

Quantifying degrees of necessity and of sufficiency in cause-effect relationships with dichotomous and survival outcomes - Supplementary material

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SUPPLEMENTARY TABLE 1: Small sample bias for estimating *DN* and *DS* for dichotomous outcome and dichotomous prognostic factor, summarized by median (quartiles), for selected scenarios in Table 3.

Scen.	P(D)	α	OR	Population		Bias (n=200)		Bias (n=500)	
				DN	DS	DN	DS	DN	DS
1	0.1	0.1	10	0.346	0.346	0.00 (-0.07; 0.07)	-0.01 (-0.08; 0.09)	0.00 (-0.04; 0.04)	0.00 (-0.05; 0.05)
2	0.1	0.5	10	0.787	0.087	0.01 (-0.10; 0.10)	0.00 (-0.02; 0.01)	0.00 (-0.06; 0.06)	0.00 (-0.01; 0.01)
3	0.1	0.9	10	0.878	0.011	0.12 (0.12; 0.12)	0.00 (0.00; 0.00)	0.12 (-0.08; 0.12)	0.00 (0.00; 0.00)
4	0.5	0.1	10	0.087	0.787	0.00 (-0.01; 0.01)	0.01 (-0.08; 0.11)	0.00 (-0.01; 0.01)	0.01 (-0.06; 0.06)
5	0.5	0.5	10	0.519	0.519	0.00 (-0.04; 0.05)	0.00 (-0.04; 0.05)	0.00 (-0.03; 0.03)	0.00 (-0.03; 0.03)
6	0.5	0.9	10	0.787	0.087	0.01 (-0.08; 0.11)	0.00 (-0.01; 0.01)	0.01 (-0.05; 0.05)	0.00 (-0.01; 0.01)
7	0.9	0.1	10	0.011	0.878	0.00 (0.00; 0.00)	0.12 (0.12; 0.12)	0.00 (0.00; 0.00)	0.12 (-0.08; 0.12)
8	0.9	0.5	10	0.087	0.787	0.00 (-0.02; 0.01)	0.00 (-0.09; 0.10)	0.00 (-0.01; 0.01)	0.00 (-0.06; 0.06)
9	0.9	0.9	10	0.346	0.346	0.01 (-0.06; 0.09)	0.01 (-0.07; 0.10)	0.00 (-0.05; 0.05)	0.01 (-0.03; 0.05)
28	0.1	0.1	1	0.000	0.000	0.14 (0.05; 0.47)	0.01 (0.01; 0.05)	0.09 (0.03; 0.29)	0.01 (0.00; 0.02)
29	0.5	0.5	1	0.000	0.000	0.05 (0.02; 0.08)	0.05 (0.02; 0.08)	0.03 (0.01; 0.05)	0.03 (0.01; 0.05)

Abbreviations: P(D), unconditional disease probability; OR, odds ratio; α , probability of harmful level of X. Results based on 500 simulated samples and presented as medians (quartiles).

SUPPLEMENTARY TABLE 2: Small sample bias for estimating DN and DS for dichotomous outcome and continuous prognostic factor, summarized by medians (quartiles), for the scenarios in Fig. 1, and additionally for $OR=1$.

P(D)	OR	Population		Bias (n=200)		Bias (n=500)	
		DN₁	DS₁	DN₁	DS₁	DN₁	DS₁
0.5	1	0.00	0.00	0.05 (0.02; 0.08)	0.05 (0.02; 0.08)	0.03 (0.01; 0.05)	0.03 (0.01; 0.05)
0.5	10	0.67	0.67	0.00 (-0.03; 0.03)	0.01 (-0.02; 0.04)	0.00 (-0.02; 0.02)	0.00 (-0.02; 0.02)
0.5	100	0.82	0.82	0.00 (-0.02; 0.03)	0.01 (-0.02; 0.03)	0.00 (-0.01; 0.02)	0.00 (-0.01; 0.01)
0.9	1	0.00	0.00	0.02 (0.01; 0.03)	0.15 (0.07; 0.23)	0.01 (0.00; 0.02)	0.08 (0.04; 0.14)
0.9	10	0.36	0.82	0.01 (-0.06; 0.06)	0.01 (-0.03; 0.04)	0.00 (-0.04; 0.03)	0.00 (-0.02; 0.02)
0.9	100	0.56	0.93	0.00 (-0.06; 0.06)	0.00 (-0.02; 0.02)	0.00 (-0.04; 0.04)	0.00 (-0.01; 0.01)

Abbreviations: $P(D)$, unconditional disease probability; OR , odds ratio. Results based on 250 simulated samples and presented as median (quartiles).

