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Can Waterlow score predict 30-day mortality and length of stay in acutely admitted medical patients (aged ≥65 years)? Evidence from a single centre prospective cohort study

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Can Waterlow score predict 30-day mortality and length of stay in acutely admitted medical patients (aged ≥65 years)? Evidence from a single centre prospective cohort study

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Waterlow Score
30-Day Mortality
Length of Stay

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2877

Abstract

Objective: To explore the potential for the Waterlow Score (WS) to be used as a predictor of 30-day mortality and length of hospital stay (LHS) in acutely admitted medical patients aged 65 years and over.

Design: Prospective observational cohort study.

Setting: United Kingdom District General Hospital.

Subjects: 834 consecutive patients aged 65 years and over admitted acutely to medical specialities between 30 May and 22 July 2014.

Methods: Admission WS (range 4-64) assessment paired with the patient's status at 30-days in terms of mortality and their length of hospital stay.

Primary Outcomes: 30-day mortality and length of inpatient stay.

Results: 834 consecutive acute medical admissions had their WS recorded. 30-day mortality was 13.1% (109 deaths). A significant difference in distribution of WS (p<0.001) was seen between those who survived (median 12) and those who died (median 16) within 30 days, particularly within respiratory (p<0.001), stroke (p<0.001), cardiology (p<0.016), non-respiratory infections (p<0.018) and trauma (p<0.044) subgroups. Odds of dying within 30 days increased 3-fold for every 10-unit increase in the WS (p<0.001). LHS was also positively linearly associated with the WS in those who survived 30 days (median=5, IQR=10; r= 0.32, p<0.01). A 5-unit increase in WS was associated with approximately 5 days increase in LHS. On the other hand, quadratic regression showed this relationship was curvilinear and negative (concave) for those who died within 30 days where a 5-unit increase in WS was associated with an approximately 10 days decrease in LHS.

Conclusion: This study demonstrates an association between a high WS and both 30-day mortality and LHS. This is particularly significant for mortality in patients in the respiratory, stroke and cardiac subcategories. The WS, a nursing led screening tool that is carried out on virtually all admissions to UK hospitals, could have additional use at the time of patient admission as a risk assessment tool for 30-day mortality as well as a predictor of length of hospital stay.

Key words: Waterlow Score, 30-day mortality, length of stay

Strengths and Limitations of this study

- This study examines the additional utility of a familiar widely used tool that is already a mandatory assessment for all United Kingdom Hospital admissions (the Waterlow Score).
- This study suggests that a higher Waterlow Score on admission is significantly correlated with higher 30-d mortality. This indicates the Waterlow Score may be used as a risk assessment tool for 30-day mortality.
- The findings indicate the Waterlow Score may be used as a predictor of length of hospital stay for patients 65 or over.
- Known high inter-observer variability inherent to Waterlow Score assessment, relating to assessor's training and experience.
- This study was an observational study done prospectively in a cohort of patients above 65 in a single district general hospital, and therefore requires further research in a wider context to confirm our observations.

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Introduction

Predictive mortality and morbidity scores are widely used to aid clinical decision making and resource allocation. Hospitals routinely collect a range of data on patient admission, however this is rarely standardised between hospitals.[1] A range of validated risk assessment models are commonly used in everyday clinical settings, usually with reference to a specific pathology e.g. the Mini-Grace risk score[2] for suspected acute coronary syndromes. These specific assessment models cannot be extrapolated beyond their validated clinical settings because of their disease specificity. The use of more general prognostic models is limited by their need to consider a broad range of clinical and demographic information due to the heterogeneous nature of patient populations. Broader prognostic models have been developed such as the APACHE III score[3], developed as a gatekeeper to Intensive Care, and frailty indices[4,5] are used in the geriatric setting. However, none are simple or practical and are limited by complexity and time constraints.

The Waterlow Score[6] (WS) is a multi-dimensional pressure ulcer risk assessment tool encompassing functional status and co-morbidity which was developed in the 1980's. It is a nursing led screening tool that is carried out on virtually every patient on admission to a United Kingdom hospital. Its components are age, nutritional status, weight, mobility, gender, smoking status, co-morbidities, use of medication and continence, all combined to generate a single numerical figure between 2 and 64. Patients with scores over 10 are considered to be at risk of pressure ulcer, whilst scores of over 15 and 20 are considered to represent high risk and very high risk statuses, respectively. [6] Other notable pressure ulcer risk assessments exist[7], but the WS remains the most widely used in the United Kingdom.

The primary aim of this study is to look at the potential role of the WS as a predictor of 30-day mortality in acute medical admissions aged 65 years and over and secondly to look at their length of hospital stay (LHS) as a marker of morbidity burden. The application of the WS is appropriate because its criteria already take into account those of other comorbidity models such as the Charlson comorbidity index.[8] The latter, a widely used comorbidity measure, is a component of the dataset used in standardised hospital mortality indicators in the United Kingdom.[9] Deriving additional utility from the WS could aid patient risk stratification, potentially improving patient care and guiding resource management. An association between the WS and outcomes has already been demonstrated in a cohort of surgical patients.[10] Furthermore, identifying additional applications of the Waterlow score would not incur any additional time resource requirements.

Methods

Study population
Following approval from a research ethics committee, the study group prospectively recruited all consecutive acute medical patients aged 65 years and over admitted during an eight-week period between 30 May and 22 July 2014 to the Lister Hospital, a 720 bed District General Hospital in Stevenage, United Kingdom. Patients who were already established on a personalised end of life care plan at the time of admission were excluded. Patients had their admission WS recorded and were followed up for 30 days or until discharge from hospital, whichever occurred first. Repeat admissions of the same patient were considered as a further episode.

Patient information was collated daily. Diagnosis and length of stay were recorded. Sources of data to determine morbidity and mortality included inpatient notes, discharge letters and submissions of death notification to the hospital bereavement office. Summary hospital mortality and death records at 30 days were checked against the bereavement office administrative service. Data were anonymised at the point of collection by the research team.

Patient and Public Involvement

No patient advocacy was sought in the design of this study. Participants have been anonymised and will therefore not be directly informed of the results. The study was designed as to have no interference with the day-to-day patient care and priorities.

Subcategory definition

Patients were classified according to the medical subspecialty managing their discharge diagnosis for subgroup analysis. These were broadly categorised according to their presenting organ specific sites, including respiratory, cardiology, stroke, haematology/ oncology, gastroenterology, trauma, neurology, renal, endocrine and other infections (which excluded respiratory infections). We defined these subcategories to reflect the common breakdown of medical specialities within a general hospital. Cross-speciality patients and patient with multiple diagnoses upon discharge were categorised according to their primary discharge diagnosis. At Lister Hospital, the electronic discharge system uses ICD-9 for coding purposes.

