

## Supplementary Online Content

Horng S, Joseph JW, Calder S, et al. Assessment of unintentional duplicate orders by emergency department clinicians before and after implementation of a visual aid in the electronic health record ordering system. *JAMA Netw Open*. 2019;2(12):e1916499. doi:10.1001/jamanetworkopen.2019.16499

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This supplementary material has been provided by the authors to give readers additional information about their work.

## **eMethods.** Statistical Methods

We modeled duplicate orders for each of the three order types with the following negative binomial regression:

$$\ln(\mu) = \beta_0 + \beta_1 Post_t + \beta_2(Post_t * T_t) + \beta_3 T_t + \beta_4 X_t$$

where  $\mu$  represents the expected number of duplicate orders of its type. The natural log of the total number of orders were included as an offset term.  $Post_t$  is an indicator variable that equals 1 after the intervention date for that order type and 0 before.  $T_t$  is the number of shifts at time  $t$  since the beginning of the sample for that order type, where the sample begins 1 year prior to the intervention date.  $X_t$  is a vector of covariates at time  $t$  at the shift-level, including average patient age, gender composition of patients, whether patients were native English speakers, percent of patients with an emergency severity index 3 or greater, percent of patients discharged, percent of patients admitted, number of patients in the department, number of patients in the waiting room, number of observation patients, number of new patients, number of ICU beds requested, number of telemetry beds requested, number of boarders, length of stay in the emergency department. Full sets of indicator variables for shift, day-of-week, and month-of-sample are included to adjust for how duplicate orders may follow a cyclical pattern throughout the day, week, and year, respectively. We estimate Newey-West standard errors, which are robust to autocorrelation and heteroskedasticity. All tests used a significance level of 0.05.

Coefficients are transformed into incidence rate ratios (IRRs) in Table 2, such that  $e^{\beta_1}$  estimates the immediate change in the incidence of duplicate orders after implementation,  $e^{\beta_2}$  estimates the gradual change in the incidence of duplicate orders over time after implementation,  $e^{\beta_3}$  estimates the pre-intervention trend in incidence over time.

We initially used a Poisson regression to model counts. However, a likelihood ratio test of  $\alpha = 0$  generated a large test statistic of 2734.83 with an associated p-value of  $< 0.001$ . This suggests that our duplicate order response variables are over-dispersed and therefore we used the more general negative binomial model.

Because we estimated three separate models for each of the types of duplicate orders, the coefficients on the covariates are estimated separately. One may be concerned that the sign and magnitude of these covariates should not be different across the three models. For instance, the association between average age and the probability of duplicate orders should likely be the same across the three types of orders as age should not be associated with the probability of a laboratory duplicate order differently than a medication or radiology duplicate order. Indeed, when testing the equality of the coefficients on the covariates across the three models, we find that they are not statistically different from each other, with the only exception being number of patients in the waitroom. Thus, we re-estimated our models with the imposed constraint that the estimated coefficients on the covariates that were not statistically different from each other are equal across models. This allows us to estimate the association between the intervention and unintentional duplicate orders of each of the order types while jointly estimating the coefficients on the covariates overall. In this alternative specification, our estimates do not change in magnitude, direction, or significance.

Stata/SE 14.2 software was used for statistical analysis.

**eTable 1. Interrupted Time Series Analysis, Full Covariate List**

Variable	Laboratory	Medications	Radiology
	IRR <sup>a</sup> (95% CI <sup>b</sup> )	IRR <sup>a</sup> (95% CI <sup>b</sup> )	IRR <sup>a</sup> (95% CI <sup>b</sup> )
Level Change <sup>d</sup>	0.51 <sup>c</sup> (0.45-0.59)	1.17 (0.52-2.61)	0.61 <sup>c</sup> (0.46-0.82)
Change in Trend <sup>e</sup>	1.00 <sup>c</sup> (1.00-1.00)	1.00 <sup>c</sup> (1.00-1.00)	1.00 <sup>c</sup> (1.00-1.00)
Pre-Intervention Trend <sup>f</sup>	1.00 (1.00-1.00)	1.00 (1.00-1.00)	1.00 (1.00-1.00)
Age	1.00 (0.99-1.01)	0.99 (0.98-1.00)	1.00 (0.99-1.01)
Gender (Female)	1.00 (1.00-1.01)	1.00 (0.99-1.02)	1.00 (0.99-1.00)
Emergency Severity Index >= 3	1.00 (1.00-1.00)	0.99 <sup>c</sup> (0.99-0.99)	1.01 <sup>c</sup> (1.00-1.01)
Primary Language (English)	1.00 (0.99-1.01)	1.00 (0.98-1.03)	0.99 (0.98-1.00)
Patients Discharged Home	1.01 (0.99-1.03)	1.00 (0.99-1.01)	1.01 (1.00-1.02)
Patients Admitted	1.01 (0.99-1.02)	1.00 (0.99-1.02)	1.01 <sup>c</sup> (1.00-1.01)
No. Patients in Department	1.00 (0.98-1.03)	0.99 (0.97-1.01)	1.00 (0.99-1.00)
No. Patients in Waitroom	0.99 (0.98-1.00)	1.05 <sup>c</sup> (1.04-1.07)	1.01 <sup>c</sup> (1.00-1.02)
No. Patients in Observation	1.01 (0.99-1.02)	1.00 (0.95-1.05)	1.02 (1.00-1.04)
No. New Patients	0.97 <sup>c</sup> (0.96-0.99)	0.83 <sup>c</sup> (0.79-0.88)	0.99 (0.94-1.03)
No. Patients Admitted to ICU	0.95 <sup>c</sup> (0.90-0.99)	1.02 (0.93-1.11)	1.04 (1.00-1.08)
No. Patients Admitted to Telemetry	0.99 (0.95-1.04)	1.02 (1.00-1.04)	1.01 (0.98-1.03)
No. Patients Admitted to Floor	0.97 <sup>c</sup> (0.96-0.99)	1.01 (0.96-1.07)	1.03 <sup>c</sup> (1.01-1.05)
No. Patients Boarding	1.05 <sup>c</sup> (1.03-1.08)	0.99 (0.97-1.02)	1.01 (0.99-1.02)
Length of Stay	1.00 (1.00-1.00)	1.00 (1.00-1.00)	1.00 (1.00-1.00)

Interrupted time series analysis was performed using a negative binomial regression with Newey-West standard errors at the shift level. All models contain a full set of indicator variables for time of day at the shift level, day-of-week, and month-of-sample.

