Supplement to: Are There Negative Cycles of Peer Victimization and Rejection Sensitivity? Testing ri-CLPMs in Two Longitudinal Samples of Young Adolescents, *Development and Psychopathology*.

Supplement 1: Overview of analyses

(ri-CLPM) Analyses done in S1 (16 in total):

- 1. Main analysis (general rejection sensitivity), constrained and unconstrained
- 2. Anxious RS, constrained and unconstrained
- 3. Mad RS, constrained and unconstrained
- 4. Peer 3 scenarios of S2, constrained and unconstrained
- 5. Peer 3 scenarios not in S2, constrained and unconstrained
- 6. CLPM, constrained and unconstrained
- 7. Adjusted outliers, constrained and unconstrained
- 8. Log-transformed, constrained and unconstrained

(ri-CLPM) Analyses done in S2 (19 in total):

- 1. Main analysis (general rejection sensitivity), constrained and unconstrained
- 2. Anxious RS, constrained and unconstrained
- 3. Mad RS, constrained and unconstrained
- 4. Peer 3 scenarios, constrained and unconstrained
- 5. Teacher 3 scenarios, constrained and unconstrained
- 6. CLPM, constrained and unconstrained
- 7. Adjusted outliers, constrained and unconstrained
- 8. Log-transformed, constrained and unconstrained
- 9. Multiple-group ri-CLPM (sex), constrained only
- 10. ri-CLPM with time invariant predictor age on the random intercepts, constrained only
- 11. ri-CLPM with time invariant predictor age on the observed scores, constrained only

Supplement 2: Correlations and Cronbach's alpha in Sample 1

Table S1

Sample 1: Internal Consistency and Correlation of Rejection Sensitivity (Based on both the Anxious and Angry Questions)

		Cronbach's α	Pearson's r		
	Three-peer- scenarios (3P-RS)	Other-three-peer- scenarios (O3P-RS)	Six-peer-scenarios (6P-RS)	3P-RS with O3P-RS	3P-RS with 6P-RS
RST T1	.78	.76	.86	.63***	.89***
RST T2	.78	.76	.86	.60***	.89***
RST T3	.76	.79	.87	.68***	.91***
RST T4	.81	.79	.88	.71***	.93***

Note. ****p* < .001

Table S2Sample 2 correlations between Peer and Teacher rejection sensitivity

							Peer				
				RSA			RSM			RST	
			T1	T2	T3	T1	T2	T3	T1	T2	T3
Teacher		T1	.52**	.34**	.29**	.40**	.25**	.20**	.52**	.33**	.27**
	RSA	T2	.45**	.81**	.57**	.35**	.59**	.42**	.45**	.77**	.55**
		T3	.25**	.34**	.58**	.18**	.31**	.41**	.24**	.36**	.55**
	RSM	T1	.35**	.27**	.21**	.55**	.32**	.28**	.50**	.32**	.26**
		T2	.17**	.40**	.34**	.32**	.63**	.36**	.27**	.56**	.38**
		Т3	.16**	.23**	.38**	.22**	.38**	.55**	.21**	.33**	.50**
		T1	.50**	.35**	.28**	.53**	.32**	.27**	.57**	.37**	.30**
	RST	T2	.35**	.68**	.51**	.38**	.69**	.44**	.40**	.75**	.52**
		T3	.23**	.33**	.55**	.22*	.39**	.54**	.25**	.39**	.59**

Note. Matched teacher and peer versions of rejection sensitivity are shown in boldface.

***p* < .01

Supplement 3: Teacher- and peer-related rejection sensitivity

For rejection sensitivity in teacher and peer interactions, the analyses were only done in S2, as in S1 there were no vignettes used that involved teacher scenarios. Three scenarios measured rejection sensitivity in teacher interactions and three scenarios measured rejection sensitivity in peer interactions. We performed exploratory factor analyses in S2 for each wave. We inserted the general rejection sensitivity score of each scenario of that wave in the analyses (three teacher and three peer scenario's). In each wave of rejection sensitivity scenarios, only one factor was found that had an eigenvalue of 1.0 or higher, the next highest value was 0.80 or lower (W1: 0.80, W2: 0.72, W3: 0.80), indicating that peer and teacher rejection sensitivity are not distinguishable from one another. To substantiate this conclusion further, we repeated the ri-CLPM analyses for the teacher and peer scenarios separately in S2.

