Title: Spatial regression analysis of MR diffusion reveals subject-specific white matter changes associated with repetitive head impacts in contact sports

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Supplementary Methods

Details of SPREAD Model used in the current study

In this study, we adapted the original SPREAD pipeline (technical details can be found in [1]) for subject-specific analysis. The statistical model for this approach is represented as follows $\mathbb{D}_{ti} = \Phi(z_i) + \beta(z_i) \cdot I\{t = post\} + \epsilon_{ti}$ for $i = 1, ..., I$ and $t = pre, post.$ (1)

Here \mathbb{D}_{ti} is a set of tensor-derived parameters (e.g., the FA map used in this study) for a subject measured at the pre- or post-scan. z_i represents the 3-dimensional spatial coordinate vector for the *i*th voxel, *i* = 1,2, ..., *I*. The pre-season image (i.e., $\mathbb{D}_{pre,i}$), is modeled as a continuous spatial function $\Phi(z_i)$ superimposed on unknown measurement errors ϵ_{ti} with joint-distribution function $F_{\epsilon}(\cdot)$. This joint-distribution function does not need to be normal and we only assume it to be invariant under temporal permutation. $I\{t = post\}$ is the indicator function of the postscan, and the difference between pre- and post-season images, $\beta(z_i)$, is also modeled as a continuous spatial function.

Spatial Regression:

The fitted FA maps were defined as

$$
FA_{ti}^{(k)} = \frac{\sum_{j=1}^{l} K_h(z_i - z_j)FA_{ti}}{\sum_{j=1}^{l} K_h(z_i - z_j)}.
$$
\n(2)

where z_i , z_j are 3-dimensional coordinate vectors associated with the *i*th and *j*th voxel, $K_h(\cdot)$ is a kernel function and h is a tuning parameter which controls its bandwidth. In this study, $K_h(\cdot)$ is the standard Gaussian Kernel function and h is the full width at half maximum (FWHM). The relationship between FWHM and the standard deviation of the Gaussian distribution is $h =$ $2\sqrt{2\ln 2}\sigma \approx 2.35\sigma$ voxels. The temporal differences were summarized by the statistic defined as $\Delta FA_i^{(k)} = |FA_{post,i}^{(k)} - FA_{pre,i}^{(k)}|.$

Hypothesis testing:

The null hypothesis is H_{i0} : $\beta(z_i) = 0$, i.e. no significant difference between pre- and post-scan images at the *i*th voxel. Under the assumption that $F_{\epsilon}(\cdot)$ is invariant for temporal permutation, permuting the time label of the observation would not affect the distribution function of FA_{ti} if $\beta(z_i) = 0$. Therefore, under the null hypothesis, the detected abnormal region can be defined as the set of voxels at which H_{i0} is rejected by a permutation-based test. Specifically, we randomly permute the time label (pre and post) of each voxel, and refit the spatial regression model (2). The fitted values and the summary of temporal differences are labeled as $FA_{ti}^{(k)}$ and $AFA_i^{(k)}$, respectively, for $k = 1, 2, ..., K$ permutations. We define the permutation-based p-value for the i^{th} voxel as $p_i = \frac{1}{\kappa}$ $\frac{1}{K} \sum_{k=1}^{K} I(\Delta FA_i^{(k)} \geq \Delta FA_i^{(0)})$. Due to the large number of hypotheses (voxels) to be tested, we must use a suitable multiple testing procedure to control for overall type I error. Two MTPs are used in this study: the Benjamini-Hochberg procedure and the Westfall and Young procedure. The first one controls the false discovery rate and the latter controls for family-wise error rate. Both procedures are consistent under positive quadrant correlation and known to work well with permutation-based inferences.

