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Risk assessment and antibiotic prescribing decisions in children presenting with cough: a vignette study

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Title:

Risk assessment and antibiotic prescribing decisions in children presenting with cough: a vignette study

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ABSTRACT

Objectives: The validated “STARWAVE” clinical prediction rule (CPR) uses seven variables to guide risk assessment and antimicrobial stewardship in children presenting with cough (**S**hort illness duration, **T**emperature, **A**ge, **R**ecession, **W**heeze, **A**sthma, **V**omiting). We aimed to compare General Practitioners’ (GPs) risk assessments and prescribing decisions to those of STARWAVE, and assess the influence of the CPR’s clinical variables.

Setting: Primary care.

Participants: 252 GPs, currently practising in the UK.

Design: GPs were randomly assigned to view four (of a possible eight) clinical vignettes online. Each vignette depicted a child presenting with cough, who was described in terms of the seven STARWAVE variables. Systematically, we manipulated patient age (20 months vs. 5 years), illness duration (3 vs. 6 days), vomiting (present vs. absent) and wheeze (present vs. absent), holding the remaining STARWAVE variables constant.

Outcome measures: Per vignette, GPs assessed risk of hospitalisation and indicated whether they would prescribe antibiotics or not.

Results: GPs overestimated risk of hospitalisation in 9% of vignette presentations (88/1008) and underestimated it in 46% (459/1008). Despite underestimating risk,

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3 they overprescribed: 78% of prescriptions were unnecessary relative to GPs' own
4 risk assessments (121/156), while 83% were unnecessary relative to STARWAVE
5 risk assessments (130/156). All four of the manipulated variables influenced risk
6 assessments, but only three influenced prescribing decisions: a shorter illness
7 duration reduced prescribing odds (OR 0.14, 95% CI 0.08-0.27, $p<0.001$), while
8 vomiting and wheeze increased them (OR_{vomit} 2.17, 95% CI 1.32-3.57, $p=0.002$;
9 OR_{wheeze} 8.98, 95% CI 4.99-16.15, $p<0.001$).

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21 **Conclusions:** Relative to STARWAVE, GPs underestimated risk of hospitalisation,
22 overprescribed, and appeared to misinterpret illness duration (prescribing for longer
23 rather than shorter illnesses). It is important to ascertain discrepancies between
24 CPRs and current clinical practice. This has implications for the integration of CPRs
25 into the electronic health record and the provision of intelligible explanations to
26 decision makers.
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STRENGTHS AND LIMITATIONS OF THIS STUDY

- This is the first study to suggest discrepancies between the STARWAVE clinical prediction rule and current clinical practice.
- Use of clinical vignettes allowed us to manipulate some variables while holding others constant; thus we could identify *causal* relationships between specific clinical variables and antibiotic prescribing decisions.
- In so doing, we bring much-needed experimental evidence to the literature, which is currently dominated by interview and observational studies.
- The disadvantage of using clinical vignettes is that our results are based on hypothetical clinical scenarios, which contained limited information.
- Moreover, we manipulated only a subset of the STARWAVE variables; future work could increase the number of clinical variables manipulated, and explore non-clinical factors too.

INTRODUCTION

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8 Combatting antimicrobial resistance is high on policy agendas internationally.[1-
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10 3] One of the key means advocated is judicious antibiotic prescribing.[1] Over 80% of
11
12 all NHS antibiotic prescriptions are issued in primary care,[4] where despite
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14 numerous campaigns, mandates and financial incentives, rates remain unacceptably
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16 high.[5] Despite strong evidence of only modest symptomatic benefits for acute
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18 respiratory tract infections (RTIs),[6-8] and even smaller effects against
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20 complications,[9, 10] RTIs are the most common justification for primary care
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22 antibiotic use [11] and a leading cause of overuse.[12] This is exacerbated in
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24 children, where perceived vulnerability and prognostic uncertainty (i.e., perceived
25
26 risk of deterioration) lead to defensive prescribing (“treat, just in case”).[12-15]
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33 To improve risk assessment and antimicrobial prescribing in children with RTIs, a
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35 clinical prediction rule (CPR) called “STARWAVE” was recently developed and
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37 validated.[12] It was based on a large prognostic cohort study, which included 8394
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39 children presenting to 247 general practices in England with acute cough and RTI
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41 symptoms.[12] Numerous characteristics were recorded at presentation, including
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43 demographic variables, parent-reported symptoms and physical examination signs.
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45 In a regression analysis, seven of these characteristics were found to predict hospital
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47 admission (for RTI) in the month following presentation: **Short illness duration** (≤ 3
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49 days), **Temperature** ($\geq 37.8^{\circ}\text{C}$), **Age** (< 2 years), **Recession**, **Wheeze**, **Asthma** and
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51 **Vomiting**. [12] This analysis gave rise to the “STARWAVE” clinical prediction rule: a
52
53 seven-item, point-of-care checklist that can distinguish children at “very low” (0.3%,
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55 with ≤ 1 characteristic), “normal” (1.5%, with 2-3 characteristics) and “high” (11.8%,
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3 with ≥ 4 characteristics) risk of hospitalisation, with good accuracy (area under the
4 receiver operating characteristic curve 0.81, 95% CI 0.76-0.85).[12] Using
5 STARWAVE, clinicians can quickly and reliably identify the “high risk” cases that
6 might warrant antimicrobial treatment. More importantly, they can identify the “very
7 low risk” and “normal risk” cases that will likely resolve on their own, and spare them
8 unnecessary treatment.[12]
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11 STARWAVE is thus a prognostic (not a diagnostic) tool. It cannot tell clinicians
12 whether an infection is bacterial or viral. This does not however invalidate it as an
13 antimicrobial prescribing aid, because overprescribing is so often driven by
14 prognostic concerns.[12-15] STARWAVE recognises this and addresses it, by
15 providing evidence-based reassurance (to clinicians and perhaps even parents) that
16 specific children are *not* at significant risk. In so doing, it can assuage the fears and
17 anxieties that are known to trigger unnecessary prescriptions.
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20 Like other CPRs and clinical risk scores (e.g., QCancer), STARWAVE could be
21 integrated into the electronic health record to guide clinicians’ risk assessments and
22 prescribing decisions. In fact, one research group has incorporated web-based
23 STARWAVE decision support into a multifaceted intervention that aims to improve
24 the management of children presenting with cough in primary care (the intervention
25 is currently undergoing clinical trial).[16] As a rule, decision support should be
26 transparent and intelligible to the decision maker;[17] a risk score is merely a number
27 and could be ignored, especially if it contradicts the decision maker’s intuitive
28 assessment of risk.[18] Thus, it is important to understand whether and how GPs’
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3 intuitive risk assessments and prescribing decisions differ from those of STARWAVE,
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5 and how GPs interpret the CPR's clinical variables.
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10 To explore this, we presented GPs with clinical vignettes describing children
11 presenting with cough. The vignettes included all seven STARWAVE variables;
12 however, only four were manipulated (i.e., varied systematically across the
13 vignettes). This was due to logistical constraints: these data were collected in
14 conjunction with another study,[19] which limited the number of vignettes that we
15 could present and thus the number of variables that we could manipulate. We chose
16 to manipulate patient age (20 months vs. 5 years), illness duration (3 days vs. 6
17 days), vomiting (present vs. absent) and wheeze (present vs. absent), holding the
18 remaining STARWAVE variables constant (temperature, asthma, recession). Fever
19 was present in all of the vignettes, as it is a common presenting feature of childhood
20 RTIs.[12] Asthma and recession are both associated with airflow obstruction, but
21 wheeze (another symptom of airflow obstruction) was more common in the
22 STARWAVE cohort;[12] therefore we chose to manipulate wheeze, and kept asthma
23 and recession constant across vignettes (always absent). Per vignette, GPs
24 assessed risk of hospitalisation (very low, normal or high) and indicated whether they
25 would prescribe antibiotics or not. We compared GPs' intuitive risk assessments and
26 prescribing decisions to STARWAVE guidelines, and assessed the influence of the
27 manipulated STARWAVE variables.
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52 53 **METHOD**

54 55 56 57 58 **Participants** 59 60

Sample size

In the STARWAVE elicitation and validation study, a young age (<2 years), a short illness duration (≤ 3 days), vomiting (present vs. absent) and wheeze (present vs. absent) were found to increase the odds of hospitalisation two- to three-fold (OR range 2.16-3.42; all $p \leq 0.004$).^[12] We powered the present study to detect effects of the same size on the decision to prescribe antibiotics. Specifically, using G*Power 3.1, we estimated that in order to detect the smallest effect (OR 2.16) in a 2-tailed logistic regression of prescribing (yes vs. no) on the four manipulated factors (with power=80% and $\alpha=0.05$), 226 responses would be required.

Recruitment

By e-mail, we invited certified and practising UK GPs that had participated in previous studies by our research group. In addition, the NIHR-CRN (National Institute for Health Research Clinical Research Network) circulated our invitation e-mail to general practices across England.