SUPPLEMENTARY TABLE 3: Small sample bias for estimating *DN* and *DS* for censored survival outcome and balanced dichotomous prognostic factor

HR	τ	% cens.	Type 1 censoring						Administrative censoring				
			Population		Bias (n=200)		Bias (n=500)		% cens.	Bias (n=200)		Bias (n=500)	
			<i>DN</i>	<i>DS</i>	<i>DN</i>	<i>DS</i>	<i>DN</i>	<i>DS</i>			<i>DN</i>	<i>DS</i>	<i>DN</i>
1	-	0	0.00	0.00	0.03 (0.04)	0.04 (0.06)	0.02 (0.03)	0.03 (0.04)					
	1.64	20	0.00	0.00	0.04 (0.04)	0.03 (0.04)	0.03 (0.03)	0.02 (0.03)	50	0.05 (0.06)	0.04 (0.04)	0.03 (0.08)	0.03 (0.06)
	0.22	80	0.00	0.00	0.10 (0.14)	0.01 (0.02)	0.07 (0.07)	0.01 (0.01)	90	0.15 (0.32)	0.02 (0.04)	0.09 (0.19)	0.01 (0.02)
2	-	0	0.22	0.31	0.00 (0.05)	0.00 (0.06)	0.00 (0.04)	0.00 (0.05)					
	2.26	23	0.25	0.20	0.01 (0.08)	0.01 (0.06)	0.00 (0.05)	0.00 (0.04)	50	0.00 (0.08)	0.00 (0.07)	0.00 (0.06)	0.00 (0.06)
	0.30	80	0.32	0.04	0.02 (0.22)	0.00 (0.02)	-0.01 (0.12)	0.00 (0.02)	90	0.00 (0.28)	0.00 (0.03)	-0.01 (0.15)	0.00 (0.02)
10	-	0	0.55	0.62	0.00 (0.04)	0.00 (0.04)	0.00 (0.03)	0.00 (0.03)					
	4.90	31	0.71	0.44	0.00 (0.06)	0.00 (0.04)	0.00 (0.04)	0.00 (0.02)	50	0.00 (0.09)	0.00 (0.06)	0.00 (0.05)	0.00 (0.04)
	0.43	81	0.80	0.09	0.00 (0.11)	0.00 (0.02)	0.00 (0.07)	0.00 (0.02)	90	0.01 (0.16)	0.00 (0.05)	0.01 (0.10)	0.00 (0.02)
100	-	0	0.67	0.69	-0.01 (0.04)	0.00 (0.04)	0.00 (0.02)	0.00 (0.02)					
	15.30	43	0.96	0.45	0.00 (0.03)	0.00 (0.03)	0.00 (0.02)	0.00 (0.02)	50	0.00 (0.03)	0.01 (0.04)	0.00 (0.02)	0.00 (0.03)
	0.47	81	0.98	0.11	0.02 (0.04)	0.00 (0.02)	0.00 (0.03)	0.00 (0.02)	90	0.02 (0.04)	0.00 (0.05)	0.02 (0.04)	0.00 (0.03)

Abbreviations: HR, hazard ratio; τ , time of maximum follow-up; % cens., percentage of censored observations. Results based on 250 simulated samples and presented as medians (inter-quartile ranges).

SUPPLEMENTARY TABLE 4: Small sample bias for estimating *DN* and *DS* for uncensored survival outcome and dichotomous prognostic factor for the scenarios of Table 5.

