WEB A	PPE	NDIX
-------	-----	------

ME	TH	OD	S

Sensitivity analyses

Cognitive Quotient

Cognitive Quotient for the Bristol cohort

Binary outcome: alive and without severe cognitive disability

Ordinal outcome: grading of cognitive disability

Adjustment for baseline covariates (by more than 10% or 0.5SDs)

Adjustment for maternal education

Imputation of missing data at ten years

Best and worst case scenarios

Subgroup analyses

RESULTS

Sensitivity analyses

eTable 1. Sensitivity Outcomes

eTable 2. Subgroup Analyses

REFERENCES

METHODS

Sensitivity analyses

To test the robustness of the primary outcome results, various sensitivity analyses were carried out. As cause of death was difficult to ascertain, different methods were used to ensure that results were consistent; these included:

Cognitive Quotient

score of <55.

It had not been anticipated that deaths would occur from disability between two and ten years of age, however, it was considered an important outcome for the trial. Of the four deaths, one child from each group had a certificate of cause of death that confirmed the death was due to disability and one child from each arm did not have traceable certification therefore, based on their developmental scores at two years, their deaths were assumed to be related to their severe disability. As a sensitivity analyses these four children were given a cognitive quotient (CQ) score of zero.

Cognitive Quotient for the Bristol cohort

The majority of patients were recruited from the Bristol center and, to check that results remained the same when looking specifically at protocol procedures in a single center, children from Glasgow, Norway and Poland were excluded.

Binary outcome: alive and without severe cognitive disability

The two-year outcomes paper [1] separated the children into those who were alive and without sensorimotor or cognitive disability versus those who had died or had severe sensorimotor or cognitive disability. The primary research objective for this follow-up study was cognitive ability, assessed using either the British Ability Scales III (BASIII) [2] or the Bayley III scales (BSIDIII) [3] age equivalent scores (Protocol: https://www.journalslibrary.nihr.ac.uk/programmes/hta/123561/#/). BSIDIII scales were used if a child's developmental age was below three years (i.e. below the range of the BASIII test). However, given the results at two years it was considered important to try and replicate the findings at ten years. Therefore, a binary outcome

2

was created that included all deaths as a negative outcome, along with a General Conceptual Ability (GCA)

Ordinal outcome: grading of cognitive disability

The team felt it was important to differentiate between the different levels of cognitive disability and therefore included this as a sensitivity analysis. We did not anticipate that certain individuals would have GCA scores greater than 85, however, this was true for 11 individuals. Therefore, the categories were expanded (post-hoc) to 1. Deceased, 2. Severe (GCA<55 or Bayley scale used), 3. Moderate (55≤GCA<70), 4. Mild (70≤GCA<85), 5. No cognitive disability (GCA≥85).

Adjustment for baseline covariates (by more than 10% or 0.5SDs)

Although sex, birthweight and grade of intraventricular hemorrhage (IVH) had been pre-specified as covariates that were imbalanced at two years the team also wanted to adjust for any imbalance between the arms of those assessed at ten years.

Adjustment for maternal education

This was not a pre-specified sensitivity analysis. The team wanted to adjust for maternal education status, as this may have influenced cognitive development after the two-year assessment. However, as this was not collected at baseline, maternal education status at ten years was used as a proxy measure.

Imputation of missing data at ten years

Similarly to Biering et al. [4] we chose to carry out five different imputation models that made different assumptions about the data, particularly death. These included each of the four permutations of including an indicator variable for death and/or imputing scores of zero for those who died, as well as a separate imputation model excluding the four deaths. As we did not have any cognitive data on those who died before two years, these were excluded from the imputation models. Baseline variables were assessed to see whether they were predictive of missingness or were appropriate predictors of the primary model, using logistic and linear regression. Any baseline variables associated with the primary outcome of interest or its missingness were added to the imputation model to inform the imputation process. Imputation by chained equations was carried out in STATA 14.1 to create 40 separate imputations that were then combined using Rubin's rules.

Best and worst case scenarios

There were two children lost to follow up in the DRIFT group where we were not aware of their survival status.

The analysis was repeated, after imputing zeros as a worst case scenario and after imputing median scores for their group as a best case scenario.

Subgroup analyses

Interaction tests were added to the primary outcome to see if the effects of the DRIFT intervention were more pronounced in certain subgroups of children; the pre-specified dichotomized subgroups included:

- Gestation (28 weeks and above vs. <28 weeks)
- Grade of IVH (Grade 3 vs. 4)
- Age of randomization (Day 1-20 vs. 21+ days)
- Unilateral vs bilateral dilatation on ultrasound scan at randomization.
- Sex
- Pre- and post-enhanced vigilance in 2006.

Additionally, maternal education was added as a post-hoc sub group analysis, categorized as 'low' if the mother left school at 16 years and 'high' if the mother carried on with further education post 16 and/or had accessed university education.