Functional norms:

In this study, we developed several global statistics based on functional norms to summarize the overall temporal changes in DTI images. Recall that after the spatial regression, FA (or other DTI derived quantities such as MD) maps for the sth individual before and after the football season are represented as real valued spatial functions $FA_{s,pre}^{(0)}(z)$ and $FA_{s,post}^{(0)}(z)$, respectively. Due to the smoothing nature of the Nadaraya-Watson kernel regression and the fact that they are defined on compact domains (the dimensions of the brain images are finite), both $FA_{s,pre}^{(0)}(z)$ and $FA_{s,post}^{(0)}(z)$ are smooth and bounded, which implies that they are members of functional spaces $L^p(\Omega, \mathbf{B}, d\lambda)$, where $\Omega \subset \mathbf{R}^3$ is the image domain, **B** is the Borel σ -algebra, and $d\lambda$ is the Lebesgue measure. The L^p norm is a generalization of Euclidean length to functional spaces, which can be defined as follows

$$
||f||_p := \left(\int_{\Omega} |f(\mathbf{z})|^p \mathrm{d}\,\lambda\right)^{\frac{1}{p}}, \quad 1 \le p < \infty.
$$

As a generalization, it is customary to define the L^{∞} norm as the essential supremum of a spatial function as follows

$$
||f||_{\infty} := \inf \{ C \ge 0 : |f(z)|, \text{for } a, e, z \in \mathbb{R}^3 \}.
$$

More details of the L^p spaces and L^p norms can be found in most functional analysis textbooks, such as [2].

Meta-analysis:

We conducted a meta-analysis based on a novel robust p-value combination test based on the Beta-distribution. Our objective was to combine p-value maps of all 28 contact athletes into a single group-level p-value map, and then select common change regions based on this single map.

Arguably, the most classical p-value combination method is Fisher's p-value combination test, which can be described as follows,

$$
Q_i := -2 \sum_{n=1}^N \ln p_{ni}, \qquad Q \sim \chi_{2N}^2, \qquad P_i := 1 - F_{\chi_{2N}^2}(Q_i).
$$

Here p_{ni} , $n = 1, 2, ... N$ represent the p-value of voxel *i* for the *n*th subject, $F_{\chi^2_{2N}}(Q_i)$ is the distribution function of Chi-squared distribution χ^2_{2N} , and P_i is the combined p-value which follows a uniform distribution on (0,1) under the null hypothesis. However, Fisher's p-value combination test is not suitable in the context of permutation test due to the granularity of permutation p-values. It's obvious that one assumption for Fisher's p-value combination test is that all p_i s must be positive, so that $\ln p_{ni}$ is well defined. For permutation tests, p_i s may be exactly zero because we cannot use infinitely many permutations, which breaks this assumption. Furthermore, this test is not robust to occasional outliers in the sample, because very small p_{ni} s have disproportionately large impact on the combined p-value: one extremely small p_{ni} (possibly due to outliers in the data) can lead to the overall significance, even if all other subjects are not significant.

As an alternative, we proposed the following robust p-value combination test based on the median p-value map, denoted by med (p_i) , where p_i is the vector of p_{ni} s pooled from all subjects. We use the following summary statistic to capture the overall significance of the data

$$
S_i := \begin{cases} \text{med}(\boldsymbol{p}_i), & n \text{ is odd} \\ \text{c}(n) \left(\text{med}(\boldsymbol{p}_i) - \frac{1}{2} \right) + \frac{1}{2}, & n \text{ is even} \end{cases} \quad c(n) := \sqrt{\frac{n+1}{n}}, \quad P_i := 1 - F_{\text{Beta}(\frac{n+1}{2}, \frac{n+1}{2})}(S_i). \tag{3}
$$

When n is an odd number, it's well known that $S_i = \text{med}(p_i)$ follows a Beta-distribution Beta $\left(\frac{n+1}{2}\right)$ $\frac{+1}{2}, \frac{n+1}{2}$ $\frac{1}{2}$) based on the theory of order statistics, therefore we can compute an overall pvalue P_i based on the distribution function of Beta $(\frac{n+1}{2})$ $\frac{+1}{2}, \frac{n+1}{2}$ $\frac{+1}{2}$). When n is even, the distribution of med(\bf{p}) is more complex but it can be proven that the adjusted median, $S_i = c(n)$ (med(\bf{p}_i) – 1 $\frac{1}{2}$ + $\frac{1}{2}$ $\frac{1}{2}$, has the same mathematical expectation as Beta $(\frac{n+1}{2})$ $\frac{+1}{2}, \frac{n+1}{2}$ $\frac{+1}{2}$), so using the quantile function of Beta $\left(\frac{n+1}{2}\right)$ $\frac{+1}{2}, \frac{n+1}{2}$ $\frac{+1}{2}$) will lead to a well approximated overall p-value in this case.