Design and materials

Study materials were eight clinical vignettes that depicted children presenting to the GP with cough. Each child was described in terms of the seven STARWAVE variables. In a 2^{4-1} fractional factorial design, we manipulated patient age (20 months vs. 5 years), illness duration (3 days vs. 6 days), vomiting (present vs. absent) and

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3 wheeze (present vs. absent), holding the remaining variables constant (presence of
4 fever, absence of asthma and recession). We chose to use a fractional factorial
5 design (rather than a full factorial design) because it delivers clear estimates of main
6 effects, using half the number of vignettes (i.e., 8 rather than 16).[20]
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14 Risk of hospitalisation ranged from “very low” (vignette 1 in Appendix 1) to “high”
15 (vignette 8 in Appendix 1), but in most cases it was “normal” (vignettes 2-7 in
16 Appendix 1). Thus, only one vignette warranted a prescription according to
17 STARWAVE (vignette 8). Each participant was randomly assigned to view four of the
18 eight vignettes.
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28 **Procedure**

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33 Interested participants were e-mailed a link to the study website, where they read an
34 information sheet and provided informed consent. Thereafter, they saw 26 clinical
35 vignettes: two pertained to this study and 24 pertained to an unrelated study
36 conducted by our research group, concerning referral for suspected cancer.[19] The
37 two antibiotics vignettes were presented after 33% and 66% of the cancer vignettes,
38 respectively. The antibiotics and cancer vignettes were comparable in length and
39 difficulty.
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51 Twenty-four hours after completing this questionnaire, participants were e-mailed a
52 link to a second questionnaire, which was structured in the same way; i.e., two
53 antibiotics vignettes were evenly dispersed among 24 cancer vignettes. Importantly,
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3 the four antibiotics vignettes seen by a given participant were selected at random
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5 and presented in a random order.
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10 Following each antibiotics vignette, GPs were asked two questions:

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12 • *In your opinion, what is the risk that this child would deteriorate, requiring hospital*
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14 *admission?*
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16 ○ *very low risk, e.g. 1 in 300*
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18 ○ *medium risk, e.g. 1 in 70 (in STARWAVE, this level of risk is labelled*
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20 *“normal”)*
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22 ○ *high risk, e.g. 1 in 8*
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26 • *In your clinical judgement, what would be the best course of action?*
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28 ○ *no antibiotics prescription*
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30 ○ *antibiotics prescription*
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32 ○ *delayed antibiotics prescription*
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35 A delayed antibiotics prescription is a forward-dated prescription, intended for use by
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37 the patient if symptoms do not improve by the specified date. Delayed prescriptions
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39 form part of the national strategy to reduce immediate prescribing.[21] They were not
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41 the focus of the present study, but were included to ensure that the options available
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43 were representative of daily practice, and that our measure of immediate prescribing
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45 was precise, i.e., not skewed by the absence of an option that is typically present.
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51 Twenty-four hours later, participants were e-mailed a link to a third questionnaire;
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53 specifically, Gerrity et al.'s Stress from Uncertainty scale, which is one of the
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55 Physicians' Reactions to Uncertainty (PRU) scales.[22] The Stress from Uncertainty
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57 scale is a self-report measure of the extent to which physicians experience anxiety
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3 due to clinical uncertainty and concern about bad outcomes.[22] We expected that
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5 GPs who experience greater Stress from Uncertainty (SfU) would also experience
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7 greater prognostic uncertainty when assessing children with RTIs, and thus be more
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9 inclined to prescribe. GPs were asked to indicate their agreement with each of the
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11 scale's eight items (presented in a random order) on a six-point Likert scale
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13 anchored at 1="strongly disagree" and 6="strongly agree" (Appendix 2).
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19 **Analyses**

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23 To investigate the effect of the manipulated factors on risk assessments and
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25 prescribing decisions, two logistic regression models were built. The first was an
26
27 ordinal logistic regression model, where patient age (0=5 years, 1=20 months),
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29 illness duration (0=6 days, 1=3 days), vomiting (0=absent, 1=present) and wheeze
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31 (0=absent, 1=present) were used to predict perceived risk of hospitalisation (0=very
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33 low, 1=medium, 2=high). The second was a binary logistic regression model, where
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35 the same independent variables were used to predict prescribing decisions (0=no
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37 prescription, 1=prescription), which we dichotomised by merging "no prescription"
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39 and "delayed prescription" into a single category (national guidelines for antimicrobial
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41 prescribing treat them interchangeably [21]). For the interested reader, results
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43 pertaining to delayed prescriptions are presented in Appendix 3.
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51 In two further logistic regression models (one ordinal and one binary), we
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53 investigated whether SfU scores (summed across items per GP) might relate to risk
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55 assessments (0=very low, 1=medium, 2=high) and prescribing decisions (0=no
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57 prescription, 1=prescription).
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5 Statistical analysis was performed using Stata/MP 13.1. Specifically, the ordinal
6 analyses were conducted using the Stata user-written program “gologit2”, [23, 24]
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8 where we computed cluster-robust standard errors to account for repeated measures
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10 (multiple responses per GP). The binary analyses were conducted using Stata’s
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12 “melogit” command, [25] where we included a random intercept for GPs.
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19 **Patient and public involvement**

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24 Patients and members of the public were not involved in the design, execution,
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26 reporting or dissemination of this research.
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30 **Ethical approval**

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35 Ethical approval for this study was obtained from the Health Research Authority
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37 (reference number 18/HRA/0021) and research sponsorship was provided by
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39 Imperial College London (JRO reference 17IC3882). All aspects of the study were
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41 conducted in the UK in 2018.
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46 **RESULTS**

47 **Descriptive statistics**

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52 We collected data from 252 GPs, with an average of 15 years’ experience in general
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54 practice post-qualification (*SD* 9.8). Half of the sample was female (52%, 131/252).
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Eighty-six per cent were recruited via direct e-mail from the research team (217/252) and 14% via the NIHR-CRN (35/252).

Each GP saw four vignettes, yielding 1008 case presentations. GPs correctly classified risk of hospitalisation in 46% of these (461/1008; Table 1). Risk was rarely overestimated (9% of responses, 88/1008; blue cells) but frequently underestimated (46% of responses, 459/1008; green cells). Specifically, medium risk patients were classified as very low risk 46% of the time (345/756), while high risk patients were classified as very low or medium risk 90% of the time (114/126).

		Risk as classified by GPs			Total
		Very low	Medium	High	
STARWAVE risk	Very low	81	44	1	126
	Medium ("normal")	345	368	43	756
	High	33	81	12	126
Total		459	493	56	1008

Table 1. Association between risk as classified by GPs and as classified by STARWAVE.

GPs classified risk as high only 6% of the time (56/1008) but prescribed immediately 15% of the time (156/1008), suggesting a dissociation between risk assessments and prescribing decisions. Indeed, 78% of prescriptions were not consistent with GPs' own risk assessments (121/156; Table 2, blue cells) and 83% were not consistent with STARWAVE risk assessments (130/156; Table 2, green cells).

		Risk as classified by GPs			STARWAVE risk			Total
		Very low	Medium	High	Very low	Medium ("normal")	High	
Prescriptions	None/delayed	445	386	21	112	640	100	852
	Immediate	14	107	35	14	116	26	156
Total		459	493	56	126	756	126	1008

Table 2. Association between risk (as classified by GPs and by STARWAVE) and prescribing decisions.

Appendix 4 presents the number and proportion of prescriptions per vignette. The case with the highest rate of prescription was not the high risk case, which received a prescription only 21% of the time (26/126; vignette 8). Rather, it was a medium risk case, describing a 5-year-old child with a 6-day illness duration who had both vomiting and wheeze (33%, 42/126; vignette 7).

Results of planned analyses

Younger patient age (20 months vs. 5 years) increased perceived risk of hospitalisation (OR 1.49, 95% CI 1.14-1.95, $p=0.003$), while a short illness duration decreased it (OR 0.54, 95% CI 0.42-0.69, $p<0.001$). Presence of vomiting and presence of wheeze were both associated with higher risk estimates (OR_{vomit} 1.92, 95% CI 1.57-2.36, $p<0.001$; OR_{wheeze} 3.33, 95% CI 2.66-4.16, $p<0.001$). Statistical tests of the proportional odds assumption revealed that all four variables met it; i.e., the effect of each independent variable was consistent for successive levels of the ordinal dependent variable (all $ps\geq 0.099$). A global Wald test confirmed that the proportional odds assumption was not violated in this model ($\chi^2(4) 4.70$, $p=0.320$).

