

Supplementary Appendix

This appendix has been provided by the authors to give readers additional information about their work.

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A clinical trial of three anticonvulsant medications for status epilepticus

The Established Status Epilepticus Treatment Trial (ESETT)

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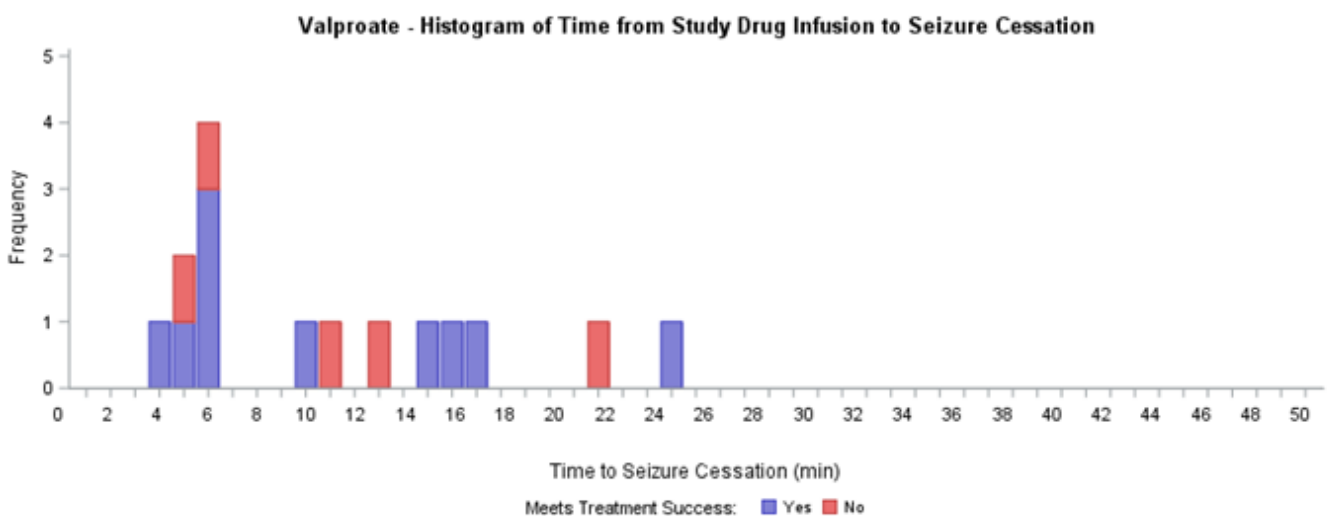
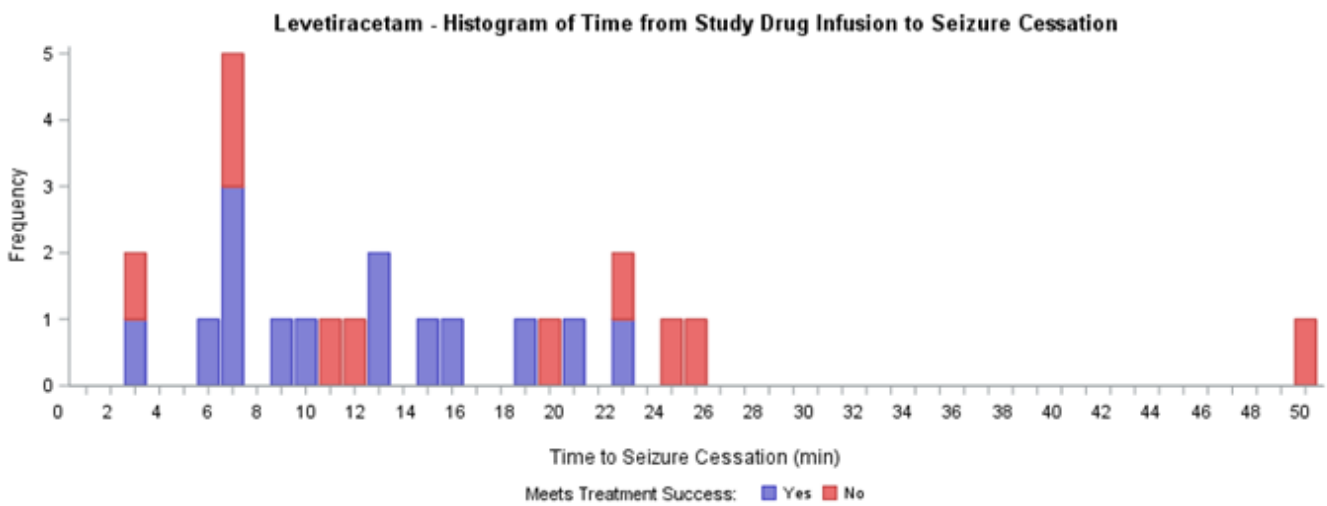
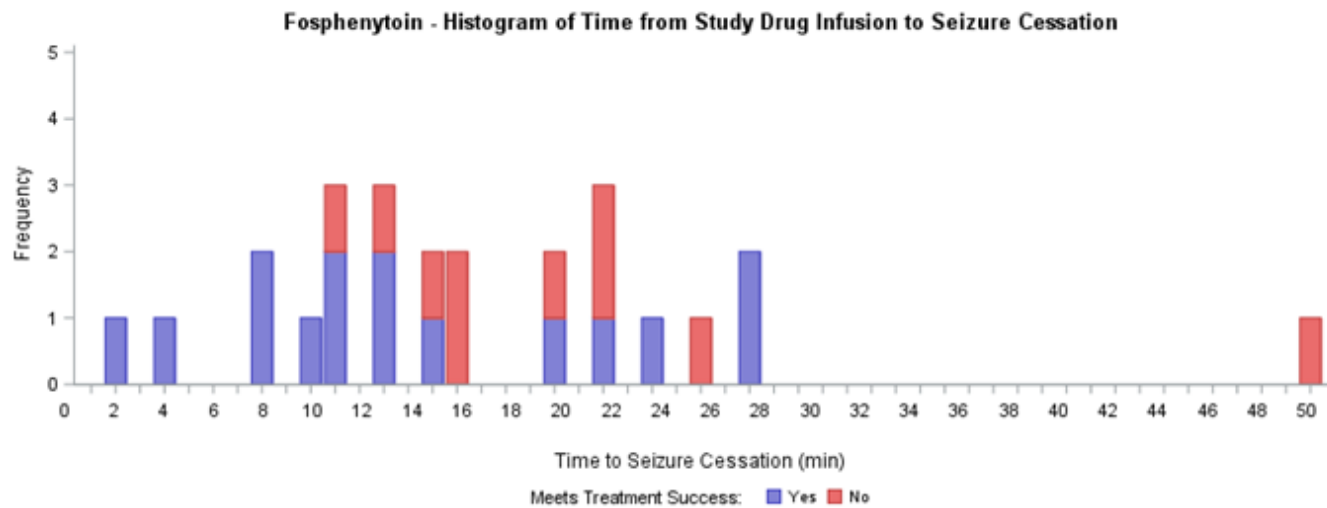


