SUPPLEMENTARY INFORMATION

SciBet a portable and fast single cell type identifier

Li et al.

SUPPLEMENTARY FIGURES

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Supplementary Figure 1. Benchmark result of feature selection methods for each of the 14 datasets. For each sub-plot, use scmap on 50 train-test instances (n=50) with different gene numbers, and measure the performance by the accuracy score. Box plot shows the center line for the median, hinges for the interquartile range and whiskers for 1.5 times the interquartile range.

Supplementary Figure 2. Benchmark result of classifiers for each of the 14 datasets. For each sub-plot, benchmark different classifiers on 50 train-test instances (n=50) with 500 genes selected by E-test, and measure the performance by the accuracy score. Box plot shows the center line for the median, hinges for the interquartile range and whiskers for 1.5 times the interquartile range.

Supplementary Figure 3. Benchmark result of classifiers for each of the 14 datasets measured by balanced accuracy score. For each sub-plot, benchmark different classifiers on 50 train-test instances (n=50) with 500 genes selected by E-test, and measure the performance by the balanced accuracy score. Box plot shows the center line for the median, hinges for the interquartile range and whiskers for 1.5 times the interquartile range.

Supplementary Figure 4. Benchmark result of classifiers for each of the 10 FPR control pairs. For each sub-plot, benchmark different classifiers on 50 train-test instances (n=50) with different FPRs, and measure the performance by the accuracy score. Box plot shows the center line for the median, hinges for the interquartile range and whiskers for 1.5 times the interquartile range.

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Supplementary Table 2. Datasets for the cross-platform benchmarks

Supplementary Table 3. Datasets for building the mock human cell atlas

Supplementary Table 4. Datasets for building the null dataset

Num	Dataset	Species	Cell types	Experimental protocol
	GSE99254 ³²	human	$\mathsf T$ cell	Smart-Seq2
	GSE10898933	human	T cell	Smart-Seq2
	GSE84799 ³⁶	human	B cell	Smart-Seq2
	GSE77940 ³⁷	human	B + macrophage + NK + T cell	TruSea
	GSE89232 ²⁴	human	DC	Smart-Seq2
	GSE70580 ¹⁰	human	$ILC + NK$	Smart-Seq2
	GSE87849 ²³	human	macrophage	SMARTer

Supplementary Table 6. Datasets of the immune datasets for the E-test case study

SUPPLEMENTARY NOTES

Note 1: Discussion on the metrics for measuring the classification performance

We collected datasets from the source of the publications with the original cell type annotation. Because the cells were annotated by the publications with the unsupervised workflow, which usually consists of clustering, differential expression and cell type identification based on marker gene of each cluster, cell numbers of different cell types rarely have equal proportions. Such unequal proportions reflect true proportions of cells within each dataset, and thus have different contributions to the global assessment of accuracy. So, we used the *accuracy score*⁴⁴ as the default metric, which equals to total number of correct classification $\frac{a}{b}$ $\frac{a}{b}$ $\frac{b}{c}$ $\frac{c}{d}$ $\frac{c}{d}$ and also equals to the micro-average recall number of all cells in the test set score of each cell type (i.e., calculating the accuracy of classification for each cell type, and then averaging them with the weight in proportion to the number of each cell type). Here is the derivation:

If we have a dataset with n cell types and for each cell type i, we have assigned a_i and b_i cells with the correct and wrong cell type in the test set, respectively: *accuracy score* = $\frac{total \ number \ of \ correct \ classification}{number \ of \ all \ cell}$ ber of correct classification
number of all cells $=$ $\frac{\sum a_i}{\sum a_i + \sum a_i}$ $\frac{2a_i}{\sum a_i + \sum b_i}$. We can also calculate the recall score for each cell type: recalli= number of correct classification for cell type i mber of correct classification for cell type i
total number of cell type i in the test set $a_i +$ a_i+b_i , and weight $:=\frac{number\ of\ cell\ type\ i}{m}\$ $\frac{ \displaystyle \textit{umber of cell type i}}{ \displaystyle \textit{number of all cells}} = \frac{\displaystyle a_i + b_i}{\displaystyle \sum a_i + \sum_i}$ $\frac{a_i+b_i}{\sum a_i+\sum b_i}$. So, the micro-average recall = $\sum (acc_i *$

weight_i)= $\sum(\frac{a_i}{a_i+1})$ $\frac{a_i}{a_i+b_i} * \frac{a_i+b_i}{\sum a_i+\sum}$ $\frac{a_i+b_i}{\sum a_i+\sum b_i} = \frac{\sum a_i}{\sum a_i+\sum}$ $\frac{\sum u_i}{\sum a_i + \sum b_i}$ accuracy score.

We also applied the balanced accuracy score 44 (the macro-average recall score) to measure the overall performance, which equals to $\frac{1}{n}\sum \text{recall}_i$, to account for the rare cell types, where each cell type in the test set has equal contribution.