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3 Patient age did not influence the odds of a prescription (OR 1.42, 95% CI 0.83-2.42,
4 $p=0.201$), but a short illness duration decreased them (OR 0.14, 95% CI 0.08-0.27,
5 $p<0.001$). Presence of vomiting and presence of wheeze both increased prescribing
6 odds (OR_{vomit} 2.17, 95% CI 1.32-3.57, $p=0.002$; OR_{wheeze} 8.98, 95% CI 4.99-16.15,
7 $p<0.001$). Presence of vomiting and presence of wheeze both increased prescribing
8 odds (OR_{vomit} 2.17, 95% CI 1.32-3.57, $p=0.002$; OR_{wheeze} 8.98, 95% CI 4.99-16.15,
9 $p<0.001$). When prescribing was treated as a 3-category ordinal variable (0=no
10 prescription, 1=delayed prescription, 2=immediate prescription), these findings did
11 not change (Appendix 3).
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22 SfU scores were unrelated to risk assessments (OR 1.00, 95% CI 0.98-1.02,
23 $p=0.935$; proportional odds assumption met with $p_{SfU}=0.406$) and prescribing
24 decisions (OR 1.00, 95% CI 0.96-1.03, $p=0.875$).
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30 DISCUSSION

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35 We compared GPs' risk assessments and antimicrobial prescribing decisions to a
36 normative model (the STARWAVE CPR), in the context of clinical vignettes that
37 varied the features (age, illness duration, vomiting, wheeze) of children presenting
38 with cough. Relative to STARWAVE, GPs frequently underestimated the patient's
39 risk of deterioration, but nonetheless overprescribed: the vast majority of their
40 prescriptions were unnecessary relative to their own risk assessments (78%) and
41 STARWAVE risk assessments (83%).
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53 All four of the manipulated variables influenced GPs' risk assessments, which
54 increased when the child was younger (20 months vs. 5 years), when illness duration
55 was longer (6 vs. 3 days) and when vomiting and/or wheeze were present (vs.
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absent). Comparing the odds ratios for these relationships to the STARWAVE model (Table 3), we note both similarities and discrepancies. Specifically, GPs' interpretations of patient age, vomiting and wheeze were consistent with the STARWAVE model, but their interpretation of illness duration was not: a shorter illness duration reduced – rather than increased – GP estimates of risk.

Predictor	OR _{GPs}	OR _{STARWAVE}
Age (<2 years)	1.49 [1.14-1.95]*	3.42 [2.12-5.58]*
Duration (≤3 days)	0.54 [0.42-0.69]*	2.77 [1.77-4.35]*
Vomiting	1.92 [1.57-2.36]*	2.56 [1.54-4.31]*
Wheeze	3.33 [2.66-4.16]*	2.16 [1.28-3.60]*
Temperature		1.99 [1.22-3.25]*
Asthma		3.93 [2.20-7.03]*
Recession		3.82 [2.23-6.62]*

Table 3. The effect of patient age, illness duration, vomiting and wheeze on risk of hospitalisation, according to present participants (OR_{GPs}) and STARWAVE (OR_{STARWAVE}). * $p \leq 0.006$. Square brackets contain 95% CIs.

Like risk perceptions, prescribing increased when illness duration was long (inverted OR 7.14) and when vomiting and/or wheeze were present (OR_{vomit} 2.17; OR_{wheeze} 8.98). Patient age had no reliable effect on prescribing (OR 1.42). Again, these findings are not entirely consistent with the STARWAVE model, but they are consistent with previous, non-experimental research. In one interview study, for example, GPs reported that they were more likely to prescribe antibiotics to children with RTIs given prolonged duration of symptoms, abnormal chest signs and (less frequently) vomiting.[13] Various observational studies have likewise identified chest

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3 abnormalities [26-30] and vomiting [29] as clinical characteristics that prompt
4 prescribing. In contrast, previous literature concerning the effect of age on
5 prescribing is mixed: two studies found that older (vs. younger) patients were more
6 likely to receive a prescription,[28, 31] but three identified no association between
7 age and prescribing.[27, 29, 30]

17 Interestingly, the one patient that warranted a prescription received one only 21% of
18 the time. Thus, we identified not only overprescription but underprescription too.
19 Underprescription has been detected in previous studies; for example, one
20 observational study (of adults presenting to their GP with cough and RTI) found that
21 16% of patients with a bacterial infection (pneumonia) did not receive a
22 prescription.[29] Presently, the rate of underprescription was considerably higher
23 (79%) and likely due to the patient's short illness duration (3 days). However,
24 conclusions cannot be drawn on the basis of a single vignette; underprescription
25 requires further investigation in a larger and more varied set of cases.

39 Risk assessments and prescribing tendencies bore no association to GPs' self-
40 reported levels of "Stress from Uncertainty". However, Grol and colleagues found
41 that greater willingness to take risks (as measured on their Attitudes to Risk Taking
42 scale) was associated with significantly fewer antibiotics prescriptions for respiratory
43 problems and URTI/common cold.[32] Attitudes toward risk – rather than attitudes
44 toward uncertainty – may thus prove a fruitful avenue for future research.

56 **Limitations and future work**

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3 The rate of prescription identified here (15%) appears lower than that observed in
4 other studies. For example, Hay et al. identified a rate of 37% in their prospective
5 cohort study of children presenting to the GP with cough.[12] Notably, present work
6 included few high risk presentations (13%), but high risk presentations were likewise
7 infrequent in the study by Hay et al. (3%).[12] The likely explanation for the lower
8 prescription rate identified here is that our vignettes ignored the complex
9 interpersonal dynamics of the clinical encounter, which are known to influence
10 prescribing behaviour.[13-15, 33, 34] For example, prescription likelihood is
11 increased by perceived pressure from patients/parents to prescribe;[14, 30, 33, 35,
12 36] by the desire to maintain good relationships with patients/parents;[13, 37, 38] by
13 fear of medicolegal problems;[13, 15, 38] and by time pressure.[13, 14, 36-38]
14 Importantly, these factors can be incorporated into clinical vignettes, as
15 demonstrated by Sirota and colleagues; these authors found that prescriptions were
16 twice as likely when patient pressure for antibiotics was present (vs. absent) from a
17 clinical vignette.[33] On the one hand, it is a limitation of our vignettes that these
18 interpersonal factors were absent; on the other, our work highlights that antibiotics
19 are overprescribed *even when* these interpersonal factors are absent. It is worrying
20 that so many GPs considered antibiotics to be the most appropriate course of action,
21 not simply the most expedient one. Qualitative research may be useful to understand
22 why GPs prescribed to patients that they deemed to be low risk, in the absence of
23 any interpersonal pressure to do so.
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54 Data for this study were collected in conjunction with another project, which limited
55 the number of STARWAVE variables that we could manipulate. A comprehensive
56 investigation of all seven STARWAVE variables would undoubtedly return new and
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3 valuable insights. Future investigations might also treat the continuous STARWAVE
4 variables (age and illness duration) as continuous (not binary), to test the
5 generalisability of the trends identified here.
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12 Despite these limitations, present work sheds light on the determinants of antibiotic
13 prescribing in child RTI presentations, bringing much-needed experimental evidence
14 to a literature that has to date relied predominantly on self-report [13-15, 31, 36, 38,
15 39] and observational [26-30] data. It also speaks to the difficulties that may be
16 encountered if STARWAVE is provided as a decision aid to GPs. Firstly, GPs'
17 classification of risk in this study was largely incompatible with STARWAVE's; GPs
18 consistently chose lower risk than STARWAVE would suggest. Still, they prescribed
19 more frequently than STARWAVE risk classification would support. Presenting GPs
20 with STARWAVE's risk classification will likely exacerbate prescribing (since GPs
21 overprescribed with their own, lower classifications of risk). Presenting them with a
22 recommendation may also be ineffective, unless the recommendation is
23 accompanied by an explanation. Explaining the recommendation in terms of the
24 variables that increase/decrease a child's risk of hospitalisation may be a way
25 forward, and enable GPs to understand why their own intuitive decision might differ
26 from the recommendation. Identifying the factors that are likely to be misinterpreted
27 by GPs is important when explaining the rationale behind recommendations.
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51 **COMPETING INTERESTS**

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56 Dr. Nurek, Dr. Delaney and Dr. Kostopoulou report grants from the NIHR Imperial
57 Patient Safety Translational Research Centre, during the conduct of the study.
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AUTHOR CONTRIBUTIONS

All authors contributed to the design of the study. MN performed the data collection; MN and OK performed the data analysis. MN drafted the manuscript; OK and BD provided critical revision.

DATA SHARING

For peer review purposes, the data are provided at https://osf.io/r3ype/?view_only=a66a7be2fcbe45a5a67a454ba5b3750a. Upon acceptance, the data will be made publicly available on the Open Science Framework under a CC-By Attribution 4.0 International Licence.

