# Technical Appendix

A model of the impact and cost-effectiveness of nonavalent HPV vaccination in the United States (2018 update)

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# 1 Overview of Model and Analytic Approach

Our deterministic and dynamic population-based model is an updated version of a model that has been used previously to examine a range of HPV vaccination strategies in the United States.<sup>1-3</sup>

## 1.1 Summary of previous model versions

In this section, we provide a brief overview of the history of the model.

## 1.1.1 2008 version of model

The first version of the model was published in 2008 and focused on quadrivalent HPV vaccination of young girls.<sup>1</sup> The original model version was static and did not account for HPV transmission dynamics. However, the published results did include some scenarios in which "herd effects" were incorporated by simple assumptions about the degree of the effect of herd immunity. Specifically, the "herd effects" scenarios were calculated by assuming an additional impact of the vaccine on non-vaccinated persons, including a reduction in genital warts in males. Cancers in males were not included.

#### 1.1.2 2011 version of model

The second version of the model was expanded to include HPV transmission dynamics, and was used in 2011 to examine quadrivalent HPV vaccination of males.<sup>2</sup> Another important change to the model was the inclusion of additional health outcomes, most importantly HPV-associated cancers in males (anal, penile, oropharyngeal). The model was also expanded to include the potential for HPV vaccination to prevent juvenile-onset recurrent respiratory papillomatosis (RRP) in the children of vaccinated mothers.

## 1.1.3 2016 version of model

The third version of the model was expanded to include the additional HPV types in the nonavalent HPV vaccine, and was used in 2016 to examine the cost-effectiveness of nonavalent HPV vaccination vs. quadrivalent HPV vaccination in the US.<sup>3</sup> This third version of the model was also applied (along with a much more complex individual-based model referred to as "HPV ADVISE") in a 2016 study of the cost-effectiveness of providing nonavalent HPV vaccine to females who had previously received the quadrivalent vaccine.<sup>4</sup>

## 1.2 Current (2018) version of the model

The model we applied in this study has the same structure as the 2016 version, except that we updated certain parameter values (vaccine cost, vaccination coverage, and medical treatment costs) as described below. We also modified the methods used for the sensitivity analyses as described in section 4.

## 1.2.1 Vaccine cost assumptions

The base-case vaccine cost per 3-dose series, including administration costs, was \$522 (range: \$372 to \$669). We assumed the vaccine cost per dose was \$116.22 (public) and \$193.63 (private) based on CDC vaccine price list

(https://www.cdc.gov/vaccines/programs/vfc/awardees/vaccine-management/price-list/)

as of March 5, 2017. The cost of administration per dose was \$8 public and \$29 private.<sup>5</sup> The base case value reflects an average of the public and private costs, and the range was calculated using the public costs (lower bound) and the private costs (the upper bound). The cost of a 2-dose series (for vaccination started through age 14 years) was two-thirds that of the 3-dose series.

#### 1.2.2 Vaccination coverage assumptions

We applied age- and sex- specific annual probabilities of vaccination based on estimated U.S. HPV vaccination coverage rates. For this 2018 application of the model, we used more recent estimates of HPV vaccination coverage than in the 2016 version.<sup>6,7</sup> See section 3.1 for details.

#### 1.2.3 Medical treatment costs

The medical treatment costs for each HPV-associated health outcome were obtained from the 2016 version of the model<sup>3</sup> (which reported costs in 2013 US dollars) and were updated to 2016 US dollars using the health care component of the Personal Consumption Expenditures price index (http://www.bea.gov).<sup>8</sup> See section 3.3 for details on the direct medical costs applied in the model.

## 1.3 Time horizon and analytic horizon

We examined the first 100 years of an HPV vaccination program (time horizon). The benefits of HPV-associated disease cases averted during the first 100 years of the vaccination program were allowed to accrue over the lifetime of the affected people (analytic horizon). The time horizon and analytic horizon are illustrated in Appendix Figure 1.

We modeled 191 birth cohorts, including the 92 cohorts between the ages of 8 years and 99 years (inclusive) in year one of the vaccine program and the subsequent 99 cohorts of incoming 8-year-olds in years 2 through 100 of the vaccine program. In each year, we focused on those aged 8 through 99 years in the given year; benefits of vaccination to those under age 8 years or over age 99 years in the given year were not included in calculations for the given year.

#### 1.4 Discounting and base year of costs

All future costs and benefits were discounted at 3% annually, consistent with US costeffectiveness recommendations<sup>9</sup> and with previous studies of the cost-effectiveness of HPV vaccination.<sup>2,10,11</sup> For additional details regarding the discounting of future costs and benefits, see sections 3.3.1 and 3.6.1.

## 1.5 Perspective

We used a health care system perspective and included costs and benefits of vaccination without regard to who incurs the costs or who receives the benefits. The costs and benefits we included were limited to the direct costs of vaccination, the direct medical costs averted by vaccination, and the QALYs saved by vaccination. Other potential costs and benefits (e.g., productivity costs, patient time and transportation costs) were beyond the scope of this analysis.

## 2 Model Description

Our model is a deterministic, dynamic population-based model. All results presented in the manuscript and in this appendix were calculated using Excel 2010 (Microsoft Corporation).

## 2.1 Overview of three main simplifying features of the model

Compared to other published models on the impact and cost-effectiveness of HPV vaccination strategies,<sup>10-14</sup> our model is relatively simple and requires fewer parameter values. Our approach uses three main simplifying features in approximating the impact of HPV vaccination: not explicitly modeling the natural history of HPV; not explicitly modeling cervical cancer screening; and using a simplified model of HPV transmission. These simplifying features are explained in more detail below.

#### 2.1.1 Approximating the impact of HPV vaccination

In estimating the impact of HPV vaccination, we did not explicitly model the natural history of HPV (e.g., the transition from HPV acquisition to HPV-associated health outcomes). Instead, the number of disease cases averted by vaccination for a given age cohort in a given year was approximated based on the percentage reduction in cumulative lifetime exposure to HPV in the given year for the given age cohort (described in more detail in section 2.3).

#### 2.1.2 Cervical cancer screening not explicitly modeled

Our approach does not explicitly model cervical cancer screening activities. Instead, we assume that the observed rates of cervical intraepithelial neoplasia (CIN) and cervical cancer applied in the model (those that have occurred in the context of current and historical cervical cancer screening practices in the US) would remain constant over time in the absence of HPV vaccination. Because we did not explicitly incorporate cervical cancer screening in our model, we cannot assess the impact of changing cervical cancer screening strategies. Thus, our model provides an assessment of the potential impact and cost-effectiveness of HPV vaccination in a scenario in which the probability of detection through screening remains constant over the duration of the HPV vaccine program.

## 2.1.3 Simplified transmission model

We used a relatively simple approach to depict the indirect effects ("herd effects") of vaccination. For example, we employed a discrete-time approach in which the impact of vaccination was modeled as a sequence of 1-year transitions among four mutually exclusive classes as described in section 2.2. As another example, we did not classify the population according to sexual activity level (e.g., rate of sex partner change). Instead, we assumed that in each year all people were subject to a sex- and age-specific probability of acquiring a specific

HPV type. As described in section 2.2.2.1, these sex- and age-specific HPV acquisition probabilities were adjusted each year in accordance with sex- and age-specific reductions in HPV in the population.

#### 2.2 Description of the model

The model is described below in detail for HPV 16. The benefits of vaccination against other HPV vaccine types (18, 31, 33, 45, 52, 58, and 6/11) were calculated in an analogous manner as summarized in section 2.3.3.

## 2.2.1 Model classes

Each age cohort was divided into four classes, based on vaccination status ("vaccinated" and "not vaccinated") and HPV 16 exposure status ("never infected" and "ever infected") as illustrated in Appendix Figure 2. The four possible classes are: (1) Not vaccinated; never acquired HPV 16, denoted as class X; (2) Vaccinated, never acquired HPV 16, denoted as class V; (3) Not vaccinated, acquired HPV 16 (ever), denoted as class Y; and (4) Vaccinated, acquired HPV 16 (ever), denoted as class Z. Movement occurs between the classes according to age-specific probabilities of acquiring HPV 16 and probabilities of being vaccinated. Vaccination reduces the probability of acquiring HPV 16 according to the vaccine efficacy assumptions.

We assumed that infection with HPV 16 provides lifelong natural immunity against HPV 16. Thus, those in the "not vaccinated, acquired HPV 16" class cannot move to the "not vaccinated, never acquired HPV 16" class, and those in the "vaccinated, acquired HPV 16" class cannot move to the "vaccinated, never acquired HPV 16" class.

## 2.2.2 Model equations

Each year, a cohort of 8-year-old boys and girls enters the model in the "not vaccinated, never acquired HPV 16" class and the cohort of 99-year-old men and women exits the model.

For each year t of the 100 years of the vaccination program, the model tracks cumulative HPV 16 exposure through year t for each age cohort (8 to 99 years) by sex. In each year, cumulative HPV 16 exposure for each age cohort in the scenario of HPV vaccination is compared to what it would have been in the absence of a vaccination program. Age-specific HPV 16 acquisition probabilities in year t+1 are adjusted proportionately according to reductions in cumulative HPV 16 exposure in year t in the opposite sex, to reflect changes in HPV prevalence in sex partners as a result of HPV vaccination.

These calculations are described in the equations below, in which k denotes sex (1 = female, 2 = male), a denotes age (in years, from ages 8 to 99), and t denotes year of vaccination program (year 0 through 100). For year 0 (the year before onset of the vaccination program in year 1), the percentage of each age cohort in the "Not vaccinated; never acquired HPV 16" class and the "Not vaccinated, acquired HPV 16 (ever)" class was calculated based on the cumulative probability of acquiring HPV 16 for sex k by age a in the absence of vaccination. The distribution of each age cohort into the four classes (X, Y, V, and Z as defined above and in Appendix Figure 2) was calculated as follows:

$$\begin{split} X_{k,a,t} &= X_{k,a-1,t-1}(1-\theta_{k,a,t})(1-\lambda_{k,a,t}) \\ Y_{k,a,t} &= X_{k,a-1,t-1}(1-\theta_{k,a,t})\lambda_{k,a,t} + Y_{k,a-1,t-1}(1-\theta_{k,a,t}) \\ V_{k,a,t} &= X_{k,a-1,t-1}(\theta_{k,a,t})(1-\lambda_{k,a,t}(1-E_{k})) + V_{k,a-1,t-1}(1-\lambda_{k,a,t}(1-E_{k})) \\ Z_{k,a,t} &= Z_{k,a-1,t-1} + V_{k,a-1,t-1}(\lambda_{k,a,t})(1-E_{k}) + Y_{k,a-1,t-1}(\theta_{k,a,t}) + X_{k,a-1,t-1}(\theta_{k,a,t})\lambda_{k,a,t}(1-E_{k}), \\ \text{and } X_{k,7,t} &= 1, \ Y_{k,7,t} = 0, \ V_{k,7,t} = 0, \ Z_{k,7,t} = 0, \end{split}$$

where  $\theta_{k,a,t}$  is the annual probability of receiving HPV vaccination for sex k at age a in year t,  $E_k$  is vaccine efficacy against HPV 16 acquisition for sex k, and  $\lambda_{k,a,t}$  is the annual probability of acquiring HPV 16 for sex k at age a in year t. The probability of acquiring HPV 16 ( $\lambda_{k,a,t}$ ) was

calculated as  $\lambda_{k,a,t} = P_{k,a,t}(1 - A_{k,a,t})$ , where  $P_{k,a}$  is the sex- and age-specific annual HPV 16 acquisition probability in the absence of vaccination (Appendix Table 33), and  $A_{k,a,t}$  is an adjustment term to account for population-level changes in HPV prevalence as described in section 2.2.2.1.