Note 2: Detailed derivation and discussion for E-test

Note 2.1: Calculation of entropy in a given cell type

We applied the strategy proposed by Splatter⁴⁵ and Saver⁴⁶ to model the observed expression Zic for gene i (ranges from 1 to m) in cell c (ranges from 1 to C) with the identical cell type as following:

Zic~Possion (λi *sc), λ i ~Gamma ($\alpha_i,\;\beta_i)$

, where λ_i represents the Gamma-distributed true expression of gene i in cell c, and s_c represents the size-factor (can be calculated by $\sum_{c} Z_{ic}$) of cell c accounting for the sequencing depth. Then we approximately estimate the Poisson rate λ_i *s_c directly by Z_{ic} using the moment estimation. Thus, for each cell we will obtain $\lambda_{ic} = \frac{Z_{ic}}{c}$ $\frac{\varepsilon_{LC}}{\varepsilon_{c}}$ as C observations for λ_{i} , which is in accordance with the size-factor normalized expression described in part Data collection and processing of section Methods.

We applied the differential entropy from information theory to measure the dispersion degree of the distribution. The Shannon entropy *Sⁱ* of the Gammadistributed true expression λ_i can be calculated by the equation⁴⁷:

$$
S_i = \alpha_i - \ln \beta_i + \ln \Gamma(\alpha_i) + (1 - \alpha_i) \psi(\alpha_i)
$$
 \n
$$
\qquad \qquad \text{---} \qquad \qquad (1)
$$

, where ψ represent the digamma function. Then we replace β_i with its maximum likelihood estimation as $\hat{\beta}_i = \frac{\alpha_i}{E(\lambda)}$ $\frac{\alpha_i}{E(\lambda_i)}$, where $E(\lambda_i)$ = $\frac{1}{C}$ $\frac{1}{c}\sum_c \lambda_{ic}$ served as the unbiased estimation of λ_i . Thus, formula (I) can be further derived into:

$$
S_i = \alpha_i - \ln \alpha_i + \ln E(\lambda_i) - \ln \Gamma(\alpha_i) + (1 - \alpha_i)\psi(\alpha_i) \qquad \qquad \text{-----}(2)
$$

We denote $h_i = \alpha_i - \ln \alpha_i + \ln \Gamma(\alpha_i) + (1 - \alpha_i)\psi(\alpha_i)$) ------------(3) as the function of the gene-specific parameter α_i , and denote the mean normalized expression over C cells as $X_i = E(\lambda_i) = \frac{1}{C}$ $\frac{1}{C}\sum_{c}\lambda_{ic}$. Thus, Supplementary Equation (2) can be derived into $S_i = \ln E(\lambda_i) + h_i = \ln X_i + h_i$ ------------(4)

Note 2.2: Calculation of entropy of different cell types

In Supplementary Note 2.1, we obtained the Supplementary Equation (4) for the calculation of entropy of gene i in a given cell type. For cell *c* (ranges from 1 to C_i) belonging to cell type j, we define $X_{ij} = \frac{1}{C}$ $\frac{1}{\mathsf{C}_j} \sum_{c \in j} \lambda_{ic}$ as the mean normalized gene expression over cells belonging to cell type j, and calculate the entropy S_{ij} of gene i in cell type j as:

$$
S_{ij} = \ln X_{ij} + h_{ij}
$$
 (5)

, where $h_{ij} = \alpha_{ij} - \ln \alpha_{ij} + \ln \Gamma(\alpha_{ij}) + (1 - \alpha_{ij})\psi(\alpha_{ij})$ -------------(6)

We then assume that $h_{ij} = h_i$, that is, h is gene-specific but not cell typespecific. Then we obtained the Equation (1) in Main Text: $S_{ij} = \ln X_{ij} + h_i$. The assumption of $h_{ij} = h_i$ is equivalent to another assumption proposed by Lun et al.⁴⁸, where they modeled the differential expression among multiple groups with fixed fold changes. Here is the derivation: We denote $\lambda_{i,c \in j}$ as the observations of the Gamma-distributed true expression λ_{ij} ~Gamma (α_{ij} , β_{ij}) of gene i in cell type j. For gene i, if we multiply λ_{ij} with a fixed constant $\mathrm{F_{i,j\rightarrow j'}}$ as the fold change of gene i, from cell type j to j', as $\lambda_{ij'}$ = $F_{i,j\to j'}$ * λ_{ij} . Thus, $\lambda_{ij'}$ will follow Gamma $(\alpha_{ij},\ \beta_{ij}/\mathrm{F}_{\mathrm{i,j}\to j'})$, according to the scaling property of the Gamma distribution (See the following lemma in Note 2.3). Here we note that the Gamma-distributed $\lambda_{ij'}$ and λ_{ij} share the identical α_{ij} . Thus, for all cell types, the shape parameter α_{ij} is not cell type-specific, as $\alpha_{ij} = \alpha_i$. According to Supplementary Equation (6), h will also be the gene-specific but not cell typespecific variable in accordance with α .