<sup>a</sup> IRR = Incidence Rate Ratios

<sup>b</sup> 95% CI = 95% Confidence Intervals

<sup>c</sup> Denotes p-value < 0.05

<sup>d</sup> Estimates the immediate change in the incidence of duplicate orders after implementation.

<sup>e</sup> Estimates the gradual change in the incidence of duplicate orders over time after implementation.

<sup>f</sup> Estimates the pre-intervention trend in incidence over time.

**eTable 2. Top 5 Unintentional Duplicate Orders**

	<b>Laboratory Duplicate Orders</b>			
	<b><u>BEFORE</u></b>		<b><u>AFTER</u></b>	
<b>Laboratory Test</b>	Rank	Count (%) <sup>*</sup>	Rank	Count (%) <sup>a</sup>
Chem-7	1	726 (16%)	1	361 (13%)
CBC/DIF	2	615 (13%)	2	269 (10%)
Coags	3	239 (5%)	4	171 (6%)
Urinalysis	4	197 (4%)	11	76 (3%)
Lactate	5	194 (4%)	6	124 (5%)
	...		...	
	<b>Total</b>	<b>4560</b>	<b>Total</b>	<b>2729</b>
	<b>Medication Duplicate Orders</b>			
	<b><u>BEFORE</u></b>		<b><u>AFTER</u></b>	
<b>Medication Name</b>	Rank	Count (%) <sup>*</sup>	Rank	Count (%) <sup>a</sup>
Hydromorphone (IV)	1	53 (14%)	1	58 (12%)
Diazepam (PO)	2	33 (9%)	2	42 (8%)
Morphine (IV)	3	26 (7%)	3	29 (6%)
Lorazepam (IV)	4	11 (3%)	9	11 (2%)
Aspirin (PO)	5	8 (2%)	33	3 (1%)
	...		...	
	<b>Total</b>	<b>380</b>	<b>Total</b>	<b>503</b>
	<b>Radiology Duplicate Orders</b>			
	<b><u>BEFORE</u></b>		<b><u>AFTER</u></b>	

<b>Radiology Test</b>	<b>Rank</b>	<b>Count (%)<sup>*</sup></b>	<b>Rank</b>	<b>Count (%)<sup>a</sup></b>
Chest X-Ray (PA & Lat)	1	149 (15%)	2	101 (13%)
CT Head w/o Contrast	2	118 (12%)	3	86 (11%)
CT Abdomen & Pelvis w/ Contrast	3	113 (12%)	1	117 (15%)
CT C-Spine w/o Contrast	4	45 (5%)	7	26 (3%)
Chest X-Ray (Portable)	5	43 (4%)	4	34 (4%)
	...		...	
	<b>Total</b>	<b>962</b>	<b>Total</b>	<b>791</b>

<sup>a</sup> Count (%) should only be compared vertically, not horizontally, as count (%) are not normalized for differences in the number of patients and total orders in each period. Please use the ranks to compare horizontally. Please refer to *Table 1* and *Table 2* for normalized results to compare between periods.

## eFigure 1. Visual Aid as Suggestion for Clinical Decision Rules

### CHADS2-VASc

clinical risk stratification schemata for predicting stroke and thromboembolism (TE) in patients with atrial fibrillation (AF)

<input checked="" type="checkbox"/> Age 65-74 y	+1 points
<input type="checkbox"/> Age $\geq$ 75 y	+2 points
<input checked="" type="checkbox"/> Female	+1 points
<input checked="" type="checkbox"/> Diabetes mellitus	+1 points
<input checked="" type="checkbox"/> Hypertension	+1 points
<input type="checkbox"/> Congestive heart failure/LV dysfunction	+1 points
<input type="checkbox"/> Prior Stroke/TIA/TE	+2 points
<input type="checkbox"/> Vascular disease <sup>i</sup>	+1 points

**Age:** 69 y/o

**Gender:** F

**HTN:** losartan-hydrochlorothiazide metoprolol tartrate

**DM:** Januvia glipizide metformin

Total Score 4. (4) 1.9% Rate of thromboembolism in one year -- Recommend warfarin



## eFigure 2. Visual Aid for Allergy Decision Support

### - Antibiotics - Parenteral

- Azithromycin 500 mg IV **Allergy**
- CeFAZolin 1 g IV **Allergy**
- CefePIME 2 g IV **Allergy**
- CefTRIAxone 1 g IV **Allergy**
- CefTRIAxone 2 g IV **Allergy**
- CefTRIAxone 250 mg IM **Allergy**
- Ciprofloxacin 400 mg IV **Allergy**
- Clindamycin 600 mg IV
- Clindamycin 900 mg IV
- Flagyl (MetRONIDAZOLE) 500 mg IV
- Levofloxacin 500 mg IV **Allergy**
- Levofloxacin 750 mg IV **Allergy**
- Zosyn 4.5 g IV **Allergy**
- Unasyn 3 g IV **Allergy**
- Vancomycin 1000 mg IV ONCE
- Vancomycin 1500 mg IV ONCE