#### 2.2.2.1 Adjustment term (A)

The adjustment term  $A_{k,a,t}$  accounts for changes in HPV prevalence in the population due to HPV vaccination and was calculated based on changes in cumulative exposure to HPV 16 in the population, where we defined cumulative exposure to HPV 16 at a given age to be the probability of having acquired HPV at or before the given age. The reduction in cumulative exposure for sex k at age a in year t ( $C_{k,a,t}$ ) was calculated as  $C_{k,a,t} = 1$  -( $\bar{e}_{k,a,t}/e_{k,a,t}$ ), where  $e_{k,a,t}$  is the cumulative exposure to HPV 16 for sex k at age a years in the absence of an HPV vaccination program, and  $\bar{e}_{k,a,t}$  is the cumulative exposure to HPV 16 for sex k at age a years in year t of the vaccination program.

The adjustment term was calculated as  $A_{k,a,t} = (1-\varepsilon)\dot{C}_{k',a,t-1} + \varepsilon\hat{C}_{k',a,t-1}$ , where  $\dot{C}_{k',a,t-1}$  is the average of  $C_{k',a-5,t-1}$  through  $C_{k',a+5,t-1}$  (that is, the average value of C for those of sex k' within 5 years of age a, excluding those younger than age 8 years or older than age 99 years),  $\hat{C}_{k',a,t-1}$  is the average of  $C_{k',a,t-1}$  for ages 8 years through 99 years (i.e., the average of  $C_{k',8,t-1}$ through  $C_{k',99,t-1}$ ), and k' refers to the opposite sex from k. The term  $\varepsilon$  was used to reflect sexual mixing across age groups, where  $\varepsilon = 1$  corresponds to random mixing by age group and  $\varepsilon = 0$ corresponds to assortative mixing by age group such that all of a person's sex partners are within 5 years of age of that person. We used  $\varepsilon = 0.1$  to reflect the fact that mixing by age group tends to be assortative.<sup>10,15,16</sup> We did not specifically vary  $\varepsilon$  in the sensitivity analyses presented in the main manuscript, but the estimated impact of vaccination was varied in the probabilistic sensitivity analyses to reflect uncertainty in a range of factors, including  $\varepsilon$ .

## 2.3 Description of calculations of vaccine impact

#### 2.3.1 Deaths from other causes

For simplicity, the same age- and sex-specific death rates were applied to all classes (X, Y, V, and Z), such that the number of people in each cohort decreased from year to year due to death, but death did not influence the age and year-specific percentage of the population in each class. We made this simplifying assumption because HPV-attributable mortality is a very small fraction of overall mortality. The death rates we applied are listed in Appendix Table 34.<sup>17</sup>

## 2.3.2 Reduction in HPV 16 related health outcomes

Reductions in HPV- 16 associated cervical cancer, for women of age a in year t of the vaccination program, were calculated as  $R_a$  (POP<sub>a,t</sub>/100,000)(ATTRIB<sub>16</sub>)C<sub>1,a-lag,t-lag</sub> where  $R_a$  is the rate of cervical cancer (per 100,000) in age group a in the absence of vaccination, POP<sub>a,t</sub> is the number of females in age group a at time t, ATTRIB<sub>16</sub> is the percentage of cervical cancer attributable to HPV 16, C<sub>1,a,t</sub> is the reduction in cumulative infection with HPV 16 due to vaccination as described above, and *lag* is a disease-specific lag term. This lag term was included to establish a minimum time between vaccination and the prevention of a given health outcome. Although protection against the HPV vaccine types was assumed to begin after completion of the vaccine series, we applied the lag term so that the adverse health outcomes averted by vaccination would accrue over a plausible time frame. For cervical and other cancers (including cancers in males), we used a minimum lag time of 5 years such that reductions in cancer for a given age cohort would not be observed in the first 5 years in which members of that cohort were vaccinated.

The number of cases of other HPV 16-related health outcomes (other cancers, CIN 1, CIN 2/3) averted by vaccination were estimated in a manner analogous to that for cervical cancer. The lag term we applied was 1 for CIN 1, 2 for CIN 2/3, and 0 for genital warts.

## 2.3.3 Reductions in health outcomes attributable to other HPV types

The reduction in the number of cases of health outcomes attributable to other high-risk HPV types (18, 31, 33, 45, 52, and 58) was estimated in a manner analogous to that of HPV- 16 related health outcomes. Similarly, the reduction in the number of cases of HPV 6- and HPV 11- related health outcomes attributable to vaccination was estimated in the same fashion, except that HPV 6 and HPV 11 were treated as if they were a single HPV type ("HPV 6/11"). To clarify, we estimated eight versions of the model described above, in order to estimate reductions in health outcomes attributable to (1) HPV 16, (2) HPV 18, (3) HPV 31, (4) HPV 33, (5) HPV45, (6) HPV52, (7) HPV 58, and (8) HPV 6/11. These eight reductions in health outcomes were combined to estimate the impact and cost-effectiveness of HPV vaccination, according to the efficacy assumptions against each type for the given HPV vaccine.

## 2.3.4 Benefits of preventing RRP

We assumed HPV vaccination would reduce juvenile-onset RRP in children of vaccinated mothers, and these potential benefits of preventing RRP were approximated as described elsewhere.<sup>18</sup> Briefly, we applied the following age-specific birth rates (per 1000 women): 10- 14 years, 0.4; 15–17 years, 14.1; 18–19 years, 51.4; 20–24 years, 83.1; 25–29 years, 106.5; 30–34 years, 97.3; 35–39 years, 48.3; 40–44 years, 10.4.<sup>19</sup> Base case values related to the cost and quality-of-life impact of RRP are described elsewhere in this appendix but are summarized here for convenience. We applied a base case value 0.735 per 100,000 (range: 0.12–2.93 per 100,000) for the annual incidence rate of RRP (per child per year from birth through age

18) in the absence of HPV vaccination.<sup>18,20,21</sup> The RRP cost per case, discounted to birth assuming an average age of onset of RRP of four years and updated to 2016 US dollars, was was \$144,200 (range: \$69,200 - \$372,000).<sup>22</sup> We assumed 1.05 QALYs (range: 0.33 - 3.05) would be lost per case of RRP (discounted to birth).<sup>18</sup> We assumed a one-year value of the lag term (described earlier) when estimating the benefits of reductions in the probability of RRP in children born to vaccinated mothers.

## 3 Model Parameters

This section describes the parameter values applied in the model of HPV vaccine cost effectiveness.

## 3.1 Vaccination coverage

## 3.1.1 Base case vaccination coverage of females

Vaccination coverage assumptions are summarized in Appendix Table 1 and Appendix Table 2. We estimated the annual probability of HPV vaccination based on reported HPV coverage rates from National Immunization Survey-Teen (NIS-Teen).<sup>6,7</sup> We assumed the probability of HPV vaccination for females through age 12 years was 29.5%, so that girls turning 13 in our model would have the same 3-dose coverage rates as reported for 13 year olds in 2015.<sup>6</sup>

Estimates of the probability of vaccination for ages 13 to 18 were based on changes in vaccination coverage from 2014 to 2015 as follows. In 2014, 3-dose HPV vaccination coverage among females was 26.2% among 13-year-olds, 35.9% among 14-year-olds, 41.2% among 15-year-olds, and 43.8% among 16-year-olds.<sup>7</sup> In 2015, 3-dose HPV vaccination coverage among

females was 37.3% among 14-year-olds, 44.1% among 15-year-olds, 44.2% among 16-year-olds, and 54.4% among 17-year-olds.<sup>6</sup>

For HPV vaccination coverage rates among females to increase from 26.2% among 13year-olds in 2014 to 37.3% among 14-year-olds in 2015, the average annual probability of vaccination in this age interval would have to be about 15.0%. For HPV vaccination coverage rates among females to increase from 35.9% among 14-year-olds in 2014 to 44.1% among 15year-olds in 2015, the average annual probability of vaccination in this age interval would have to be about 12.8%. For HPV vaccination coverage rates among females to increase from 41.2% among 15-year-olds in 2014 to 44.2% among 16-year-olds in 2015, the average annual probability of vaccination in this age interval would have to be about 5.1%. For HPV vaccination coverage rates among females to increase from 43.8% among 16-year-olds in 2014 to 54.4% among 17-year-olds in 2015, the average annual probability of vaccination in this age interval would have to be about 18.9%. We used 12.9% (the average of 15.0%, 12.8%, 5.1%, and 18.9%) as the average annual probability of vaccination for females from ages 13 to 18 years.

We assumed the annual probability of vaccination for females from age 19 to 26 was 20% the annual probability from ages 13 to 18 years, as HPV vaccine uptake rates in adults are relatively low.<sup>23</sup> The 20% adjustment was applied so that the resulting probability of vaccination from aged 19-26 would be consistent with available data. With a 2.6% annual probability of vaccination, average 3-dose coverage among females aged 19-26 who were not vaccinated prior to age 19 years will be 11.0%, which is consistent with estimated 1-dose coverage of 11.8% among females aged 19-26 who were not vaccinated prior to age 19 years.<sup>23</sup>

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#### 3.1.2 Base case vaccination coverage of males

We estimated the annual probability of HPV vaccination among males in the same manner as for females. We assumed the probability of HPV vaccination for males at age 12 years was 24.9%, so that boys turning 13 in our model would have the same 3-dose coverage rates as reported for 13 year olds in 2015.