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APPENDIX 1

Characteristics of vignettes

Vignette identifier	Vignette Text	STARWAVE risk assessment	STARWAVE recommendation
1	A 5-year-old child is brought by their parent with a 6-day history of cough. They have no history of asthma or vomiting. On examination, they are pyrexial, and have neither wheeze nor intercostal recession.	Very low (1 risk factor)	No immediate prescription
2	A 20-month-old child is brought by their parent with a 3-day history of cough. They have no history of asthma or vomiting. On examination, they are pyrexial, and have neither wheeze nor intercostal recession.	Normal (3 risk factors)	No immediate prescription
3	A 20-month-old child is brought by their parent with a 6-day history of cough. They have no history of asthma and have vomited twice in the last 24 hours. On examination, they are pyrexial, and have neither wheeze nor intercostal recession.	Normal (3 risk factors)	No immediate prescription
4	A 5-year-old child is brought by their parent with a 3-day history of cough. They have no history of asthma and have vomited twice in the last 24 hours. On examination, they are pyrexial, and have neither wheeze nor intercostal recession.	Normal (3 risk factors)	No immediate prescription
5	A 20-month-old child is brought by their parent with a 6-day history of cough. They have no history of asthma or vomiting. On examination, they are pyrexial, and have wheeze on chest auscultation. There is no intercostal recession.	Normal (3 risk factors)	No immediate prescription

6	A 5-year-old child is brought by their parent with a <u>3-day</u> history of cough. They have no history of asthma or vomiting. On examination, they are <u>pyrexial</u> , and <u>have wheeze</u> on chest auscultation. There is no intercostal recession.	Normal (3 risk factors)	No immediate prescription
7	A 5-year-old child is brought by their parent with a 6-day history of cough. They have no history of asthma and <u>have vomited</u> twice in the last 24 hours. On examination, they are <u>pyrexial</u> , and <u>have wheeze</u> on chest auscultation. There is no intercostal recession.	Normal (3 risk factors)	No immediate prescription
8	A <u>20-month-old</u> child is brought by their parent with a <u>3-day</u> history of cough. They have no history of asthma and <u>have vomited</u> twice in the last 24 hours. On examination, they are <u>pyrexial</u> , and <u>have wheeze</u> on chest auscultation. There is no intercostal recession.	High (5 risk factors)	Consider an immediate prescription

Note: STARWAVE risk factors are underlined for salience. They were not underlined when vignettes were presented to GPs.

APPENDIX 2

Stress from Uncertainty (SfU) scale

1. I usually feel anxious when I am not sure of a diagnosis.
2. I find the uncertainty involved in patient care disconcerting.
3. Uncertainty in patient care makes me uneasy.
4. I am quite comfortable with the uncertainty in patient care.*
5. The uncertainty of patient care often troubles me.
6. When I am uncertain of a diagnosis, I imagine all sorts of bad scenarios -- patient dies, patient sues, etc.
7. I fear being held accountable for the limits of my knowledge.
8. I worry about malpractice when I do not know a patient's diagnosis.

Note: items 1-5 measure the construct "Anxiety due to Uncertainty" (Cronbach's alpha=0.86); items 6-8 measure the construct "Concern About Bad Outcomes" (Cronbach's alpha=0.73). *Reverse-scored item.

APPENDIX 3

Delayed prescriptions

Delayed prescriptions were administered 12% of the time (121/1008). The association between delayed prescribing and risk (as classified by GPs and by STARWAVE) is displayed below (yellow cells).

		Risk as classified by GPs			STARWAVE risk			Total
		Very low	Medium	High	Very low	Medium ("normal")	High	
Prescriptions	None	420	294	17	95	551	85	731
	Delayed	25	92	4	17	89	15	121
	Immediate	14	107	35	14	116	26	156
Total		459	493	56	126	756	126	1008

To investigate the effect of the manipulated factors on both delayed and immediate prescribing, we regressed the 3-category prescribing variable (0=no prescription, 1=delayed prescription, 2=immediate prescription) on patient age (0=5 years, 1=20 months), illness duration (0=6 days, 1=3 days), vomiting (0=absent, 1=present) and wheeze (0=absent, 1=present). This ordinal logistic regression analysis was conducted using the Stata user-written program "gologit2".[23, 24] Statistical tests of the proportional odds assumption revealed that two variables met it ($p_{\text{age}}=0.124$ and $p_{\text{vomit}}=0.522$) and two did not ($p_{\text{duration}}=0.034$ and $p_{\text{wheeze}}=0.003$). Put differently: the respective effects of age and vomiting were consistent for successive levels of the ordinal dependent variable, while those of duration and wheeze were not. Thus, we constructed a partial proportional odds (PPO) model, where two coefficients were fixed (age and vomiting) and two were allowed to vary (duration and wheeze). A global Wald test confirmed that the proportional odds assumption was not violated in this PPO model ($\chi^2(2) 2.63, p=0.268$).

Results are tabulated below. The model progresses in two steps: the first step compares “no prescription” (coded 0) to “delayed prescription” and “immediate prescription” (both coded 1); the second compares “no prescription” and “delayed prescription” (both coded 0) to “immediate prescription” (coded 1). Trends were consistent across steps, and consistent with those reported in the main text. Specifically, patient age did not influence the odds of a prescription ($p=0.569$) and short illness duration decreased them ($p<0.001$). Presence of vomiting and presence of wheeze both increased prescribing odds (both $ps<0.001$). Two coefficients were allowed to vary across steps (duration and wheeze): in both cases, effects grew stronger from step 1 to step 2.

	STEP 1: no prescription (coded 0) vs. delayed/immediate (coded 1)	STEP 2: no/delayed prescription (coded 0) vs. immediate (coded 1)
Age (<2 years)	0.92 [0.69-1.23]	0.92 [0.69-1.23]
Duration (≤ 3 days)	0.46 [0.34-0.62]*	0.34 [0.24-0.49]*
Vomiting	1.49 [1.24-1.80]*	1.49 [1.24-1.80]*
Wheeze	2.50 [1.91-3.28]*	3.89 [2.66-5.69]*

* $p<0.001$. Cells contain odds ratios; square brackets contain 95% CIs. Step 2 of the model (no/delayed prescription vs. immediate) is akin to the model reported in the main text; differences in coefficients may be attributed to different estimation procedures (e.g., the ordinal model estimates all parameters simultaneously).[23]

APPENDIX 4**Risk assessments and prescribing decisions per vignette**

Vignette identifier	GP risk assessments			GP prescribing decisions		
	Very low	Medium	High	None	Delayed	Immediate
1	64% (81/126)	35% (44/126)	1% (1/126)	75% (95/126)	14% (17/126)	11% (14/126)
2	71% (89/126)	29% (37/126)	0% (0/126)	94% (118/126)	5% (6/126)	2% (2/126)
3	42% (52/125)	53% (66/125)	6% (7/125)	74% (93/125)	13% (16/125)	13% (16/125)
4	59% (75/127)	40% (51/127)	1% (1/127)	81% (103/127)	14% (18/127)	5% (6/127)
5	25% (32/127)	62% (79/127)	13% (16/127)	59% (75/127)	11% (14/127)	30% (38/127)
6	53% (66/125)	44% (55/125)	3% (4/125)	77% (96/125)	14% (17/125)	10% (12/125)
7	25% (31/126)	64% (80/126)	12% (15/126)	52% (66/126)	14% (18/126)	33% (42/126)
8	26% (33/126)	64% (81/126)	10% (12/126)	67% (85/126)	12% (15/126)	21% (26/126)

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Title:

Risk assessment and antibiotic prescribing decisions in children presenting to UK primary care with cough: a vignette study

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ABSTRACT

Objectives: The validated “STARWAVE” clinical prediction rule (CPR) uses seven variables to guide risk assessment and antimicrobial stewardship in children presenting with cough (**S**hort illness duration, **T**emperature, **A**ge, **R**ecession, **W**heeze, **A**sthma, **V**omiting). We aimed to compare General Practitioners’ (GPs) risk assessments and prescribing decisions to those of STARWAVE, and assess the influence of the CPR’s clinical variables.

Setting: Primary care.

Participants: 252 GPs, currently practising in the UK.

Design: GPs were randomly assigned to view four (of a possible eight) clinical vignettes online. Each vignette depicted a child presenting with cough, who was described in terms of the seven STARWAVE variables. Systematically, we manipulated patient age (20 months vs. 5 years), illness duration (3 vs. 6 days), vomiting (present vs. absent) and wheeze (present vs. absent), holding the remaining STARWAVE variables constant.

Outcome measures: Per vignette, GPs assessed risk of hospitalisation and indicated whether they would prescribe antibiotics or not.

Results: GPs overestimated risk of hospitalisation in 9% of vignette presentations (88/1008) and underestimated it in 46% (459/1008). Despite underestimating risk, they

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3 overprescribed: 78% of prescriptions were unnecessary relative to GPs' own risk
4 assessments (121/156), while 83% were unnecessary relative to STARWAVE risk
5 assessments (130/156). All four of the manipulated variables influenced risk
6 assessments, but only three influenced prescribing decisions: a shorter illness
7 assessments, but only three influenced prescribing decisions: a shorter illness
8 duration reduced prescribing odds (OR 0.14, 95% CI 0.08-0.27, $p<0.001$), while
9 vomiting and wheeze increased them (OR_{vomit} 2.17, 95% CI 1.32-3.57, $p=0.002$;
10 OR_{wheeze} 8.98, 95% CI 4.99-16.15, $p<0.001$).

