

Supplementary Information for

Systematic analysis of differential rhythmic liver gene expression mediated by the circadian clock and feeding rhythms

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SI References

Other supplementary materials for this manuscript include the following:

Datasets S1 to S3

Supplementary Information Text

SI Materials and Methods

RNA extraction and RNA-Seq. ~100 mg of frozen tissue was grinded with a Polytron PT 2500 E homogenizer in extraction buffer consisting of 3.9 M guanidium thiocyanate, 0.03 M sodium citrate, 0.2 M sodium acetate, and 1% (vol/vol) 2-mercaptoethanol. An equal volume of 0.5 volume of chloroform/isoamyl alcohol (49:1 [vol/vol]) and phenol (saturated in H₂O) was added to the homogenate. The mixture was vortexed and then centrifuged for 20 min at 12,000 × *g* at 4 °C for phase separation. Precipitation of RNA of the aqueous phase was performed in an equal volume of isopropanol at −20 °C for at least 20 min before precipitate was pelleted by centrifugation for 15 min at 12,000 × *g* at 4 °C. The pellet was resuspended in 4 M lithium chloride and then repelleted by centrifugation for 15 min at 12,000 × *g* at 4 °C. Subsequently, the pellet was washed with 75% ethanol for 15 min at 12,000 × *g* at 4 °C, dried at room temperature, and finally dissolved in RNase/DNase free water. Quantification of total RNA was done with Ribogreen (Life Technologies), and RNA quality was assessed with a Fragment Analyzer (Advances Analytical).

Libraries of the *Bmal1* KO and *Cry1/2* KO animal series were prepared using the TruSeq Stranded mRNA Library Prep protocol (Illumina) following the manufacturer's protocol. We used 250 ng of total RNA as starting material and performed 11 PCR cycles for library amplification. The optimal number of PCR cycles was assessed by qPCR (Kapa BioSystems). All 96 samples (Fig. 1B) were pooled, denatured, spiked with 3% of the PhiX Control v3 Library and loaded onto a paired-end read flow cell v3 at a final concentration of 9 pM for a paired-end 126 cycles run on a HiSeq 2500 (Illumina). Sequencing depth was equivalent to six samples per lane. For the data on *Hlf/Dbp/Tef* KO mice, sequencing libraries were prepared from 3 μg of total RNA using the TruSeq Stranded mRNA Library Prep Kit following the manufacturer's protocol. Purified libraries were quantified with Qubit DNA HS (Thermo Fisher Scientific), and profile analyses were retrieved on a TapeStation TS4200 (Agilent). Paired-end sequencing with 75 cycles were performed on a NextSeq 500 following Illumina's recommendations. We mapped the reads on the Ensembl *M. musculus* genome (GRCm38/mm10) using STAR-2.7.3a (1) with the annotation from Ensembl release 98 (2). To count uniquely mapped reads per gene, we used the quantMode GeneCounts option of STAR.

Modeling Temporal Gene Expression Profiles across Multiple Conditions. To assess rhythmicity and mean differences of gene expression in RNA-Seq count data, we developed a model selection framework based on generalized linear models (GLMs) (*SI Appendix*, Fig. S1A). As proposed (3), we modeled counts Y_{gs} that have been uniquely mapped to a gene g in a sample s as a negative binomial (NB) with a fitted mean μ_{gs} and a gene-specific but sample-independent dispersion parameter θ_g .

$$Y_{gs} \sim NB(\mu_{gs}, \theta_g)$$

The fitted mean is proportional to the quantity q_{gs} of fragments that correspond to a gene g in a sample s scaled by a sample-specific scaling factor λ_s (3). This scaling factor depends on the sampling depth of each library and was estimated using the median-of-ratios method of DESeq2 (3).

$$\mu_{gs} = \lambda_s q_{gs}$$

A gene-specific distribution θ_g was estimated using empirical Bayes shrinkage (4).

The fit was performed using a GLM with a logarithmic link function implemented in DESeq2. Sample specific size factor (λ_s) was defined as an offset. The full GLM was defined as follows:

$$\log_2(\mu_{gbc_s}) = m_{gc}^1 + m_{gb}^2 + \alpha_{gc} \cos(\omega t(s)) + \beta_{gc} \sin(\omega t(s)) + \log_2(\lambda_s)$$