Estimates of the probability of vaccination for males ages 13 to 18 were based on changes in vaccination coverage in males from 2014 to 2015, as follows. In 2014, 3-dose HPV vaccination coverage among males was 16.2% among 13-year-olds, 20.9% among 14-year-olds, 24.9% among 15-year-olds, and 22.9% in 16-year-olds.<sup>24</sup> In 2015, 3-dose HPV vaccination coverage among males was 27.7% among 14-year-olds, 28.6% among 15-year-olds, 30.6% among 16-year-olds, and 28.8% among 17-year-olds.<sup>25</sup>

For HPV vaccination coverage rates among males to increase from 16.2% among 13year-olds in 2014 to 27.7% among 14-year-olds in 2015, the average annual probability of vaccination in this age interval would have to be about 13.7%. For HPV vaccination coverage rates among males to increase from 20.9% among 14-year-olds in 2014 to 28.6% among 15year-olds in 2015, the average annual probability of vaccination in this age interval would have to be about 9.7%. For HPV vaccination coverage rates among males to increase from 24.9% among 15-year-olds in 2014 to 30.6% among 16-year-olds in 2015, the average annual probability of vaccination in this age interval would have to be about 7.6%. For HPV vaccination coverage rates among males to increase from 22.9% among 16-year-olds in 2014 to 28.8% among 17-year-olds in 2015, the average annual probability of vaccination in this age interval would have to be about 7.7%. We used 9.7% (the average of 13.7%, 9.7%, 7.6%, and 7.7%) as the average annual probability of vaccination for males from ages 13 to 18 years.

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We assumed the probability of receiving HPV vaccine for males from age 19 to 26 was 20% the annual probability from ages 13 to 18 years, based on the assumptions applied for females as described above.

#### 3.1.3 Lower and higher coverage scenarios

We also examined a lower and higher coverage scenario. For the lower coverage scenario, we reduced the base case probability of vaccination for ages 13 to 18 years by 40% for females and 82% for males, so that the implied coverage among the 13- 17 year age groups was consistent with reported 3-dose coverage rates for this age group in 2015 (41.9% for females and 28.1% for males).<sup>25</sup> Whereas our base case probabilities reflect "vaccination incidence" rates extrapolated from NIS-Teen data, our lower bound probabilities were calculated to yield "vaccination prevalence" rates among 13 to 17 year olds consistent with NIS-Teen data.

For the higher coverage assumption, we followed the same approach used for the base case except that we examined annual changes in 1-dose coverage from 2014 to 2015 instead of annual changes in 3-dose coverage. The higher coverage scenario thus reflects 1-dose vaccine uptake rates instead of 3-dose vaccine uptake rates.

#### 3.2 Vaccine efficacy and cost

For females and males, vaccine efficacy was assumed to be 95% (range: 85% to 100%) for protection against infection with each of the HPV vaccine types.<sup>26</sup> As described in Section 1.2.1, the base-case vaccine cost per 3-dose series, including administration costs, was \$522 (range: \$372 to \$669).

## 3.3 Costs of HPV-associated health outcomes

Appendix Table 3 provides the lifetime, discounted, direct medical treatment costs we applied per health outcome, in 2016 US dollars. As noted in Section 1.2.3, these costs were obtained from the 2016 version of the model<sup>3</sup> (which reported costs in 2013 US dollars) and were updated to 2016 US dollars using the health care component of the Personal Consumption Expenditures price index (http://www.bea.gov).<sup>8</sup>

## 3.3.1 A note on discounting of averted medical costs

We multiplied the number of outcomes averted in year t by the estimate of the discounted lifetime medical cost of the outcome. This yielded the lifetime medical costs saved by the outcomes averted in year t, discounted to year t. In order to discount these averted medical costs to the onset of the vaccination program, we discounted these averted medical costs by an additional t- 1 years.

## 3.4 Disease incidence rates

The age-specific incidence rates of CIN, genital warts, and cervical and other cancers we applied in the model are listed in Appendix Table 4 through Appendix Table 15. These incidence rates were applied in our model as those that would be expected in the absence of HPV vaccination, and we calculated the reductions in these outcomes after onset of vaccination as described in section 2.

## 3.4.1 Incidence of CIN

Incidence rates for CIN 1 and CIN 2/3 (Appendix Table 4 to Appendix Table 5) are based on data from a 2009 study by Henk and colleagues using medical claims<sup>27</sup> and on data from a 2004 study by Insinga and colleagues using health plan administrative and laboratory data.<sup>28</sup> For CIN 1, base case values were based on the Henk study through age 59 years and the Insinga study for ages 60 to 79 years. The Henk study provided confidence intervals for ages 20 to 29 years and for ages 30 to 39 years. For ages 40 to 59 years, we approximated confidence intervals based on the confidence intervals for ages 30 to 39 years (relative to the base case value for ages 30 to 39 years). For ages 15 to 19 years, and for ages 60 years and over, the lower bound value was set to 0 and the upper bound value was set to twice the base case value.

For CIN 2/3, base case values were based on the Henk study through age 69 years and the Insinga study for ages 70 to 79 years. Similar to the approach above for CIN 1, for ages 40 to 69 years we approximated confidence intervals based on the confidence intervals for ages 30 to 39 years (relative to the base case value for ages 30 to 39 years). For ages 15 to 19 years, and for ages 70 and over, the lower bound value was set to 0 and the upper bound value was set to twice the base case value.

For CIN 1 and CIN 2/3, the incidence rates described above (including the lower and upper bounds) were reduced by 10% to account for lower utilization of cervical cancer screening services in the general US population as compared to women in the Henk and Insinga studies.<sup>28</sup>

## 3.4.2 Incidence of genital warts

Incidence rates for genital warts (Appendix Table 6 to Appendix Table 7) were based on reported incidence rates among a commercially-insured population.<sup>29</sup> To obtain lower bound values, we estimated 95% confidence intervals based on the age-specific incidence rates in that study and our conservative approximation of the age-specific sample sizes in that study.<sup>29</sup> The upper bound values we applied are estimates of genital warts prevalence (rather than incidence) rates among members of a privately insured population.<sup>30</sup>

## 3.4.3 Incidence of cervical and other HPV-associated cancers

Cancer incidence rates (Appendix Table 8 to Appendix Table 15) were obtained from population-based cancer registries that participate in the Centers for Disease Control and Prevention's National Program of Cancer Registries (NPCR) and the National Cancer Institute's Surveillance, Epidemiology, and End Results (SEER) Program.<sup>31,32</sup> The annual incidence rates we applied reflect the average annual rate over the period 2006 - 2010. Incidence rates and cancer case counts are suppressed if there are fewer than 16 cases. In such instances, we assumed a rate of 0. The International Classification of Diseases for Oncology (ICD-O- 3) codes that were used for the various cancer sites are listed in Appendix Table 16, along with other details of the cancer registry data.

## 3.4.4 Incidence of recurrent respiratory papillomatosis (RRP)

We assumed an annual incidence rate (per 100,000) of RRP of 0.735 (range: 0.12 to 0.293) for children through age 18 years.<sup>18,20,21</sup> Although we calculated the probability that a child would have RRP at any time from birth to age 18, for simplicity the costs of RRP and quality of life impact of RRP were calculated assuming that all cases of RRP occurred at age 4 years, as was assumed in one of the source studies for our RRP cost and quality of life estimates.<sup>18</sup> See section 2.3.4 for additional details of our assessment of vaccine impact on RRP.

3.5 Percentages of health outcomes attributable to HPV vaccine types

## 3.5.1 Percent of CIN 1 and CIN 2/3 attributable to HPV vaccine types

Appendix Table 17 provides the estimated percentages of CIN 1 and CIN 2/3 attributable to the nonavalent HPV vaccine types. Base case values and ranges for CIN 1 were obtained from a systematic review of the prevalence and attribution of HPV types in cervical precancers and cancers in the US.<sup>33</sup> The ranges reflect the 95% confidence intervals reported in that study.

Estimates for CIN 2 /3 were based on data from the HPV vaccine impact monitoring project (HPV-IMPACT).<sup>34</sup> The ranges we applied for CIN 2/3 represent the extreme values across four 5-year age groups (20- 24, 25- 29, 30- 34, and 35- 39 years). These ranges are notably greater than the 95% confidence intervals suggested by the HPV-IMPACT study and allow more uncertainty in the percent of CIN 2/3 attributable to each HPV type.

## 3.5.2 Percent of genital warts and RRP attributable to HPV 6 and 11

Appendix Table 18 provides the assumptions regarding the percentage of genital warts and RRP attributable to HPV 6 and 11. We assumed that HPV 6 and 11 account for 90% of genital warts (range: 70% to 100%)<sup>35,36</sup> and 90% of RRP (range: 70% to 100%). The values we applied for RRP were the same as for genital warts, based on evidence that a maternal history of genital warts in pregnancy is the strongest reported risk factor for RRP in the child.<sup>37</sup>

## 3.5.3 Percent of cancers attributable to high-risk nonavalent HPV types

The percent of cancers attributable to the nonavalent HPV types (Appendix Table 19 and Appendix Table 20) was based on a study of HPV typing of cancers in the US, which included data from seven cancer registries (Kentucky, Michigan, Louisiana, Florida, Hawaii, Iowa, and Los Angeles).<sup>38</sup> In these data HPV was detected in 91% of cases of cervical cancer.<sup>38</sup> This study provides grouped results for HPV 31,33,45,52, and 58. For these types, the base case values and confidence intervals from this study were provided by Trevor Thompson (personal communication, April 21, 2014). As described in Section 4, in the sensitivity analyses we modified the upper bound values when necessary so that the sum of the attributable percentages for each vaccine type was capped at 100%.

## 3.6 QALYs lost per HPV-related health outcome

Estimates of the age-specific, expected number of discounted lifetime quality-adjusted life years (QALYs) lost per HPV-related health outcome are presented in Appendix Table 21 through Appendix Table 29. When estimating the quality of life impact of HPV-associated health outcomes, we took into account the quality of life in the absence of these HPV-associated health outcomes (Appendix Table 30).<sup>39</sup> The methods we used to develop these QALY estimates are described below.

## 3.6.1 Discounting of QALYs lost per health outcome

We multiplied the number of outcomes averted in year t by the appropriate age- and sexspecific estimate of the discounted lifetime number of QALYs lost per health outcome. This yielded the number of QALYs saved by the outcomes averted in year t, discounted to year t. In order to discount these QALYs saved to the onset of the vaccination program, we discounted these QALYs by an additional t- 1 years.