Statistical analysis

The analysis was based on determining if the WS at admission was different for those patients who were alive or dead at 30 days. The WS exhibited approximately normal distribution and hence as a conservative approach a non-parametric Mann-Whitney U test was used to compare the distribution of WS between two groups. The result was further explained using a box plot. Subgroup analyses were also performed using Mann-Whitney U test. Binary logistic regression was used to assess the predictability of WS for 30-day mortality. ROC analysis is normally carried out as a diagnostic test. We calculated the area under the curve (AUC) to see the accuracy of WS for the 30-day mortality/survival.

The second part of the analysis was to assess if there was a relationship between the WS and LHS. Inspection of the median LHS and WS indicated a large degree of variance. To reduce the effect of outliers, the median length of stay was calculated for each of five index WS categories (4-9, 10-14, 15-19, 20-24, 25+) and displayed using a clustered bar chart. LHS is distributed as positively skewed. Therefore, non-parametric Spearman rank correlation was used between LHS and the WS separately for alive and dead patients. Non-linear quadratic regression was used to determine the predictability of LHS using categorical WS. Both logistic and quadratic regression models were built without consideration of other parameters (univariable) as calculation of the WS included all other variables and hence multivariable adjustments were not needed.

P-values of less than 0.05 was considered statistically significant. IBM SPSS Statistics for Windows V 25.0 (Armonk,NY: IBM Corp) was used to conduct statistical analyses.

Results

Patients studied

A total of 834 patients were recruited during the allocated period of whom 460 (55.2%) were female. Of the population 207 (24.8%) were aged 65-74, 198 (23.7%) were aged 75-80 and 429 (51.4%) were aged ≥81. 55 (6.6%) episodes were re-admissions, of which 7 patients had a third re-admission. All patients had a WS calculated on admission. A total of 714 (85.6%) of patients had a categorical pathology. The remaining 120 patients, where no firm diagnosis was made or where the diagnosis was based on symptoms alone (e.g. chest pain), were grouped into a miscellaneous subcategory for analysis. These subcategories showed a wide variation in patient numbers. As would be expected in acute medical admissions, respiratory, cardiac and non-respiratory infections were the leading subcategories, see Table 1. There were 109 deaths (13.1%) recorded at 30 days.

Table 1. The median Waterlow scores at admission for patients who were alive or dead at 30 days.

| Patient Subcategory | 30-day survivors | | 30-day mortality | | | Mann Whitney |
	Median Waterlow score	N	Median Waterlow score	N	Total Number of patients	P value
Respiratory	13	106	16	30	136	<0.001
Stroke	15	66	19.5	12	78	<0.001
Cardiology	10	103	17	11	114	0.016
Non-respiratory infections	13	101	16	9	110	0.018
Trauma	14	39	20	4	43	0.044
Miscellaneous	10	107	12	13	120	0.067
Renal	13	28	19	7	35	0.143
Gastroenterology	10	59	14	5	64	0.352
Endocrine	15	20	17	3	23	0.590
Haematology/ Oncology	11	57	12	15	72	0.627
Neuro	12	39	-	-	39	No deaths
All groups	12	725	16	109	834	<0.001

Relationship between Waterlow Score and survival at 30 days

There was considerable variation in the WS within the 2 groups (those who survived and those deceased at 30 days). Mann-Whitney U test was used to assess the difference in distribution of the WS in terms of 30-day mortality. A significant difference in the distribution of WS (p<0.001) was seen between those who survived (median 12) and those who died (median 16) at 30 days as (see Table 1) demonstrated in Fig 1. On subgroup analysis, a significant difference in the distribution of WS between those who survived and those deceased at 30 days was also found in the following clinical subcategories: respiratory (p<0.001), stroke (p<0.001), cardiology (p<0.016), non-respiratory infections (p<0.018), and trauma (p<0.044), see Table 1.

Logistic regression model (Table 2) showed that the odds of dying within 30-days was increased 3-fold (exp (1.12)) for every 10 unit increase in WS(p<0.001). Cox & Snell pseudo R-square was 0.043, indicating only a limited variation in the 30-day mortality can be explained by the WS.

Table 2. Results of the binary logistic regression model.

Variable	coefficient	Wald	df	P-value	Odds Ratio
Constant	-3.46	133.49	1	<0.001	.03
Waterlow Score	0.11	36.16	1	<0.001	1.12

Dependent variable: 30-day mortality. Cox & Snell R-square=0.043

We also, studied how well the WS can separate patients in terms of 30-day mortality using ROC curve, see Fig 2. The ROC curve demonstrates a relationship with an area under the curve (AUC) of 0.69 (95% CI: 0.64-0.74) which may be considered borderline (between poor and moderate). The actual cut off value for WS depends on the required sensitivity and specificity in the context. An appropriate cut-off value of WS to separate high risk patients (based on our ROC analysis, not presented) may be ≥12, which gives a sensitivity of 73% and a specificity of 54%.

Relationship between Waterlow Score and Length of Stay

There was a wide variation in the WSs in terms of LHS. To reduce the effect of outliers, median LHS for five index WS categories was used. Figure 3 shows that the length of stay linearly increases with an increase of the WS for those patients who were alive at 30-days (median=5, IQR=10; r=0.32, p<0.01). However, the scenario was opposite for the patients who died within 30-days (median=5, IQR=9; r=-0.20, p=0.04), meaning LHS had a negative relationship with the WS.

LHS was not normally distributed and quadratic regression fitted relatively better with LHS as a dependent variable. Therefore, in Table 3 we fitted quadratic regression with WS as a categorical variable (as done in the Figure 3). As expected, it showed that WS had a significant linear relationship with LHS (because the quadratic term was not significant, *p*=0.27) for those patients who were alive at 30-days. LHS increases by 4.7 days (5.1-.41) if WS increases by 5 units. On the other hand, for the patients who died within 30-days, WS had a significant curvilinear (concave) relationship with LHS (linear term was negative and quadratic term was positive and they were both significant). LHS decreases by approximately 9.8 days (-11.45+1.65) with a 5 unit increase in the WS. However, this decrease reduces gradually with the increase of the WS to form a concave shape. This indicates that WS may be used as a predictor of LHS for l patients 65 or over.

Table 3. Results of quadratic regression analyses.

Dependent Variable:	Length of hospital stay						
	Model Summary					Parameter Estimates	
	R Square	F	df1	df2	Model P-value	Constant	Linear coefficient (p-value)	Quadratic coefficient (p-value)
Alive within 30-days	.06	23.67	2	726	<0.001	1.12	5.10 (p=0.01)	-.41 (p=0.27)
Died within 30-days	.14	8.58	2	106	<0.001	26.13	-11.45 (p=0.001)	1.65 (p=0.01)
The independent variable is Waterlow score categories (4-9,10-14,15-19,20-24,25+).							