where μ_{gbc_s} is the mean of the NB distribution for gene g and condition c in sample s at Zeitgeber/circadian time $t(s)$. The index b refers to a batch of samples. The parameters α_{gc} and β_{gc} are coefficients of the cosine and sine functions, respectively. m_{gc}^1 is a coefficient to describe a condition specific mean expression level. We included a batch specific m_{gb}^2 to account for technical batch effects when detected in the data. Specifically, we included m_{gb}^2 to account for a batch effect observed for one series of *Cry1/2* KO mice under *ad libitum* feeding regimen (GSM4037515 – GSM4037520). To select an optimal gene-specific model, we proceeded in two steps. First, we assessed rhythmicity across the different conditions, where the parameters m_{gc}^1 and m_{gb}^2 were free but the parameters α_{gc} and β_{gc} had constraints depending on the particular model considered. To this end, we defined different models for four (e.g., *Bmal1* WT, *Bmal1* KO, *Cry1/2* WT, *Cry1/2* KO) or two conditions (i.e., *Hlf/Dbp/Tef* KO and *Hlf/Dbp/Tef* WT). Models were defined to have either zero (nonrhythmic pattern) or nonzero (rhythmic pattern) α and β coefficients for each analyzed condition. Moreover, for some models, the values of α and β can be also shared within any combination of the four conditions (*SI Appendix*, Fig. S1 E and F). For example, for four conditions, there are 52 such models (*SI Appendix*, Fig. S1A). The coefficients α and β were used to calculate the phase [$\text{atan2}(\beta, \alpha)$] and amplitude (\log_2 fold-change peak-to-trough; $2\sqrt{\alpha^2 + \beta^2}$) of a gene. Bayesian information criterion (BIC) based model selection was employed to account for model complexity (5) using the following formula:

$$\text{BIC}_j = k \log(n) - 2\log(\hat{L}_j).$$

\hat{L}_j is defined as the log-likelihood of the model j from the regression, n is the number of data points, and k is the number of parameters. To assess the confidence of the selected model j , we calculated the Schwarz weight (BICW):

$$\text{BICW}_j = \frac{e^{\frac{1}{2}\Delta\text{BIC}_j}}{\sum_{m=1}^M e^{\frac{1}{2}\Delta\text{BIC}_m}},$$

where ΔBIC_j is defined as the difference in BIC between model j and the minimum BIC value in the entire model set BIC_{m^*} :

$$\Delta\text{BIC}_j = \text{BIC}_j - \text{BIC}_{m^*}.$$

BICW_j is interpreted as the probability for model j . The model with the highest BIC was selected as the optimal model within the set of all defined models. In the second step, we analyzed the mean levels (*SI Appendix*, Fig. S1A) and thus set the coefficient α and β to the values of the selected model in the first regression. A new regression was performed where the parameter m_{gc}^1 was either free or subject to constraints between conditions based on several competing models. We considered all possible combinations for the mean coefficient m_{gc}^1 with differing or shared means between conditions. For example, for four conditions, there are 15 such models (*SI Appendix*, Fig. S1 A and D). Each model was solved using generalized linear regression, and each gene was assigned to a preferred model based on the BICW as described above for the first iteration. Cook's distance is an estimate of how much a data point influences the fitted coefficients for a gene. A large value of Cook's distance is considered to be an outlier. Thus, genes that did not reach at least a BICW above 0.4 for the preferred model or genes that had at least a Cook's distance of more than 1 were categorized as "ambiguous" (model 0). In the case of two conditions (i.e., PARbZip KO dataset) the threshold on BICW was set to 0.95. The difference in the two BICW thresholds reflected that the number of possible models is much larger for four compared with two conditions. Also, we considered only genes with a \log_2 amplitude of at least 0.25 to be rhythmic in the corresponding condition. To assess temporal variation of normally distributed measurements (e.g., relative liver size, nuclear abundance of

PPAR α), we modified model selection approach described above to deal with gaussian noise and therefore used linear models. A function handling normally distributed data is implemented in the *dryR* package. To assess differences in mean expression levels between *Bmal1* and *Cry1/2* KOs, we calculated a double-difference score that is the difference of the log₂ fold-changes in mean levels of *Bmal1* KO vs. *Bmal1* WT and *Cry1/2* KO vs. *Cry1/2* WT:

$$\Delta\Delta(Bmal1_{KO}, Cry1/2_{KO}, Bmal1_{WT}, Cry1/2_{WT}) = (m_{Bmal1\ KO} - m_{Bmal1\ WT}) - (m_{Cry1/2\ KO} - m_{Cry1/2\ WT}),$$

where m is the condition specific log₂ mean expression level retrieved from *dryR*.

Functional and Gene Set-Enrichment Analysis. We tested enrichment for annotated gene sets of several sources including Gene Ontology (GO) (6), Molecular Signatures Database (MSigDB) (7), Kyoto Encyclopedia of Genes and Genomes (KEGG) (8) and Reactome Pathway database (9). For phase-shifted rhythmic genes in *Cry1/2* and *Bmal1* KO animals, we performed an untargeted ChIP-enrichment analysis as described in ref. (10) to screen for potentially interesting transcription factors. Target genes for circadian clock regulators and PPAR α were defined as outlined below. We used GOseq 3.1 (11) to assess overrepresentation of all gene sets for genes assigned to a statistical model. Temporal-resolved enrichments were calculated using a Fisher's exact test. We defined foreground genes as genes with a peak phase within in a sliding 4-h window. The gene universe was defined as all expressed KO genes. For each gene set, the window was moved by 0.1 h to calculate the P value.