#### 3.6.2 QALYs lost per case of genital warts

The number of QALYs lost per case of genital warts was calculated based on an average of two published studies. Drolet et al. (2011) provide estimates of the number of QALYs lost per case of genital warts, based on a study of 272 Canadian patients with genital warts.<sup>40</sup> Drolet et al. assessed quality of life impacts using the EuroQol EQ- 5D, a visual analog scale (VAS) and the Short-Form (SF)- 12 and estimated the loss in QALYs per episode of genital warts to be 0.017 to 0.041.<sup>40</sup> Woodhall et al. (2011) applied the EQ- 5D questionnaire to a sample of 370 patients with genital warts in England and Northern Ireland and found the loss in QALYs per episode of genital warts to be 0.018 (range: 0.008 to 0.031).<sup>41</sup> Combined, these two studies suggest that each episode of genital warts impose an average loss of about 0.024 QALYs per

episode, with a range of 0.008 to 0.041. This range is generally consistent with the range of 0.0014 to 0.039 estimated by Woodhall et al. (2009),<sup>42</sup> based on data on the average duration of genital warts among approximately 200 patients in the United Kingdom combined with previous estimates of the impact of genital warts on the disutility associated with genital warts by Woodhall et al. (2008).<sup>43</sup> To examine a wide range of plausible values for the quality of life loss to genital warts, we applied an upper bound value of the number of QALYs lost per case of genital warts of 0.10, which is consistent with assumptions of a relative loss in quality of life of 0.09 over a duration of about 1.1 years.<sup>2,10</sup> We do not use age-specific values for the number of QALYs lost per case of genital warts.

## 3.6.3 QALYS lost per case of CIN

The number of QALYs lost per case of CIN was based on two published studies. For our base case value, we used estimates from Drolet et al. (2012),<sup>44</sup> which assessed quality of life impacts of abnormal cervical smear results using the EuroQol EQ- 5D, a visual analog scale (VAS) and the Short-Form (SF)- 12. Their study included 952 Canadian women, of which 492 had an abnormal cervical screening result and 460 had a normal result. The loss in QALYs was about 0.007 per case of LSIL and 0.010 per case of HSIL. For our upper bound value, we used estimates from Insinga et al. (2007),<sup>45</sup> which combined information regarding the duration of various health states related to CIN with information on quality of life from an earlier patient preference study.<sup>46</sup> This approach suggested that 0.105 QALYs are lost per case of CIN1 and 0.115 QALYs are lost per case of CIN 2- 3. Owing to the considerable uncertainty in the impact of CIN on quality of life, we assigned a lower bound value of 0. We do not use age-specific values for the number of QALYs lost per case CIN.

#### 3.6.4 QALYS lost per case of RRP

We assumed 1.05 QALYs (range: 0.33 - 3.05) would be lost per case of RRP (discounted to birth), based on a study of how the inclusion of prevention of RRP can affect the estimated cost-effectiveness of HPV vaccination.<sup>18</sup>

## 3.6.5 QALYs lost per case of HPV-associated cancer

For each HPV-associated cancer, the age-specific number of QALYs lost was calculated based on quality of life in the absence of cancer (Appendix Table 30), quality of life detriments as a result of cancer (Appendix Table 31), and cancer survival probabilities (Appendix Table 32). Cancer survival probabilities were obtained for two age groups: those under age 50 years and those 50 years and older. In order to allow for a more gradual change with age in the probability of cancer survival, the probability of survival for those under age 45 was adjusted linearly through age 55 years, rather than being applied abruptly at age 50 years.

Quality of life detriments associated with HPV-associated cancers were obtained from Jit el al. (2011).<sup>47</sup> As described in more detail by Jit et al. (2011), the quality of life weights for cervical cancer treatment were obtained from a time-tradeoff study,<sup>46</sup> and were consistent with results obtained from applying the Health and Limitations Index (HALex) instrument to data from a nationally representative survey (National Health Interview Surveys [NHIS], 1987 to 1992).<sup>39</sup> Similarly, the quality of life weights used by Jit et al. (2011) for treatment of vulvar, vaginal, and anal cancers were based on the NHIS data and HALex instrument.<sup>39</sup> Jit et al. (2011) based their quality of life weights for oropharyngeal cancer treatment on a study that administered the EuroQol EQ- 5D survey by mail to oral and oropharyngeal cancer patients after primary surgery.<sup>48</sup> The quality of life weight for penile cancer treatment used by Jit et al. (2011) was based on expert opinion.<sup>49</sup> Jit et al. (2011) assumed a permanent, post-treatment reduction in quality of life of 0.0305 (range: 0 to 0.061) among cancer survivors, based on studies of cervical cancer survivors that used various health-related quality of life survey instruments.<sup>50-52</sup> We made the same assumption except that we applied an upper bound value of 0.15 rather than 0.061, based on 2010 NHIS data which suggested that cancer survivors were about 15 percentage points more likely to report poor physical health-related quality of life than adults without cancer.<sup>53</sup>

The number of QALYs lost per case of HPV-associated cancer was estimated by assuming that everyone with cancer would be subject to the treatment-related detriment to quality of life for exactly two years. After these two years, survivors would be subject to the permanent reduction in quality of life, and non-survivors would lose all of their remaining QALYs.

#### 3.6.5.1 Example: QALYs lost due to cervical cancer at age x years

The discounted number of QALYs lost per case of cervical cancer for a women diagnosed at age x years was calculated as follows. First, we calculated the potential QALY loss over the first two years as  $0.285*Q_x + (0.285*Q_{x+1}*[1-D_x])/(1+r)$ , where  $Q_x$  is the number of QALYs for a woman at age x years in the absence of cervical cancer (Appendix Table 30), 0.285 is the detriment to quality of life during treatment for cervical cancer as described above and in Appendix Table 31,  $D_x$  is the annual all-cause probability of death at age x years as in Appendix Table 34, and r is the discount rate (3%). The term  $D_x$  is included to account for the probability of death due to background mortality at age x years, so that no QALY losses are attributed to cervical cancer beyond age x years for those who would have died at age x years due to causes unrelated to cervical cancer.

Second, we calculated the QALY loss over the remaining years of life according to the probability of survival. The QALY loss at age x + 2 years due to cancer at age x years was calculated as  $(0.0305*Q_{x+2}*[1-D_x]*[1-D_{x+1}])/(1+r)^2$  for cancer survivors (for whom we assumed a residual loss in quality of life of 0.0305) and  $(Q_{x+2}*[1-D_x]*[1-D_{x+1}])/(1+r)^2$  for cancer non-survivors. Similarly, the QALY loss at age x + 3 years due to cancer at age x years was calculated as  $(0.0305*Q_{x+3}*[1-D_x]*[1-D_{x+1}]*[1-D_{x+2}])/(1+r)^3$  for cancer survivors and  $(Q_{x+3}*[1-D_x]*[1-D_x+1]*[1-D_x+2])/(1+r)^3$  for cancer survivors and  $(Q_{x+3}*[1-D_x]*[1-D_x+1]*[1-D_x+2])/(1+r)^3$  for cancer non-survivors. QALY losses in all remaining years (ages x + 4 years and beyond up to the maximum potential age of 99 years) were calculated in an analogous manner.

This example shows how we calculated the base-case, age-specific estimates of the number of QALYs lost per case of cervical cancer. Calculations for the other cancers were performed in an analogous manner.

## 3.6.5.2 Upper and lower bound values for QALY losses due to cancer

The lower bound values of the discounted number of QALYs lost due to cancer were calculated by applying the upper bound value of cancer survival (Appendix Table 32) and the lower bound values of the QALY detriments (Appendix Table 31). The upper bound values of the discounted number of QALYs lost due to cancer were calculated by applying the lower bound value of cancer survival and the upper bound values of the QALY detriments.

## 3.7 HPV incidence (annual HPV acquisition probabilities) in the absence of vaccination

Appendix Table 33 lists the type-specific annual probabilities of HPV acquisition by age that we applied in the model for ages 8- 60 years. These values represent the probability of acquisition of a given HPV type at a given age, provided no acquisition of that HPV type had occurred previously. The table provides values through age 60 years. We assumed the typespecific probability of HPV acquisition decreased by 10% in each year of age after age 60 years.

The base case values for the probability of acquisition of each HPV type were calculated as follows. First, the age-specific probability of acquisition of any HPV type was estimated as the average of the probabilities in two previously published models by Myers et al. (2000)<sup>54</sup> and Canfell et al. (2004).<sup>55</sup> The youngest age at which acquisition probabilities were provided was 15 years by Myers et al. (2000) and 16 years by Canfell et al. (2004). To calculate HPV acquisition probabilities for these younger ages, we assumed that HPV acquisition was possible beginning at age 13 years, and assigned probabilities of HPV acquisition by assuming that the probability of acquiring HPV at age x- 1 years was 0.25 that of age x years. That is, we calculated HPV acquisition probabilities for age 15 years, and we calculated HPV acquisition probabilities for age 15 years, and we calculated HPV acquisition probabilities for age 15 years, and we calculated HPV acquisition probabilities for age 15 years, and we calculated HPV acquisition probabilities for age 16 years.

Second, we smoothed the HPV acquisition probabilities noted above (which were provided by age group) to allow for gradual changes in the probability of HPV acquisition with age. In this smoothing process, the probability of HPV acquisition was held constant for age 12 years and age 60 years. For intermediate years, the smoothed probability of HPV acquisition at age x years, was set equal to the average of the unadjusted probability of HPV acquisition at age x- 1 years, age x years, and age x+1 years.

Third, we estimated type-specific acquisition probabilities by multiplying by the all-type acquisition probabilities by type-specific adjustment term. The adjustment for a given HPV type was selected manually so that the resulting HPV acquisition probabilities for the given type would be consistent with the observed prevalence of the given type in the US.<sup>56</sup> To do so, we

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calculated the age-specific HPV prevalence rates implied by our HPV acquisition rates under the following two assumptions: (1) the probability that an infection would be persistent was 8% for HPV 16 and 3.5% for other HPV types;<sup>13</sup> and (2) the average probability of clearance per year was 45% for high-risk HPV types and 75% for HPV 6/11.<sup>57</sup>

Our base case values of the age-specific HPV acquisition probabilities were based on models of HPV in females.<sup>54,55</sup> In the base case, we assumed that males and females have the same age-specific HPV incidence rates, owing to a lack of data on type-specific HPV incidence and prevalence among males.

#### 3.8 Issues regarding model fit, calibration and validation

In our model, we assess how HPV vaccination can reduce the burden of HPV-associated health outcomes, including cervical and other cancers, CIN, and genital warts. As described above, we did not specifically model the transition from HPV acquisition to HPV-associated diseases. A key simplification of our approach is that we estimated the percentage reduction in HPV-associated outcomes that can be achieved by vaccination, and then applied these percentage reductions to the existing burden of these HPV-associated outcomes in the absence of vaccination. The burden of HPV-associated outcomes in the absence of vaccination is based on the best data currently available: SEER/NPCR data for the incidence of HPV-associated cancers and medical claims data for the incidence of CIN and genital warts. Our model is constructed using age-specific disease incidence rates in the absence of HPV vaccination as listed in Appendix Table 4 to Appendix Table 15. Thus, it is of no use to show how well our model "fits" the disease incidence data in these tables, because in the absence of vaccination, our model would, by its simple design, predict exactly the same incidence rates as in Appendix Table 4 to Appendix Table 15.

## 4 Sensitivity Analyses

We conducted one-way and multi-way sensitivity analyses.