	HR	Population		Bias (n=200)		Bias (n=500)	
		<i>DN</i>	<i>DS</i>	<i>DN</i>	<i>DS</i>	<i>DN</i>	<i>DS</i>
$\alpha = 0.1$	1	0.00	0.00	0.02 (0.01; 0.08)	0.04 (0.02; 0.18)	0.01 (0.01; 0.05)	0.02 (0.01; 0.09)
	2	0.05	0.47	0.00 (-0.01; 0.01)	0.01 (-0.07; 0.06)	0.00 (-0.01; 0.01)	-0.01 (-0.05; 0.03)
	10	0.16	0.85	0.00 (-0.02; 0.02)	0.00 (-0.03; 0.02)	0.00 (-0.01; 0.01)	0.00 (-0.01; 0.01)
	100	0.24	0.94	0.01 (-0.02; 0.04)	-0.01 (-0.02; 0.00)	0.00 (-0.02; 0.03)	0.00 (-0.01; 0.00)
$\alpha = 0.5$	1	0.00	0.00	0.03 (0.01; 0.05)	0.04 (0.02; 0.08)	0.02 (0.01; 0.04)	0.03 (0.01; 0.05)
	2	0.22	0.31	0.00 (-0.02; 0.03)	0.00 (-0.03; 0.03)	0.00 (-0.02; 0.02)	0.00 (-0.03; 0.02)
	10	0.55	0.61	0.00 (-0.02; 0.02)	0.00 (-0.02; 0.02)	0.00 (-0.02; 0.01)	0.00 (-0.02; 0.01)
	100	0.68	0.68	-0.01 (-0.03; 0.01)	0.00 (-0.02; 0.02)	0.00 (-0.01; 0.01)	0.00 (-0.01; 0.01)
$\alpha = 0.9$	1	0.00	0.00	0.02 (0.01; 0.08)	0.04 (0.01; 0.13)	0.02 (0.01; 0.06)	0.02 (0.01; 0.08)
	2	0.36	0.08	-0.01 (-0.09; 0.07)	0.00 (-0.03; 0.02)	-0.01 (-0.06; 0.05)	0.00 (-0.02; 0.01)
	10	0.80	0.21	0.00 (-0.04; 0.03)	-0.01 (-0.03; 0.02)	0.00 (-0.03; 0.02)	0.00 (-0.02; 0.01)
	100	0.93	0.25	-0.01 (-0.03; 0.00)	0.00 (-0.03; 0.02)	0.00 (-0.01; 0.01)	0.00 (-0.02; 0.01)

Abbreviations: HR, hazard ratio; α , probability of harmful level of X. Results based on 250 simulated samples and presented as medians (quartiles).

Programming Code

In the following we provide a SAS macro and an R function to calculate DN and DS for dichotomous outcomes. See <http://cemsii.meduniwien.ac.at/en/kb/science-research/software/statistical-software> for latest versions as well as for additional functions for survival.

SAS Macro

```
*****;
*
* NECSUFF
* =====
*
* SAS-macro for computation of degrees of necessity and sufficiency
*
* Gleiss, A. & Schemper, M., Quantifying degrees of necessity and
* sufficiency in cause-effect relationships with dichotomous and
* survival outcome
* submitted to Stat.Med.
*
* Author:  Andreas Gleiss
* Version: 1.0
* Date:    8 August 2018
*
* Macro parameters:
* =====
*
* data      name of SAS data set
* y         name of dichotomous outcome variable
* x         name of independent variable (or linear predictor)
* refcat   reference category of y (first=default or last)
* inpred   dataset containing predictions
* inpredvar name of variable in inpred dataset containing
*          predictions
* odssel   control output from proc logistic
*          (default: ResponseProfile)
* print    =1 to print results (default), =0 (e.g., for simulations)
*
*****;

%macro necsuff(data, y, x, refcat=first, inpred=, inpredvar=,
              odssel=ResponseProfile, print=1);
data _work;
  set &data;
  if missing(&x) or missing(&y) then delete;
run;
%if &inpred= %then %do;
  ods select &odssel;
  proc logistic data=_work outest=_betas;
    class &y / param=ref ref=&refcat;
    model &y=&x;
    output out=_est1 pred=p_i_dach;
  run;
  data _betas;
    set _betas(rename=(intercept=beta0 &x=beta1));
    _dummy=1;
    keep _dummy beta0 beta1;
%end;
endmacro necsuff;
```

```

run;
%end;
%else %do;
data _est1;
set &inpred;
p_i_dach=&inpredvar;
run;
data _betas;
_dummy=1; beta0=.; betal=.; output;
run;
%end;
proc sort data=_est1;
by descending p_i_dach;
run;
data _est1;
set _est1;
_dummy=1;
_row=_n_;
run;
proc means data=_work noprint;
var &y;
output out=_my1 mean=p_bar max=_dummy;
run;
data _est1_;
merge _est1
_my1(keep=p_bar _dummy _freq_ rename=( _freq_=n))
_betas;
by _dummy;
retain dn1_sum ds1_sum dn2_sum ds2_sum n_smaller n_larger (0 0 0 0 0 0);

smaller=(p_i_dach<p_bar);
larger=(p_i_dach>p_bar);
n_smaller=n_smaller+smaller;
n_larger=n_larger+larger;

if p_bar~=0 then do;
if smaller then dn1_sum=dn1_sum + ((p_bar - p_i_dach)/p_bar)**2;
if smaller then dn2_sum=dn2_sum + ((p_bar - p_i_dach)/p_bar);
end;
if p_bar~=1 then do;
if larger then ds1_sum=ds1_sum + ((p_i_dach - p_bar)/(1-
p_bar))**2;
if larger then ds2_sum=ds2_sum + ((p_i_dach - p_bar)/(1-p_bar));
end;
run;

data _necsuff_;
set _est1_;
where _row=n;
if n_smaller~=0 then do;
dn_1=dn1_sum/n_smaller;
dn_2=dn2_sum/n_smaller;
end;
if n_larger~=0 then do;
ds_1=ds1_sum/n_larger;
ds_2=ds2_sum/n_larger;
end;
weight_dn=p_bar/(1-p_bar) * n_smaller/n;