Discussion

Prognostic scoring tends to be cumbersome and resource intensive when trying to account for a heterogeneous population with multiple variables. Introduction of new tools also require additional resources in the form of training and support. The WS was developed in the 1980s as a focus for education, intervention and resource management in the prevention of pressure ulcers.[6,11,12] It is widely used in the United Kingdom as part of standard nursing practice across all hospital admissions. Thus, an infrastructure is already in place for this assessment and the results are easily accessible for use as a potential prognostic tool. The high compliance rate in our study (100%) confirms previous findings, as expected in a mandatory assessment[10]. We have shown that the WS, a routinely collected tool which is used throughout the UK at the point of hospital admission to assess pressure ulcer risk, is actually a good predictor of both 30-day survival (OR=1.12, AUC=0.69) and length of hospital stay.

Our study captured the heterogeneous nature of acute medical admissions, as illustrated by the patient subcategories (see Table 1). Despite this heterogeneous cohort, our results demonstrate a statistically significant relationship between the WS and 30-day mortality. This was also confirmed by Sampson[13]. Our data confirm that there is a difference in the distribution of WS for those who survived compared to those who died, suggesting a score above 12 (Figure 2) to be associated with a higher mortality risk, although further analysis in wider context would be required to confirm this. WS

categories have a positive correlation with 30-d mortality categories (see Figure 1). The respiratory, stroke, cardiology, non-respiratory infections and trauma subcategories showed a statistically significant difference in distribution of the WS between those who survived and died. This was not seen in other subcategories, which generally had fewer patients. In acute medical admissions, as seen in our study, respiratory, cardiac and septic patients predominate. A further larger prospective study would be necessary to determine if the WS had additional predictive value in some clinical subcategories compared to others.

With regard to LHS, there is clearly a complex relationship between it and the WS. A high WS is both correlated with increased LHS in those who survive past 30-days, and conversely also correlated with a decreased LHS in those who were deceased at 30-days. Although this information has minimal value retrospectively in managing acute admissions, we can prospectively draw the conclusion that a patient with high admission WS will likely be a more resource intensive admission – either from high mortality risk (where acute deterioration and death can be inferred from the shorter LHS) or a protracted stay. The WS is scored includes an assessment of chronic comorbidities, and the poorer baseline in the high WS patient could account for both of these associations.

There are limitations to the WS itself, which was designed to be a tool to aid clinical risk assessment and is not superior to clinical judgement.[14] In its original role in predicting the development of pressure ulcers, the WS is highly sensitive (82.4%) but has a low specificity (27.4%) which was confirmed in subsequent studies.[12,15] Despite its widespread use, inter observer variability in score calculation has been confirmed in several studies, relating to the assessor's clinical experience and training.[16,17,18]

The utility of the WS has been explored in various other cohorts. Thorn et al[10] looked at in patient mortality in a high-risk surgical group, including elective and emergency admissions, and found a high WS was a predictor of mortality. A positive correlation has been shown between surgical patients with WS above 15 and mortality[19]. It also correlated with rates of post-operative infections in a study in patients with fractured neck of femur[20] and in predicting outcomes in patients with acute pancreatitis, with an AUC score of 0.73 for mortality[21]. Furthermore, it has specifically been found to be an independent predictor of mortality among acute medical patients suffering from dementia[22]. Our study population was limited to patients aged 65 years and over and thus cannot be extrapolated to a younger population. However, acute hospital attendances today among the elderly are increasing[23], 50% of our patients were aged over 81 years. Further studies would be required to confirm our findings in other acutely admitted patient populations. This study also took place at only one centre, and therefore more work needs to be done to validate the findings in other hospitals before it can be generalised nationally.

Conclusion

The findings suggest that the admission WS could have further potential use beyond its role as a predictor of risk of pressure ulcer development. This study shows that it could have an additional role as a predictor of 30-day mortality and length of stay in older, acutely admitted medical patients. The findings support further study to test this hypothesis.

Figure 1. The box plot for median Waterlow scores for patients who were alive or dead at 30 days post admission

Figure 2. ROC curve analysis for Waterlow scores in terms of 30-day Mortality (AUC 0.69, CI 95% 0.64-0.74)

Figure 3. Clustered bar chart for median length of stay for all patients within each category of Waterlow score

Contributorship Statement

All authors provided input in all aspects of the study. There were additional dedicated contributions beyond this as follows:

JW - Principle Investigator and lead author. Design and execution of study, analysis and final write-up. TH & ML - Co-investigators, contributed to design, analysis and final write up. AO, CB, SE - Co-investigators, data collection and analysis. CCT, SS & PS - Statistical analysis and final write-up.

Competing Interests
None declared.

Funding
No funding was applied for this study.

Data Sharing Statement
De-identified participant data is available in the form of excel database upon request from the lead author for a year following publication for use in collaborative research.

Ethics Approval
Ethical approval for the study was given by the NRES Committee Yorkshire &The Humber – Bradford Leeds, REC reference: 14/YH/0074.

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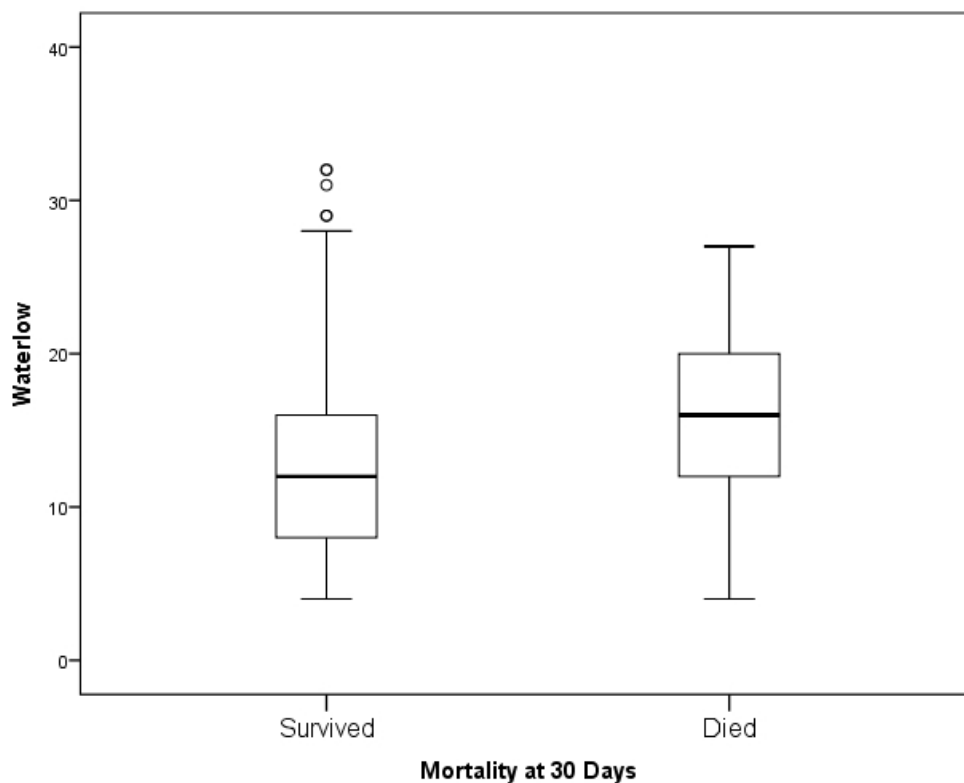16
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Figure 1. The box plot for median Waterlow scores for patients who were alive or dead at 30 days post admission