The algorithm used in our meta-analysis can be summarized as follows

- 1. Apply affine registration on the fitted difference maps, which is implemented by R package **RNiftyReg**
- 2. Record the affine transformation matrices and then apply them to the p-value maps.
- 3. Calculate sample median p-values from all 28 unadjusted p-value maps.
- 4. Use adjusted Beta-distribution to compute the combined p-value maps.
- 5. Apply a suitable multiple testing adjustment, such as Holm-Bonferroni procedure which controls familywise error rate, to get the adjusted combined p-value map. Of note, while the WY procedure also controls for the family-wise error rate, it is a permutation-based MTP which is not directly applicable to p-value combination test.
- 6. Select significant voxels based on this adjusted combined p-value map, which is presented in Fig. 3.

Supplementary Fig. 1: Examples of SPREAD Parameter Combinations

(A) A visualization of the actual small synthesized signal created in the simulation study. (B) Region of significantly changed voxels detected by SPREAD at a smoothing bandwidth of 5 and using the Benjamini-Hochberg procedure. (C) A visualization of the actual small synthesized signal created in the simulation study. (D) Region of significantly changed voxels detected by SPREAD at a smooth bandwidth of 5 and using the Westfall-Young procedure. (E) A visualization of the actual medium synthesized signal created in the simulation study. (F) Region of significantly changed voxels detected by SPREAD at a smoothing bandwidth of 10 and using the Benjamini-Hochberg procedure. (G) A visualization of the actual medium synthesized signal created in the simulation study. (H) Region of significantly changed voxels detected by SPREAD

at a smooth bandwidth of 10 and using the Westfall-Young procedure. (I) A visualization of the actual large synthesized signal created in the simulation study. (J) Region of significantly changed voxels detected by SPREAD at a smoothing bandwidth of 19 and using the Benjamini-Hochberg procedure. (K) A visualization of the actual large synthesized signal created in the simulation study. (L) Region of significantly changed voxels detected by SPREAD at a smooth bandwidth of 19 and using the Westfall-Young procedure. (M) A visualization of the small synthesized signal from the simulation study. (N) A visualization of the medium synthesized signal from the simulation study. (O) A visualization of the large synthesized signal from the simulation study.

Supplementary Fig. 2. All athletes' individual raw p-value map at registered slice 32.

The yellow/red highlighted regions of the brain are voxels with associated raw p-values of <0.002 when comparing the pre-season to post-season scans. The red voxels are associated with a lower raw p-value than the yellow ones.

Supplementary Fig. 3. Significant ($p < 0.002$) temporal changes of FA values (post-season – pre-season) for all athletes at registered slice 32. Blue color is used to represent decreased FA values and red represents increased FA values. The overwhelming majority (93.3%) of significant voxels have decreased FA values after the football season.

Supplementary Fig. 4: Removal of the ringing artifacts

An illustration of the ringing artifacts and our strategy of removing them. First, we detect the foreground and background of the FA map by a pre-specified intensity threshold such as 0.01. Next, we enlarge the background mask by a pre-specified Euclidean distance such as two voxels in all 26 spatial directions. The final foreground mask is defined as the set complement of the enlarged background map. This procedure is implemented as function MountDoom() in our software package.

Supplementary Table 1. Basic demographics for all contact athletes and controls

* PIP is poor DTI image processing

Supplementary Table 2: Summary of helmet-based impact measures. LA: linear acceleration; RA: rotational acceleration; HIC15: head impact criterion 15; GSI: Gadd Severity Index; HITsp: helmet impact technology severity profile.