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21 **Conclusions:** Relative to STARWAVE, GPs underestimated risk of hospitalisation,
22 overprescribed, and appeared to misinterpret illness duration (prescribing for longer
23 rather than shorter illnesses). It is important to ascertain discrepancies between CPRs
24 and current clinical practice. This has implications for the integration of CPRs into the
25 electronic health record and the provision of intelligible explanations to decision
26 makers.
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STRENGTHS AND LIMITATIONS OF THIS STUDY

- This is the first study to suggest discrepancies between the STARWAVE clinical prediction rule and current clinical practice.
- Use of clinical vignettes allowed us to manipulate some variables while holding others constant; thus we could identify *causal* relationships between specific clinical variables and antibiotic prescribing decisions.
- In so doing, we bring much-needed experimental evidence to the literature, which is currently dominated by interview and observational studies.
- The disadvantage of using clinical vignettes is that our results are based on hypothetical clinical scenarios, which contained limited information.
- Moreover, we manipulated only a subset of the STARWAVE variables; future work could increase the number of clinical variables manipulated, and explore non-clinical factors too.

INTRODUCTION

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8 Combatting antimicrobial resistance is high on policy agendas internationally.[1-
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10 3] One of the key means advocated is judicious antibiotic prescribing.[1] Over 80% of
11
12 all NHS antibiotic prescriptions are issued in primary care,[4] where despite numerous
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14 campaigns, mandates and financial incentives, rates remain unacceptably
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16 high.[5] Despite strong evidence of only modest symptomatic benefits for acute
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18 respiratory tract infections (RTIs),[6-8] and even smaller effects against
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20 complications,[9, 10] RTIs are the most common justification for primary care antibiotic
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22 use [11] and a leading cause of overuse.[12] This is exacerbated in children, where
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24 perceived vulnerability and prognostic uncertainty (i.e., perceived risk of deterioration)
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26 lead to defensive prescribing (“treat, just in case”).[12-15]
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33 To improve risk assessment and antimicrobial prescribing in children with RTIs, a
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35 clinical prediction rule (CPR) called “STARWAVE” was recently developed and
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37 validated.[12] It was based on a large prognostic cohort study, which included 8394
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39 children presenting to 247 general practices in England with acute cough and RTI
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41 symptoms.[12] Numerous characteristics were recorded at presentation, including
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43 demographic variables, parent-reported symptoms and physical examination signs. In
44
45 a regression analysis, seven of these characteristics were found to predict hospital
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47 admission (for RTI) in the month following presentation: **Short illness duration** (≤ 3
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49 days), **Temperature** ($\geq 37.8^{\circ}\text{C}$), **Age** (< 2 years), **Recession**, **Wheeze**, **Asthma** and
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51 **Vomiting**. [12] This analysis gave rise to the “STARWAVE” clinical prediction rule: a
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53 seven-item, point-of-care checklist that can distinguish children at “very low” (0.3%,
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55 with ≤ 1 characteristic), “normal” (1.5%, with 2-3 characteristics) and “high” (11.8%,
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3 with ≥ 4 characteristics) risk of hospitalisation, with good accuracy (area under the
4 receiver operating characteristic curve 0.81, 95% CI 0.76-0.85).[12] Using
5 STARWAVE, clinicians can quickly and reliably identify the “high risk” cases that might
6 warrant antimicrobial treatment. More importantly, they can identify the “very low risk”
7 and “normal risk” cases that will likely resolve on their own, and spare them
8 unnecessary treatment.[12]
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19 STARWAVE is thus a prognostic (not a diagnostic) tool. It cannot tell clinicians whether
20 an infection is bacterial or viral. This does not however invalidate it as an antimicrobial
21 prescribing aid, because overprescribing is so often driven by prognostic
22 concerns.[12-15] STARWAVE recognises this and addresses it, by providing
23 evidence-based reassurance (to clinicians and perhaps even parents) that specific
24 children are *not* at significant risk. In so doing, it can assuage the fears and anxieties
25 that are known to trigger unnecessary prescriptions.
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38 Like other CPRs and clinical risk scores (e.g., QCancer), STARWAVE could be
39 integrated into the electronic health record to guide clinicians’ risk assessments and
40 prescribing decisions. In fact, one research group has incorporated web-based
41 STARWAVE decision support into a multifaceted intervention that aims to improve the
42 management of children presenting with cough in primary care (the intervention is
43 currently undergoing clinical trial).[16] As a rule, decision support should be
44 transparent and intelligible to the decision maker;[17] a risk score is merely a number
45 and could be ignored, especially if it contradicts the decision maker’s intuitive
46 assessment of risk.[18] Thus, it is important to understand whether and how GPs’
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3 intuitive risk assessments and prescribing decisions differ from those of STARWAVE,
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5 and how GPs interpret the CPR's clinical variables.
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10 To explore this, we presented GPs with clinical vignettes describing children
11 presenting with cough. The vignettes included all seven STARWAVE variables;
12 however, only four were manipulated (i.e., varied systematically across the vignettes).
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14 This was due to logistical constraints: these data were collected in conjunction with
15 another study, which limited the number of vignettes that we could present and thus
16 the number of variables that we could manipulate. We chose to manipulate patient
17 age (20 months vs. 5 years), illness duration (3 days vs. 6 days), vomiting (present vs.
18 absent) and wheeze (present vs. absent), holding the remaining STARWAVE
19 variables constant (temperature, asthma, recession). Fever was present in all of the
20 vignettes, as it is a common presenting feature of childhood RTIs.[12] Asthma and
21 recession are both associated with airflow obstruction, but wheeze (another symptom
22 of airflow obstruction) was more common in the STARWAVE cohort;[12] therefore we
23 chose to manipulate wheeze, and kept asthma and recession constant across
24 vignettes (always absent). Per vignette, GPs assessed risk of hospitalisation (very low,
25 normal or high) and indicated whether they would prescribe antibiotics or not. We
26 compared GPs' intuitive risk assessments and prescribing decisions to STARWAVE
27 guidelines, and assessed the influence of the manipulated STARWAVE variables.
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51 **METHOD**

52 53 54 55 56 **Participants** 57 58 59 60

Sample size

In the STARWAVE elicitation and validation study, a young age (<2 years), a short illness duration (≤ 3 days), vomiting (present vs. absent) and wheeze (present vs. absent) were found to increase the odds of hospitalisation two- to three-fold (OR range 2.16-3.42; all $ps \leq 0.004$).^[12] We powered the present study to detect effects of the same size on the decision to prescribe antibiotics. Specifically, using G*Power 3.1, we estimated that in order to detect the smallest effect (OR 2.16) in a 2-tailed logistic regression of prescribing (yes vs. no) on the four manipulated factors (with power=80% and $\alpha=0.05$), 226 responses would be required.

Recruitment

By e-mail, we invited certified and practising UK GPs that had participated in previous studies by our research group. In addition, the NIHR-CRN (National Institute for Health Research Clinical Research Network) circulated our invitation e-mail to general practices across England.

Design and materials

Study materials were eight clinical vignettes that depicted children presenting to the GP with cough. Each child was described in terms of the seven STARWAVE variables. In a 2^{4-1} fractional factorial design, we manipulated patient age (20 months vs. 5 years), illness duration (3 days vs. 6 days), vomiting (present vs. absent) and wheeze (present vs. absent), holding the remaining variables constant (presence of fever,

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3 absence of asthma and recession). We chose to use a fractional factorial design
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5 (rather than a full factorial design) because it delivers clear estimates of main effects,
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7 using half the number of vignettes (i.e., 8 rather than 16).[19]
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12 Risk of hospitalisation ranged from “very low” (vignette 1 in Appendix 1) to “high”
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14 (vignette 8 in Appendix 1), but in most cases it was “normal” (vignettes 2-7 in Appendix
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16 1). Thus, only one vignette warranted a prescription according to STARWAVE
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18 (vignette 8). Each participant was randomly assigned to view four of the eight
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20 vignettes.
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23 24 25 26 **Procedure** 27

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30 Interested participants were e-mailed a link to the study website, where they read an
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32 information sheet and provided informed consent. Thereafter, they saw 26 clinical
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34 vignettes: two pertained to this study and 24 pertained to an unrelated study conducted
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36 by our research group, concerning referral for suspected cancer. The two antibiotics
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38 vignettes were presented after 33% and 66% of the cancer vignettes respectively, and
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40 were introduced as follows: “*We understand that this is somewhat monotonous, so*
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42 *here is something quite different to help you re-engage attention*”. The antibiotics and
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44 cancer vignettes were comparable in length and difficulty.
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52 Twenty-four hours after completing this questionnaire, participants were e-mailed a
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54 link to a second questionnaire, which was structured in the same way; i.e., two
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56 antibiotics vignettes were evenly dispersed among 24 cancer vignettes. Importantly,
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3 the four antibiotics vignettes seen by a given participant were selected at random and
4 presented in a random order.
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10 Following each antibiotics vignette, GPs were asked two questions:

