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Supplemental Information

Immune Cell Associations with Cancer Risk

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Supplementary figures legends

Fig. S1. Evaluation of immune/stromal cell tissue content estimates in relation to two other methods. Related to Figure 1.

(A) Heatmap showing the correlations (Spearman's ρ) between Consensus^{TME}based values and analogous TIMER cell type estimates.

(B) Heatmap showing the correlations (Spearman's ρ) between Consensus^{TME}based values and analogous MCP-counter cell type estimates.

Fig. S2. Evaluation of immune/stromal cell tissue content estimates in relation to independent leukocyte estimates. Related to Figure 1. Heatmap showing the correlations (Spearman's ρ) between Consensus^{TME}-based values and independent leukocyte estimates using the approach of Taylor et al. (2018).

Fig. S3. Evaluation of immune/stromal cell tissue content estimates in relation to aneuploidy scores. Related to Figure 1.

(A) Heatmap showing the correlations (Spearman's ρ) between Consensus^{TME}based values and aneuploidy scores (Taylor et al., 2018) across major cancer types.

(B) Heatmap showing the correlations (Spearman's ρ) between Consensus^{TME}based values and aneuploidy scores (Taylor et al., 2018) across in BRCA subtypes, which show positive correlations in claudin-low.

Fig. S4. Differences of inferred immune/stromal cell content between primary tumors with low and high levels of *CD274/PDL1* **expression.**

Related to Figure 1. The graphs show the median cell content value in each group and the significance of the difference (Wilcoxon test *P* value).

Fig. S5. Differences of inferred immune/stromal cell content between primary tumors with low and high levels of *CD279/PDCD1* **expression. Related to Figure 1.** The graphs show the median cell content value in each group and the significance of the difference (Wilcoxon test *P* value).

Fig. S6. Correlations between inferred blood immune cell contents and measures from fluorescence-activated cell sorting in blood samples. Related to Figure 1. Forest plot showing correlation estimates and 95% CIs of each inferred cell type (data from whole blood samples of healthy adults; $n = 12$, GEO GSE127813).

Fig. S7. Correlations between inferred immune/stromal cell tissue contents and single cells used to generate pseudo-bulk breast tumors. Related to Figure 1. Each panel shows the correlation between immune cell signature scores (Y-axis) and the number of cells (X-axis) used to generate 100 pseudo-bulk breast tumors (data from Gene Expression Omnibus reference GSE75688). The trend lines, Spearman's correlations and *P* values are shown.

Fig. S8. Correlations between immune/stromal cell tissue contents and expression of immune benchmark genes. Related to Figure 1. Top panel, distribution of PCCs using data from normal TCGA tissue. Bottom panel,

distribution of PCCs using data from primary tumors of TCGA. Mean PCCs and 95% CIs are shown.

Fig. S9. Correlations between immune cell signatures and pathway signaling-inferred activities. Related to Figure 1. Unsupervised clustering of the correlation coefficients between inferred cell contents (Y-axis) and KEGG pathway activities (X-axis). Differentiated clusters in normal tissue are marked by red-outlined rectangles.

Fig. S10. Gene targets of eQTL recognized in isQTLs are frequently correlated with the corresponding immune/stromal cell signatures. Related to Figure 2. Distributions of random gene sets (same gene set size and equivalent comparisons for each signature and TCGA setting) relative to the number of significant correlations between eQTL-target and immune/stromal signatures. Left- and right-hand panels show results for the first and second isQTL sets presented in the main text, respectively. Empirical test probabilities are shown.

Fig. S11. Minimal correlation estimates to detect significant signature-PRS associations. Related to Figure 3. Left and right panels show the lowest correlations required in each normal and primary tumor setting, respectively, to detect nominal (*P* < 0.05) associations given the TCGA sample sizes.

Fig. S12. LUAD and LUSC PRS correlations with NK cell content. Related

to Figure 3. Top panels, positive correlations between NK cell content in primary tumors of LUAD and LUSC, and the corresponding PRSs. The adjusted-R2 and *P* values of the linear regression model are shown. Bottom panels, correlation trends of patients stratified by smoking status, as depicted in the insets. The estimate for LUAD cases classified as current smokers was found to be significantly less than zero $(r = -0.12, P = 0.012)$.