Figure S1. Time from start of study drug infusion to seizure cessation in those with and without success on primary outcome.

This figure shows the number of enrollments that had seizures cessation at each minute within the first hour from the start of study drug infusion. Exact time of seizure cessation was only determined from the subset of enrollments for which clinical teams marked the event in real time on the study voice data logger. (n=64)

Table S1. Serious adverse events by MedDRA preferred term

MedDRA preferred term	# EVENTS				# SUBJECTS				% SUBJECTS			
	LEV	FOS	VPA	Total	LEV	FOS	VPA	Total	LEV	FOS	VPA	Total
Total	98	78	72	248	64	57	46	167	42.7%	45.6%	36.8%	41.8%
Convulsion	32	26	23	81	30	25	23	78	20.0%	20.0%	18.4%	19.5%
Depressed level of consciousness	15	12	9	36	15	12	9	36	10.0%	9.6%	7.2%	9.0%
Respiratory depression	10	16	8	34	10	15	8	33	6.7%	12.0%	6.4%	8.3%
Hypotension	4	7	6	17	4	7	6	17	2.7%	5.6%	4.8%	4.3%
Pneumonia	4	2	2	8	4	2	2	8	2.7%	1.6%	1.6%	2.0%
Respiratory failure	1	2	4	7	1	2	4	7	0.7%	1.6%	3.2%	1.8%
Sepsis	1	2	2	5	1	2	2	5	0.7%	1.6%	1.6%	1.3%
Encephalopathy	4	0	1	5	4	0	1	5	2.7%	0.0%	0.8%	1.3%
Cardiac arrest	2	1	0	3	2	1	0	3	1.3%	0.8%	0.0%	0.8%
Rhabdomyolysis	3	0	0	3	3	0	0	3	2.0%	0.0%	0.0%	0.8%
Cerebral infarction	0	2	1	3	0	2	1	3	0.0%	1.6%	0.8%	0.8%
Conversion disorder	2	1	0	3	2	1	0	3	1.3%	0.8%	0.0%	0.8%
Deep vein thrombosis	1	0	2	3	1	0	2	3	0.7%	0.0%	1.6%	0.8%
Atrial fibrillation	0	1	1	2	0	1	1	2	0.0%	0.8%	0.8%	0.5%
Device malfunction	0	0	2	2	0	0	2	2	0.0%	0.0%	1.6%	0.5%
Hypothermia	0	0	2	2	0	0	2	2	0.0%	0.0%	1.6%	0.5%
Septic shock	1	0	1	2	1	0	1	2	0.7%	0.0%	0.8%	0.5%
Hypoglycemia	1	0	1	2	1	0	1	2	0.7%	0.0%	0.8%	0.5%
Intra-cardiac thrombus	0	0	1	1	0	0	1	1	0.0%	0.0%	0.8%	0.3%
Supraventricular tachycardia	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Dysphagia	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Gastrointestinal hemorrhage	0	1	0	1	0	1	0	1	0.0%	0.8%	0.0%	0.3%
Intestinal ischemia	0	1	0	1	0	1	0	1	0.0%	0.8%	0.0%	0.3%
Retroperitoneal hemorrhage	0	0	1	1	0	0	1	1	0.0%	0.0%	0.8%	0.3%
Small intestinal obstruction	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Infusion site extravasation	0	0	1	1	0	0	1	1	0.0%	0.0%	0.8%	0.3%
Hepatic failure	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Bacteremia	0	1	0	1	0	1	0	1	0.0%	0.8%	0.0%	0.3%
Clostridium difficile colitis	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Endocarditis	0	1	0	1	0	1	0	1	0.0%	0.8%	0.0%	0.3%
Respiratory tract infection	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Accidental overdose	0	1	0	1	0	1	0	1	0.0%	0.8%	0.0%	0.3%
Liver function test abnormal	0	0	1	1	0	0	1	1	0.0%	0.0%	0.8%	0.3%
Hypokalemia	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Muscle hemorrhage	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Cerebrovascular accident	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%

Cerebrovascular spasm	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Hemorrhage intracranial	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Hypoxic-ischemic encephalopathy	0	0	1	1	0	0	1	1	0.0%	0.0%	0.8%	0.3%
Agitation	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Delirium	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Urinary retention	0	0	1	1	0	0	1	1	0.0%	0.0%	0.8%	0.3%
Aspiration	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Epistaxis	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Hypoxia	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%
Obstructive airways disorder	0	0	1	1	0	0	1	1	0.0%	0.0%	0.8%	0.3%
Pleural effusion	0	1	0	1	0	1	0	1	0.0%	0.8%	0.0%	0.3%
Stridor	1	0	0	1	1	0	0	1	0.7%	0.0%	0.0%	0.3%