#### 4.1 One-way sensitivity analyses

In the one-way sensitivity analyses, one parameter value (or set of parameter values, as explained below) was varied at a time, holding all other parameters at their base case values. Specifically, we calculated the cost-effectiveness of 9vHPV vaccination strategies would change when varying one of the follow parameter values at a time: vaccine price per series; vaccine efficacy; the medical cost per case of the HPV-associated health outcomes; the number of QALYs lost per case of each health outcome; the incidence rates of the health outcomes in the absence of vaccination; and the percentages of the health outcomes attributable to the HPV vaccine types.

When varying vaccine efficacy, all vaccine efficacy parameters (HPV 16 efficacy, HPV 18 efficacy, HPV 6/11 efficacy, HPV 31 efficacy, and so on) were treated as a set and varied in the same manner. That is, when varying vaccine efficacy values, all efficacy parameters were varied together such that all were set to their lower bound value or all were set to their upper bound value (e.g., we did not examine scenarios in which HPV 16 efficacy was set to its upper bound value while HPV 18 efficacy was set to its lower bound value). The same approach was used for the remaining parameter groups (cost of HPV health outcomes, number of QALYs lost per HPV health outcome, and the percent of each health outcome attributable to the HPV vaccine types). That is, when varying the cost of HPV health outcomes, all cost parameters were varied together such that all were set to their lower bound value or all were set to their upper bound value (e.g., we did not examine scenarios in which the cost per case of cervical cancer was set to value (e.g., we did not examine scenarios in which the cost per case of cervical cancer was set to

its lower bound value while the cost per case of penile cancer was set to its upper bound value). Likewise, the number of QALYs lost per health outcome was varied as a group for all health outcomes, and the incidence rates of the health outcomes in the absence of vaccination were varied as a group for all health outcomes. Likewise, the percentages of the health outcomes attributable to the HPV vaccine types were varied as a group (e.g., we did not examine scenarios in which the percent of cervical cancer attributable to HPV 16 was set to its lower bound while the percent of cervical cancer attributable to HPV 18 was set to its upper bound; and we did not examine scenarios in which the percent of cervical cancer attributable to HPV 16 was set to its upper bound. When setting the percent of cancers attributable to HPV 16 was set to its upper bound value, the total percentage of cancers attributable to the 9 HPV vaccine types exceeded 100% for vaginal cancer, anal cancer in women, and anal cancer in men (the sums were 126%, 119%, and 124%, respectively). We therefore capped these sums at 100%, by reducing each upper bound value in a proportion manner.

## 4.2 Multi-way sensitivity analyses

We conducted multi-way sensitivity analyses to examine how the cost-effectiveness of 9vHPV vaccination strategies would change when numerous parameter values (or set of parameter values, as explained above) were varied simultaneously. Specifically, we conducted a probabilistic sensitivity analysis consisting of 5,000 model simulations. In each simulation, four sets of parameter values were varied (treatment costs per case of each health outcome, number of QALYs lost per case of each health outcome, the incidence rates of the health outcomes in the absence of vaccination, and the percentages of the health outcomes attributable to the HPV

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vaccine types). Parameter values within each parameter set were varied as a group as described above for the one-way sensitivity analyses.

We used the lognormal distribution for cost parameters because this is a common practice in health economic studies. That is, the lognormal distribution is often used to capture uncertainty in cost parameters, given that the cost estimates cannot be negative and cost estimates are typically right-skewed.<sup>13,58</sup> We used the lognormal distribution for the number of QALYs lost per case of each health outcome and for the incidence rates of the health outcomes, because these values are also constrained to be non-negative. For the percentage of each health outcome attributable to the HPV vaccine types, we assumed a uniform distribution between the lower and upper bound values.

We followed the methods of Elbasha and Dasbach (2010) to estimate the lognormal distribution parameters.<sup>13</sup> Specifically, we calculated the parameter  $\mu$  as ln(BASECASE) – 0.5ln[1+(SE<sup>2</sup>/BASECASE<sup>2</sup>)], where BASECASE is the base case value, ln indicates the natural log, and SE is the standard error (approximated as the upper bound minus the lower bound, divided by 2\*1.96). We calculated the parameter  $\sigma$  as the square root of ln[1+(SE<sup>2</sup>/BASECASE<sup>2</sup>)].

The 5,000 simulations were done using the base case assumptions regarding vaccine efficacy, vaccination coverage, and vaccine price. Even when assuming fixed values for vaccine efficacy, coverage, and price, the model is subject to uncertainty beyond that which is reflected by the four parameter sets listed above (treatment costs per case of each health outcome, number of QALYs lost per case of each health outcome, the incidence rates of the health outcomes in the absence of vaccination, and the percentages of the health outcomes attributable to the HPV vaccine types). To account for additional uncertainty, we applied an adjustment factor to each

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simulation to account for uncertainty in the model predictions regarding the percentage reduction in each HPV-associated health outcome. Specifically, before calculating cost-effectiveness ratios in each simulation, the total number of QALYs gained and the costs averted by vaccination were both multiplied by an impact adjustment factor which ranged from 0.75 to 1.25 and was assumed to follow a uniform distribution. This adjustment factor in effect allows for us to include scenarios in which the impact of vaccination (in terms of the percentage reduction in HPV-attributable health outcomes) is up to 25% less or 25% greater than suggested by the model, in addition to the effects of varying the percentage of health outcomes attributable to the HPV types, the incidence rates of the health outcomes included in the analysis, and assumptions regarding the cost and number of QALYs lost per case of disease.

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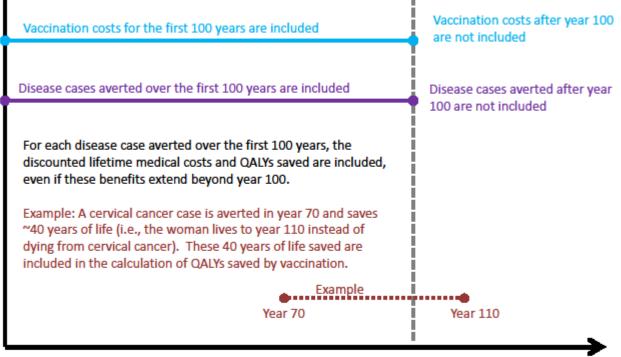
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## 6 Figures

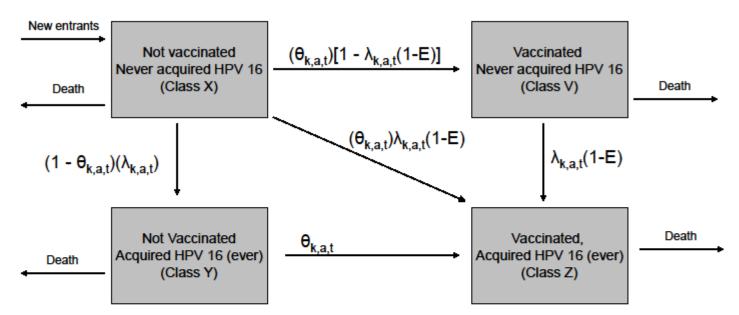
**Appendix Figure 1: Illustration of time horizon and analytic horizon** 



Year 1 of vaccination program

Year 100 of vaccination program

This figure illustrates the time horizon and analytic horizon of our analysis. We examined the first 100 years of an HPV vaccination program (time horizon). The benefits of HPV-associated disease cases averted during the first 100 years of the vaccination program were accrued over the lifetimes of the affected people (analytic horizon).



## **Appendix Figure 2: Illustration of model of cumulative, lifetime probability of exposure to HPV 16**

A cohort of susceptible 8-year-olds enters the population each year. In each year of the vaccine program (from year 1 to 100), there is a probability of death, a probability of acquiring HPV 16 ( $\lambda$ ) for those never infected with HPV 16 previously, and a probability of receiving HPV vaccination ( $\theta$ ). Vaccine efficacy against infection with HPV 16 is given by E. The subscripts k, a, and t denote sex, age, and year of vaccine program, respectively. The age- and sex-specific probabilities of HPV 16 infection ( $\lambda$ ) were adjusted for each year t to reflect changes in HPV exposure in sex partners as a result of HPV vaccination.

The reduction in HPV 16-related heath outcomes due to vaccination for a given age cohort in a given year was assumed to be proportional to the reduction in cumulative HPV 16 infection in that age cohort attributable to vaccination.

Reductions in health outcomes attributable to other HPV types were estimated in an analogous manner.

## 7 Tables

## 7.1 Tables of vaccination coverage assumptions

•••				· · ·		0
Age	Lower cov	verage	Base cove	rage	Higher co	verage
(years)	scenario		scenario		scenario	
	Female	Male	Female	Male	Female	Male
12	0.295	0.249	0.295	0.249	0.564	0.487
13 to 18	0.077	0.017	0.129	0.097	0.143	0.142
19 to $26^*$	0.015	0.003	0.026	0.019	0.029	0.028

## **Appendix Table 1: Probabilities of vaccination by age under three coverage scenarios**

# Appendix Table 2: Approximate cumulative vaccination coverage implied by vaccine probability assumptions

Age (years)	Lower cov scenario	verage	Base cove scenario	erage	Higher co scenario	verage
	Female	Male	Female	Male	Female	Male
13 to 17	41.9%	28.1%	49.0%	41.1%	69.6%	64.1%
17	52.9%	31.2%	64.7%	54.9%	79.8%	76.1%
26	61.6%	34.3%	75.0%	65.2%	86.3%	83.7%

## 7.2 Table of costs of health outcomes

## Appendix Table 3: Base case estimates and ranges of the cost per case of HPV-related health outcomes (2016 US dollars)

Health outcome	Base case value	Lower bound	Upper bound
CIN 1	\$1,340	\$930	\$1,750
CIN 2/3	\$2,470	\$1,030	\$4,010
Genital warts	\$660	\$330	\$750
Cervical cancer	\$42,000	\$33,200	\$56,900
Anal cancer	\$39,200	\$18,900	\$75,900
Vaginal cancer	\$29,300	\$21,900	\$36,900
Vulvar cancer	\$25,500	\$16,800	\$34,300
Oropharyngeal cancer	\$46,700	\$21,700	\$66,200
Penile cancer	\$21,400	\$10,600	\$42,000
RRP	\$144,200	\$69,200	\$372,000

CIN: cervical intraepithelial neoplasia. RRP: recurrent respiratory papillomatosis.

Cervical cancer screening costs were not included because we did not explicitly model cervical cancer screening, and these costs were assumed to be incurred regardless of HPV vaccination strategy.

The medical treatment costs for each HPV-associated health outcome were obtained from the 2016 version of the model<sup>3</sup> (which reported costs in 2013 US dollars) and were updated to 2016 US dollars using the health care component of the Personal Consumption Expenditures price index (<u>http://www.bea.gov</u>).<sup>8</sup> See section 3.3 for details on the direct medical costs applied in the model.