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Figure 2. ROC curve analysis for Waterlow scores in terms of 30-day Mortality (AUC 0.69, CI 95% 0.64-0.74)

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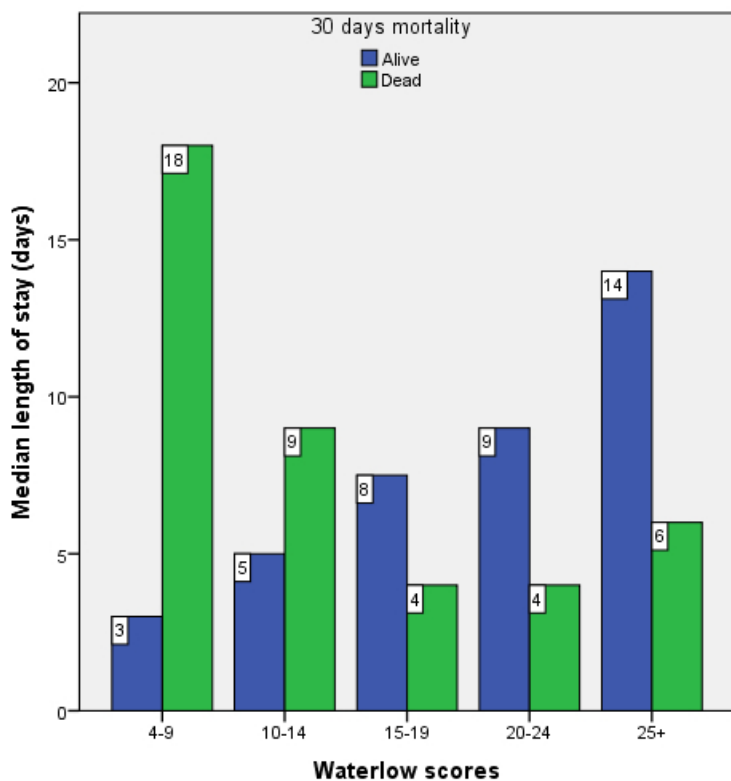18
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Figure 3. Clustered bar chart for median length of stay for all patients within each category of Waterlow score

STROBE Statement—Checklist of items that should be included in reports of *cohort studies*

	Item No	Recommendation	Page No
Title and abstract	1	(*a*) Indicate the study's design with a commonly used term in the title or the abstract	1
		(*b*) Provide in the abstract an informative and balanced summary of what was done and what was found	
Introduction			
Background/rationale	2	Explain the scientific background and rationale for the investigation being reported	2
Objectives	3	State specific objectives, including any prespecified hypotheses	2
Methods			
Study design	4	Present key elements of study design early in the paper	2
Setting	5	Describe the setting, locations, and relevant dates, including periods of recruitment, exposure, follow-up, and data collection	2
Participants	6	(*a*) Give the eligibility criteria, and the sources and methods of selection of participants. Describe methods of follow-up	2
		(*b*) For matched studies, give matching criteria and number of exposed and unexposed	2
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders, and effect modifiers. Give diagnostic criteria, if applicable	2
Data sources/ measurement	8*	For each variable of interest, give sources of data and details of methods of assessment (measurement). Describe comparability of assessment methods if there is more than one group	2
Bias	9	Describe any efforts to address potential sources of bias	2, 3
Study size	10	Explain how the study size was arrived at	2
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If applicable, describe which groupings were chosen and why	3
Statistical methods	12	(*a*) Describe all statistical methods, including those used to control for confounding	3
		(*b*) Describe any methods used to examine subgroups and interactions	3
		(*c*) Explain how missing data were addressed	
		(*d*) If applicable, explain how loss to follow-up was addressed	
		(*e*) Describe any sensitivity analyses	
Results			
Participants	13*	(a) Report numbers of individuals at each stage of study—eg numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed	3
		(b) Give reasons for non-participation at each stage	
		(c) Consider use of a flow diagram	
Descriptive data	14*	(a) Give characteristics of study participants (eg demographic, clinical, social) and information on exposures and potential confounders	
		(b) Indicate number of participants with missing data for each variable of interest	3, 4
		(c) Summarise follow-up time (eg, average and total amount)	
Outcome data	15*	Report numbers of outcome events or summary measures over time	4

Main results	16	(*a*) Give unadjusted estimates and, if applicable, confounder-adjusted estimates and their precision (eg, 95% confidence interval). Make clear which confounders were adjusted for and why they were included	4, 5
		(*b*) Report category boundaries when continuous variables were categorized	
		(*c*) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period	
Other analyses	17	Report other analyses done—eg analyses of subgroups and interactions, and sensitivity analyses	5

Discussion

Key results	18	Summarise key results with reference to study objectives	5
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias	6
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations, multiplicity of analyses, results from similar studies, and other relevant evidence	6
Generalisability	21	Discuss the generalisability (external validity) of the study results	6

Other information

| Funding | 22 | Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on which the present article is based | 7 |

*Give information separately for exposed and unexposed groups.

Note: An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at http://www.plosmedicine.org/, Annals of Internal Medicine at http://www.annals.org/, and Epidemiology at http://www.epidem.com/). Information on the STROBE Initiative is available at http://www.strobe-statement.org.

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Can Waterlow score predict 30-day mortality and length of stay in acutely admitted medical patients (aged ≥65 years)? Evidence from a single centre prospective cohort study

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Can Waterlow score predict 30-day mortality and length of stay in acutely admitted medical patients (aged ≥65 years)? Evidence from a single centre prospective cohort study

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Key Words:
Waterlow Score
30-Day Mortality
Length of Stay

Word Count:
2877

Abstract

Objective: To explore the potential for the Waterlow Score (WS) to be used as a predictor of 30-day mortality and length of hospital stay (LHS) in acutely admitted medical patients aged 65 years and over.

Design: Prospective observational cohort study.

Setting: United Kingdom District General Hospital.

Subjects: 834 consecutive patients aged 65 years and over admitted acutely to medical specialities between 30 May and 22 July 2014.