B

Spearman's ρ **0.5 1**

PDCD1 **expression**

PDCD1 **expression**

Figure S7

Immune cell tissue content correlations with defined immune gene benchmarks

eQTL-gene target correlations with immune/stromal cell signatures

Minimal correlation value to detect a significant PRS-cell signature associations (*P* **< 0.05)**

Transparent Methods

TCGA data

Clinical and gene expression (RNA-seq fragments per kilobase of transcript per million mapped reads (FPKM) upper quartile normalized (UQ)) data from The Cancer Genome Atlas (TCGA) projects were obtained from the Genomic Data Commons Data Portal (https://portal.gdc.cancer.gov) and from the corresponding publications. Genetic data at the individual level were obtained following approval by the dbGaP Data Access Committee (project #11689). Metastases and recurrent tumors were excluded from this study, making normal tissue (blood or solid tissue) and primary tumor samples the focus of the analyses. The cancer types are named using the corresponding TCGA study abbreviations (https://gdc.cancer.gov/resources-tcga-users/tcga-code-tables/tcga-studyabbreviations). For normal tissue, according to the TCGA protocols, these samples were collected > 2 cm from the tumor margin and/or did not contain tumor identified by histopathological review. The protein expression measures of CD26 and TFCR corresponded to those obtained by TCGA using reverse-phase protein arrays (RPPAs; level 4 data, https://tcpaportal.org/tcpa/). The COAD subtypes were defined based genomic/genetic alterations (chromosomal instability (CIN), genomic stable (GS), and microsatellite instability (MSI) tumors) and on molecular features (consensus molecular subtypes, CMS1-4) (Guinney et al., 2015).

Cancer risk variants

The variants were compiled from the GWAS Central (Beck et al., 2014) and GWAS Catalog (Buniello et al., 2019) databases, and by literature searches using the PubMed MeSH terms "GWAS", "association", "cancer", and "risk". The variants are listed in **Table S1**. The UK Biobank GWAS results were taken from the public repository at http://www.nealelab.is/uk-biobank.

Benchmark immune genes

These genes were compiled from The Immunological Genome Project (ImmGen) (Shay and Kang, 2013) and CellMarker (Zhang et al., 2019) databases, and by a literature search using the MeSH terms corresponding to the specific immune cell types represented by the gene expression signatures. The benchmarks and their cell type assignments are included in **Table S3**.

Genotype data and imputation

Bulk genotyping data corresponding to the Affymetrix Genome-Wide Human SNP 6.0 Array were downloaded from the TCGA legacy archive (https://gdcportal.nci.nih.gov/legacy-archive/). Of the initial normal tissue and primary tumor samples ($n = 16,599$), those corresponding to individuals of self-reported non-white origin ($n = 4,770$), and those of non-European origin based on principal component analysis using variants intersected in the 1000 Genome Project phase III ($n = 2,598$) were excluded from subsequent analyses; these filters were applied because summary

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statistics of the GWASs used in this study are strongly biased towards populations of European origin. Normal and tumor samples were then examined separately for duplicates and up to third-degree relatives (kinship cutoff $= 0.05$), which resulted in the exclusion of an additional 672 samples. In the joint dataset, 765 samples were also excluded because they showed a gender mismatch in an analysis of pseudoautosomal genomic regions. Considering genetic variants, 108 samples that deviated by four or more standard deviations from the mean heterozygosity rate were also excluded. For imputation, variants were excluded if they fulfilled any of the following criteria: they mapped to chromosome Y, pseudoautosomal regions or the mitochondrial genome; they had a call rate < 100%; their minor allele frequency was < 0.01; they departed from Hardy–Weinberg equilibrium (P < 5x10⁻⁶); or they mapped to AT-CG sites. Finally, 7,686 samples (4,154 normal, comprising 3,287 blood-derived and 867 solid-tissue samples; and 3,532 primary tumors, of which 94.4% were paired) and 589,101 variants were retained for subsequent analyses. Imputation was performed using the Shape-IT V2 (Delaneau et al., 2008) and IMPUTE2 (Howie et al., 2009) algorithms, and the 1000 Genome Project Phase III panel as reference. Poorly imputed variants (accuracy score < 0.7) were excluded from subsequent analyses. A standard cutoff dose was applied to calculate genotypes using a hard-call threshold of 0.1 (i.e., $0 - 0.1$, $0.9 - 1.1$, $1.9 - 2.0$ for reference homozygote, heterozygous and alternative homozygous genotypes, respectively).