Sample 2

Teacher-related rejection sensitivity

The constrained ri-CLPM on these teacher scenarios had a good fit (RMSEA < .001, CFI > .999, SRMR = .018, $\chi^2(5) = 4.64$, p = .461). Similar to the general rejection sensitivity model the betweenperson covariance was positive ($\beta = .51$, p < .001), as were the within-person auto-regressive effects for both victimization (T1-T2: $\beta = .18$, p = .018; T2-T3: $\beta = .22$, p = .038) and rejection sensitivity (T1-T2: $\beta = .35$, p < .001; T2-T3: $\beta = .42$, p < .001). No within-person cross-lagged effects were present, neither from victimization to rejection sensitivity (T1-T2: $\beta = .02$, p = .634; T2-T3: $\beta = .03$, p = .635), nor from rejection sensitivity to victimization (T1-T2: $\beta = .07$, p = .225; T2-T3: $\beta = .08$, p = .223). Last, within-person concurrent covariances were mostly significant (T1: $\beta = .18$, p = .019; T2: $\beta = .11$, p = .083; T3: $\beta = .18$, p = .002).

Peer-related rejection sensitivity

The constrained ri-CLPM on these peer scenarios had a good fit (RMSEA < .001, CFI > .999, SRMR = .020, $\chi^2(5) = 4.47$, p = 484). Similar to the general rejection sensitivity model, the between-person covariance was positive ($\beta = .44$, p < .001), as were the within-person auto-regressive effects for both victimization (T1-T2: $\beta = .18$, p = .021; T2-T3: $\beta = .21$, p = .041) and rejection sensitivity (T1-T2: $\beta = .20$, p = .005; T2-T3: $\beta = .23$, p = .016). No within-person cross-lagged effects were present, neither from victimization to rejection sensitivity (T1-T2: $\beta = .09$, p = .124; T2-T3: $\beta = .11$, p = .124), nor from rejection sensitivity to victimization (T1-T2: $\beta = .08$, p = .165; T2-T3: $\beta = .09$, p = .171). Last, within-person concurrent covariances were all positive (T1: $\beta = .28$, p < .001; T2: $\beta = .28$, p < .001; T3: $\beta = .27$, p < .001).

Sample 1

Using the three-peer scenarios of S2

As a check we replicated the peer-related rejection sensitivity analyses in S1, with rejection sensitivity calculated with only the three peer scenarios that were used in S2. The difference from the original analyses is that rejection sensitivity in these analyses is based only on the three peer scenarios that were also used in S2. The constrained model had an acceptable to good fit (RMSEA = .038, CFI = .987, SRMR = .071, $\chi^2(17) = 22.70$, p = .159). The unconstrained model had a good fit (RMSEA = .038, CFI = .038, CFI = .993, SRMR = .049, $\chi^2(9) = 12.09$, p = .208), yet did not outperform the constrained model ($\chi^2(8) = 10.60$, p = .225).

In the constrained model was a positive between-person covariance for victimization and rejection sensitivity ($\beta = .37$, p < .001). The within-person auto-regressive effects were not significant. For rejection sensitivity these betas were also small (T1-T2: $\beta = .05$, p = .622; T2-T3: $\beta = .06$, p = .626; T3-T4: $\beta = .03$, p = .636), whereas they indicated a possible within-person auto-regressive effect for victimization (T1-T2: $\beta = .27$, p = .065; T2-T3: $\beta = .22$, p = .105; T3-T4: $\beta = .19$, p = .094). There were no indications of existing within-person cross-lagged effects from victimization to peer rejection sensitivity (T1-T2: $\beta = .07$, p = .451; T2-T3: $\beta = .07$, p = .449; T3-T4: $\beta = .05$, p = .441) or from peer rejection sensitivity to victimization (T1-T2: $\beta = .08$, p = .223; T2-T3: $\beta = .08$, p = .204; T3-T4: $\beta = .06$, p = .230). The within-person concurrent covariances were positive and mostly significant (T1: $\beta = .31$, p = .017; T2: $\beta = .28$, p = .003; T3: $\beta = .26$, p = .006; T4: $\beta = .16$, p = .056).

