# **Supplementary Online Content**

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This supplementary material has been provided by the authors to give readers additional information about their work.

## eMethods. Supplementary Methods

**Data Matching.** An age-sex-matched data set of no-to-low and heavy drinkers was created using the maximum bipartite matching algorithm  $1, 2$ . The method first constructed a bipartite graph such that the first set of nodes represented 291 no-to-low drinkers and the second set of nodes represented 160 heavy drinkers. An edge was connected between a no-to-low drinker and a heavy drinker if they were of the same gender and had a gap of subject-age (average age over visits of no-to-low/heavy drinking, eFigure 1b) less than 6 months. A Ford-Fulkerson algorithm  $3$  was then applied to select a maximum number of matching pairs. This process resulted in a matched dataset of 78 no-to-low and 78 heavy drinkers.

**Participants.** Of all the 831 NCANDA participants, 782 (94% of the participants) had at least 2 usable DTI scans by the  $5<sup>th</sup>$  year of the study, 699 (84%) has at least 3 scans, 588 (71%) had at least 4 scans, 373 (45%) had all 5 scans.

eTable describes the demographics of the NCANDA participants used in the study. The whole cohort is composed 451 participants who had at least 2 usable DTI scans by the  $5<sup>th</sup>$  year and were labelled as no-to-low drinking (adjusted Cahalan  $4 = 0$ ) throughout the study or as heavy drinking (adjusted Cahalan  $\geq$  2) for at least two consecutive visits. Two age-sex matched groups of 78 no-to-low and 78 heavy drinkers were selected from the whole cohort. 63 out of the 160 heavy drinkers were classified as transitioners as they remained no-to-low drinking for at least two visits before initiating heavy drinking.

**MRI Preprocessing.** The structural and diffusion data of all NCANDA participants were preprocessed using the publicly available longitudinal NCANDA pipeline <sup>5</sup>. Skull stripping and aligning with the SRI 24 atlas  $6$  were performed by registering the T1w of each visit to the baseline T1w and registering the baseline to the atlas with ANTS  $^7$ . For each visit, the b0 scan of the DTI sequence was aligned to the T1w MRI by aligning the b0 to the T2w scan via ANTS and aligning the T2w to the T1w via CMTK <sup>8</sup>. Besides skull stripping, the pipeline performed removal of bad single shots, echo-planar structural distortion, Eddy-current distortion correction, fractional anisotropy (FA) estimation by CAMINO<sup>9</sup>, and FA skeleton estimation by Tract-Based Spatial Statistics (TBSS) <sup>10</sup>. The average FA value over the whole-brain TBSS skeleton and with respect to the parcellation defined by the Johns Hopkins University (JHU) DTI atlas <sup>11</sup> was computed for each scan and corrected for manufacturer difference based on humanphantom data <sup>5</sup>.

**Slope Estimation.** The group-level DTI trajectory of the 291 no-to-low drinking youths was estimated by a mixed effects model <sup>12</sup>, which used a stepwise regression <sup>13</sup> to determine the highest order of the aging effect (eFigure 1a). The model was initialized as a linear function of age (with other covariates) and higher-order terms were successively added until the coefficient was not significantly different from 0 (*p*-value > 0.05) according to the t-test performed within the GLM. In doing so, the model used a cubic function of age as the fixed effect and incorporated a linear random effect of age (a random intercept and a random slope) for each participant. Confounders including sex, ethnicity, supratentorial volume, pubertal development, socioeconomic status and

manufacturer type were considered as covariates in the mixed effects model. For each participant, confounding effects (not including age) were residualized from the FA measures based on the fixed effects estimated within the mixed effects model. To compute the slope measure for each no-to-low drinker, a linear model was fitted between age and residualized FA across all visits (eFigure 1b). For each heavy drinker, the slope was computed with respect to the visits during which the participants engaged in heavy drinking.

### eResults. Supplementary Results

**Analysis on the Slopes of Transitioners.** An Analysis of Covariance (ANCOVA) model compared the slopes among the 291 no-to-low drinkers, the 97 heavy drinkers excluded in the within-subject analysis, and the 63 transitioners before and after drinking onset, while controlling age as a covariate. eFigure 3a indicates that the slopes of the 291 no-to-low drinkers were not significantly different from those of the transitioners prior to drinking onset (95% CI of group difference [-0.0010,0.0002], two tailed  $p = 0.09$ ) but significantly larger than the slopes of transitioners after drinking onset (95% CI of group difference [0.0020,0.0037]*, p* < 0.001). On the other hand, the slopes of the 97 heavy drinkers were not significantly different from those of the transitioners after drinking onset (95% CI of group difference [-0.0003,0.0023], two tailed  $p = 0.13$ ) but significantly smaller than the slopes of transitioners before drinking onset (95% CI of group difference [-0.0048,-0.0023]*, p* < 0.001).





(a) The group-level developmental trajectory of 291 no-to-low drinking youths; (b) the slope quantifies the developmental change of FA across visits.

**eFigure 2.** Extension of Figure 2d in the Main Text



Slopes of the 291 no-to-low drinkers vs. slopes of the 97 heavy drinkers excluded in the within-subject analysis vs. slopes of the 63 transitioners before and after drinking onset.

### **Estimating Transition Age via Piecewise Linear Model.** For each of the 63

transitioners**,** a piecewise linear model <sup>14</sup> was fitted to the residualized FA across visits.

The model considered two linear segments and used the `residual sum of squares'

algorithm <sup>14</sup> to automatically determine the breakpoint. This breakpoint was compared to the 'actual' age of drinking onset, which was defined as the average age between the two consecutive visits that the participant transitioned from no-to-low to heavy drinking. A paired *t*-test resulted in a two-tailed *p*-value of 0.53, indicating that the data-driven breakpoints (change of slope) were not significantly different from the 'actual' age of transition to heavy drinking.

**eFigure 3.** Piecewise Linear Regression to Estimate Drinking Onset



Breakpoints of the 63 transitioners estimated by piecewise linear model vs. the actual age of transition from no-to-low to heavy drinking.

**eFigure 4.** Correlation Between Confounding Variables and Slope of FA



The slope of FA did not significantly correlate with the log of nicotine use, log of marijuana use, or the number of visits of the 156 age-sex-matched subjects ( $p > 0.1$ ).

**Age-alcohol Interaction on the Whole Cohort.** On the entire cohort of 451 youth, a general linear model (GLM) examined the subject-age-by-drinking-group interaction on the slope measures with additional covariates being subject-age, sex, and drinkinggroup. The coefficient associated with the interaction term was statistically significant (two-tailed *p*<0.001).

**eFigure 5.** Extension of Figure 3b in the Main Text



The age-drinking-group interaction on the slope in the 291 no-to-low and 160 heavy drinkers.

## **Differential Alcohol Effects in Younger and Older Adolescents.**



**eFigure 6.** Extension of Figure 4 in the Main Text

Alcohol effects on the slope within the corticospinal and fasciculi tracts <sup>11</sup> were significant in the younger cohort (age < 19.3 years) but not in the older cohort.



# **eTable.** Demographics of the NCANDA Participants

**\*** Distribution is characterized by mean±std

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