Immune/stromal cell signatures

Immune/stromal cell gene expression signatures for each TCGA cancer setting were computed using the Consensus^{™E} method (Jiménez-Sánchez et al., 2019), which was provided available as an R package (https://github.com/cansysbio/ConsensusTME). Ten single-cell breast cancer signatures (Azizi et al., 2018) were included in the TCGA BRCA analyses. Therefore, 18 signatures were examined in each normal tissue and primary tumor setting, except for normal breast and breast cancer tissue, for which a total of 28 signatures were analyzed. The signature scores were computed using the single-sample Gene Set Expression Analysis (ssGSEA) algorithm calculated within the Gene Set Variation Analysis (GSVA) software (Hänzelmann et al., 2013). These scores were calculated for normal tissue and primary tumors, but not for blood samples, since data from blood are limited to germline genotypes. Genes whose expression was uninformative in more than half the samples in a given setting were excluded from the signature calculations; otherwise, missing data were assigned the average value of the informative samples. Evaluation of signature scores computed by two different methods —ssGSEA and summing normalized gene expression Z-scores— revealed global coherence, whereby Pearson correlation coefficients (PCCs) were > 0.80 in 99% (571/578) of the score comparisons. To select independent signatures in each normal and cancer setting, we performed a principal component analysis using the *prcomp* function in R. Components with eigenvalues > 1 were retained to study quantitative trait loci (subsequent sections). Estimates of immune-related pathway activities were calculated using directed graphs from the Kyoto Encyclopedia of Genes and Genomes (KEGG, https://www.genome.jp/kegg/). Briefly, gene expression profiles were converted into pathway module activity scores by taking into account the chain of reactions from a defined molecular input to a specific molecular output (Cubuk et al., 2018). The 84-gene signature linked to SH2B3 included the genes differentially expressed in *Sh2b3*-null cells and that participate in genetic and/or protein interactions to this gene/protein (Huan et al., 2015); *SH2B3* was excluded from this signature for subsequent analyses.

Pseudo-bulk breast tumors

To generate 100 pseudo-bulk breast tumors, we used the single-cell RNA-seq data from the Gene Expression Omnibus (GEO) reference GSE75688 (Chung et al., 2017) and aggregated read counts using the aggregateData function in R (https://github.com/HelenaLC/muscat). Each simulated sample of 100 cells was forced to include >50% tumor cells (average 75.3%, 95% CI 72.53 – 77.93%). For non-tumoral cells, 10% of them were fixed as stromal (bulk average 7.22% , 95% CI 6.22 – 8.36%), while the other 90% were a random combination of B cells (average 5.16%, 95% CI 4.28 – 6.28%), T cells (average 6.21%, 95% CI 5.05 – 7.48%), and myeloid cells (average 6.11%, 95% CI 5.05 – 7.39%). Most of the myeloid cells were originally assigned to macrophages (Chung et al., 2017).

Quantitative trait loci of immune/stromal cell tissue content

The germline genetic calls corresponded to genotype data obtained from blood or normal tissue samples. For cases with both types of sample, the variants with discordant

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calls were excluded from subsequent analyses. As specified above, the somatic genetic calls corresponded to primary tumors only. The immune/stromal cell-content quantitative trait loci (isQTL) were analyzed using the *R/qtl2* package in R (Broman et al., 2019). These analyses included the covariates of gender (when informative), age at diagnosis, tumor stage and histology. The Haley–Knott regression method was used to compute the log odds (LOD) of the associations between genetic variants and immune/stromal cell scores. One thousand permutations were performed in each setting to obtain significance thresholds (Manichaikul et al., 2007) and the variant-signature associations with empirical values of *P* < 0.05 were considered significant isQTL. The gene targets were defined according to the genomic location of the identified variants. Additional targets were identified by analyzing all variants correlated ($r^2 > 0.8$, 1000 Genomes phase 3, version 5) with each isQTL and intersect them with various functional genomic data, including promoter capture Hi-C (Javierre et al., 2016), annotated enhancers (Hnisz et al., 2013, p.), and eQTL (Schmiedel et al., 2018) from B cells, monocytes, and CD4+ and CD8+ T cells. In addition, correlated variants were queried using the Ensembl Variant Effect Predictor (McLaren et al., 2016) for potential effects on protein coding sequences.