Using the peer-scenarios that were not used in S2

As a check, we also repeated the analyses in S1 with the rejection sensitivity score based on the peer scenarios that were *not* used in S2. The constrained model had a good fit (RMSEA = .018, CFI = .997, SRMR = .066, $\chi^2(17) = 18.23$, p = .372). The results were stronger here. There was a positive between-person effect ($\beta = .43$, p < .001), like the previous analysis. Unlike the previous analysis, the within-person auto-regressive effects were near significance for victimization (T1-T2: $\beta = .30$, p = .049; T2-T3: $\beta = .23$, p = .085; T3-T4: $\beta = .20$, p = .077), yet clearly significant for rejection sensitivity (T1-T2: $\beta = .29$, p < .001; T2-T3: $\beta = .25$, p < .001; T3-T4: $\beta = .32$, p = .001). Similar to the previous analysis, the within-person cross-lagged effects were not significant nor large in effect size from victimization to rejection sensitivity (T1-T2: $\beta = .06$, p = .419) nor vice versa from rejection sensitivity to victimization (T1-T2: $\beta = .04$, p = .612; T2-T3: $\beta = .03$, p = .615; T3-T4: $\beta = - .03$, p = .609), and the within-person concurrent covariances were positive and mostly significant (T1: $\beta = .24$, p = .062; T2: $\beta = .24$, p = .001; T3: $\beta = .04$, p = .619; T4: $\beta = .29$, p = .004).

Supplement 4: Subtypes of rejection sensitivity

Sample 1

Anxious rejection sensitivity

The constrained ri-CLPM on anxious rejection sensitivity and victimization had an acceptable fit (RMSEA = .051, CFI = .978, SRMR = .077, $\chi^2(17) = 27.33$, p = .053), whereas the unconstrained model had good fit (RMSEA < .001, CFI > .999, SRMR = .044, $\chi^2(9) = 8.86$, p = .451), and outperformed the constrained model ($\chi^2(8) = 20.79$, p = .008).

The unconstrained model showed a positive between-person covariance of victimization and anxious rejection sensitivity ($\beta = .32$, p < .001). There were no consistent within-person auto-regressive effects. Individual deviations from expected victimization at wave 1 predicted subsequent deviations from expected victimization at wave 2, but no other within-person auto-regressive effects were found for victimization (T1-T2: $\beta = .35$, p = .010; T2-T3: $\beta = .03$, p = .903; T3-T4: $\beta = -0.01$, p = .952). Somewhat similar, individual deviations from expected anxious rejection sensitivity at wave 3 predicted individual deviations from expected rejection sensitivity at wave 4, but we found no other within-person auto-regressive effects for anxious rejection sensitivity (T1-T2: $\beta = -0.08$, p = .658; T2-T3: $\beta = .06$, p = .748; T3-T4: $\beta = .42$, p = .002). Similar to general rejection sensitivity to peer victimization (T1-T2: $\beta = .06$, p = .583; T2-T3: $\beta = -0.08$, p = .522; T3-T4: $\beta = .06$, p = .362), nor from victimization to anxious rejection sensitivity (T1-T2: $\beta = .13$, p = .362; T2-T3: $\beta = -0.04$, p = .805; T3-T4: $\beta = -0.09$, p = .360). Last, in line with the general rejection sensitivity results, we found positive within-person concurrent covariances at nearly each wave (T1: $\beta = .27$, p = .020; T2: $\beta = .029$, p = .038; T3: $\beta = .05$, p = .748; T4: $\beta = .23$, p = .024).

Angry rejection sensitivity

The constrained ri-CLPM on angry rejection sensitivity and victimization had an acceptable to good fit (RMSEA < .001, CFI > .999, SRMR = .059, $\chi^2(17) = 15.62$, p = .551), whereas the unconstrained model had a good fit (RMSEA < .001, CFI > .999, SRMR = .034, $\chi^2(9) = 6.74$, p = 664). However, the unconstrained model did not outperform the constrained model ($\chi^2(8) = 9.00$, p = .342).

