) subjects lost to follow up unlikely to introduce bias - small number lost

- _<20%____ or description provided of those lost ?*

c) follow up rate _>20%_lost and no description of those lost.

d) no statement.

Registration of protocol.

Date of registration: 6 Dec 2011

Registration no. CRD42011001810 at http://www.crd.york.ac.uk/PROSPERO.

SUPPLEMENTAL MATERIALS REFERENCES

- Pereira MA. Dietary glycemic index and glycemic load in diabetes prevention--what can we learn from observational studies? Nat Clin Pract Endocrinol Metab 2008;4:430-1.
- Sahyoun NR, Anderson AL, Tylavsky FA, Lee JS, Sellmeyer DE, Harris TB, ; Health A, and Body Composition Study. Dietary glycemic index and glycemic load and the risk of type 2 diabetes in older adults. Am J Clin Nutr 2008;87:126-31.
- Hu FB, Manson JE, Stampfer MJ, Colditz G, Liu S, Solomon CG, Willett WC. Diet, lifestyle, and the risk of type 2 diabetes mellitus in women. N Engl J Med 2001;345:790-7.
- Salmeron J, Hu FB, Manson JE, Stampfer MJ, Colditz GA, Rimm EB, Willett WC. Dietary fat intake and risk of type 2 diabetes in women. Am J Clin Nutr 2001;73:1019-26.
- Mohan V, Radhika G, Sathya RM, Tamil SR, Ganesan A, Sudha V. Dietary carbohydrates, glycaemic load, food groups and newly detected type 2 diabetes among urban Asian Indian population in Chennai, India (Chennai Urban Rural Epidemiology Study 59). Br J Nutr 2009;102:1498-506.
- Schulz M, Liese AD, Fang F, Gilliard TS, Karter AJ. Is the association between dietary glycemic index and type 2 diabetes modified by waist circumference? Diabetes Care 2006;29:1102-4.
- Mayer-Davis EJ, Dhawan A, Liese AD, Teff K, Schulz M. Towards understanding of glycaemic index and glycaemic load in habitual diet: associations with measures of glycaemia in the Insulin Resistance Atherosclerosis Study. Br J Nutr 2006;95:397-405.
- Fung TT, Hu FB, Pereira MA, Liu S, Stampfer MJ, Colditz GA, Willett WC. Whole-grain intake and the risk of type 2 diabetes: a prospective study in men. Am J Clin Nutr 2002;76:535-40.

- Salmerón J, Ascherio A, Rimm EB, Colditz GA, Spiegelman D, Jenkins DJ, Stampfer MJ, Wing AL, Willett WC. Dietary fiber, glycemic load, and risk of NIDDM in men. Diabetes Care 1997;20:545-50.
- Barclay AW, Flood VM, Rochtchina E, Mitchell P, Brand-Miller JC. Glycemic index, dietary fiber, and risk of type 2 diabetes in a cohort of older Australians. Diabetes Care 2007;30:2811-3.
- Yu R, Woo J, Chan R, Sham A, Ho S, Tso A, Cheung B, Lam TH, Lam K. Relationship between dietary intake and the development of type 2 diabetes in a Chinese population: the Hong Kong Dietary Survey. Public Health Nutr 2011;14:1133-41.
- Woo J, Leung SSF, Ho SC, Lam TH, Janus ED. A food frequency questionnaire for use in the Chinese population in Hong Kong : description and examination of validity. Nutrition Research 1977;17:1633-41.
- Liu S, Chou EL. Dietary glycemic load and type 2 diabetes: modeling the glucose-raising potential of carbohydrates for prevention. Am J Clin Nutr 2010;92:675-7.
- Barclay AW, Petocz P, McMillan-Price J, Flood VM, Prvan T, Mitchell P, Brand-Miller JC. Glycemic index, glycemic load, and chronic disease risk--a metaanalysis of observational studies. Am J Clin Nutr 2008;87:627-37.
- Patel AV, McCullough ML, Pavluck AL, Jacobs EJ, Thun MJ, Calle EE.
 Glycemic load, glycemic index, and carbohydrate intake in relation to pancreatic cancer risk in a large US cohort. Cancer Causes Control 2007;18:287-94.
- Salmerón J, Manson JE, Stampfer MJ, Colditz GA, Wing AL, Willett WC. Dietary fiber, glycemic load, and risk of non-insulin-dependent diabetes mellitus in women. JAMA 1997;277:472-7.
- Meyer KA, Kushi LH, Jacobs DRJ, Slavin J, Sellers TA, Folsom AR.
 Carbohydrates, dietary fiber, and incident type 2 diabetes in older women. Am J Clin Nutr 2000;71:921-30.
- Stevens J, Ahn K, Juhaeri, Houston D, Steffan L, Couper D. Dietary fiber intake and glycemic index and incidence of diabetes in African-American and white adults: the ARIC study. Diabetes Care 2002;25:1715-21.