RRP costs reflect the expected lifetime costs of RRP discounted to birth, and we therefore discounted RRP cost estimates by 4 years to reflect the approximate average age at onset of RRP.

The base-case vaccine cost per 3-dose series, including administration costs, was \$522 (range: \$372 to \$669) as described in section 1.2.1.

## 7.3 Tables of disease incidence rates

#### 7.3.1 CIN incidence tables

## Appendix Table 4: Annual CIN 1 incidence rates (per person)

Age (years)	Base case	Lower bound	Upper bound
0-14	0	0	0
15-19	0.00036	0	0.00072
20-24	0.00297	0.00189	0.00414
25-29	0.00297	0.00189	0.00414
30-34	0.00261	0.00144	0.00423
35-39	0.00261	0.00144	0.00423
40-44	0.00171	0.00094	0.00277
45-49	0.00171	0.00094	0.00277
50- 54	0.00171	0.00094	0.00277
55- 59	0.00171	0.00094	0.00277
60- 64	0.00036	0	0.00072
65-69	0.00036	0	0.00072
70-74	0.00018	0	0.00036
75-79	0.00018	0	0.00036
80-84	0	0	0
85+	0	0	0

CIN: cervical intraepithelial neoplasia. See notes to Appendix Table 5.

## **Appendix Table 5: Annual CIN 2/3 incidence rates (per person)**

Age (years)	Rate	Lower bound	Upper bound
0-14	0	0	0
15-19	0.00018	0	0.00036
20-24	0.00324	0.00216	0.00441
25-29	0.00324	0.00216	0.00441
30-34	0.00243	0.00126	0.00396
35-39	0.00243	0.00126	0.00396
40-44	0.00054	0.00028	0.00088
45-49	0.00054	0.00028	0.00088
50- 54	0.00054	0.00028	0.00088
55- 59	0.00054	0.00028	0.00088
60- 64	0.00054	0.00028	0.00088
65-69	0.00054	0.00028	0.00088
70- 74	0.00009	0	0.00018
75-79	0.00009	0	0.00018
80-84	0	0	0
85+	0	0	0

CIN: cervical intraepithelial neoplasia

Incidence rates for CIN 1 and CIN 2/3 were based on data from a 2009 study by Henk and colleagues using medical claims data<sup>27</sup> and a 2004 study by Insinga and colleagues using health plan administrative and laboratory data.<sup>28</sup> We reduced their incidence estimates by 10% to account for lower intensity of resource use in the general population.<sup>28</sup>

#### 7.3.2 Genital warts incidence tables

Age (years)	Rate	Lower bound	Upper bound
10-14	0.00013	0.00004	0.00043
15-19	0.00223	0.00176	0.00287
20-24	0.00459	0.00356	0.00620
25-29	0.00272	0.00195	0.00394
30-34	0.00150	0.00119	0.00265
35-39	0.00150	0.00119	0.00199
40-44	0.00108	0.00081	0.00139
45-49	0.00108	0.00081	0.00144
50- 54	0.00073	0.00052	0.00092
55- 59	0.00073	0.00052	0.00086
60- 64	0.00062	0.00035	0.00076
65-69	0.00062	0.00029	0.00055
70-74	0.00045	0.00018	0.00055
75-79	0.00045	0.00018	0.00055
80-84	0.00016	0.00001	0.00055
85+	0.00016	0.00001	0.00055

## **Appendix Table 6: Annual genital warts incidence rates, females (per person)**

See notes in the following table.

#### **Appendix Table 7: Annual genital warts incidence rates, males (per person)**

Age (years)	Rate	Lower bound	Upper bound
10-14	0.00011	0.00003	0.00041
15-19	0.00074	0.00051	0.00065
20-24	0.00236	0.00176	0.00293
25-29	0.00272	0.00207	0.00501
30-34	0.00223	0.00183	0.00388
35-39	0.00223	0.00183	0.00252
40-44	0.00118	0.00094	0.00189
45-49	0.00118	0.00094	0.00128
50- 54	0.00092	0.00071	0.00118
55- 59	0.00092	0.00071	0.00086
60- 64	0.00048	0.00028	0.00100
65-69	0.00048	0.00024	0.00087
70- 74	0.00043	0.00020	0.00087
75-79	0.00043	0.00020	0.00087
80-84	0.00024	0.00008	0.00087
85+	0.00024	0.00008	0.00087

Based on reported incidence rates among a commercially-insured population.<sup>29</sup> Lower bound values reflect approximate 95% confidence intervals and the upper bound values are estimates of genital warts prevalence rates.<sup>30</sup>

#### 7.3.3 Cancer incidence tables

Age (years)	Rate	Lower bound	Upper bound
0-14	0.0	0.0	0.0
15-19	0.1	0.1	0.2
20-24	1.2	1.1	1.3
25-29	5.2	5.0	5.4
30-34	10.3	10.0	10.6
35-39	14.1	13.8	14.4
40-44	14.8	14.4	15.1
45-49	13.6	13.3	13.9
50-54	12.2	11.9	12.5
55- 59	12.2	11.9	12.5
60-64	12.0	11.7	12.4
65-69	12.0	11.7	12.5
70-74	11.2	10.7	11.6
75-79	9.6	9.1	10.0
80-84	9.0	8.6	9.5
85+	7.4	7.0	7.8

#### **Appendix Table 8: Annual cervical cancer incidence rates (per 100,000)**

Base case values and 95% confidence intervals were obtained from population-based cancer registries that participate in the National Program of Cancer Registries (NPCR) and the Surveillance, Epidemiology, and End Results (SEER) Program, 2006- 2010.<sup>31,32</sup> Incidence rates and cancer case counts are suppressed if there are fewer than 16 cases. In such instances, we assumed a rate of 0.

Age (years)	Rate	Lower bound	Upper bound
0-14	0.0	0.0	0.0
15-19	0.0	0.0	0.0
20-24	0.0	0.0	0.1
25-29	0.1	0.1	0.2
30-34	0.4	0.3	0.5
35-39	0.9	0.8	1.0
40-44	1.6	1.5	1.8
45-49	2.5	2.4	2.7
50- 54	2.9	2.8	3.1
55- 59	3.3	3.1	3.4
60-64	3.9	3.7	4.1
65-69	4.7	4.4	4.9
70-74	6.4	6.1	6.8
75-79	8.4	8.0	8.8
80-84	10.6	10.1	11.1
85+	12.8	12.3	13.3

#### **Appendix Table 9: Annual vulvar cancer incidence rates (per 100,000)**

Age (years)	Rate	Lower bound	Upper bound
0-14	0.0	0.0	0.0
15-19	0.0	0.0	0.0
20-24	0.0	0.0	0.0
25-29	0.0	0.0	0.0
30-34	0.1	0.0	0.1
35-39	0.1	0.1	0.2
40-44	0.3	0.2	0.3
45-49	0.4	0.4	0.5
50- 54	0.6	0.6	0.7
55- 59	0.8	0.7	0.9
60- 64	1.0	0.9	1.2
65-69	1.3	1.2	1.4
70- 74	1.7	1.6	1.9
75-79	2.1	1.9	2.3
80-84	2.5	2.3	2.8
85+	2.9	2.6	3.2

#### **Appendix Table 10: Annual vaginal cancer incidence rates (per 100,000)**

For more information, see Appendix Table 8 and Appendix Table 16.

#### **Appendix Table 11: Annual penile cancer incidence rates (per 100,000)**

Age (years)	Rate	Lower bound	Upper bound
0-14	0.0	0.0	0.0
15-19	0.0	0.0	0.0
20-24	0.0	0.0	0.0
25-29	0.0	0.0	0.1
30-34	0.1	0.1	0.2
35-39	0.2	0.2	0.2
40-44	0.4	0.3	0.5
45-49	0.5	0.5	0.6
50- 54	0.8	0.7	0.9
55- 59	1.2	1.1	1.3
60- 64	1.8	1.6	1.9
65-69	2.6	2.4	2.8
70- 74	3.5	3.3	3.8
75-79	4.4	4.1	4.8
80-84	5.1	4.7	5.5
85+	6.3	5.8	6.9

Age (years)	Rate	Lower bound	Upper bound
0-14	0.0	0.0	0.0
15-19	0.0	0.0	0.0
20-24	0.0	0.0	0.0
25-29	0.1	0.1	0.1
30-34	0.2	0.2	0.2
35-39	0.6	0.5	0.7
40-44	1.5	1.4	1.6
45-49	2.1	2.0	2.2
50- 54	2.4	2.3	2.5
55- 59	2.5	2.4	2.7
60-64	2.6	2.4	2.8
65-69	2.9	2.7	3.2
70- 74	2.8	2.5	3.0
75-79	2.8	2.5	3.0
80-84	2.9	2.6	3.3
85+	2.6	2.3	3.0

Appendix Table 12:	Annual anal	cancer incidence	e rates, males	(per 100,000	)
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For more information, see Appendix Table 8 and Appendix Table 16.

Age (years)	Rate	Lower bound	Upper bound
0-14	0.0	0.0	0.0
15-19	0.0	0.0	0.0
20-24	0.0	0.0	0.0
25-29	0.0	0.0	0.0
30-34	0.2	0.1	0.2
35-39	0.5	0.4	0.6
40-44	1.3	1.2	1.4
45-49	2.8	2.6	2.9
50- 54	4.3	4.1	4.5
55- 59	4.9	4.7	5.1
60-64	4.9	4.7	5.1
65-69	4.9	4.7	5.2
70-74	5.3	5.0	5.6
75-79	5.1	4.8	5.4
80-84	5.1	4.8	5.5
85+	4.5	4.2	4.8

## Appendix Table 13: Annual anal cancer incidence rates, females (per 100,000)

Age	Rate	Lower bound	Upper bound
0-14	0.0	0.0	0.0
15-19	0.0	0.0	0.0
20-24	0.0	0.0	0.1
25-29	0.1	0.1	0.1
30-34	0.3	0.2	0.3
35-39	1.2	1.1	1.3
40-44	3.9	3.8	4.1
45-49	9.8	9.5	10.0
50- 54	17.4	17.0	17.7
55- 59	23.6	23.1	24.0
60- 64	24.8	24.3	25.3
65-69	23.4	22.8	23.9
70- 74	20.3	19.7	21.0
75-79	15.9	15.3	16.6
80-84	13.0	12.4	13.7
85+	8.5	7.9	9.2

Appendix Table 14: Annual oropharyngeal cancer incidence rates, males (per 100,000)

For more information, see Appendix Table 8 and Appendix Table 16.