In the constrained model there was a positive between-person covariance of victimization and angry rejection sensitivity ($\beta = .41, p < .001$). Similar to general rejection sensitivity, there were consistent positive within-person auto-regressive effects for angry rejection sensitivity (T1-T2: $\beta = .30, p = .002$; T2-T3: $\beta = .25, p = .001$; T3-T4: $\beta = .25, p = .007$) and marginally significant effects for victimization (T1-T2: $\beta = .28, p = .055$; T2-T3: $\beta = .22, p = .097$; T3-T4: $\beta = .19, p = .088$), yet no within-person cross-lagged effects from angry rejection sensitivity to victimization (T1-T2: $\beta = .02, p = .790$; T2-T3: $\beta = .01, p = .788$; T3-T4: $\beta = .02, p = .789$), nor from victimization to angry rejection sensitivity (T1-T2: $\beta = -0.03, p = .770$; T2-T3: $\beta = -0.02, p = .770$; T3-T4: $\beta = -0.20, p = .772$). The within-person concurrent covariances between victimization and rejection sensitivity were almost all

positive and significant (T1: β = .40, *p* = .006; T2: β = .32, *p* < .001; T3: β = .16, *p* = .133; T4: β = .21, *p* = .018).

Sample 2

Anxious rejection sensitivity

The constrained ri-CLPM on anxious rejection sensitivity and victimization had a good fit (RMSEA = .045, CFI = .992, SRMR = .025, $\chi^2(5) = 12.04$, p = .034), as had the unconstrained model (RMSEA = .050, CFI = .998, SRMR = .014, $\chi^2(1) = 2.81$, p = 094), yet the unconstrained model did not outperform the constrained model ($\chi^2(4) = 9.22$, p = .056).

In the constrained model there was a positive between-person covariance of victimization and anxious rejection sensitivity ($\beta = .39$, p < .001). Similar to the general rejection sensitivity results, the within-person auto-regressive effects for anxious rejection sensitivity were present (T1-T2: $\beta = 0.34$, p < .001; T2-T3: $\beta = .39$, p < .001). However, for victimization the within-person auto-regressive effects were only significant from the first to the second wave (T1-T2: $\beta = .17$, p = .037; T2-T3: $\beta = .20$, p = .065). The within-person cross-lagged effects from anxious rejection sensitivity to victimization (T1-T2: $\beta = .13$, p = .052; T2-T3: $\beta = .14$, p = .047) and the within-person cross-lagged effects from victimization to anxious rejection sensitivity were marginally significant (T1-T2: $\beta = .10$, p = .075; T2-T3: $\beta = .12$, p = .060). A Wald test showed that the within-person cross-lagged effects from victimization to rejection sensitivity and from rejection sensitivity to victimization did not statistically differ from one another ($\chi^2(1) = 3.29$, p = .070). The within-person concurrent covariances were all positive, similar to the models with general rejection sensitivity (T1: $\beta = .26$, p = .001; T2: $\beta = .29$, p < .001; T3: $\beta = .25$, p < .001).

Angry rejection sensitivity

The constrained ri-CLPM on angry rejection sensitivity and victimization had a good fit (RMSEA < .001, CFI > .999, SRMR = .014, $\chi^2(5) = 1.42$, p = .923), as had the unconstrained model (RMSEA < .001, CFI > .999, SRMR = .001, $\chi^2(1) = 0.02$, p = .887), which did not outperform the constrained model ($\chi^2(4) = 1.37$, p = .850).

We found a positive between-person covariance of victimization and angry rejection sensitivity ($\beta = .51$, p < .001). Similar to the general rejection sensitivity results, the within-person auto-regressive effects for angry rejection sensitivity were present (T1-T2: $\beta = 0.23$, p < .001; T2-T3: $\beta = .30$, p < .001). However, for victimization the within-person auto-regressive effects were only significant for the first period and marginally significant for the second period (T1-T2: $\beta = .17$, p =.032; T2-T3: $\beta = .21$, p = .066). The within-person cross-lagged effects from angry rejection sensitivity to victimization were not present (T1-T2: $\beta = .03$, p = .529; T2-T3: $\beta = .04$, p = .533), nor were the within-person cross-lagged effects from victimization to angry rejection sensitivity (T1-T2: β = .00, p = .977; T2-T3: $\beta = .00$, p = .977). The within-person concurrent covariances were roughly similar to general rejection sensitivity, with exception of the second wave (T1: β = .22, *p* = .002; T2: β = .11, *p* = .132; T3: β = .21, *p* = .001).

Supplement 5: Exploring sex differences

We examined influence of sex on between- and within-person effects of peer victimization and general rejection sensitivity only in S2 because the sample size of S1 was too small for multigroup analyses.