Methods: Admission WS (range 4-64) assessment paired with the patient's status at 30-days in terms of mortality and their length of hospital stay.

Primary Outcomes: 30-day mortality and length of inpatient stay.

Results: 834 consecutive acute medical admissions had their WS recorded. 30-day mortality was 13.1% (109 deaths). A significant difference in distribution of WS (p<0.001) was seen between those who survived (median 12) and those who died (median 16) within 30 days, particularly within respiratory (p<0.001), stroke (p<0.001), cardiology (p<0.016), non-respiratory infections (p<0.018) and trauma (p<0.044) subgroups. Odds of dying within 30 days increased 3-fold for every 10-unit increase in the WS (p<0.001, 95% CI:2.1-4.3). LHS was also positively linearly associated with the WS in those who survived 30 days (median=5, IQR=10; r= 0.32, p<0.01). A 5-unit increase in WS was associated with approximately 5 days increase in LHS. On the other hand, quadratic regression showed this relationship was curvilinear and negative (concave) for those who died within 30 days where a 5-unit increase in WS was associated with an approximately 10 days decrease in LHS.

Conclusion: This study demonstrates an association between a high WS and both 30-day mortality and LHS**.** This is particularly significant for mortality in patients in the respiratory, stroke and cardiac subcategories. The WS, a nursing led screening tool that is carried out on virtually all admissions to UK hospitals, could have additional use at the time of patient admission as a risk assessment tool for 30-day mortality as well as a predictor of length of hospital stay.

Key words: Waterlow Score, 30-day mortality, length of stay

Strengths and Limitations of this study

- This prospective cohort study examines the additional utility of the Waterlow Score, a widely used tool that is already a mandatory assessment for all patients admitted to hospital in the United Kingdom.
- This study uses logistic regression and non-linear quadratic regression models for statistical prediction.
- The study captures the heterogenous nature of patients admitted acutely to hospital, as illustrated by the patient subcategories.
- This is a study of the potential value of the Waterlow Score in evaluating risk of death and predicting length of hospital stay in patients over 65 years of age.
- This study was based on a single district general hospital, and therefore requires further research in a wider context to confirm our observations.

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Introduction

Predictive mortality and morbidity scores are widely used to aid clinical decision making and resource allocation. Hospitals routinely collect a range of data on patient admission, however this is rarely standardised between hospitals.[1] A range of validated risk assessment models are commonly used in everyday clinical settings, usually with reference to a specific pathology e.g. the Mini-Grace risk score[2] for suspected acute coronary syndromes. These specific assessment models cannot be extrapolated beyond their validated clinical settings because of their disease specificity. The use of more general prognostic models is limited by their need to consider a broad range of clinical and demographic information due to the heterogeneous nature of patient populations. Broader prognostic models have been developed such as the APACHE III score[3], developed as a gatekeeper to Intensive Care, and frailty indices[4,5] are used in the geriatric setting. However, none are simple or practical and are limited by complexity and time constraints.

The Waterlow Score[6] (WS) is a multi-dimensional pressure ulcer risk assessment tool encompassing functional status and co-morbidity which was developed in the 1980's. It is a nursing led screening tool that is carried out on virtually every patient on admission to a United Kingdom hospital. Its components are age, nutritional status, weight, mobility, gender, smoking status, co-morbidities, use of medication and continence, all combined to generate a single numerical figure between 2 and 64. Patients with scores over 10 are considered to be at risk of pressure ulcer, whilst scores of over 15 and 20 are considered to represent high risk and very high risk statuses, respectively. [6] Other notable pressure ulcer risk assessments exist[7], but the WS remains the most widely used in the United Kingdom.

The primary aim of this study is to look at the potential role of the WS as a predictor of 30-day mortality in acute medical admissions aged 65 years and over and secondly to look at their length of hospital stay (LHS) as a marker of morbidity burden. The application of the WS is appropriate because its criteria already take into account those of other comorbidity models such as the Charlson comorbidity index.[8] The latter, a widely used comorbidity measure, is a component of the dataset used in standardised hospital mortality indicators in the United Kingdom.[9] Deriving additional utility from the WS could aid patient risk stratification, potentially improving patient care and guiding resource management. An association between the WS and outcomes has already been demonstrated in a cohort of surgical patients.[10] Furthermore, identifying additional applications of the Waterlow score would not incur any additional time resource requirements.

Methods

Study population
Following approval from a research ethics committee, the study group prospectively recruited all consecutive acute medical patients aged 65 years and over admitted during an eight-week period between 30 May and 22 July 2014 to the Lister Hospital, a 720 bed District General Hospital in Stevenage, United Kingdom. Patients who were already established on a personalised end of life care plan at the time of admission were excluded. Patients had their admission WS recorded and were followed up for 30 days or until discharge from hospital, whichever occurred first. Repeat admissions of the same patient were considered as a further episode.

Patient information was collated daily. Diagnosis and length of stay were recorded. Sources of data to determine morbidity and mortality included inpatient notes, discharge letters and submissions of death notification to the hospital bereavement office. Summary hospital mortality and death records at 30 days were checked against the bereavement office administrative service. Data were anonymised at the point of collection by the research team.

Patient and Public Involvement

No patient advocacy was sought in the design of this study. Participants have been anonymised and will therefore not be directly informed of the results. The study was designed as to have no interference with the day-to-day patient care and priorities.

Subcategory definition

Patients were classified according to the medical subspecialty managing their discharge diagnosis for subgroup analysis. These were broadly categorised according to their presenting organ specific sites, including respiratory, cardiology, stroke, haematology/ oncology, gastroenterology, trauma, neurology, renal, endocrine and other infections (which excluded respiratory infections). We defined these subcategories to reflect the common breakdown of medical specialities within a general hospital. Cross-speciality patients and patient with multiple diagnoses upon discharge were categorised according to their primary discharge diagnosis. At Lister Hospital, the electronic discharge system uses ICD-9 for coding purposes.

Statistical analysis

The analysis was based on determining if the WS at admission was different for those patients who were alive or dead at 30 days. The WS exhibited approximately normal distribution and hence as a conservative approach a non-parametric Mann-Whitney U test was used to compare the distribution of WS between two groups. The result was further explained using a box plot. Subgroup analyses were also performed using Mann-Whitney U test. Binary logistic regression was used to assess the predictability of WS for 30-day mortality. ROC analysis is normally carried out as a diagnostic test. We calculated the area under the curve (AUC) to see the accuracy of WS for the 30-day mortality/survival.

The second part of the analysis was to assess if there was a relationship between the WS and LHS. Inspection of the median LHS and WS indicated a large degree of variance. To reduce the effect of outliers, the median length of stay was calculated for each of five index WS categories (4-9, 10-14, 15-19, 20-24, 25+) and displayed using a clustered bar chart. LHS is distributed as positively skewed. Therefore, non-parametric Spearman rank correlation was used between LHS and the WS separately for alive and dead patients. Non-linear quadratic regression was used to determine the predictability of LHS using categorical WS. Both logistic and quadratic regression models were built without consideration of other parameters (univariable) as calculation of the WS included all other variables and hence multivariable adjustments were not needed.