Simulated Count Data and Comparing *dryR* with Other Methods. We simulated rhythmic count data from a NB distribution for four conditions and 5,000 features in an interval of 4 h using Symphony (12). The ratio of rhythmic genes in each condition was set to 25%. The period was 24 h with a minimum amplitude (log₂ peak-to-trough) of 0.25 and a maximum of 6. The mean was between 10⁰ to 10⁶ counts. To test the aspect of rhythm detection across four conditions, we considered rhythm detection as true positive if all conditions were correctly detected as rhythmic or nonrhythmic. We compared *dryR* to the following alternative methods for rhythmicity detection algorithm: Jonckheere-Terpstra-Kendall (JTK) (13), Lomb-Scargle (14) and meta2d (15) on log-transformed count data. To compare *dryR* to a harmonic regression without using the BIC model selection, we fitted the full GLM model outlined in the previous paragraph. Subsequently, we employed a Wald test to test for significance of the α or β coefficients in a negative binomial GLM using DEseq2 (4). The retrieved P values were corrected for multiple testing according to Benjamini and Hochberg (16). The simulated data to test the performance for differential rhythmicity analysis was generated similarly as described above with minor modifications. We simulated data for 2 conditions with only rhythmic genes. Two datasets were generated with a phase shift of 0 or 4 h and an amplitude change of log₂ fold-change of 0 or 1 between the two conditions. The standard datasets had a sampling interval of 4 h and consisted of two replicates. Differences in amplitude and phase were assessed with *dryR*. We compared the performance with CircaCompare (17), LimoRhyde (18) and DODR (19) on log-transformed data. CircaCompare was run with either default settings or an alpha threshold of 1 (test for presence of rhythmicity). *dryR* was run either with or without a threshold for a minimum phase difference of 2 h or minimum log₂ peak-to-trough of 0.25. For the extended analysis using simulated datasets, we varied the sampling interval (i.e., 1, 2, and 6 h), or we changed the number of replicates (i.e., one, three, four, and eight replicates) while keeping the other parameter at the standard value. We also generated a standard dataset with one, two, or three missing samples in each condition.

Data resources and analysis. Target genes for clock genes and PPAR α in mouse liver have been identified by published ChIP-Seq. ChIP-Seq data was retrieved from the National Center for Biotechnology Information (NCBI) Gene Expression Omnibus (GEO) (20) for BMAL1, CLOCK, NPAS2, CRY1, CRY2, PER1 and PER2, which are available under GEO Series accession no. [GSE39977](#) (21); NR1D1 and NR1D2 are available under GEO Series accession no. [GSE34020](#) (22). ChIP-

Seq data for NFIL3 and DBP were retrieved from the European Nucleotide Archive (ENA) (23) under accession no. [PRJDB7796](#) (24). We mapped reads on the Ensembl *M. musculus* genome (GRCm38/mm10) using STAR-2.7.3a (1). Peaks were called using macs/2-2.1.1 (25) and peak annotation was performed with CHIPseeker 3.10 (26) using the genome annotation of Ensembl release 98 (2). Binding sites were visualized using the Integrative Genomics Viewer (27). Proteomics data were retrieved from the study of ref. (28), which are available in the ProteomeXchange Consortium (29) under accession number [PXD003818](#). RNA-Seq data for *Nfil3* KO mice series in liver were retrieved from ENA (accession no. [PRJDB7789](#)) (24). Mapping and downstream analysis of the published reads followed the same procedures outlined in section RNA extraction and RNA-Seq.

Web application. We built a web application (<https://clockprofile.epfl.ch>), providing a user-friendly interface of the datasets and the statistical analysis. The application was built using the Ruby-on-Rails framework. The backend consists in a PostgreSQL relational database in which analysis results are stored. In the user interface, dynamic plots are produced with the Plotly.js library.