We constrained the within-person auto-regressive and within-person cross-lagged associations in the ri-CLPM to be equal, similar to the main analyses. We ran a constrained multiplegroup ri-CLPM which had a good fit (RMSEA = .019, CFI = .998, SRMR = .033, $\chi^2(10) = 11.23$, p = .240). We examined the within-person cross-lagged, auto-regressive and concurrent associations, as well as the between-person association of victimization and rejection sensitivity. The overall model of these associations did not differ significantly between boys and girls ($W_T(8) = 3.62$, p = .890). Therefore, we did not interpret the path coefficients of boys and girls separately.

Supplement 6: Exploring age differences

We also examined the influence of age at baseline on peer victimization and general rejection sensitivity only in S2. First, we examined whether age was related to observed scores of victimization and rejection sensitivity. We constrained the within-person auto-regressive and within-person crosslagged associations in the ri-CLPM to be equal, similar to the main analyses. The constrained ri-CLPM with time invariant predictor on observed scores had a good fit (RMSEA < .001, CFI > .999, SRMR = .017, $\chi^2(9) = 8.96$, p = .441). None of the observed scores were significantly predicted by baseline age ($\beta_{PV-T1} = .01$, p = .747; $\beta_{PV-T2} = -.04$, p = .328; $\beta_{PV-T3} = -.05$, p = .215; $\beta_{RST-T1} = -.04$, p = .352; $\beta_{RST-T2} = -.03$, p = .489; $\beta_{RST-T3} = -.07$, p = .064).

Second, we examined whether age was related to the random intercepts of peer victimization and rejection sensitivity. We constrained the within-person auto-regressive and within-person crosslagged associations in the ri-CLPM to be equal, similar to the main analyses. The constrained ri-CLPM with time invariant predictor on the random intercepts had a good fit (RMSEA < .001, CFI > .999, SRMR = .020). Age did not relate significantly to the peer victimization random intercept (β = -.03, *p* = .495), nor to the rejection sensitivity random intercept (β = -.10, *p* = .134).

Supplement 7: Adjusted outliers (general rejection sensitivity)

Sample 1

As the Sample contained outliers in all waves for the victimization and rejection sensitivity (total) variables (2.52% of all possible scores), the constrained ri-CLPM was repeated with the outliers adjusted (Q₁ - 1.5 × IQR; Q₃ + 1.5 × IQR). The constrained model had acceptable to good fit (RMSEA = .030, CFI = .994, SRMR = .061, $\chi^2(17) = 20.51$, p = .249). Results were similar to analyses without adjusted outliers (see Fig. S1): There was positive between-person covariance for victimization and rejection sensitivity ($\beta = .39$, p < .001), nearly all concurrent within-person covariances of victimization and rejection sensitivity were positive as well (T1: $\beta = .21$, p = .086; T2: $\beta = .31$, p < .001; T3: $\beta = .21$, p = .020; T4: $\beta = .22$, p = 0.17), and within-person auto-regressive effects were present, for victimization (T1-T2: $\beta = .27$, p < .001; T2-T3: $\beta = .40$, p < .001; T3-T4: $\beta = .36$, p < .001) and rejection sensitivity (T1-T2: $\beta = .24$, p = .007; T2-T3: $\beta = .23$, p = .004; T3-T4: $\beta = .21$, p = .022). There were no within-person cross-lagged effects, not from victimization to rejection sensitivity, nor from rejection sensitivity to victimization (all betas near zero and non-significant).

Sample 2

As there were also outliers in S2 (2.21% of all possible scores), we repeated the main analyses with the outliers adjusted (Q₁ - 1.5 × IQR; Q₃ + 1.5 × IQR) for the variables victimization and total rejection sensitivity of the six scenarios (3P3T). The constrained model had acceptable fit (RMSEA = 0.054, CFI = 0.989, SRMR = 0.036, $\chi^2(5) = 15.49$, p = .009). The unconstrained model had a good fit (RMSEA = 0.024, CFI > .999, SRMR = 0.009, $\chi^2(1) = 1.42$, p = .233). and outperformed the constrained model ($\chi^2(4) = 14.73$, p = .005). The between-person covariance was positive ($\beta = .49$, p < .001), as were the within-person concurrent covariances of victimization and rejection sensitivity (T1: $\beta = .27$, p = .001; T2: $\beta = .25$, p = .005; T3: $\beta = .26$, p < .001). The within-person auto-regressive effects present for rejection sensitivity (T1-T2: $\beta = .27$, p = .001; T2: $\beta = .42$, p < .001), but, unlike the original analyses, not for victimization (T1-T2: $\beta = -0.03$, p = .829; T2-T3: $\beta = .16$, p = .083). Most within-person cross-lagged effects were non-significant, from victimization to rejection sensitivity to victimization (T1-T2: $\beta = .08$, p = .326; T2-T3: $\beta = .04$, p = .526) and from rejection sensitivity to victimization (T1-T2: $\beta = .09$). Individual deviation in rejection sensitivity at T2 did positively correlate with individual deviation in victimization at T3 (T2-T3: $\beta = .16$, p = .025), see Fig. S2.