P-values of less than 0.05 was considered statistically significant. IBM SPSS Statistics for Windows V 25.0 (Armonk,NY: IBM Corp) was used to conduct statistical analyses.

Results

Patients studied

A total of 834 patients were recruited during the allocated period of whom 460 (55.2%) were female. Of the population 207 (24.8%) were aged 65-74, 198 (23.7%) were aged 75-80 and 429 (51.4%) were aged $\geq$81. 55 (6.6%) episodes were re-admissions, of which 7 patients had a third re-admission. All patients had a WS calculated on admission. A total of 714 (85.6%) of patients had a categorical pathology. The remaining 120 patients, where no firm diagnosis was made or where the diagnosis was based on symptoms alone (e.g. chest pain), were grouped into a miscellaneous subcategory for analysis. These subcategories showed a wide variation in patient numbers. As would be expected in acute medical admissions, respiratory, cardiac and non-respiratory infections were the leading subcategories, see Table 1. There were 109 deaths (13.1%) recorded at 30 days.

Table 1. The median Waterlow scores at admission for patients who were alive or dead at 30 days.

Patient Subcategory	30-day survivors		30-day mortality			Mann Whitney
	Median Waterlow score	N	Median Waterlow score	N	Total Number of patients	P value
Respiratory	13	106	16	30	136	<0.001
Stroke	15	66	19.5	12	78	<0.001
Cardiology	10	103	17	11	114	0.016
Non-respiratory infections	13	101	16	9	110	0.018
Trauma	14	39	20	4	43	0.044
Miscellaneous	10	107	12	13	120	0.067
Renal	13	28	19	7	35	0.143
Gastroenterology	10	59	14	5	64	0.352
Endocrine	15	20	17	3	23	0.590
Haematology/ Oncology	11	57	12	15	72	0.627
Neuro	12	39	-	-	39	No deaths
All groups	12	725	16	109	834	<0.001

Relationship between Waterlow Score and survival at 30 days

There was considerable variation in the WS within the 2 groups (those who survived and those deceased at 30 days). Mann-Whitney U test was used to assess the difference in distribution of the WS in terms of 30-day mortality. A significant difference in the distribution of WS (p<0.001) was seen between those who survived (median 12) and those who died (median 16) at 30 days as (see Table 1) demonstrated in Fig 1. On subgroup analysis, a significant difference in the distribution of WS between those who survived and those deceased at 30 days was also found in the following clinical subcategories: respiratory (p<0.001), stroke (p<0.001), cardiology (p<0.016), non-respiratory infections (p<0.018), and trauma (p<0.044), see Table 1.

Logistic regression model (Table 2) showed that the odds of dying within 30-days was increased 3-fold (exp (1.1)) for every 10 unit increase in WS (p<0.001, 95% CI:2.1-4.3). Cox & Snell pseudo R-square was 0.043, indicating only a limited variation in the 30-day mortality can be explained by the WS.

Table 2. Results of the binary logistic regression model.

Variable	coefficient	Wald	df	P-value	Odds Ratio
Constant	-3.46	133.49	1	<0.001	.03
Waterlow Score	0.11	36.16	1	<0.001	1.12

Dependent variable: 30-day mortality. Cox & Snell R-square=0.043

We also, studied how well the WS can separate patients in terms of 30-day mortality using ROC curve, see Fig 2. The ROC curve demonstrates a relationship with an area under the curve (AUC) of 0.69 (95% CI: 0.64-0.74) which may be considered borderline (between poor and moderate). The actual cut off value for WS depends on the required sensitivity and specificity in the context. An appropriate cut-off value of WS to separate high risk patients (based on our ROC analysis, not presented) may be ≥12, which gives a sensitivity of 73% and a specificity of 54%.

Relationship between Waterlow Score and Length of Stay

There was a wide variation in the WSs in terms of LHS. To reduce the effect of outliers, median LHS for five index WS categories was used. Figure 3 shows that the length of stay linearly increases with an increase of the WS for those patients who were alive at 30-days (median=5, IQR=10; r=0.32, p<0.01). However, the scenario was opposite for the patients who died within 30-days (median=5, IQR=9; r=-0.20, p=0.04), meaning LHS had a negative relationship with the WS.

LHS was not normally distributed and quadratic regression fitted relatively better with LHS as a dependent variable. Therefore, in Table 3 we fitted quadratic regression with WS as a categorical variable (as done in the Figure 3). As expected, it showed that WS had a significant linear relationship with LHS (because the quadratic term was not significant, *p*=0.27) for those patients who were alive at 30-days. LHS increases by 4.7 days (5.1-.41) if WS increases by 5 units. On the other hand, for the patients who died within 30-days, WS had a significant curvilinear (concave) relationship with LHS (linear term was negative and quadratic term was positive and they were both significant). LHS decreases by approximately 9.8 days (-11.45+1.65) with a 5 unit increase in the WS. However, this decrease reduces gradually with the increase of the WS to form a concave shape. This indicates that WS may be used as a predictor of LHS for l patients 65 or over.

Table 3. Results of quadratic regression analyses.

Dependent Variable:	Length of hospital stay						
	Model Summary					Parameter Estimates	
	R Square	F	df1	df2	Model P-value	Constant	Linear coefficient (p-value)	Quadratic coefficient (p-value)
Alive within 30-days	.06	23.67	2	726	<0.001	1.12	5.10 (p=0.01)	-.41 (p=0.27)
Died within 30-days	.14	8.58	2	106	<0.001	26.13	-11.45 (p=0.001)	1.65 (p=0.01)

The independent variable is Waterlow score categories (4-9,10-14,15-19,20-24,25+).

Discussion

Prognostic scoring tends to be cumbersome and resource intensive when trying to account for a heterogeneous population with multiple variables. Introduction of new tools also require additional resources in the form of training and support. The WS was developed in the 1980s as a focus for education, intervention and resource management in the prevention of pressure ulcers.[6,11,12] It is widely used in the United Kingdom as part of standard nursing practice across all hospital admissions. Thus, an infrastructure is already in place for this assessment and the results are easily accessible for use as a potential prognostic tool. The high compliance rate in our study (100%) confirms previous findings, as expected in a mandatory assessment[10]. We have shown that the WS, a routinely collected tool which is used throughout the UK at the point of hospital admission to assess pressure ulcer risk, is actually a good predictor of both 30-day survival (OR=1.12, AUC=0.69) and length of hospital stay.