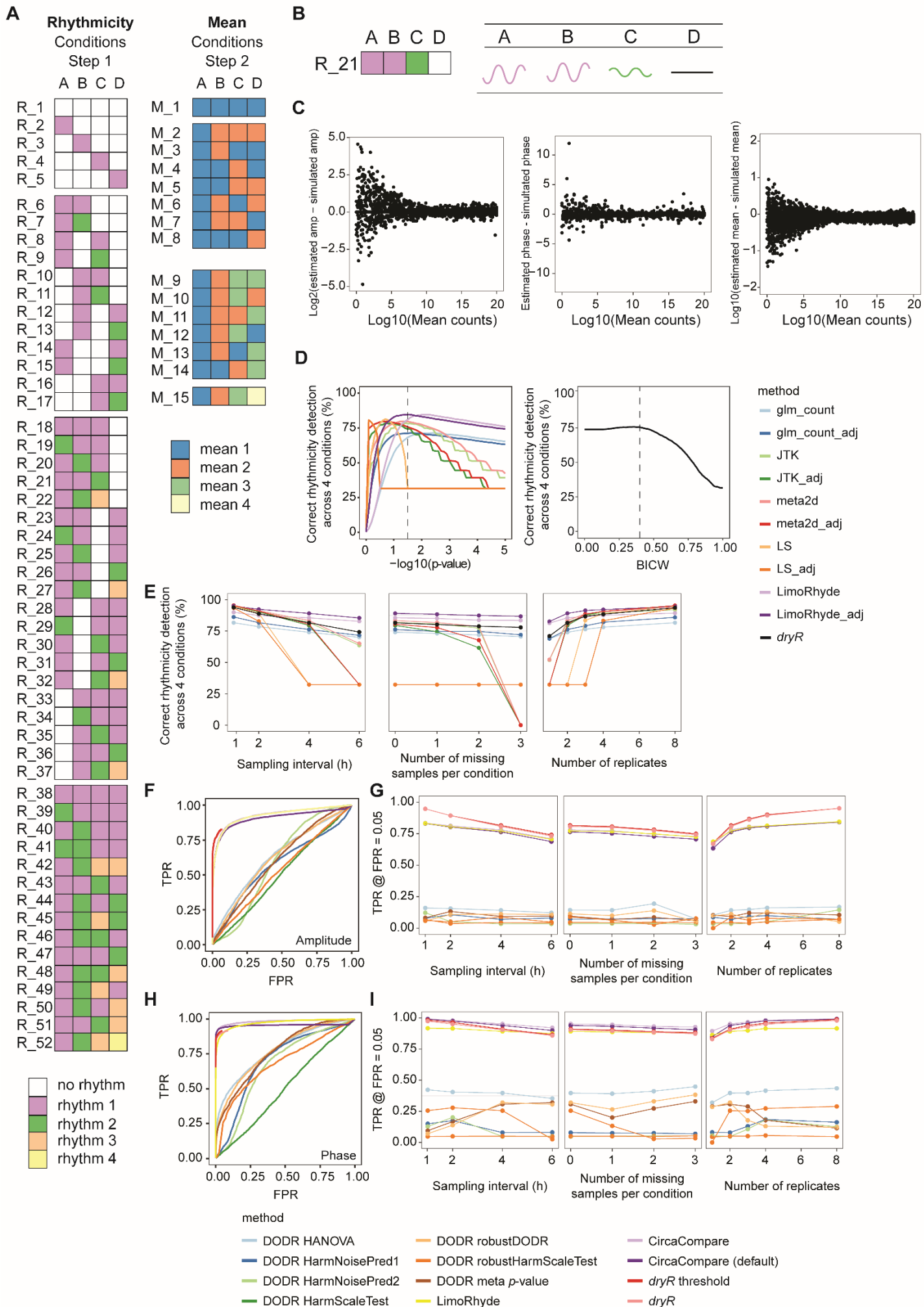


Fig. S1. Benchmarking of *dryR* to analyze rhythmic datasets that consist of multiple conditions.

A. Overview of all rhythmic and mean models that are fitted in case the data consists of four conditions. Same color indicates shared mean levels and rhythmic parameters between indicated conditions for mean and rhythmic models, respectively.

B. Schematic definition of an example rhythmic model. For model R_21 rhythmic parameters (*i.e.*, amplitude and phase) are shared between condition A and B. Condition C shows an altered rhythmicity pattern compared to condition A and B. No rhythm is detected for condition D.

C. Scatterplots to assess the differences of predicted and simulated values in function of mean expression for amplitude (left), phase (middle) and mean levels (right).

D. Rhythm detection across four conditions in a simulated dataset using *dryR* and other methods including JTK, Lomb-Scargle (LS), ARSER, meta2d and LimoRhyde on log transformed count data and a cosin fit on count data using generalized linear model (GLM_cos). *P* values were adjusted for multiple testing using Benjamini-Hochberg correction if indicated (*_adj*). The dashed lines represent a typical threshold used for rhythm detection (*P* value ≤ 0.05 and BICW ≥ 0.4).

E. Percentage of correctly detected rhythms across 4 condition at a typical threshold (*P* value ≤ 0.05 and BICW ≥ 0.4) for datasets with altered sampling intervals (left), number of missing samples in each of the 4 conditions (middle) and number of replicates (right).

F-G. ROC curves illustrating false positive rate (FPR, 1-specificity) and true positive rate (TPR, sensitivity) to detect a 2-fold change in amplitude (F) and a 4 hours phase difference (G) between two conditions. We compared *dryR* with CircaCompare (default settings and an alpha threshold of 1), LimoRhyde and DODR. For *dryR*_threshold, only features with differences in amplitude ≥ 0.25 and phase ≥ 2 h were considered, respectively. In contrast to DODR and CircaCompare which use *P* values to indicate differential rhythmicity of a certain feature, *dryR* provides likelihoods for a set of fitted models, which can then be converted to probabilities using the Bayesian Information Criterion weight (BICW, see methods). The model with the highest BICW is selected. ROC curves for *dryR* differ in shape from the other methods as loosening the decision threshold for the *p*-value based methods will ultimately result in a TPR/FPR ratio of 1, as true positives and false positives will reach their maximum number. A loosening of BICW threshold will still result in a selection of a model and therefore will not result in a TPR/FPR of 1 as the detection of false positive and true positive feature will reach a plateau.

H-I. Level of TPR at an FPR of 0.05 for differential rhythmicity analysis for amplitude (H) and phase (I) in synthetic rhythmic datasets with altered sampling intervals (left), number of missing samples in each of the four conditions (middle) and number of replicates (right).