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12 • *In your opinion, what is the risk that this child would deteriorate, requiring hospital*
13 *admission?*
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15 ○ *very low risk, e.g. 1 in 300*
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17 ○ *medium risk, e.g. 1 in 70 (in STARWAVE, this level of risk is labelled*
18 *“normal”)*
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20 ○ *high risk, e.g. 1 in 8*
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26 • *In your clinical judgement, what would be the best course of action?*
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28 ○ *no antibiotics prescription*
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30 ○ *antibiotics prescription*
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32 ○ *delayed antibiotics prescription*
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35 A delayed antibiotics prescription is a forward-dated prescription, intended for use by
36 the patient if symptoms do not improve by the specified date. Delayed prescriptions
37 form part of the national strategy to reduce immediate prescribing.[20] They were not
38 the focus of the present study, but were included to ensure that the options available
39 were representative of daily practice, and that our measure of immediate prescribing
40 was precise, i.e., not skewed by the absence of an option that is typically present.
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51 Twenty-four hours later, participants were e-mailed a link to a third questionnaire;
52 specifically, Gerrity et al.'s Stress from Uncertainty scale, which is one of the
53 Physicians' Reactions to Uncertainty (PRU) scales.[21] The Stress from Uncertainty
54 scale is a self-report measure of the extent to which physicians experience anxiety
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3 due to clinical uncertainty and concern about bad outcomes.[21] We expected that
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5 GPs who experience greater Stress from Uncertainty (SfU) would also experience
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7 greater prognostic uncertainty when assessing children with RTIs, and thus be more
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9 inclined to prescribe. GPs were asked to indicate their agreement with each of the
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11 scale's eight items (presented in a random order) on a six-point Likert scale anchored
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13 at 1="strongly disagree" and 6="strongly agree" (Appendix 2).
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19 **Analyses**

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23 To investigate the effect of the manipulated factors on risk assessments and
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25 prescribing decisions, two logistic regression models were built. The first was an
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27 ordinal logistic regression model, where patient age (0=5 years, 1=20 months), illness
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29 duration (0=6 days, 1=3 days), vomiting (0=absent, 1=present) and wheeze
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31 (0=absent, 1=present) were used to predict perceived risk of hospitalisation (0=very
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33 low, 1=medium, 2=high). The second was a binary logistic regression model, where
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35 the same independent variables were used to predict prescribing decisions (0=no
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37 prescription, 1=prescription), which we dichotomised by merging "no prescription" and
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39 "delayed prescription" into a single category (national guidelines for antimicrobial
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41 prescribing treat them interchangeably [20]). For the interested reader, results
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43 pertaining to delayed prescriptions are presented in Appendix 3.
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52 In two further logistic regression models (one ordinal and one binary), we investigated
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54 whether SfU scores (summed across items per GP) might relate to risk assessments
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56 (0=very low, 1=medium, 2=high) and prescribing decisions (0=no prescription,
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58 1=prescription).
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5 Statistical analysis was performed using Stata/MP 13.1. Specifically, the ordinal
6 analyses were conducted using the Stata user-written program “gologit2”, [22, 23]
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8 where we computed cluster-robust standard errors to account for repeated measures
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10 (multiple responses per GP). The binary analyses were conducted using Stata’s
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12 “melogit” command, [24] where we included a random intercept for GPs.
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19 **Patient and public involvement**

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24 Patients and members of the public were not involved in the design, execution,
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26 reporting or dissemination of this research.
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30 **Ethical approval**

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35 Ethical approval for this study was obtained from the Health Research Authority
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37 (reference number 18/HRA/0021) and research sponsorship was provided by Imperial
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39 College London (JRO reference 17IC3882). All aspects of the study were conducted
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41 in the UK in 2018.
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46 **RESULTS**

47 **Descriptive statistics**

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52 We collected data from 254 GPs. Of these, two gave only partial data and thus were
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54 excluded from the analyses. The final sample comprised 252 GPs, with an average of
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15 years' experience in general practice post-qualification (*SD* 9.8). Half of the sample was female (52%, 131/252). Eighty-six per cent were recruited via direct e-mail from the research team (217/252) and 14% via the NIHR-CRN (35/252).

Each GP saw four vignettes, yielding 1008 case presentations. GPs correctly classified risk of hospitalisation in 46% of these (461/1008; Table 1). Risk was rarely overestimated (9% of responses, 88/1008; blue cells) but frequently underestimated (46% of responses, 459/1008; green cells). Specifically, medium risk patients were classified as very low risk 46% of the time (345/756), while high risk patients were classified as very low or medium risk 90% of the time (114/126).

		Risk as classified by GPs			Total
		Very low	Medium	High	
STARWAVE risk	Very low	81	44	1	126
	Medium ("normal")	345	368	43	756
	High	33	81	12	126
Total		459	493	56	1008

Table 1. Association between risk as classified by GPs and as classified by STARWAVE.

GPs classified risk as high only 6% of the time (56/1008) but prescribed immediately 15% of the time (156/1008), suggesting a dissociation between risk assessments and prescribing decisions. Indeed, 78% of prescriptions were not consistent with GPs' own risk assessments (121/156; Table 2, blue cells) and 83% were not consistent with STARWAVE risk assessments (130/156; Table 2, green cells).

		Risk as classified by GPs			STARWAVE risk			Total
		Very low	Medium	High	Very low	Medium ("normal")	High	
Prescriptions	None/delayed	445	386	21	112	640	100	852
	Immediate	14	107	35	14	116	26	156
Total		459	493	56	126	756	126	1008

Table 2. Association between risk (as classified by GPs and by STARWAVE) and prescribing decisions.

Appendix 4 presents the number and proportion of prescriptions per vignette. The case with the highest rate of prescription was not the high risk case, which received a prescription only 21% of the time (26/126; vignette 8). Rather, it was a medium risk case, describing a 5-year-old child with a 6-day illness duration who had both vomiting and wheeze (33%, 42/126; vignette 7).

Results of planned analyses

Younger patient age (20 months vs. 5 years) increased perceived risk of hospitalisation (OR 1.49, 95% CI 1.14-1.95, $p=0.003$), while a short illness duration decreased it (OR 0.54, 95% CI 0.42-0.69, $p<0.001$). Presence of vomiting and presence of wheeze were both associated with higher risk estimates (OR_{vomit} 1.92, 95% CI 1.57-2.36, $p<0.001$; OR_{wheeze} 3.33, 95% CI 2.66-4.16, $p<0.001$). Statistical tests of the proportional odds assumption revealed that all four variables met it; i.e., the effect of each independent variable was consistent for successive levels of the ordinal dependent variable (all $ps\geq 0.099$). A global Wald test confirmed that the proportional odds assumption was not violated in this model ($\chi^2(4) 4.70$, $p=0.320$).

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3 Patient age did not influence the odds of a prescription (OR 1.42, 95% CI 0.83-2.42,
4 $p=0.201$), but a short illness duration decreased them (OR 0.14, 95% CI 0.08-0.27,
5 $p<0.001$). Presence of vomiting and presence of wheeze both increased prescribing
6 odds (OR_{vomit} 2.17, 95% CI 1.32-3.57, $p=0.002$; OR_{wheeze} 8.98, 95% CI 4.99-16.15,
7 $p<0.001$). Presence of vomiting and presence of wheeze both increased prescribing
8 odds (OR_{vomit} 2.17, 95% CI 1.32-3.57, $p=0.002$; OR_{wheeze} 8.98, 95% CI 4.99-16.15,
9 $p<0.001$). When prescribing was treated as a 3-category ordinal variable (0=no
10 prescription, 1=delayed prescription, 2=immediate prescription), these findings did not
11 change (Appendix 3).
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22 SfU scores were unrelated to risk assessments (OR 1.00, 95% CI 0.98-1.02, $p=0.935$;
23 proportional odds assumption met with $p_{SfU}=0.406$) and prescribing decisions (OR
24 1.00, 95% CI 0.96-1.03, $p=0.875$).
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30 DISCUSSION

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35 We compared GPs' risk assessments and antimicrobial prescribing decisions to a
36 normative model (the STARWAVE CPR), in the context of clinical vignettes that varied
37 the features (age, illness duration, vomiting, wheeze) of children presenting with
38 cough. Relative to STARWAVE, GPs frequently underestimated the patient's risk of
39 deterioration, but nonetheless overprescribed: the vast majority of their prescriptions
40 were unnecessary relative to their own risk assessments (78%) and STARWAVE risk
41 assessments (83%).
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54 This is not the first study to observe a disconnect between physicians' risk
55 assessments and antimicrobial prescribing decisions. In one study, for example, an
56 educational intervention was successful in reducing physicians' overestimations of the
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3 likelihood of a bacterial infection, but unsuccessful in reducing antibiotic
4 prescribing.[25] In another, patient expectations for antibiotics increased physicians'
5 rates of antibiotic prescribing, but did not influence their probability estimates of a
6 bacterial infection.[26] Presently, a dissociation between risk assessments and
7 antibiotic prescribing decisions suggests that the former may not be the sole
8 determinant of the latter. It is also possible that explicit risk ratings (as elicited in this
9 type of study) do not reflect physicians' intuitive assessments of risk.

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12 All four of the manipulated variables influenced physicians' (explicit) risk assessments,
13 which increased when the child was younger (20 months vs. 5 years), when illness
14 duration was longer (6 vs. 3 days) and when vomiting and/or wheeze were present
15 (vs. absent). Comparing the odds ratios for these relationships to the STARWAVE
16 model (Table 3), we note both similarities and discrepancies. Specifically, GPs'
17 interpretations of patient age, vomiting and wheeze were consistent with the
18 STARWAVE model, but their interpretation of illness duration was not: a shorter illness
19 duration reduced – rather than increased – GP estimates of risk.