Figure S1. Sample 1: Longitudinal within- and between-person effects of victimization and rejection sensitivity with the outliers adjusted *Note.* In the model the stability and cross-lagged paths are constrained to be equal. In the figure the standardized coefficients are shown. The model had an acceptable fit (RMSEA = 0.030, CFI = 0.994, SRMR = 0.061). *p < .042, **p < .01, ***p < .001.



Figure S2. Sample 2: Longitudinal within- and between-person effects of victimization and rejection sensitivity with the outliers adjusted Note. In the model the stability and cross-lagged paths were not constrained. In the figure the standardized coefficients are shown. The model had a good fit (RMSEA = 0.024, CFI > .999, SRMR = 0.009). *p < .042, **p < .01, ***p < .001.

Supplement 8: Log-transformation

Sample 1

Peer victimization and rejection sensitivity had, on each timepoint, a skewed distribution. Therefore, they were log-transformed and the main analyses repeated. The constrained model had a good fit (RMSEA < .001, CFI > .999, SRMR = .044, $\chi^2(17) = 14.63$, p = .623) and outperformed the unconstrained model (RMSEA = .013, CFI = .999, SRMR = .034, $\chi^2(9) = 9.38$, p = .403, loglikelihood test: $\chi^2(8) = 5.22$, p = .734). Results were similar to the main analyses (see Fig. S3). There was positive between-person covariance for victimization and rejection sensitivity ($\beta = .38$, p < .001), nearly all concurrent within-person covariances of victimization and rejection sensitivity were significant (T1: $\beta = .25$, p = .011; T2: $\beta = .25$, p = .002; T3: $\beta = .17$, p = .050; T4: $\beta = .21$, p = 0.008), and the within-person auto-regressive effects were also present, both for victimization (T1-T2: $\beta = .29$, p = .016; T2-T3: $\beta = .26$, p = .035; T3-T4: $\beta = .23$, p = .040) and rejection sensitivity (T1-T2: $\beta = .25$, p = .002; T2-T3: $\beta = .22$, p = .004; T3-T4: $\beta = .22$, p = .009). Yet, no within-person cross-lagged effects, not from victimization to rejection sensitivity, nor from rejection sensitivity to victimization (all betas near zero and non-significant).

Sample 2

In S2 the variables peer victimization and rejection sensitivity were skewed as well. Hence, the main analyses were repeated with log-transformed variables. The constrained model had a good fit (RMSEA = .028, CFI = .997, SRMR = .029, $\chi^2(5) = 7.78$, p = .169) and outperformed the unconstrained model (RMSEA < .001, CFI > .999, SRMR = .003, $\chi^2(1) = 0.20$, p = .655, loglikelihood test: $\chi^2(4) = 7.96$, p = .093). Results clearly diverged from the main analyses (see Fig. S4). In the log-transformed variable analyses, all within-person cross-paths were significant. Individual fluctuations in peer victimization at one time-point was positively related to fluctuations in rejection sensitivity half a year later (T1-T2: $\beta = .12$, p = .034; T2-T3: $\beta = .14$, p = .030), and vice versa, individual fluctuations in rejection sensitivity were positively related to fluctuations in peer victimization half a year later (T1-T2: $\beta = .15$, p = .010; T2-T3: $\beta = .16$, p = .010). The other effects were similar to the main analyses. There was positive between-person covariance for victimization and rejection sensitivity ($\beta = .40, p < .001$), and all concurrent within-person covariances of victimization and rejection sensitivity were positive as well (T1: β = .31, p < .001; T2: β = .26, p < .001; T3: $\beta = .27$, p < .001). The within-person auto-regressive effects were also present, both for victimization (T1-T2: $\beta = .20$, p = .007; T2-T3: $\beta = .22$, p = .016) and rejection sensitivity (T1-T2: $\beta =$ $.25, p = .001; T2-T3: \beta = .27, p = .004).$



Figure S3. Sample 1: ri-CLPM with longitudinal effects of log-transformed victimization and rejection sensitivity Note. In the model the stability and cross-lagged paths are constrained to be equal. In the figure the standardized coefficients are shown. The model had an acceptable fit (RMSEA < .001, CFI > .999, SRMR = .044). *p = .05, *p < .042, **p < .01, ***p < .001.