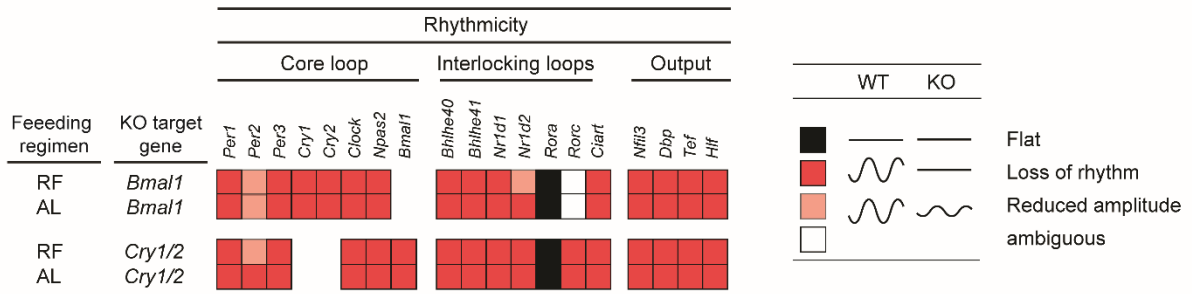
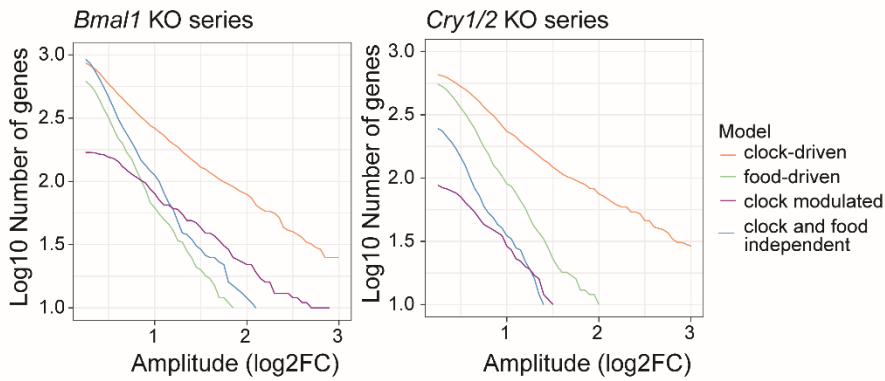
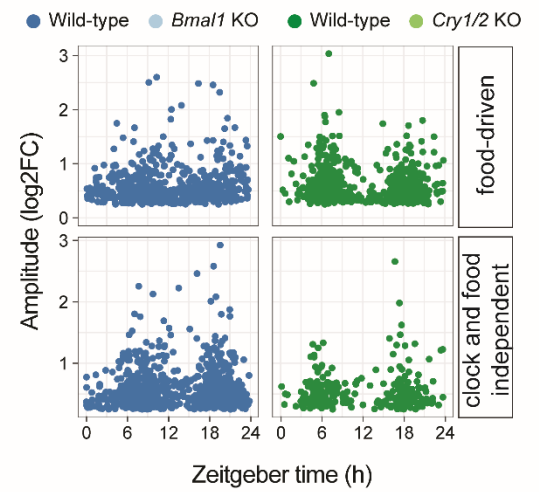
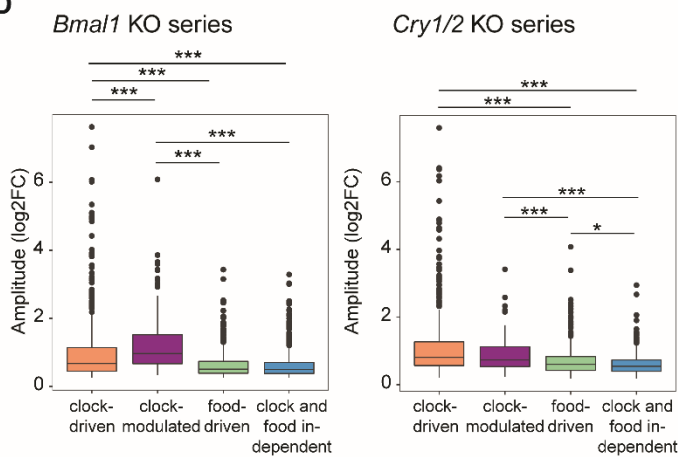
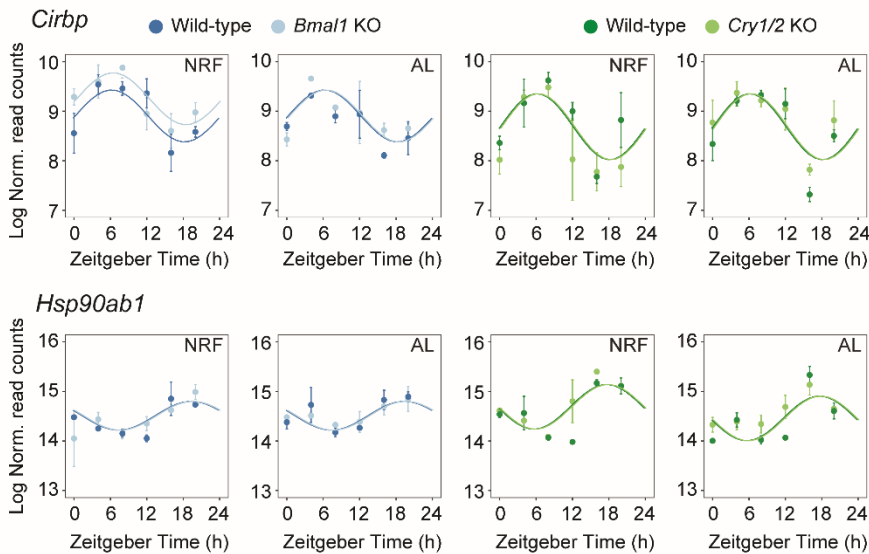
A**B****C****D****E**