RESULTS

Sensitivity analyses

All sensitivity analyses gave results that were consistent with the primary analysis (eTable 1). Adjustment for center and maternal education (not concurrently) weakened the effect size slightly but models additionally adjusting for sex, birthweight and IVH grade were still in favour of DRIFT. Adjusting for factors that showed imbalance at baseline gave very similar results to the primary adjusted model as the variables imbalanced at baseline were sex and birthweight; two of the three covariates used in the primary model. Results from each of the five multiple imputation models had slightly tighter confidence intervals but similar effect sizes. After imputing scores of zero for those lost to follow up in the DRIFT group as a 'worst case scenario', the effect estimate was weakened to produce a treatment difference consistent with chance, however, still in favor of DRIFT. The best case scenario gave similar results to the primary analysis, after imputing scores of zero for those who died post two years.

Results from the binary outcome of death or severe cognitive disability gave even stronger evidence to suggest a benefit from DRIFT. Including all deaths between randomization and ten years as a negative outcome left 21/32 (66%) of children alive and without severe cognitive disability in the DRIFT group, compared with 11/31 (35%) in the standard treatment group; adjusted odds ratio 7.69 (95% CI 1.96, 30.11), p=0.003. For the four deaths that occurred after two years, the trial team confirmed death due to severe disability from the cause of death certification for two children. For the two deaths where certification could not be obtained the assumption was made that the causes of death were also attributable to disability based on their severe disability grading at their two-year assessments. Deaths prior to two-year follow-up were handled with caution as they mainly occurred in the first months of life when accurate ascertainment of presence and extent of disability is difficult. Prior to two years only four certificates of cause of death could be obtained of which three were deemed unrelated to neurological causes (1 DRIFT and 2 standard treatment) and 1 (standard treatment) was attributable to neurological causes. That left four deaths before two years where the cause could not be obtained (2 DRIFT and 2 standard treatment). Excluding deaths before two years of age resulted in 21/29 (72%) of children alive and without cognitive disability in the DRIFT group compared with 11/26 (42%) in the standard treatment group; adjusted odds ratio 9.96 (95% CI 2·12, 46·67), p=0·004. There was no evidence to suggest that there were any subgroup interactions with the treatment effect (eTable 2).

eTable 1. Sensitivity outcomes

		DRIFT	Standard	Unadj. difference (95% C.I); P value	Adj. difference (95% C.I); P value ^a
	n(D:S)	n(%)/	n(%)/		
		Mean(SD)	Mean(SD)		
Cognitive Quotient (CQ; B)	ristol cohori	t only)			
Score	23:19	71-76 (27-42)	57-83 (34-78)	13.93 (-5.46, 33.33); 0.15	24.88 (6.82, 42.94); 0.008
Score ^b	24:20	68-77 (30-56)	54.94 (36.24)	13.84 (-6.48, 34.15); 0.18	23-27 (4-65, 41-88); 0-016
Cognitive disability categor	у				
Deceased		2 (7%)	2 (8%)		
Severe disability		6 (21%)	13 (50%)		
Moderate disability	29:25	7 (24%)	2 (8%)	2.04 (0.77, 5.42); 0.15	3.63 (1.21, 10.90); 0.02
Mild disability		8 (28%)	4 (15%)		
No cognitive disability		6 (21%)	5 (19%)		
Adjustments					
Adj. for centre	27:24	69-33 (30-06)	53.68 (35.70)	13.76 (-4.45, 31.97); 0.14	22.0 (5.7, 38.3); 0.009
Adj. for centre	27:24	69-33 (30-06)	53.68 (35.70)	14.54 (-3.78, 32.87); 0.12	23.19 (6.35, 40.04); 0.008
(binary: Bristol vs. other)	21.24	09-33 (30-00)	33.00 (33.70)	14.34 (-3.78, 32.87), 0.12	25.19 (0.33, 40.04); 0.008
Adj. for maternal	27:23	69.33 (30.06)	55.90 (34.77)	11.50 (-6.86, 29.87); 0.21	20.08 (2.96, 37.21); 0.02
education (post-hoc)	27.23	07 33 (30 00)	33 70 (34 77)	11 30 (-0 60, 27 67), 0 21	20 00 (2)0, 37 21), 0 02
Adj. for imbalance	27:24	69-33 (30-06)	53.68 (35.70)	24.58 (6.69, 42.46); 0.008	-
Multiple imputation ^c					
Assumption 1 ^d	36:33	65-24 (5-63)	50.81 (6.23)	14.43 (-2.10, 30.96); 0.09	21.17 (5.66, 36.68); 0.008
Assumption 2e	36:33	65.42 (5.45)	50.87 (6.31)	14.54 (-1.98, 31.07); 0.08	21-42 (6-21, 36-64); 0-007
Assumption 3 ^f	36:33	62.95 (5.80)	49.41 (6.44)	13.55 (-3.84, 30.93); 0.12	20.53 (4.49, 36.56); 0.013
Assumption 4g	36:33	62.80 (5.91)	49.58 (6.47)	13-22 (-4-49, 30-93); 0-14	20.08 (3.79, 36.38); 0.017
Assumption 5 ^h	34:32	66.85 (5.42)	53.70 (6.43)	13·14 (-3·67, 29·96); 0·12	20-47 (4-62, 36-31); 0-012
Case scenarios (for those with unknown survival status)					
Best case scenarioi	31:26	65.04 (32.94)	49.55 (37.22)	15-49 (-3-13, 34-12); 0-101	20.67 (3.68, 37.65); 0.018
Worst case scenarioj	31:26	60.38 (36.62)	49.55 (37.22)	10.83 (-8.83, 30.50); 0.274	15.28 (-3.72, 34.29); 0.113