LEV is levetiracetam. FOS is fosphenytoin. VPA is valproate

Table S2. Pairwise treatment group differences in the primary outcome

Pairwise Comparison	Absolute Difference in proportion (95% credible intervals)	Probability treatment groups are different
LEV – FOS	1.9% (-10.0%, 13.9%)	Pr(LEV > FOS) = 0.62
VPA - FOS	1.3% (-11.1%, 13.8%)	Pr(VPA > FOS) = 0.58
LEV – VPA	0.06% (-11.3%, 12.5%)	Pr(LEV > VPA) = 0.54

LEV is levetiracetam. FOS is fosphenytoin. VPA is valproate

Table S3. Etiology of established status epilepticus as adjudicated by the phenomenology core

Precipitant of Enrolling Episode, n (%)	Levetiracetam (N=145)	Fosphenytoin (N=118)	Valproate (N=121)	Total
Unprovoked	50 (34.4)	42 (35.5)	34 (28.0)	126 (32.8)
Other *	18 (12.4)	22 (18.6)	21 (17.3)	61 (15.8)
Febrile illness	21 (14.5)	13 (11.0)	25 (20.7)	59 (15.4)
Anti-epileptic drug withdrawal / noncompliance	24 (16.6)	15 (12.7)	16 (13.2)	55 (14.3)
Toxic (alcohol/drug withdrawal, poisoning, etc.)	12 (8.3)	8 (6.8)	8 (6.6)	28 (7.3)
Insufficient information to determine - idiopathic	7 (4.8)	3 (2.5)	9 (7.4)	19 (4.9)
Acute stroke / hemorrhage	6 (4.1)	8 (6.8)	3 (2.5)	17 (4.4)
CNS tumor	2 (1.4)	4 (3.4)	2 (1.7)	8 (2.1)
CNS infection	3 (2.1)	1 (0.8)	2 (1.7)	6 (1.6)
Metabolic (hypoglycemia, hyponatremia, etc.)	2 (1.4)	2 (1.7)	1 (0.8)	5 (1.3)

* "Other" most frequently included afebrile and non-CNS infections, combinations of etiology, subacute stroke or hemorrhage, vasculitis, other encephalopathy, ventricular-peritoneal shunt failure, or sleep deprivation

Table S4. Analysis of primary outcome and interaction with baseline covariates

Covariate	p-value for test of all interactions*	p-value for test that random effect is 0
Sex (Male or Female)	0.76	N/A
Race (White, Black, Other)	0.98	N/A
Ethnicity (Hispanic, Not Hispanic)	0.92	N/A
Time from Onset (minutes)	0.33	N/A
Etiology (Non-epileptic spell, Seizure/Status epilepticus, Unable to determine)	0.60	N/A
Site (random effect)	N/A	1

Baseline covariates were evaluated individually in logistic regression models of the primary outcome that included treatment group indicators, the main effect of the covariate, and interaction terms with treatment. Site was handled as a random effect in a generalized linear mixed model with logit link with two indicators for treatment groups.

*Joint test of all interaction terms with treatment groups is a 2 degrees of freedom chi-squared test except for etiology and race which are 4 degree of freedom tests.

Table S5, Criteria for determination of primary outcome

Treatment	Additional anticonvulsant medications (including intubations)?,	Clinically apparent seizures at 60 minute assessment?,	Improving responsiveness at 60 minute assessment?,
Levetiracetam n=145	Y 52**	Y 14	Y 4
			N 10
		N 37	Y 15
			N 22
	N 93	Y 6	Y 1
			N 5
	N 87	Y 68	
		N 19	
Fosphenytoin n=118	Y 46	Y 6	Y 1
			N 5
		N 40	Y 13
			N 27
	N 72	Y 2	Y 0
			N 2
	N 70	Y 53	
		N 17	
Valproate n=121	Y 46	Y 8	Y 3
			N 5
		N 38	Y 15
			N 23
	N 75	Y 3	Y 1
			N 2
	N 72	Y 56	
		N 16	

Y = yes, N = no, **One subject receiving levetiracetam received additional anticonvulsant medications prior to 60 minutes, and the clinically apparent seizures and improving responsiveness at 60 minutes assessment was marked as "Unknown". Blue background indicates treatment success on primary outcome.