Supplementary Table 3. Correlations between helmet impact metrics and L¹ bandwidth combinations. LA: linear acceleration; RA: rotational acceleration; HIC15: head impact criterion

		Bandwidth 3		Bandwidth 5		Bandwidth 10		Bandwidth 15	
Metric	HIM	c.c.	adjusted p-value	c.c.	adjusted p-value	c.c.	adjusted p-value	c.c.	adjusted p-value
Mean	LA	0.0454	1.0000	0.1138	0.5859	0.1232	0.5308	0.1368	0.4858
	RA	0.0668	1.0000	0.2414	0.3898	0.3525	0.1573	0.3870	0.1942
	HIC15	0.0213	1.0000	0.1478	0.5415	0.1888	0.3715	0.1779	0.4039
	GSI	0.0564	1.0000	0.1642	0.5026	0.2080	0.3586	0.2058	0.3651
	HITsp	0.0312	1.0000	0.2217	0.3898	0.3695	0.1573	0.4193	0.1942
Peak	LA	0.0274	1.0000	0.1128	0.5859	0.1609	0.4259	0.1593	0.4309
	RA	0.0000	1.0000	0.0925	0.6385	0.1686	0.4173	0.1675	0.4207
	HIC15	0.0044	1.0000	0.1237	0.5859	0.2014	0.3632	0.1932	0.3877
	GSI	0.0148	1.0000	0.1352	0.5667	0.1888	0.3715	0.1866	0.3924
	HITsp	0.0881	0.9351	0.2222	0.3898	0.2813	0.2178	0.2846	0.2365
CUW	LA	0.3711	0.3077	0.3612	0.2441	0.3443	0.1573	0.3262	0.1942
	RA	0.3519	0.3077	0.3503	0.2441	0.3574	0.1573	0.3481	0.1942
	HIC15	0.3897	0.3077	0.4034	0.2441	0.3968	0.1573	0.3908	0.1942
	GSI	0.4149	0.3077	0.4149	0.2441	0.3935	0.1573	0.3875	0.1942
	HITsp	0.3558	0.3077	0.3503	0.2441	0.3476	0.1573	0.3355	0.1942
TBH	LA	0.2742	0.3715	0.3383	0.2441	0.3695	0.1573	0.3421	0.1942
	RA	0.2315	0.4696	0.2989	0.2520	0.3519	0.1573	0.3361	0.1942
	HIC15	0.2895	0.3680	0.3870	0.2441	0.4023	0.1573	0.3530	0.1942
	GSI	0.2720	0.3715	0.3629	0.2441	0.3815	0.1573	0.3415	0.1942
	HITsp	0.2540	0.4104	0.3333	0.2441	0.3727	0.1573	0.3563	0.1942
TUA	LA	0.3662	0.3077	0.3273	0.2441	0.3158	0.2037	0.2917	0.2365
	RA	0.3290	0.3077	0.3010	0.2520	0.3010	0.2052	0.2781	0.2388
	HIC15	0.2901	0.3680	0.3082	0.2520	0.3443	0.1573	0.3279	0.1942
	GSI	0.3246	0.3077	0.3333	0.2441	0.3563	0.1573	0.3426	0.1942
	HITsp	0.3262	0.3077	0.2961	0.2520	0.2983	0.2052	0.2846	0.2365
TBH+TUA	LA	0.1554	0.7806	0.2167	0.3898	0.2775	0.2178	0.2682	0.2388
	RA	0.1494	0.7806	0.2206	0.3898	0.3071	0.2052	0.3032	0.2337
	HIC15	0.1275	0.8151	0.2042	0.4036	0.2677	0.2292	0.2534	0.2625
	GSI	0.1111	0.8581	0.1856	0.4473	0.2507	0.2576	0.2348	0.2975
	HITsp	0.1423	0.7806	0.2140	0.3898	0.2802	0.2178	0.2709	0.2388

15; GSI: Gadd Severity Index; HITsp: helmet impact technology severity profile. "c.c" is the Spearman rank correlation coefficient (ρ).

Supplementary Table 4. Correlations between helmet impact metric with L^2 bandwidth combinations. LA: linear acceleration; RA: rotational acceleration; HIC15: head impact criterion 15; GSI: Gadd Severity Index; HITsp: helmet impact technology severity profile. "c.c" is the Spearman rank correlation coefficient (ρ). Bold values signify significant correlations (adjusted p-value < 0.05).

References

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