Age	Rate	Lower bound	Upper bound
0-14	0.0	0.0	0.0
15-19	0.0	0.0	0.0
20-24	0.0	0.0	0.1
25-29	0.1	0.0	0.1
30- 34	0.2	0.1	0.2
35-39	0.4	0.4	0.5
40-44	0.9	0.8	1.0
45-49	1.9	1.8	2.0
50- 54	3.2	3.0	3.3
55- 59	4.1	3.9	4.3
60- 64	4.8	4.6	5.0
65-69	5.3	5.1	5.6
70- 74	5.3	5.0	5.6
75-79	5.0	4.7	5.3
80-84	4.2	3.9	4.5
85+	3.2	3.0	3.5

## **Appendix Table 15: Annual oropharyngeal cancer incidence rates, females (per 100,000)**

	Site variable	Histologic type		Diagnostic
<b>Cancer site</b>	description	ICD-O- 3	Site code	Confirmation
Cervix	"Cervix Uteri"	8010- 8671,8940- 8941	C530-C539	Microscopically confirmed
Anus	"Rectum", "Anus, Anal Canal and Anorectum"	8050- 8084,8120- 8131	C210-C212, C218, C209	Microscopically confirmed
Penis	"Penis"	8050- 8084,8120- 8131	C600-C609	Microscopically confirmed
Vagina	"Vagina"	8050- 8084,8120- 8131	C529	Microscopically confirmed
Vulva	"Vulva"	8050- 8084,8120- 8131	C510-C519	Microscopically confirmed
Oropharynx	Primary Site=19,24,28,90- 91,98- 99,102,108- 109,140,142,148	8050- 8084,8120- 8131	(as listed)	Microscopically confirmed

## **Appendix Table 16: Details of cancer incidence data by cancer site**

ICD-O-: International Classification of Diseases for Oncology. Cancer incidence rates were calculated for the United States by including all states meeting United States Cancer Statistics (USCS) publication criteria for all years 2006- 2010, which covers approximately 94.8% of the US population.

Source: Meg Watson, personal communication, May 4, 2015.

## 7.4 Tables of HPV type attribution

HPV type	CIN 1	CIN 2/3
6/11	6.9 (2.7 - 17.0)	0
16	8.6 (5.3 - 13.7)	45.6 (32.4 - 50.4)
18	4.9 (2.5 - 9.2)	3.8 (1.6 - 8.3)
31	6.4 (3.6 - 11.1)	9.6 (7.0 - 15.0)
33	3.3 (1.5 - 7.1)	2.6 (1.3 - 5.5)
45	3.7 (1.7 - 7.6)	1.5 (0.4 - 3.4)
52	6.2 (3.5 - 10.8)	7.4 (4.9 - 12.9)
58	2.4 (1.0 - 6.0)	4.2 (2.0 - 5.5)

Appendix Table 17: Percent of cervical intraepithelial neoplasia (CIN) 1- 3 attributable to HPV types

Values are in percent. Estimates for CIN 1 were obtained from a systematic review of the prevalence and attribution of HPV types among cervical precancers and cancers in the United States.<sup>33</sup> Estimates for CIN 2/3 were based on data from the HPV vaccine impact monitoring project (HPV-IMPACT).<sup>34</sup> The ranges we applied for CIN 2/3 represent the extreme values across four 5-year age groups (20- 24, 25- 29, 30- 34, and 35- 39 years).

## Appendix Table 18: Percent of genital warts and recurrent respiratory papillomatosis (RRP) attributable to HPV 6 & 11

Genital warts         90.0         70.0         100           RRP         90.0         70.0         100	Health outcome	Base case	Lower bound	Upper bound
RRP 90.0 70.0 100	Genital warts	90.0	70.0	100
100	RRP	90.0	70.0	100

Values are in percent, and are based on several several sources<sup>35-37</sup> as described in sections 3.5.1 and 3.5.2.

HPV	Cervical	Vulvar	Vaginal	Penile
type				
16	50.1 (46.6 - 53.6)	48.1 (40.8 - 55.4)	53.4 (40.9 - 59.5)	45.2 (34.7 - 56.1)
18	16.1 (13.7 - 18.8)	0.6 (0.1 - 3.2)	1.7 (0.3 - 5.4)	2.7 (0.8 - 9.0)
31	2.1 ( 1.3 - 3.3)	1.1 (0.3 - 4.0)	0.0	0.0
33	3.5 ( 2.4 - 5.0)	9.3 (5.8 - 14.5)	11.6 (5.7 - 17.0)	5.1 (2.0 - 12.3)
45	5.5 ( 4.1 - 7.3)	0.6 (0.1 - 3.1)	3.3 (0.9 - 7.4)	2.7 (0.8 - 9.0)
52	1.8 (1.1 - 3.1)	2.7 (1.1 - 6.2)	1.7 (0.3 - 5.4)	1.3 (0.2 - 6.8)
58	1.8 ( 1.1 - 3.0)	0.6 (0.1 - 3.1)	1.7 (0.3 - 5.4)	0.0

Appendix Table 19: Percent of cervical, vulvar, vaginal, and penile cancers attributable to HPV types

Values are in percent. Values were obtained from a study of prevaccine type-specific prevalence of HPV–associated cancers in the United States.<sup>38</sup> This prevalence study provided grouped results for HPV 31,33,45,52, and 58. For these types, the base case values and confidence intervals from the study were provided by Trevor Thompson (personal communication, April 21, 2014). For vaginal cancer, anal cancer in females, and anal cancers in males, the sum of the upper bound percentage-attributable values exceed 100%. These were scaled down so that their sum was capped at 100%. The values shown here and in the subsequent table are the adjusted values.

HPV	Anal		Oroph	aryngeal
type	Male	Female	Male	Female
16	75.3 (62.3 - 79.3)	78.5 (69.1 - 80.9)	61.6 (56.9 - 66.0)	48.4 (40.6 - 56.4)
18	3.8 (1.1 - 7.5)	1.1 (0.2 - 2.7)	1.8 (0.9 - 3.6)	2.4 (0.9 - 6.2)
31	0.0	1.2 (0.2 - 2.8)	0.0 (0.0 - 0.9)	0.7 (0.1 - 3.7)
33	1.9 (0.3 - 5.2)	8.4 (4.3 - 10.9)	2.8 (1.6 - 4.8)	8.8 (5.3 - 14.5)
45	0.0	0.0	0.7 (0.2 - 2.0)	0.0
52	1.9 (0.3 - 5.2)	0.0	0.7 (0.2 - 2.0)	0.0
58	0.0 (0.0 - 2.8)	1.1 (0.2 - 2.7)	0.2 (0.0 - 1.3)	0.0

Appendix Table 20: Percent of anal and oropharyngeal cancers attributable to HPV types

Values are in percent. See notes to previous table.

#### 7.5 Tables of QALY losses per HPV-associated health outcome

Age	Base case	Lower bound	Upper bound
Genital warts	0.024	0.008	0.100
CIN 1	0.007	0.0	0.105
CIN 2/3	0.01	0.0	0.115
RRP	1.05	0.33	3.05

Appendix Table 21: Number of QALYs lost per case of genital warts, cervical intraepithelial neoplasia (CIN), and recurrent respiratory papillomatosis (RRP)

For references and additional information, see section 3.6: QALYs lost per HPV-related health outcome.

Age	Base case	Lower bound	Upper bound
15 to 19	6.32	5.48	8.76
20 to 24	6.08	5.28	8.42
25 to 29	5.82	5.04	8.04
30 to 34	5.51	4.77	7.61
35 to 39	5.17	4.48	7.13
40 to 44	4.80	4.16	6.61
45 to 49	5.00	4.41	6.59
50 to 54	5.89	5.37	7.18
55 to 59	5.97	5.51	7.03
60 to 64	5.20	4.79	6.12
65 to 69	4.39	4.04	5.16
70 to 74	3.55	3.26	4.17
75 to 79	2.76	2.53	3.25
80 to 84	2.06	1.88	2.42
85 to 89	1.56	1.41	1.83
90 to 94	1.26	1.14	1.48
95 +	0.57	0.50	0.67

#### Appendix Table 22: Number of QALYs lost per case of cervical cancer

Age	Base case	Lower bound	Upper bound
15 to 19	7.95	5.13	12.50
20 to 24	7.64	4.93	12.02
25 to 29	7.30	4.71	11.49
30 to 34	6.91	4.45	10.88
35 to 39	6.48	4.17	10.20
40 to 44	6.01	3.86	9.46
45 to 49	6.04	4.20	8.98
50 to 54	6.68	5.29	8.84
55 to 59	6.59	5.51	8.24
60 to 64	5.74	4.77	7.21
65 to 69	4.84	4.01	6.11
70 to 74	3.92	3.21	4.98
75 to 79	3.05	2.48	3.90
80 to 84	2.28	1.81	2.95
85 to 89	1.72	1.34	2.27
90 to 94	1.39	1.05	1.87
95 +	0.64	0.42	0.93

Appendix Table 23: Number of QALYs lost per case of vaginal cancer

## **Appendix Table 24: Number of QALYs lost per case of vulvar cancer**

Age	Base case	Lower bound	Upper bound
15 to 19	4.81	3.20	8.29
20 to 24	4.63	3.08	7.98
25 to 29	4.44	2.94	7.64
30 to 34	4.21	2.79	7.25
35 to 39	3.96	2.62	6.81
40 to 44	3.69	2.43	6.34
45 to 49	4.02	2.90	6.35
50 to 54	5.08	4.13	6.95
55 to 59	5.29	4.48	6.82
60 to 64	4.62	3.89	5.98
65 to 69	3.92	3.27	5.09
70 to 74	3.19	2.63	4.18
75 to 79	2.49	2.04	3.29
80 to 84	1.88	1.50	2.52
85 to 89	1.44	1.12	1.96
90 to 94	1.18	0.89	1.64
95 +	0.58	0.37	0.87

Age	Base case	Lower bound	Upper bound
15 to 19	6.62	4.32	10.63
20 to 24	6.34	4.14	10.18
25 to 29	6.04	3.94	9.69
30 to 34	5.69	3.72	9.12
35 to 39	5.30	3.46	8.48
40 to 44	4.87	3.18	7.78
45 to 49	4.66	3.22	7.14
50 to 54	4.70	3.63	6.58
55 to 59	4.41	3.58	5.88
60 to 64	3.79	3.07	5.05
65 to 69	3.17	2.56	4.22
70 to 74	2.56	2.06	3.41
75 to 79	1.99	1.59	2.65
80 to 84	1.51	1.19	2.01
85 to 89	1.17	0.92	1.56
90 to 94	0.99	0.77	1.31
95 +	0.51	0.37	0.67