Table S6. Computations of the futility analysis

Look	Predictive probability that an arm is identified as best / worst at maximum sample size*			Predictive probability that any arm Wins**
	Levetiracetam	Fosphenytoin	Valproate	
Analysis after 400 ⁿ Enrollment (N=384 unique subjects)	.0013 / .0008	.002 / .0027	.0022 / .0013	0.01

* Maximum sample size was assumed to be 720 unique subjects for calculation of the predictive probabilities.

** This represents the sum of the predictive probabilities arm is best/worst at the maximum sample size for each of the 3 groups. If this sum is < 5%, the trial stops for futility.

Table S7. Post hoc analysis of time to seizure cessation by adjudication and real time response to query

	Levetiracetam	Fosphenytoin	Valproate
Seizure cessation within 20 minutes of trial drug initiation among patients with treatment success, n (%) [*]	53 (77.9%) n=68	43 (81.1%) n=53	43 (78.2%) n=55 [*]

^{*} Seizure cessation within 20 minutes was unable to be determined for 1 patient

Summary Description of Statistical Methods for Primary Analysis

The primary objective of ESETT was to identify the most effective and/or the least effective treatment among three commonly used second-line therapies for status epilepticus within an emergency department setting. A Bayesian approach was used to estimate the probability that each treatment is the best or worst, with a prior assumption that all three treatments were equally likely to be the best. We chose this approach, rather than a Bayes factor or a frequentist testing approach, because it allowed us to directly answer the question of clinical interest for this comparative effectiveness trial: Which treatment is best? Or, if there is no single best treatment, is there a worst treatment?

Prior to observing study data all three treatments were considered to be equally likely to be the most effective or least effective treatment. Each of the three treatment arms is modeled independently. Using Bayes theorem, the prior, likelihood, and data, we found a posterior distribution for each treatment arm. We assumed a noninformative Uniform(0,1) prior, (i.e. Beta(1,1)) and updated based upon the observed binomial data using a conjugate beta-binomial model. From these three posterior distributions, the probability that each treatment is the most effective (least effective) treatment was calculated. In brief, we randomly drew from the three posteriors, repeatedly (106 iterations), and counted the number of times a treatment was better than the other two, out of all the iterations, to calculate the probability that a given treatment was the best. The same approach was taken to calculate the probability that a given treatment was the worst. To control the false positive rate, we pre-specified criteria for trial success. Specifically, the probability that a treatment was the best (or worst) had to be greater than 0.975, in order to claim we have identified the best (or worst).

The threshold of 0.975 was chosen by convention (analogous to alpha of 0.025 one-sided comparison), and because an extensive simulation study demonstrated that, with this threshold and study design, the overall probability of a Type I error for identifying a best or worst treatment was less than 5%. The type I error probability (false positive rate) of this trial is the probability of incorrectly identifying a most effective treatment and/or incorrectly identifying a least effective treatment. In the simulations we empirically determined the type I error rate for this design by the number of times that one treatment was identified as the best and/or one treatment was identified as the worst (i.e. the probability that a treatment was the best (worst) was > 0.975) when in fact it was not the best/worst treatment, divided by the total number of iterations. Unlike a study in which a success can be achieved in numerous different ways (e.g. multiple doses versus a control or primary multiple endpoints), here one and only one treatment could be identified as best and it cannot also simultaneously be identified as the worst treatment. However, the multiple interim analyses would lead to multiplicity. Starting adaptive randomization prior to the possibility of early stopping likely may have decreased the probability of a Type I error.