Computation of PRSs

The PRSs were compiled from the literature and computed by summing the products of the per-allele LOD ratio assigned to each risk variant, and the corresponding allele dosage, for the total number of variants initially defined for each PRS. There was no

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previous evidence of significant interactions or deviations from a log-additive model in BRCA PRSs (Mavaddat et al., 2019), but it is not known for other cancers. In the analyses of BRCA, OV (no normal tissue data available), and PRAD PRSs, two sets were analyzed, both based on GWAS-identified variants: set #1 (hereafter PRSs-1), which corresponded to scores derived from large collections of GWAS cohorts and validated in independent studies (Mavaddat et al., 2019); and set #2 (hereafter PRSs-2), which corresponded to scores derived from a phenome-wide longitudinal study using electronic health records collected by the Michigan Genomics Initiative (Fritsche et al., 2018). In both sets, PRSs were developed for all BRCA patients, and separately for the estrogen receptor (ER)-positive and ER-negative subtypes. The number of initial variants in these BRCA PRSs and those included in our study, based on available genotypes and obtained imputations were 307 and 185 for PRSs-1, and 3,820 and 3,629 for PRSs-2. As expected, the PRSs from the two sets were found to be positively correlated using germline or primary tumor data: BRCA PRSs PCCs = 0.60 – 0.66, *P* < 10⁻⁵; OV tumors PRSs PCC = 0.72, $P < 10^{-25}$ (serous PCC = 0.72); and PRAD PRSs PCCs = 0.23 – 0.99, *P* < 0.01. The Michigan Genomics Initiative also provided PRSs for seven other cancer types, and the number of variants originally included and analyzed in this study were, respectively: 103 and 21 for PRAD; 42 and 41 for COAD; 16 and 16 for BLCA and SKCM; 15 and 15 for OV; 9 and 9 for GBM, LUAD and LUSC; 8 and 7 for THCA; and 7 and 6 for KIRC.

Cell signature associations with PRSs

The *bestNormalize* package in R (https://github.com/petersonR/bestNormalize) was used to normalize the cell signature values. The transformation that produced the lowest value from the Pearson's statistic divided by the degrees of freedom was taken to indicate the best function. The error distributions of the models and Q-Q plots were examined individually. The parameters of each signature transformation are provided in **Table S10**. Outliers were identified using the interquartile range rule and excluded from subsequent analyses; these were < 5% in all settings. Normalized signature values were used as dependent variables in a linear regression analysis relative to the PRSs. Stepwise analyses including covariates of gender, tumor stage and histology were performed, and the best model was selected based on the Akaike information criterion (AIC). For normal tissue, only those studies with at least 50 informative samples were analyzed. The small number of samples in each setting meant that these analyses could only detect significant (nominal $P < 0.05$) correlation estimates > 0.27 and > 0.09 in normal breast tissue and BRCA, and stronger correlations would be required in all other settings if nominal significance were to be reached (**Fig. S11**). The significance of the associations was corrected for multiple testing using the false-discovery rate (FDR) method.

Cell signature associations with age at diagnosis

The associations between the cell signature scores (dependent variables) and age at diagnosis were evaluated by multiple linear regression, including gender and tumor

stage as covariates, the best model being determined from an AIC-based stepwise selection algorithm. The statistical significances of the associations were corrected for multiple testing separately in normal tissue and primary tumor analyses (since the expected effects were the opposite of what they proved to be) using the FDR method.

Breast cancer risk in *BRCA1/2***-mutation carriers**

Analyses were performed using data from the OncoArray and Collaborative Oncological Gene-environment Study (iCOGS) consortiums with the participation of the Consortium of Investigators of BRCA1/2 Modifiers (CIMBA). The OncoArray and iCOGS designs, quality controls, and statistical analyses have been described previously (Milne et al., 2017). Summary statistics from the retrospective likelihood method are reported.

Analysis of blood cell parameters and age at diagnosis of breast cancer

Clinical and histopathological data from breast cancer patients were compiled through manual curation of hospital records of the Catalan Institute of Oncology, L'Hospitalet del Llobregat (Barcelona, Catalonia, Spain). Patients were randomly selected from health records collected between 2009 and 2014. The compiled data included date of birth, age, gender (only women selected), date at diagnosis, tumor stage, subtype and/or ER status, and date at initial-diagnostic blood test. The blood test parameters analyzed were the normalized numbers (x10⁹/L) of basophils, eosinophils, leucocytes, lymphocytes, monocytes, neutrophils, and platelets. Linear regressions of each of these parameters

on age at diagnosis, including tumor stage and subtype as covariates, were performed.

The IDIBELL's Research Ethics Committee approved this study (reference PR066/20).

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