Fig. S2. Rhythmicity in *Bmal1* and *Cry1/2* KO mice in *ad libitum* (AL) or night restricted feeding (NRF) conditions.

A. Assignment of circadian clock genes to their corresponding temporal expression pattern (right).

B. Cumulative number of genes that show a larger amplitude than the value on the x-axis in the indicated model: circadian clock-driven genes (orange) have higher amplitudes than food-driven (green), clock and food independent rhythmic (blue), or clock-modulated genes (purple).

C. Amplitude vs. phase of rhythmic genes of indicated rhythmic model.

D. Amplitude represented as boxplot for the indicated model.

E. Temporal expression pattern of the cold inducible RNA-binding protein (*Cirbp*) and the heat shock protein *Hsp90ab1*. Rhythmic gene expression is food and clock independent for these genes as rhythmicity is kept under all conditions.

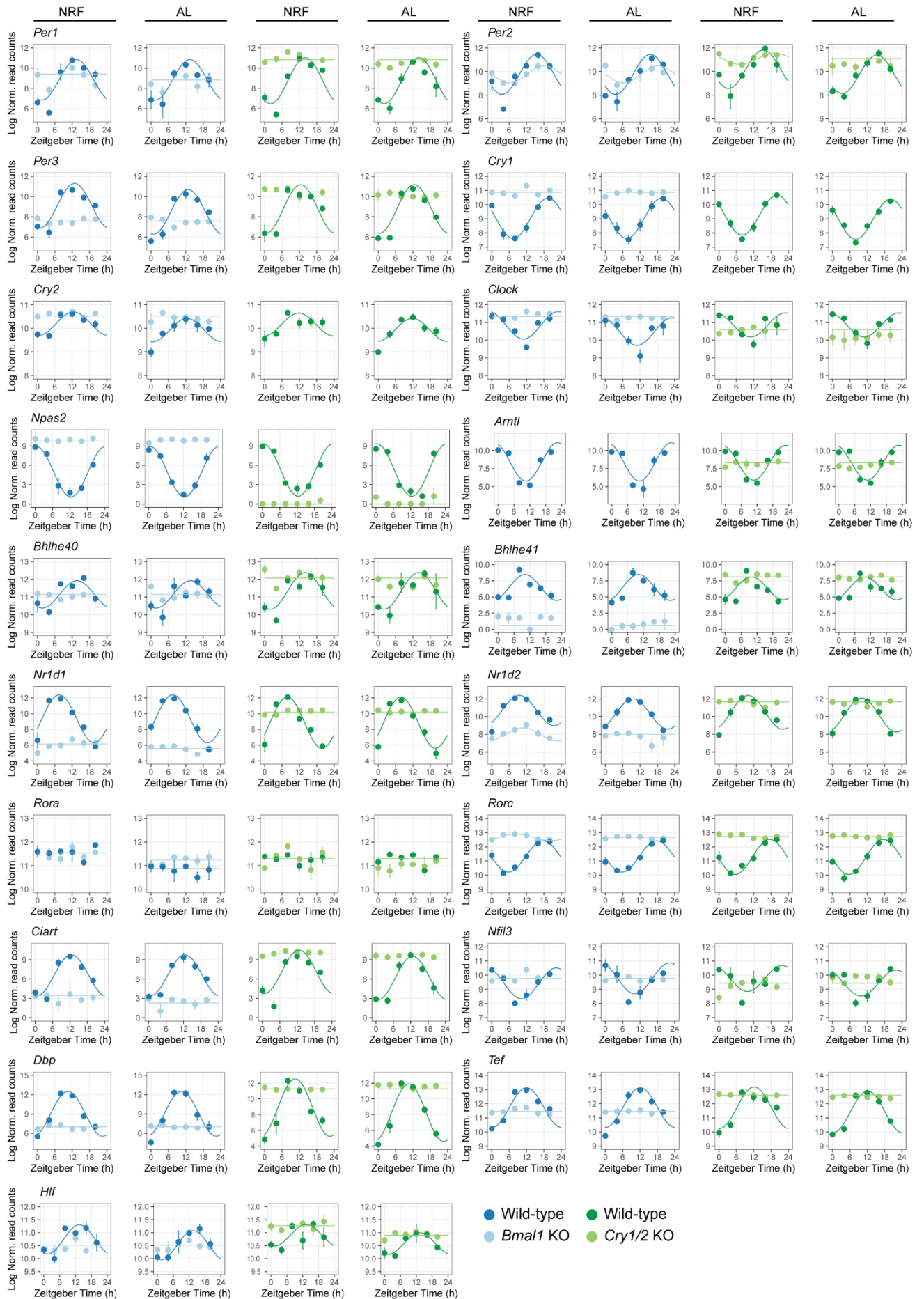


Fig. S3. Rhythmic expression of clock genes is blunted in the liver of *Bmal1* KO and *Cry1/2* KO mice, independent of feeding patterns. Circadian clock gene expression in *Bmal1* WT, *Bmal1* KO, *Cry1/2* WT, and *Cry1/2* KO under an *ad libitum* (AL) and night restricted feeding (NRF) regimen.

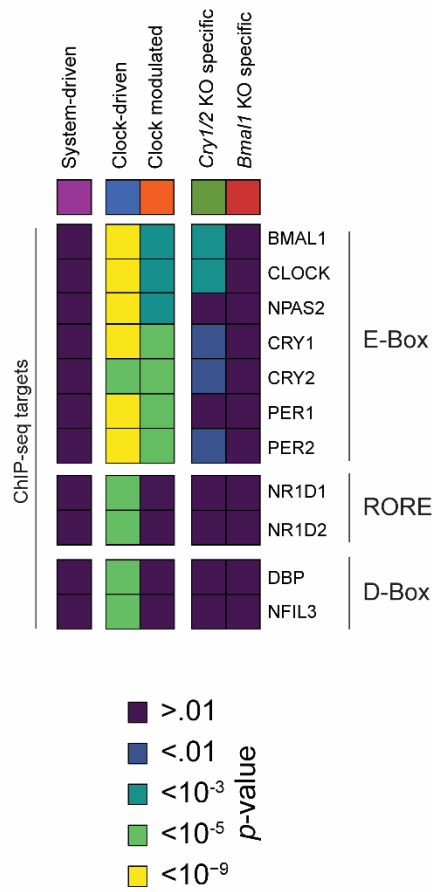
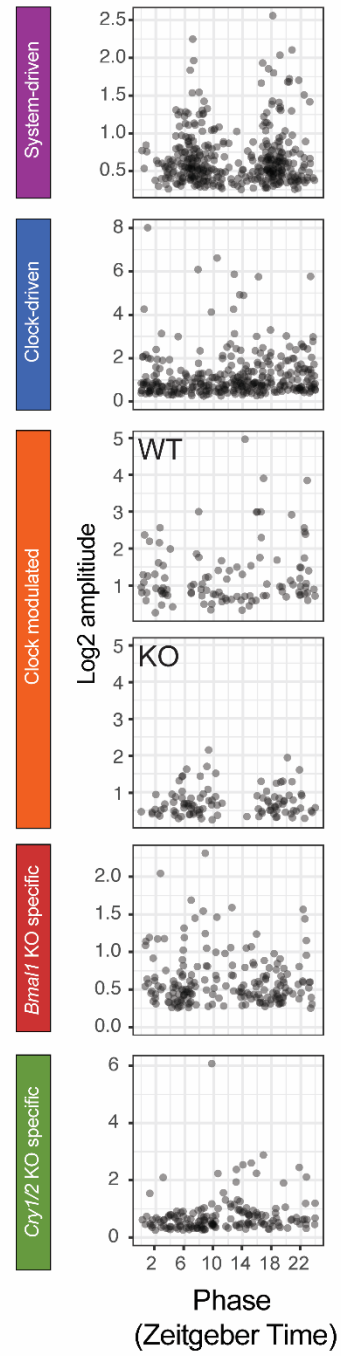
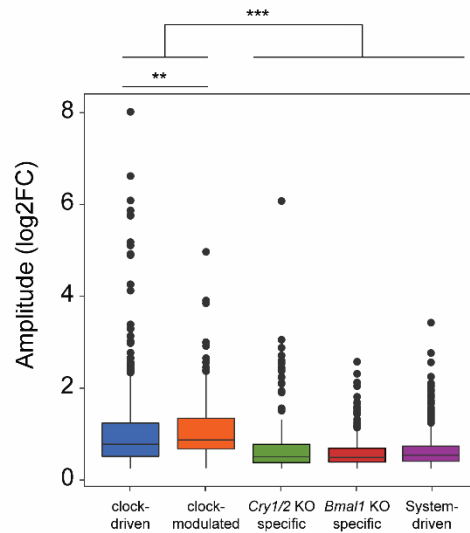
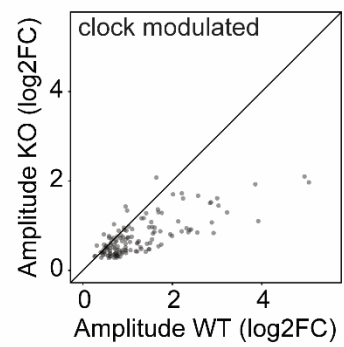
A**B****C****D**

Fig. S4. System-driven and clock-driven hepatic gene expression under conditions of night restricted feeding.