```

```

weight_ds=(1-p_bar)/p_bar * n_larger/n;
ev_indir=weight_dn*dn_1_ + weight_ds*ds_1_;
dn_1=sqrt(dn_1_);
ds_1=sqrt(ds_1_);
alpha=n_larger/n;
or=exp(beta1);
progn_fact="&x";
keep progn_fact p_bar alpha or ev_indir dn_1 ds_1 dn_2 ds_2;
run;
ods select all;
%if &print=1 %then %do;
  title "Degrees of necessity and sufficiency";
  title2 "Outcome = &y";
  proc print data=_necsuff_ noobs label;
    var progn_fact p_bar alpha or ev_indir dn_1 ds_1 dn_2 ds_2;
    format or p_bar alpha ev_indir dn: ds: f5.3;
    label
      progn_fact="Prognostic factor"
      or="OR" ev_indir="EV" p_bar="est.P(D)"
      dn_1="DN1" ds_1="DS1" dn_2="DN2" ds_2="DS2";
  run;
  title; title2;
%end;
%mend;

*%necsuff(data=sh_tab3.hl_prost, y=capsule, x=psa);
*%necsuff(data=sh_tab3.hl_prost, y=capsule, x=gleason, odssel=none);

data lungca; * Swedish cohort study (Nilsson et al, 2001);
  do i=1 to 8120;
    x=0; y=0; output;
  end;
  do i=1 to 36;
    x=0; y=1; output;
  end;
  do i=1 to 4331;
    x=1; y=0; output;
  end;
  do i=1 to 177;
    x=1; y=1; output;
  end;
run;
%necsuff(data=lungca, y=y, x=x);

```

R function

```
#####  
#  
# NecSuff  
# =====  
#  
# R-function for computing degrees of necessity and sufficiency  
#  
# Gleiss, A. & Schemper, M. Quantifying degrees of necessity and  
# sufficiency in cause-effect relationships with dichotomous and  
# survival outcome,  
# submitted to Stat.Med.  
#  
# Author:  Andreas Gleiss  
# Version: 1.0  
# Date:    09 Aug 2018  
#  
# Arguments:  
# =====  
#  
# pred      name of variable containing predictions  
#  
#####  
  
NecSuff<-function(pred){  
  p_bar<-mean(pred)  
  smaller<-(pred<p_bar)  
  larger<-(pred>p_bar)  
  DN1<-sqrt(mean(((p_bar-pred[smaller])/p_bar)^2))  
  DS1<-sqrt(mean(((pred[larger]-p_bar)/(1-p_bar))^2))  
  DN2<-mean((p_bar-pred[smaller])/p_bar)  
  DS2<-mean((pred[larger]-p_bar)/(1-p_bar))  
  
  w_dn<-p_bar/(1-p_bar) * sum(smaller)/length(pred)  
  w_ds<-(1-p_bar)/p_bar * sum(larger)/length(pred)  
  EV<-w_dn*DN1^2 + w_ds*DS1^2  
  
  cat('\nest.P(D) =',round(p_bar,3))  
  cat('\nDN1 =',round(DN1,3),', DS1 =',round(DS1,3))  
  cat('\nDN2 =',round(DN2,3),', DS1 =',round(DS2,3))  
  cat('\nEV =',round(EV,3))  
}  
  
xx<-rnorm(1000)  
pp<-1/(1+exp(-5-log(5)*xx))  
yy<-1*(runif(1000)<=pp)  
plot(yy~xx)  
points(cbind(xx,pp),col="red")  
res.glm<-glm(yy~xx, family=binomial)
```



```
summary(res.glm)
pred<-predict(res.glm, type="response")
NecSuff(pred)

# Swedish cohort (Nilsson et al, 2001)
xx<-c(rep(0,36),rep(1,177),rep(0,8120),rep(1,4331))
yy<-c(rep(1,36),rep(1,177),rep(0,8120),rep(0,4331))
table(yy,xx)
res.glm<-glm(yy~xx, family=binomial)
summary(res.glm)
pred<-predict(res.glm, type="response")
NecSuff(pred)
```