Our study captured the heterogeneous nature of acute medical admissions, as illustrated by the patient subcategories (see Table 1). Despite this heterogeneous cohort, our results demonstrate a statistically significant relationship between the WS and 30-day mortality. This was also confirmed by Sampson[13]. Our data confirm that there is a difference in the distribution of WS for those who survived compared to those who died, suggesting a score above 12 (Figure 2) to be associated with a higher

mortality risk, although further analysis in wider context would be required to confirm this. WS categories have a positive correlation with 30-d mortality categories (see Figure 1). The respiratory, stroke, cardiology, non-respiratory infections and trauma subcategories showed a statistically significant difference in distribution of the WS between those who survived and died. This was not seen in other subcategories, which generally had fewer patients. In acute medical admissions, as seen in our study, respiratory, cardiac and septic patients predominate. A further larger prospective study would be necessary to determine if the WS had additional predictive value in some clinical subcategories compared to others.

With regard to LHS, there is clearly a complex relationship between it and the WS. A high WS is both correlated with increased LHS in those who survive past 30-days, and conversely also correlated with a decreased LHS in those who were deceased at 30-days. Although this information has minimal value retrospectively in managing acute admissions, we can prospectively draw the conclusion that a patient with high admission WS will likely be a more resource intensive admission – either from high mortality risk (where acute deterioration and death can be inferred from the shorter LHS) or a protracted stay. The WS is scored includes an assessment of chronic comorbidities, and the poorer baseline in the high WS patient could account for both of these associations.

There are limitations to the WS itself, which was designed to be a tool to aid clinical risk assessment and is not superior to clinical judgement.[14] In its original role in predicting the development of pressure ulcers, the WS is highly sensitive (82.4%) but has a low specificity (27.4%) which was confirmed in subsequent studies.[12,15] Despite its widespread use, inter observer variability in score calculation has been confirmed in several studies, relating to the assessor's clinical experience and training.[16,17,18]

The utility of the WS has been explored in various other cohorts. Thorn et al[10] looked at in patient mortality in a high-risk surgical group, including elective and emergency admissions, and found a high WS was a predictor of mortality. A positive correlation has been shown between surgical patients with WS above 15 and mortality[19]. It also correlated with rates of post-operative infections in a study in patients with fractured neck of femur[20] and in predicting outcomes in patients with acute pancreatitis, with an AUC score of 0.73 for mortality[21]. Furthermore, it has specifically been found to be an independent predictor of mortality among acute medical patients suffering from dementia[22]. Our study population was limited to patients aged 65 years and over and thus cannot be extrapolated to a younger population. However, acute hospital attendances today among the elderly are increasing[23], 50% of our patients were aged over 81 years. Further studies would be required to confirm our findings in other acutely admitted patient populations. This study also took place at only one centre, and therefore more work needs to be done to validate the findings in other hospitals before it can be generalised nationally.

Conclusion
The findings suggest that the admission WS could have further potential use beyond its role as a predictor of risk of pressure ulcer development. This study shows that it could have an additional role as a predictor of 30-day mortality and length of stay in older, acutely admitted medical patients. The findings support further study to test this hypothesis.

Figure 1. The box plot for median Waterlow scores for patients who were alive or dead at 30 days post admission

Figure 2. ROC curve analysis for Waterlow scores in terms of 30-day Mortality (AUC 0.69, CI 95% 0.64-0.74)

Figure 3. Clustered bar chart for median length of stay for all patients within each category of Waterlow score

Contributorship Statement
All authors provided input in all aspects of the study. There were additional dedicated contributions beyond this as follows:
JW - Principle Investigator and lead author. Design and execution of study, analysis and final write-up. TH & ML - Co-investigators, contributed to design, analysis and final write up. AO, CB, SE - Co-investigators, data collection and analysis. CCT, SS & PS - Statistical analysis and final write-up.

Competing Interests
None declared.

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No funding was applied for this study.

Data Sharing Statement
Extra data can be accessed via the Dryad data repository at http://datadryad.org/ with the doi: 10.5061/dryad.8931zcrkp

Ethics Approval
Ethical approval for the study was given by the NRES Committee Yorkshire &The Humber – Bradford Leeds, REC reference: 14/YH/0074.

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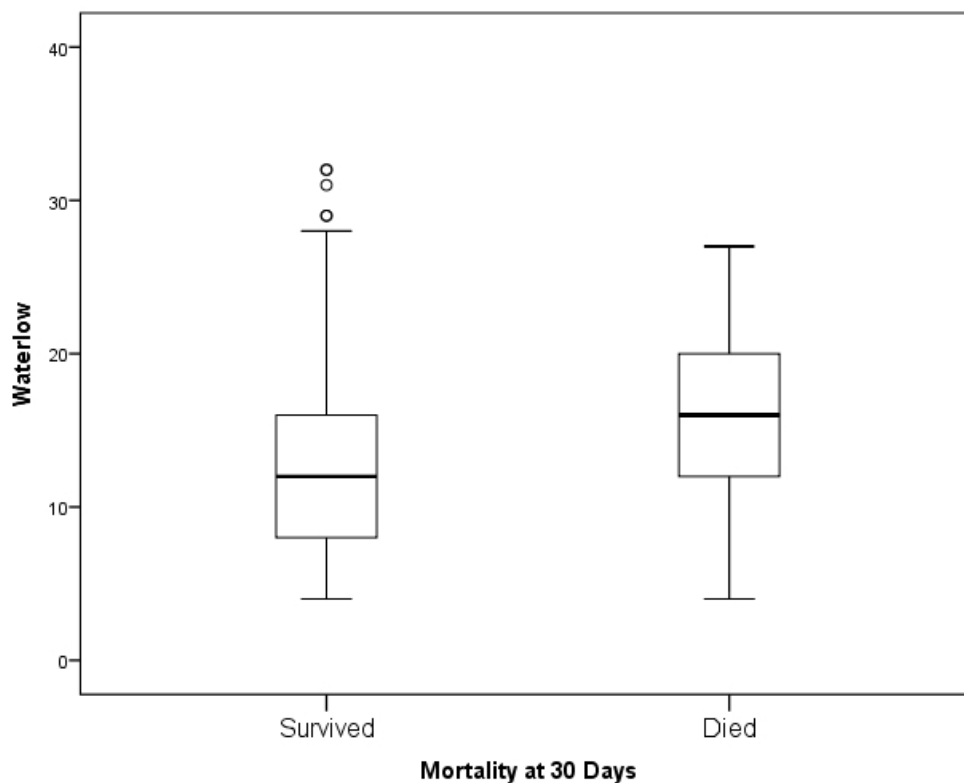47
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Figure 1. The box plot for median Waterlow scores for patients who were alive or dead at 30 days post admission

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Figure 2. ROC curve analysis for Waterlow scores in terms of 30-day Mortality (AUC 0.69, CI 95% 0.64-0.74)