Appendix Table 25: Nun	iber of QALYs lost p	per case of penile cancer

Age	Base case	Lower bound	Upper bound
15 to 19	6.73	4.80	10.30
20 to 24	6.49	4.62	9.92
25 to 29	6.22	4.41	9.52
30 to 34	5.91	4.18	9.05
35 to 39	5.56	3.92	8.52
40 to 44	5.19	3.64	7.95
45 to 49	4.93	3.53	7.41
50 to 54	4.83	3.63	6.97
55 to 59	4.50	3.46	6.33
60 to 64	3.98	3.02	5.63
65 to 69	3.43	2.56	4.87
70 to 74	2.86	2.08	4.09
75 to 79	2.28	1.63	3.29
80 to 84	1.81	1.23	2.64
85 to 89	1.45	0.94	2.15
90 to 94	1.25	0.78	1.88
95 +	0.76	0.38	1.19

## Appendix Table 26: Number of QALYs lost per case of anal cancer, females

Age	Base case	Lower bound	Upper bound
15 to 19	9.79	7.88	12.95
20 to 24	9.39	7.54	12.43
25 to 29	8.95	7.17	11.87
30 to 34	8.44	6.74	11.21
35 to 39	7.87	6.26	10.46
40 to 44	7.25	5.74	9.66
45 to 49	6.50	5.13	8.70
50 to 54	5.66	4.44	7.62
55 to 59	4.90	3.82	6.62
60 to 64	4.26	3.27	5.79
65 to 69	3.59	2.73	4.91
70 to 74	2.96	2.19	4.08
75 to 79	2.34	1.70	3.26
80 to 84	1.84	1.27	2.60
85 to 89	1.48	0.98	2.13
90 to 94	1.29	0.82	1.88
95 +	0.76	0.39	1.18

**Appendix Table 27: Number of QALYs lost per case of anal cancer, males** 

## Appendix Table 28: Number of QALYs lost per case of oropharyngeal cancer, female

Age	Base case	Lower bound	Upper bound
15 to 19	9.42	7.74	12.34
20 to 24	9.05	7.44	11.86
25 to 29	8.64	7.10	11.31
30 to 34	8.16	6.71	10.69
35 to 39	7.64	6.28	10.00
40 to 44	7.07	5.81	9.25
45 to 49	6.68	5.60	8.58
50 to 54	6.51	5.72	7.99
55 to 59	6.05	5.43	7.24
60 to 64	5.26	4.72	6.29
65 to 69	4.43	3.97	5.30
70 to 74	3.57	3.19	4.27
75 to 79	2.77	2.47	3.31
80 to 84	2.05	1.82	2.46
85 to 89	1.54	1.36	1.84
90 to 94	1.23	1.08	1.48
95 +	0.54	0.46	0.64

Age	Base case	Lower bound	Upper bound
15 to 19	6.81	5.73	9.34
20 to 24	6.52	5.48	8.95
25 to 29	6.21	5.22	8.51
30 to 34	5.84	4.91	8.01
35 to 39	5.44	4.57	7.44
40 to 44	4.99	4.19	6.83
45 to 49	4.81	4.12	6.41
50 to 54	4.95	4.42	6.24
55 to 59	4.69	4.26	5.74
60 to 64	4.02	3.64	4.92
65 to 69	3.35	3.03	4.09
70 to 74	2.69	2.42	3.28
75 to 79	2.08	1.87	2.54
80 to 84	1.56	1.39	1.90
85 to 89	1.20	1.06	1.46
90 to 94	1.00	0.88	1.21
95 +	0.48	0.40	0.58

Appendix Table 29: Number of QALYs lost per case of oropharyngeal cancer, male

#### 7.5.1 Tables of inputs used to calculate QALY losses for cancer

Age (years)	Male	Female	
17 and younger	0.93	0.93	
18 to 24	0.92	0.91	
25 to 34	0.92	0.91	
35 to 44	0.90	0.89	
45 to 54	0.87	0.86	
55 to 64	0.81	0.80	
65 to 74	0.76	0.78	
75 and older	0.69	0.70	

Appendix Table 30: Quality of life weights in absence of HPV-associated disease

Obtained from Gold et al. (1998).<sup>39</sup>

#### **Appendix Table 31: Quality of life detriments for treatment for HPV-associated cancers**

	Base case	Lower bound	Upper bound
Cervical cancer	0.285	0.24	0.33
Vaginal cancer	0.32	0.16	0.52
Vulvar cancer	0.32	0.16	0.52
Penile cancer	0.29	0.20	0.38
Anal cancer	0.51	0.21	0.83
Oropharyngeal cancer	0.25	0.20	0.30

All values in the above table are the same as applied by Jit et al.  $(2011)^{47}$  except the lower and upper bound values of 0.20 and 0.30 for oropharyngeal cancer, which we selected based on Jit et al.'s use of a normal distribution with a mean of 0.25 and a standard error of 0.02.

	Age < 5	Age < 50 years			Age 50 years and over		
	Base	Lower	Upper	Base	Lower	Upper	
	case	bound	bound	case	bound	bound	
Cervical	0.774	0.764	0.783	0.571	0.557	0.585	
Vaginal	0.705	0.592	0.792	0.524	0.470	0.575	
Vulvar	0.844	0.805	0.875	0.632	0.605	0.658	
Penile	0.755	0.662	0.826	0.667	0.621	0.708	
Anal, women	0.774	0.732	0.810	0.723	0.698	0.747	
Anal, men	0.629	0.584	0.670	0.654	0.618	0.688	
Oropharyngeal, women	0.634	0.580	0.683	0.560	0.533	0.586	
Oropharyngeal, men	0.743	0.721	0.764	0.635	0.622	0.647	

## **Appendix Table 32: Relative five-year cancer survival probabilities**

Obtained from Surveillance, Epidemiology, and End Results (SEER) Program data provided by Meg Watson and Jessica Blythe King, CDC (personal communication, April 18, 2014).

## 7.6 HPV acquisition probabilities and other data

••					•	·	<b>^</b>	
Age				HPV	' type			
(years)	6 & 11	16	18	31	33	45	52	58
12	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
13	0.0020	0.0016	0.0007	0.0009	0.0006	0.0008	0.0016	0.0005
14	0.0083	0.0067	0.0029	0.0037	0.0025	0.0034	0.0068	0.0022
15	0.0176	0.0142	0.0062	0.0079	0.0052	0.0071	0.0145	0.0047
16	0.0268	0.0215	0.0095	0.0120	0.0079	0.0108	0.0220	0.0072
17	0.0328	0.0263	0.0116	0.0146	0.0097	0.0133	0.0269	0.0088
18	0.0364	0.0293	0.0129	0.0163	0.0108	0.0147	0.0299	0.0098
19	0.0466	0.0374	0.0165	0.0208	0.0138	0.0189	0.0383	0.0125
20	0.0536	0.0431	0.0190	0.0239	0.0159	0.0217	0.0440	0.0144
21	0.0586	0.0470	0.0207	0.0261	0.0174	0.0237	0.0481	0.0157
22	0.0559	0.0449	0.0198	0.0250	0.0166	0.0226	0.0459	0.0150
23	0.0523	0.0420	0.0185	0.0233	0.0155	0.0212	0.0429	0.0140
24	0.0444	0.0357	0.0157	0.0198	0.0132	0.0180	0.0365	0.0119
25	0.0366	0.0294	0.0130	0.0163	0.0109	0.0148	0.0301	0.0098
26	0.0314	0.0252	0.0111	0.0140	0.0093	0.0127	0.0258	0.0084
27	0.0314	0.0252	0.0111	0.0140	0.0093	0.0127	0.0258	0.0084
28	0.0314	0.0252	0.0111	0.0140	0.0093	0.0127	0.0258	0.0084
29	0.0244	0.0196	0.0087	0.0109	0.0072	0.0099	0.0201	0.0065
30	0.0175	0.0141	0.0062	0.0078	0.0052	0.0071	0.0144	0.0047
31	0.0106	0.0085	0.0038	0.0047	0.0031	0.0043	0.0087	0.0028
32	0.0106	0.0085	0.0038	0.0047	0.0031	0.0043	0.0087	0.0028
33	0.0106	0.0085	0.0038	0.0047	0.0031	0.0043	0.0087	0.0028
34	0.0093	0.0075	0.0033	0.0042	0.0028	0.0038	0.0077	0.0025
35	0.0081	0.0065	0.0029	0.0036	0.0024	0.0033	0.0066	0.0022
36	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
37	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018

Appendix Table 33: Base case values of annual probability of HPV acquisition

Age				HPV	type			
(years)	6 & 11	16	18	31	33	45	52	58
38	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
39	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
40	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
41	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
42	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
43	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
44	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
45	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
46	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
47	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
48	0.0068	0.0055	0.0024	0.0030	0.0020	0.0027	0.0056	0.0018
49	0.0059	0.0048	0.0021	0.0027	0.0018	0.0024	0.0049	0.0016
50	0.0051	0.0041	0.0018	0.0023	0.0015	0.0021	0.0042	0.0014
51	0.0043	0.0034	0.0015	0.0019	0.0013	0.0017	0.0035	0.0011
52	0.0043	0.0034	0.0015	0.0019	0.0013	0.0017	0.0035	0.0011
53	0.0043	0.0034	0.0015	0.0019	0.0013	0.0017	0.0035	0.0011
54	0.0043	0.0034	0.0015	0.0019	0.0013	0.0017	0.0035	0.0011
55	0.0043	0.0034	0.0015	0.0019	0.0013	0.0017	0.0035	0.0011
56	0.0043	0.0034	0.0015	0.0019	0.0013	0.0017	0.0035	0.0011
57	0.0043	0.0034	0.0015	0.0019	0.0013	0.0017	0.0035	0.0011
58	0.0043	0.0034	0.0015	0.0019	0.0013	0.0017	0.0035	0.0011
59	0.0043	0.0034	0.0015	0.0019	0.0013	0.0017	0.0035	0.0011
60	0.0043	0.0034	0.0015	0.0019	0.0013	0.0017	0.0035	0.0011

Based on estimates of overall HPV acquisition probability by age applied in cervical cancer screening models by Myers et al. (2000)<sup>54</sup> and Canfell et al. (2004)<sup>55</sup>, scaled for each HPV type to be consistent with prevalence of HPV types observed in US <sup>56</sup> as described in section 3.7.

Age (years)	Male	Female
8 - 9	12.8	10.1
10 - 14	16.3	12.1
15 - 19	69.6	28.1
20 - 24	126.4	44.8
25 - 29	135.7	55.7
30 - 34	147.7	72.6
35 - 39	175.4	102.6
40 - 44	248.4	154.3
45 - 49	401.0	248.9
50 - 54	613.5	374.5
55 - 59	911.2	524.5
60 - 64	1,269.2	781.7
65 - 69	1,871.3	1,222.0
70 - 74	2,831.9	1,926.9
75 - 79	4,493.7	3,151.9
80 - 84	7,358.2	5,319.8
85+	15,414.3	13,219.2

Appendix Table 34: Annual death rates (per 100,000) applied in model

Source: National Vital Statistics Reports, Vol. 61, No. 4, May 8, 2013