Predictor	OR _{GPs}	OR _{STARWAVE}
Age (<2 years)	1.49 [1.14-1.95]*	3.42 [2.12-5.58]*
Duration (≤3 days)	0.54 [0.42-0.69]*	2.77 [1.77-4.35]*
Vomiting	1.92 [1.57-2.36]*	2.56 [1.54-4.31]*
Wheeze	3.33 [2.66-4.16]*	2.16 [1.28-3.60]*
Temperature		1.99 [1.22-3.25]*
Asthma		3.93 [2.20-7.03]*
Recession		3.82 [2.23-6.62]*

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3 Table 3. The effect of patient age, illness duration, vomiting and wheeze on risk of
4 hospitalisation, according to present participants (OR_{GPs}) and STARWAVE
5 ($OR_{STARWAVE}$). $*p \leq 0.006$. Square brackets contain 95% CIs.
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12 Like risk assessments, prescribing increased when illness duration was long (inverted
13 OR 7.14) and when vomiting and/or wheeze were present (OR_{vomit} 2.17; OR_{wheeze}
14 8.98). Patient age had no reliable effect on prescribing (OR 1.42). Again, these
15 findings are not entirely consistent with the STARWAVE model, but they are consistent
16 with previous, non-experimental research. In one interview study, for example, GPs
17 reported that they were more likely to prescribe antibiotics to children with RTIs given
18 prolonged duration of symptoms, abnormal chest signs and (less frequently)
19 vomiting.[13] Various observational studies have likewise identified chest
20 abnormalities [27-31] and vomiting [30] as clinical characteristics that prompt
21 prescribing. In contrast, previous literature concerning the effect of age on prescribing
22 is mixed: two studies found that older (vs. younger) patients were more likely to receive
23 a prescription,[29, 32] but three identified no association between age and
24 prescribing.[28, 30, 31]
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45 Interestingly, the one patient that may have warranted a prescription received one only
46 21% of the time. This appears low, but in fact only 27% of hospitalised children in the
47 STARWAVE cohort had a discharge diagnosis suggestive of a bacterial infection.[12]
48 Consequently, STARWAVE does not argue (or prove) that all high risk children require
49 immediate antimicrobial treatment; rather, it recommends close monitoring and urgent
50 follow-up with a view to prescribe if needed.[12] Viewed thus, the rate of prescription
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3 that we observed in high risk cases (21%) seems not low, but well-calibrated to the
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5 epidemiological landscape (27%).
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10 Risk assessments and prescribing tendencies bore no association to GPs' self-
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12 reported levels of "Stress from Uncertainty". However, Grol and colleagues found that
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14 greater willingness to take risks (as measured on their Attitudes to Risk Taking scale)
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16 was associated with significantly fewer antibiotics prescriptions for respiratory
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18 problems and URTI/common cold.[33] Attitudes toward risk – rather than attitudes
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20 toward uncertainty – may thus prove a fruitful avenue for future research.
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26 **Limitations and future work**

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30 This is the first study to identify discrepancies between the STARWAVE clinical
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32 prediction rule and current clinical practice. There are several possible reasons for
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34 these discrepancies. Firstly, GPs may be unaware of the STARWAVE rule, which was
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36 published only four years ago; if so, then dissemination and training may be needed.
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38 Alternatively, GPs may be aware of the rule but fail to deploy it at the point of care; in
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40 this case, automated STARWAVE support (e.g., incorporation of STARWAVE metrics
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42 into the electronic health record) could increase uptake. Even so, the rule is intended
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44 to "...supplement, not supplant, clinical judgment" (p. 908)[12] and thus – thirdly – GPs
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46 may choose to override it for sound clinical reasons. To illustrate: the major factor
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48 triggering prescriptions in the present study was a long illness duration (6 vs. 3 days).
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50 This is inconsistent with STARWAVE, but could form part of GPs' strategy to reduce
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52 prescriptions, if the alternative is to prescribe early in the illness (i.e., a "wait-and-see"
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54 approach). Nonetheless, a more evidence-based strategy is not to prescribe at all in
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3 simple RTI, which is likely to last longer than 6 days in any case.[20, 34] Finally, it is
4 also possible that methodological aspects of the present study contributed to the
5 discrepancies observed. For example, the distribution of risk in our vignettes (13%
6 very low, 75% medium, 13% high) was not representative of the patient population
7 (67% very low, 30% medium, 3% high [12]) – an unavoidable consequence of our
8 fractional factorial design. In the “real world”, GPs see many more very low risk cases
9 (67% rather than 13%) and fewer medium and high risk cases (30% rather than 75%
10 medium; 3% rather than 13% high). This may have hurt GPs’ performance by being
11 ecologically invalid (i.e., mismatched to true base rates) and could explain their
12 tendency to underestimate risk in the present study.
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28 A more representative set of vignettes would enhance not only the external validity of
29 the study but also the clinical significance of the findings. Our findings speak mostly to
30 the medium risk group (because we employed mostly medium risk cases) but very low
31 risk cases are twice as common in clinical practice, and indeed account for two-thirds
32 of child RTI presentations in primary care.[12] They are also the focal point of the
33 STARWAVE rule, which aims primarily to rule out prescriptions in very low risk cases.
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35 The present study employed only one very low risk case and identified a prescription
36 rate of 11%; further work is needed to assess the stability of this estimate in a larger
37 and more varied set of very low risk cases.
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51 While GPs overprescribed relative to STARWAVE guidelines, the rate of prescription
52 identified here (15% across cases) is lower than that observed in other studies. For
53 example, Hay et al. identified a rate of 37% in their prospective cohort study of children
54 presenting to the GP with cough.[12] Notably, present work included few high risk
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3 presentations (13%), but high risk presentations were likewise infrequent in the study
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5 by Hay et al. (3%).[12] If our finding is reflective of real-world practice, then this
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7 reduced rate of prescribing is promising indeed. However, it could also reflect the
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9 limitations of our vignettes, which ignored the complex interpersonal (doctor-patient)
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11 dynamics that are known to influence prescribing behaviour.[13-15, 26, 35] For
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13 example, prescription likelihood is increased by perceived pressure from
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15 patients/parents to prescribe;[14, 26, 31, 36, 37] by the desire to maintain good
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17 relationships with patients/parents;[13, 38, 39] by fear of medicolegal problems;[13,
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19 15, 39] and by time pressure.[13, 14, 37-39] Importantly, these factors can be
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21 incorporated into clinical vignettes, as demonstrated by Sirota and colleagues; these
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23 authors found that prescriptions were twice as likely when patient pressure for
24
25 antibiotics was present (vs. absent) from a clinical vignette.[26] On the one hand, it is
26
27 a limitation of our vignettes that these interpersonal factors were absent; on the other,
28
29 our work highlights that antibiotics are overprescribed *even when* these interpersonal
30
31 factors are absent. It is worrying that so many GPs considered antibiotics to be the
32
33 most appropriate course of action, not simply the most expedient one. Qualitative
34
35 research may be useful to understand why GPs prescribed to patients that they
36
37 deemed to be low or medium risk, in the absence of any interpersonal pressure to do
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39 so.
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49 Data for this study were collected in conjunction with another project, which limited the
50
51 number of STARWAVE variables that we could manipulate. A comprehensive
52
53 investigation of all seven STARWAVE variables would undoubtedly return new and
54
55 valuable insights. Future investigations might also treat the continuous STARWAVE
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3 variables (age and illness duration) as continuous (not binary), to test the
4
5 generalisability of the trends identified here.
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10 A second consequence of collecting data in conjunction with another project is that
11
12 the antibiotics vignettes ($n = 8$) were interspersed among many cancer-related
13
14 vignettes ($n = 48$). We cannot exclude the possibility that the cancer vignettes
15
16 influenced performance on the antibiotics task. For example, the cancer vignettes may
17
18 have primed a hyper-cautious attitude (cancer being a serious, “can’t-miss” diagnosis)
19
20 that lowered the threshold for intervention (prescription) in the antibiotics task.
21
22 Threshold for intervention could also be lowered by response fatigue, which
23
24 participants may well have experienced in assessing so many vignettes. Cognizant
25
26 of this, we were careful to present the antibiotics vignettes in a random order. Random-
27
28 ordering would not preclude the cancer vignettes from influencing antibiotics
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30 responding; it simply ensured that any such influence was “spread equally” among the
31
32 antibiotics vignettes.
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40 Despite these limitations, present work sheds light on the determinants of antibiotic
41
42 prescribing in child RTI presentations, bringing much-needed experimental evidence
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44 to a literature that has to date relied predominantly on self-report [13-15, 32, 37, 39,
45
46 40] and observational [27-31] data. It also speaks to the difficulties that may be
47
48 encountered if STARWAVE is provided as a decision aid to GPs. Firstly, GPs’
49
50 classification of risk in this study was largely incompatible with STARWAVE’s; GPs
51
52 consistently chose lower risk than STARWAVE would suggest. Still, they prescribed
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54 more frequently than STARWAVE risk classification would support. Presenting GPs
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56 with STARWAVE’s risk classification will likely exacerbate prescribing (since GPs
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3 overprescribed with their own, lower classifications of risk). Presenting them with a
4 recommendation may also be ineffective, unless the recommendation is accompanied
5 by an explanation. Explaining the recommendation in terms of the variables that
6 increase/decrease a child's risk of hospitalisation may be a way forward, and enable
7 GPs to understand why their own intuitive decision might differ from the
8 recommendation. Identifying the factors that are likely to be misinterpreted by GPs is
9 important when explaining the rationale behind recommendations.
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21 **COMPETING INTERESTS**