^aadjusted for sex, birthweight & grade of IVH

^bgiving children who have died post 2 years a cognitive quotient (CQ) score of 0

cstandard errors replace standard deviations here

dimputing CQ for those who died post 2 years or were lost to follow up with no indicator for death

eimputing CQ for those who died post 2 years or were lost to follow up with an indicator for death

fimputing CQ for those who were lost to follow-up with no indicator for death and replace CQ with 0 for those who died post 2 years

gimputing CQ for those who were lost to follow-up with an indicator for death and replace CQ with 0 for those who died post 2 years

himputing CQ for those who were lost to follow-up only with no indicator for death

ⁱassuming the 2 children in the DRIFT group were all alive and without severe cognitive disability (with the median score for their group) at 10 years

jassuming the 2 children in the DRIFT group had died

eTable 2. Subgroup analyses

		Subgrou	p specific	Interaction	Interaction ^a (95% C.I.); P value		
	N (D:S)	DRIFT Mean(SD)	Standard Mean(SD)	(95% C.I.); P value			
Cognitive Quotient scores at school age (treatment-subgroup interaction)							
Gestation (<28 weeks) Gestation (≥28 weeks)	15:10 12:14	62·5 (28·4) 77·8 (31·1)	42·9 (30·7) 61·4 (38·1)	3.20 (-33.7, 40.1); 0.892	-18.9 (-54.7, 17.0); 0.295		
Grade of IVH (Grade 3) Grade of IVH (Grade 4)	14:13 13:11	75·6 (26·1) 62·6 (33·6)	71·1 (36·3) 33·1 (22·0)	25.0 (-9.2, 59.2); 0.149	15.7 (-19.8, 51.1); 0.379		
Age ^b (<21 days) Age ^b (≥21 days)	15:15 12:9	75·2 (30·1) 62·0 (29·6)	60·7 (37·3) 42·0 (31·3)	-5·5 (-42·9, 32·0); 0·770	-5·0 (-38·3, 28·2); 0·762		
Unilateral dilatation ^c Bilateral dilatation ^c	4:4 23:20	64·0 (23·7) 70·3 (31·4)	35·4 (16·9) 57·3 (37·6)	15.7 (-35.4, 66.8); 0.539	6.8 (-40.0, 53.6); 0.771		
Sex: Male Sex: Female	22:15 5:9	65·3 (31·6) 86·9 (12·5)	47·6 (34·4) 63·9 (37·6)	-5·3 (-47·7, 37·2); 0·803	-3·6 (-42·1, 34·8); 0·851		
Pre enhanced vigilance ^d Post enhanced vigilance ^d	22:23 5:1	67·7 (33·1) 76·7 (6·4)	51·5 (34·8) 104·0 (0)	-43·5 (-117·8, 30·9); 0·246	-15·2 (-84·5, 54·0); 0·660		
Maternal educ. (Low) ^e Maternal educ. (High) ^e	10:11 17:12	64·2 (37·8) 72·4 (25·3)	52·1 (31·8) 59·4 (38·4)	-0.9 (-39.0, 37.1); 0.961	-15.8 (-50.8, 19.1); 0.366		

^aadjusted for sex, birthweight & grade of IVH,

bage at randomization,

^cdilatation on ultrasound scan at randomization,

din 2006 the trial was temporarily halted as there were concerns about secondary hemorrhages in the DRIFT group, after this time 7 more patients were recruited during an 'enhanced vigilance' period [5]

^eMaternal education was collected at 10 years and therefore only classed as an indicator of education at baseline.

This was classed as 'low' if the mother left school at 16 and 'high' if the mother carried on with further education post 16 and/or went to university.

REFERENCES

- Whitelaw, A., et al., Randomized Trial of Drainage, Irrigation and Fibrinolytic Therapy for Premature Infants with Posthemorrhagic Ventricular Dilatation: Developmental Outcome at 2 years. Pediatrics, 2010. 125(4): p. E852-E858.
- 2. Elliot, C. and P. Smith, British Ability Scales 3rd edition. 2012, London GL Assessment.
- 3. Bayley, N., *Bayley Scales of Infant & Toddler Development 3rd edition*. The Psychological Corporation: San Antonio, TX.
- 4. Biering, K., N.H. Hjollund, and M. Frydenberg, *Using multiple imputation to deal with missing data* and attrition in longitudinal studies with repeated measures of patient-reported outcomes. Clin Epidemiol, 2015. 7: p. 91-106.
- 5. Whitelaw, A., et al., Randomized clinical trial of prevention of hydrocephalus after intraventricular hemorrhage in preterm infants: Brain-washing versus tapping fluid. Pediatrics, 2007. 119(5): p. E1071-E1078.