- Schulze MB, Liu S, Rimm EB, Manson JE, Willett WC, Hu FB. Glycemic index, glycemic load, and dietary fiber intake and incidence of type 2 diabetes in younger and middle-aged women. Am J Clin Nutr 2004;80:348-56.
- 20. Hodge AM, English DR, O'Dea K, Giles GG. Glycemic index and dietary fiber and the risk of type 2 diabetes. Diabetes Care 2004;27:2701-6.
- 21. Zhang C, Liu S, Solomon CG, Hu FB. Dietary fiber intake, dietary glycemic load, and the risk for gestational diabetes mellitus. Diabetes Care 2006;29:2223-30.
- 22. Villegas R, Liu S, Gao YT, Yang G, Li H, Zheng W, Shu XO. Prospective study of dietary carbohydrates, glycemic index, glycemic load, and incidence of type 2 diabetes mellitus in middle-aged Chinese women. Arch Intern Med 2007;167:2310-6.
- 23. Krishnan S, Rosenberg L, Singer M, Hu FB, Djousse L, Cupples LA, Palmer JR. Glycemic index, glycemic load, and cereal fiber intake and risk of type 2 diabetes in US black women. Arch Intern Med 2007;167:2304-9.
- 24. Mosdol A, Witte DR, Frost G, Marmot MG, Brunner EJ. Dietary glycemic index and glycemic load are associated with high-density-lipoprotein cholesterol at baseline but not with increased risk of diabetes in the Whitehall II study. Am J Clin Nutr 2007;86:988-94.
- 25. Halton TL, Liu S, Manson JE, Hu FB. Low-carbohydrate-diet score and risk of type 2 diabetes in women. Am J Clin Nutr 2008;87:339-46.
- 26. Hopping BN, Erber E, Grandinetti A, Verheus M, Kolonel LN, Maskarinec G. Dietary fiber, magnesium, and glycemic load alter risk of type 2 diabetes in a multiethnic cohort in Hawaii. J Nutr 2010;140:68-74.
- 27. Sluijs I, van der Schouw YT, van der AD, Spijkerman AM, Hu FB, Grobbee DE, Beulens JW. Carbohydrate quantity and quality and risk of type 2 diabetes in the European Prospective Investigation into Cancer and Nutrition-Netherlands (EPIC-NL) study. Am J Clin Nutr 2010;92:905-11.
- Simila ME, Valsta LM, Kontto JP, Albanes D, Virtamo J. Low-, medium- and high-glycaemic index carbohydrates and risk of type 2 diabetes in men. Br J Nutr 2011;105:1258-64.

- 29. Sakurai M, Nakamura K, Miura K, Takamura T, Yoshita K, Morikawa Y, Ishizaki M, Kido T, Naruse Y, Suwazono Y, et al. Dietary glycemic index and risk of type
 2 diabetes mellitus in middle-aged Japanese men. Metabolism 2011.
- 30. van Woudenbergh GJ, Kuijsten A, Sijbrands EJ, Hofman A, Witteman JC, Feskens EJ. Glycemic index and glycemic load and their association with Creactive protein and incident type 2 diabetes. J Nutr Metab 2011;2011:623076.
- Mekary RA, Rimm EB, Giovannucci E, Stampfer MJ, Willett WC, Ludwig DS, Hu FB. Joint association of glycemic load and alcohol intake with type 2 diabetes incidence in women. Am J Clin Nutr 2011;94:1525-32.
- Livesey G. Joint association of glycemic load and alcohol intake with type 2 diabetes incidence in women. Am J Clin Nutr 2012;95:983.
- Frost G, Leeds AA, Dore CJ, Madeiros S, Brading S, Dornhorst A. Glycaemic index as a determinant of serum HDL-cholesterol concentration. Lancet 1999;353:1045-8.
- 34. Brunner E, Stallone D, Juneja M, Bingham S, Marmot M. Dietary assessment in Whitehall II: comparison of 7 d diet diary and food-frequency questionnaire and validity against biomarkers. Br J Nutr 2001;86:405-14.
- 35. Howarth NC, Murphy SP, Wilkens LR, Henderson BE, Kolonel LN. The association of glycemic load and carbohydrate intake with colorectal cancer risk in the Multiethnic Cohort Study. Am J Clin Nutr 2008;88:1074-82.
- 36. Pietinen P, Hartman AM, Haapa E, Rasanen L, Haapakoski J, Palmgren J, Albanes D, Virtamo J, Huttunen JK. Reproducibility and validity of dietary assessment instruments. I. A self-administered food use questionnaire with a portion size picture booklet. Am J Epidemiol 1988;128:655-66.
- 37. Wells G, Shea S, O'Connell D, Robertson J, Peterson P, Welch V, Losos M, Tugwell P. The Newcastle-Ottawa Scale (NOS) for assessing the quality of nonrandomised studies in meta-analyses. URL http://www.evidencebasedpublichealth.de/download/Newcastle_Ottowa_Scale_P ope_Bruce.pdf. Accessed 28th August 2009.