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42 submit the paper for publication.
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53 **AUTHOR CONTRIBUTIONS**

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3 All authors contributed to the design of the study. MN performed the data collection;
4
5 MN and OK performed the data analysis. MN drafted the manuscript; OK and BD
6
7 provided critical revision and approved the final version.
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10 11 12 **DATA SHARING**

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14 The data are publicly available on the Open Science Framework under a CC-By
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16 Attribution 4.0 International Licence: <https://osf.io/r3ype/>.
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19 20 21 **LICENSE STATEMENT**

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APPENDIX 1**Characteristics of vignettes**

Vignette identifier	Vignette Text	STARWAVE risk assessment	STARWAVE recommendation
1	A 5-year-old child is brought by their parent with a 6-day history of cough. They have no history of asthma or vomiting. On examination, they are <u>pyrexial</u> , and have neither wheeze nor intercostal recession.	Very-low (1 risk factor)	No immediate prescription
2	A <u>20-month-old</u> child is brought by their parent with a <u>3-day</u> history of cough. They have no history of asthma or vomiting. On examination, they are <u>pyrexial</u> , and have neither wheeze nor intercostal recession.	Normal (3 risk factors)	No immediate prescription
3	A <u>20-month-old</u> child is brought by their parent with a 6-day history of cough. They have no history of asthma and <u>have vomited</u> twice in the last 24 hours. On examination, they are <u>pyrexial</u> , and have neither wheeze nor intercostal recession.	Normal (3 risk factors)	No immediate prescription
4	A 5-year-old child is brought by their parent with a <u>3-day</u> history of cough. They have no history of asthma and <u>have vomited</u> twice in the last 24 hours. On examination, they are <u>pyrexial</u> , and have neither wheeze nor intercostal recession.	Normal (3 risk factors)	No immediate prescription
5	A <u>20-month-old</u> child is brought by their parent with a 6-day history of cough. They have no history of asthma or vomiting. On examination, they are <u>pyrexial</u> , and <u>have wheeze</u> on chest auscultation. There is no intercostal recession.	Normal (3 risk factors)	No immediate prescription

6	A 5-year-old child is brought by their parent with a <u>3-day</u> history of cough. They have no history of asthma or vomiting. On examination, they are <u>pyrexial</u> , and <u>have wheeze</u> on chest auscultation. There is no intercostal recession.	Normal (3 risk factors)	No immediate prescription
7	A 5-year-old child is brought by their parent with a 6-day history of cough. They have no history of asthma and <u>have vomited</u> twice in the last 24 hours. On examination, they are <u>pyrexial</u> , and <u>have wheeze</u> on chest auscultation. There is no intercostal recession.	Normal (3 risk factors)	No immediate prescription
8	A <u>20-month-old</u> child is brought by their parent with a <u>3-day</u> history of cough. They have no history of asthma and <u>have vomited</u> twice in the last 24 hours. On examination, they are <u>pyrexial</u> , and <u>have wheeze</u> on chest auscultation. There is no intercostal recession.	High (5 risk factors)	Consider an immediate prescription

Note: STARWAVE risk factors are underlined for salience. They were not underlined when vignettes were presented to GPs.

APPENDIX 2

Stress from Uncertainty (SfU) scale

1. I usually feel anxious when I am not sure of a diagnosis.
2. I find the uncertainty involved in patient care disconcerting.
3. Uncertainty in patient care makes me uneasy.
4. I am quite comfortable with the uncertainty in patient care.*
5. The uncertainty of patient care often troubles me.
6. When I am uncertain of a diagnosis, I imagine all sorts of bad scenarios -- patient dies, patient sues, etc.
7. I fear being held accountable for the limits of my knowledge.
8. I worry about malpractice when I do not know a patient's diagnosis.

Note: items 1-5 measure the construct "Anxiety due to Uncertainty" (Cronbach's alpha=0.86); items 6-8 measure the construct "Concern About Bad Outcomes" (Cronbach's alpha=0.73).

*Reverse-scored item.

APPENDIX 3

Delayed prescriptions

Delayed prescriptions were administered 12% of the time (121/1008). The association between delayed prescribing and risk (as classified by GPs and by STARWAVE) is displayed below (yellow cells).

		Risk as classified by GPs			STARWAVE risk			Total
		Very-low	Medium	High	Very-low	Medium ("normal")	High	
Prescriptions	None	420	294	17	95	551	85	731
	Delayed	25	92	4	17	89	15	121
	Immediate	14	107	35	14	116	26	156
Total		459	493	56	126	756	126	1008

To investigate the effect of the manipulated factors on both delayed and immediate prescribing, we regressed the 3-category prescribing variable (0=no prescription, 1=delayed prescription, 2=immediate prescription) on patient age (0=5 years, 1=20 months), illness duration (0=6 days, 1=3 days), vomiting (0=absent, 1=present) and wheeze (0=absent, 1=present). This ordinal logistic regression analysis was conducted using the Stata user-written program "gologit2".^[22, 23] Statistical tests of the proportional odds assumption revealed that two variables met it ($p_{\text{age}}=0.124$ and $p_{\text{vomit}}=0.522$) and two did not ($p_{\text{duration}}=0.034$ and $p_{\text{wheeze}}=0.003$). Put differently: the respective effects of age and vomiting were consistent for successive levels of the ordinal dependent variable, while those of duration and wheeze were not. Thus, we constructed a partial proportional odds (PPO) model, where two coefficients were fixed (age and vomiting) and two were allowed to vary (duration and wheeze). A global Wald test confirmed that the proportional odds assumption was not violated in this PPO model ($\chi^2(2) 2.63, p=0.268$).

Results are tabulated below. The model progresses in two steps: the first step compares “no prescription” (coded 0) to “delayed prescription” and “immediate prescription” (both coded 1); the second compares “no prescription” and “delayed prescription” (both coded 0) to “immediate prescription” (coded 1). Trends were consistent across steps, and consistent with those reported in the main text. Specifically, patient age did not influence the odds of a prescription ($p=0.569$) and short illness duration decreased them ($p<0.001$). Presence of vomiting and presence of wheeze both increased prescribing odds (both $ps<0.001$). Two coefficients were allowed to vary across steps (duration and wheeze): in both cases, effects grew stronger from step 1 to step 2.

	STEP 1: no prescription (coded 0) vs. delayed/immediate (coded 1)	STEP 2: no/delayed prescription (coded 0) vs. immediate (coded 1)
Age (<2 years)	0.92 [0.69-1.23]	0.92 [0.69-1.23]
Duration (≤ 3 days)	0.46 [0.34-0.62]*	0.34 [0.24-0.49]*
Vomiting	1.49 [1.24-1.80]*	1.49 [1.24-1.80]*
Wheeze	2.50 [1.91-3.28]*	3.89 [2.66-5.69]*

* $p<0.001$. Cells contain odds ratios; square brackets contain 95% CIs. Step 2 of the model (no/delayed prescription vs. immediate) is akin to the model reported in the main text; differences in coefficients may be attributed to different estimation procedures (e.g., the ordinal model estimates all parameters simultaneously).[22]

APPENDIX 4**Risk assessments and prescribing decisions per vignette**

Vignette identifier	GP risk assessments			GP prescribing decisions		
	Very-low	Medium	High	None	Delayed	Immediate
1	64% (81/126)	35% (44/126)	1% (1/126)	75% (95/126)	14% (17/126)	11% (14/126)
2	71% (89/126)	29% (37/126)	0% (0/126)	94% (118/126)	5% (6/126)	2% (2/126)
3	42% (52/125)	53% (66/125)	6% (7/125)	74% (93/125)	13% (16/125)	13% (16/125)
4	59% (75/127)	40% (51/127)	1% (1/127)	81% (103/127)	14% (18/127)	5% (6/127)
5	25% (32/127)	62% (79/127)	13% (16/127)	59% (75/127)	11% (14/127)	30% (38/127)
6	53% (66/125)	44% (55/125)	3% (4/125)	77% (96/125)	14% (17/125)	10% (12/125)
7	25% (31/126)	64% (80/126)	12% (15/126)	52% (66/126)	14% (18/126)	33% (42/126)
8	26% (33/126)	64% (81/126)	10% (12/126)	67% (85/126)	12% (15/126)	21% (26/126)