Figure S4. Sample 2: ri-CLFM with longitudinal effects of log-transformed victimization and rejection sensitivity

Note. In the model the stability and cross-lagged paths are constrained to be equal. In the figure the standardized coefficients are shown. The model had an acceptable fit (RMSEA = .028, CFI = .997, SRMR = .029). *p < .042 **p < .01, ***p < .001.

Supplement 9: CLPMs

Sample 1

To point out the effect of disentangling between- and within-person effects, a traditional CLPM was estimated for victimization and the six peer scenarios of rejection sensitivity (total). The CLPM with auto-regressive and cross-lagged paths constrained to be equal over time fitted the data borderline acceptable (RMSEA = .079, CFI = .934, SRMR = .086, $\chi^2(20) = 50.45$, p < .001). The unconstrained CLPM had unacceptable fit (RMSEA = .105, CFI = .933, SRMR = .070, $\chi^2(17) = 202.06$, p < .001), and did not outperform the constrained model ($\chi^2(9) = 8.43$, p = .491).

In the constrained CLPM (see Fig. S5), auto-regressive paths of both victimization (T1-T2: β = .58, p < .001; T2-T3: β = .54, p < .001; T3-T4: β = .47, p < .001) and rejection sensitivity (T1-T2: β = .67, p < .001; T2-T3: β = .68, p < .001; T3-T4: β = .68, p < .001) were significant and positive. Furthermore, the concurrent relations were also positive (T1: β = .32, p < .001; T2: β = .34, p < .001; T3: β = .17, p = .003; T4: β = .29, p < .001 respectively). The cross-lagged paths were all non-significant and the beta values were smaller than 0.10.

Sample 2

Similarly, a traditional CLPM was estimated for victimization and the six scenarios of rejection sensitivity (total) in S2. The CLPM with auto-regressive and cross-lagged paths constrained to be equal over time fitted the data borderline acceptable (RMSEA = .074, CFI = .960, SRMR = .041, $\chi^2(8) = 39.54$, p < .001). The unconstrained model also had an unacceptable fit (RMSEA = .114, CFI = .954, SRMR = .039, $\chi^2(4) = 40.69$, p < .001) and did not outperform the constrained model in a loglikelihood test ($\chi^2(5) = 4.00$, p = .666).

In the constrained model (see Fig. S6), auto-regressive paths of both victimization (T1-T2: β = .53, p < .001; T2-T3: β = .57, p < .001) and rejection sensitivity (T1-T2: β = .50, p < .001; T2-T3: β = .58, p < .001) were significant and positive. Furthermore, the cross-lagged paths from victimization to rejection sensitivity were significant and positive (T1-T2: β = .08, p = .001; T2-T3: β = .09, p = .002), as were the cross-lagged paths from rejection sensitivity to victimization (T1-T2: β = .09, p < .001; T2-T3: β = .09, p < .001; and the concurrent associations between rejection sensitivity and victimization (T1: β = .35, p < .001; T2: β = .23, p < .001; T3: β = .25, p < .001).



Figure S5. Sample 1: CLPM with longitudinal effects of victimization and rejection sensitivity

Note. In the model the stability and cross-lagged paths are constrained to be equal. In the figure the standardized coefficients are shown. The model had an acceptable fit (RMSEA = 0.079, CFI = 0.934, SRMR = 0.086). *p < .042, ***p < .001.



Figure S6. Sample 2: CLPM with longitudinal effects of victimization and rejection sensitivity

Note. In the model the stability and cross-lagged paths are constrained to be equal. In the figure the standardized coefficients are shown. The model had an acceptable fit (RMSEA = 0.074, CFI = 0.960, SRMR = 0.041). *p < .042, **p < .01, ***p < .001.