Note 2.3: The scaling property of the Gamma distribution

Definition: if X ∼ Gamma (α, β), then Y = kX ∼ Gamma (α, β/k) for any k>0.

Proof: The probability density function for X is defined as:

$$
f_X(x) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} x^{\alpha - 1} e^{-\beta x}, \text{ for all } x > 0
$$

We then let k as a positive constant, and thus the Jacobian of the transformation will be $\frac{dX}{dY} = \frac{1}{k}$ $\frac{1}{k}$.

So, $f_Y(y) = \frac{dX}{dy}$ $rac{\mathrm{dX}}{\mathrm{d}Y} * f_X\left(\frac{y}{k}\right)$ $\left(\frac{y}{k}\right) = \frac{1}{k}$ $\frac{1}{k} * \frac{\beta^{\alpha}}{\Gamma(\alpha)}$ $\frac{\beta^{\alpha}}{\Gamma(\alpha)} * (\frac{y}{k})$ $(\frac{y}{k})^{\alpha-1} * e^{-\frac{\beta y}{k}} = \frac{(\frac{\beta}{k})^{\alpha}}{\Gamma(\alpha)}$ $\frac{\beta}{k}$ ^{\α} $\frac{(\frac{\rho}{k})^{\alpha}}{\Gamma(\alpha)} * y^{\alpha-1} e^{-\frac{\beta}{k}}$ $\frac{P}{k}$ *^y, for all y>0. Now $f_Y(y)$ is exactly the probability density function of Gamma (α, β/k).

Note 2.4: Permutation test for obtaining the significance of

 $\varDelta S_i,$ the total entropy difference of gene i is defined as Equation (3) in the main text, as $\Delta S_i = \sum_{j=1}^n (S_{i0} - S_{ij}) = \sum_{j=1}^n (ln X_{i0} + h_i - ln X_{ij} - h_i)$ $j=1$ \boldsymbol{n} $\sum_{j=1}^n (S_{i0} - S_{ij}) = \sum_{j=1}^n (ln X_{i0} + h_i - ln X_{ij} - h_i)$, where the significance of ΔS can be approximated by the permutation test (randomly permutate the cell group labels, calculate ΔS for each permutation and find the percentile of the actual ΔS), which may take long time. Here we show a parametric method to accelerate this process by replacing the permutating step, as following:

Under the null hypothesis that all cells from the pre-defined groups are randomly sampled from the the same cell population, each cell type j will have the identical mean size-factor normalized expression X_{ij} , which follows a normal distribution $N(\mu_i, \sigma_i)$ according to the central limit theorem. Here we can denote the unbiased estimation for the parameters as $\hat{\mu}_i = \frac{1}{n}$ $\frac{1}{n} * \sum_{j=1}^{n} X_{ij}$ and $\hat{\sigma}_i = \frac{1}{n-1}$ $\frac{1}{n-1} *$ $\sum_{j=1}^n(X_{ij}-\hat{\mu}_\mathrm{i})^2.$ Notably, $\,\hat{\mu}_\mathrm{i}\,$ is exact the $\,X_{i0}\,$, the mean size-factor normalized

expression of the null group 0, as we defined in the main text. For each round

of permutation, we can generate n X_i observations from $N(\mu_i, \sigma_i)$ with the constraint that $\hat{\mu}_i = \frac{1}{n}$ $\frac{1}{n} * \sum_{j=1}^{n} X_{ij}$ (this process equals to generating n-1 X observations and calculate another by the n-1 observations and $\hat{\mu}_i$). Then, we can calculate ΔS according to formula (***).

Note 3: Discussion on whether to use the cell type prior probability in SciBet model training

The proportion of each type of cell in the data measured by a single cell sequencing does not necessarily and correctly reflect the prior probability of appearance. For example, if a piece of tissue is sequenced without any sorting, then the proportion of each cell type in the results of single-cell sequencing can reflect the prior probability. However, if certain artificial filtering (such as Fluorescence-activated cell sorting to select cells highly expressing certain surface protein) is performed, or the dataset is integrated from different batches or studies, then the final cell type ratio at this time cannot correctly reflect the prior probability of the appearance of such cell types. This latter situation is more common and considered to be more appropriate for maximum likelihood estimation (i.e., the prior probabilities of each class are considered equal in Bayesian decision making), which is used as the default option.

If users choose to consider the prior probabilities of different cell types, they can replace the strategy for making decision (formular (****) in part Supervised cell type prediction by SciBet, section METHODS) with the following strategy of Bayes Decision Rule $j = argmax_j(P(y|j) * P(j)) = argmax_j(\prod_i (p_{ij}^{y_i}) * P(j)),$ to make decisions. And in most cases, users can estimate the prior probability according to the proportion of cell types in the training set.

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