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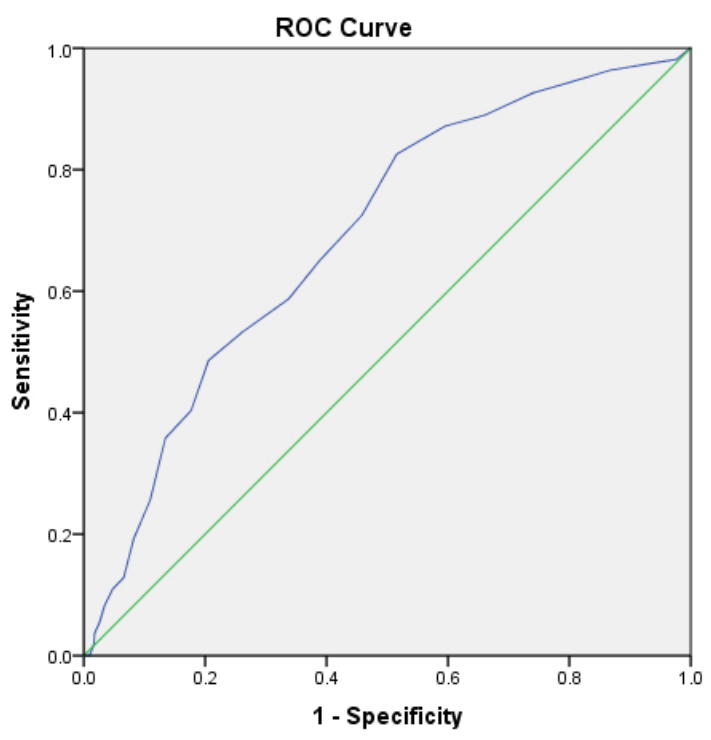41
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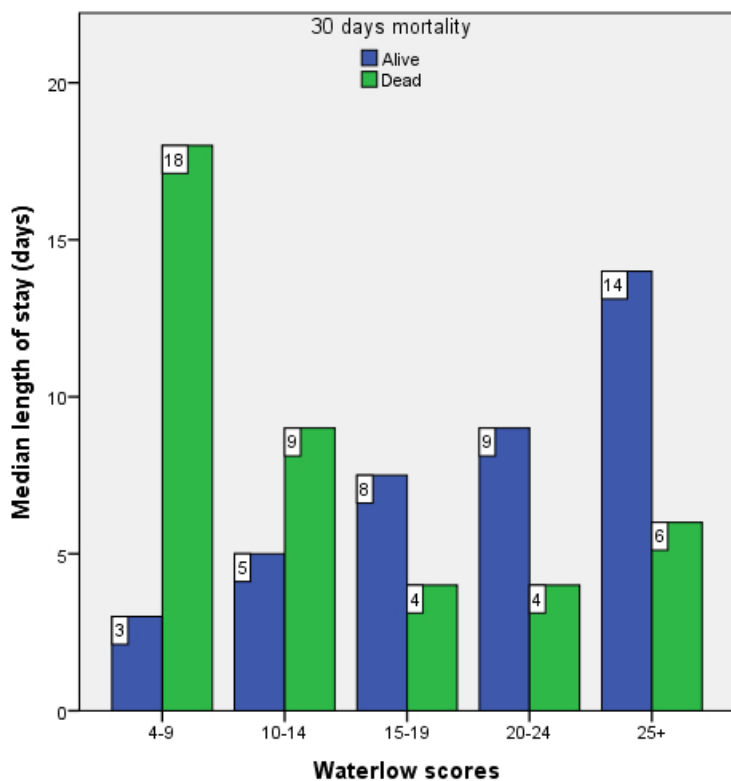47
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Figure 3. Clustered bar chart for median length of stay for all patients within each category of Waterlow score

STROBE Statement—Checklist of items that should be included in reports of *cohort studies*

	Item No	Recommendation	Page No
Title and abstract	1	(*a*) Indicate the study's design with a commonly used term in the title or the abstract (*b*) Provide in the abstract an informative and balanced summary of what was done and what was found	1
Introduction			
Background/rationale	2	Explain the scientific background and rationale for the investigation being reported	2
Objectives	3	State specific objectives, including any prespecified hypotheses	2
Methods			
Study design	4	Present key elements of study design early in the paper	2
Setting	5	Describe the setting, locations, and relevant dates, including periods of recruitment, exposure, follow-up, and data collection	2
Participants	6	(*a*) Give the eligibility criteria, and the sources and methods of selection of participants. Describe methods of follow-up (*b*) For matched studies, give matching criteria and number of exposed and unexposed	2 2
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders, and effect modifiers. Give diagnostic criteria, if applicable	2
Data sources/ measurement	8*	For each variable of interest, give sources of data and details of methods of assessment (measurement). Describe comparability of assessment methods if there is more than one group	2
Bias	9	Describe any efforts to address potential sources of bias	2, 3
Study size	10	Explain how the study size was arrived at	2
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If applicable, describe which groupings were chosen and why	3
Statistical methods	12	(*a*) Describe all statistical methods, including those used to control for confounding (*b*) Describe any methods used to examine subgroups and interactions (*c*) Explain how missing data were addressed (*d*) If applicable, explain how loss to follow-up was addressed (*e*) Describe any sensitivity analyses	3 3
Results			
Participants	13*	(a) Report numbers of individuals at each stage of study—eg numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed (b) Give reasons for non-participation at each stage (c) Consider use of a flow diagram	3
Descriptive data	14*	(a) Give characteristics of study participants (eg demographic, clinical, social) and information on exposures and potential confounders (b) Indicate number of participants with missing data for each variable of interest (c) Summarise follow-up time (eg, average and total amount)	 3, 4
Outcome data	15*	Report numbers of outcome events or summary measures over time	4

Main results	16	(*a*) Give unadjusted estimates and, if applicable, confounder-adjusted estimates and their precision (eg, 95% confidence interval). Make clear which confounders were adjusted for and why they were included	4, 5
		(*b*) Report category boundaries when continuous variables were categorized	
		(*c*) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period	
Other analyses	17	Report other analyses done—eg analyses of subgroups and interactions, and sensitivity analyses	5

Discussion

Key results	18	Summarise key results with reference to study objectives	5
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias	6
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations, multiplicity of analyses, results from similar studies, and other relevant evidence	6
Generalisability	21	Discuss the generalisability (external validity) of the study results	6

Other information

| Funding | 22 | Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on which the present article is based | 7 |

*Give information separately for exposed and unexposed groups.

Note: An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at http://www.plosmedicine.org/, Annals of Internal Medicine at http://www.annals.org/, and Epidemiology at http://www.epidem.com/). Information on the STROBE Initiative is available at http://www.strobe-statement.org.