A. Enrichment analysis of core circadian clock targets for genes classified into the indicated rhythmic model. Target genes have been identified using published chromatin immunoprecipitation sequencing (ChIP-Seq) data in mouse liver.

B. Amplitude vs. phase of rhythmic genes of indicated rhythmic model.

C. Amplitude represented as boxplots for indicated rhythmic model.

D. Most genes that show a consistent change in rhythmic expression in both KO model (clock modulated) show a declined amplitude in the absence of a functional circadian clock.

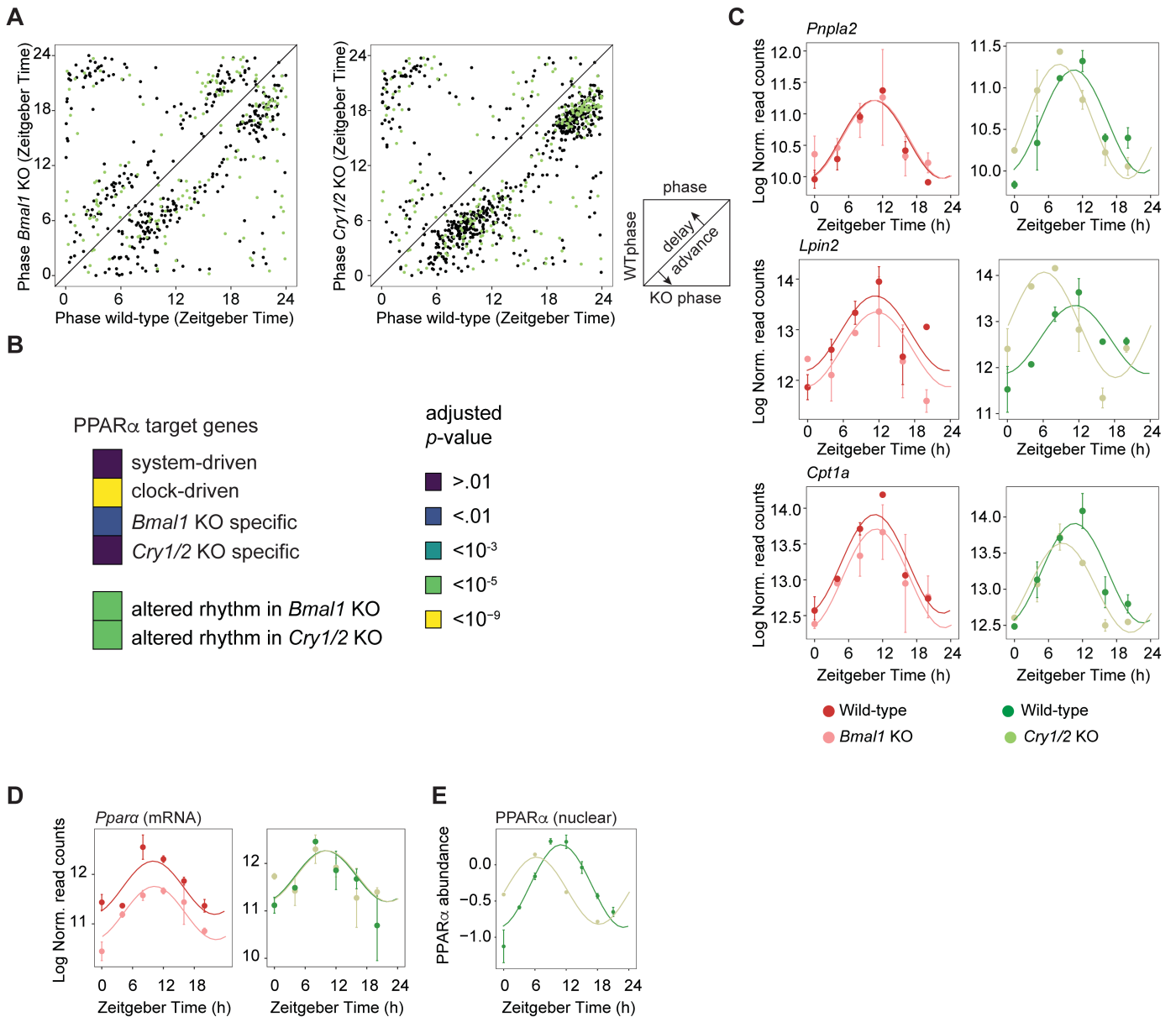


Fig. S5. *Cry1/2* KO but not *Bmal1* KO mice display a phase advance under night restricted feeding in liver gene expression

A. Phase distribution of differentially rhythmically expressed genes in *Bmal1* KO (left) and *Cry1/2* KO (right). *Bmal1* KO show a less pronounced phase advance than *Cry1/2* KO mice despite a matching NRF regimen.

B. Enrichment analysis of PPAR α target genes that display differential rhythmicity in *Bmal1* KO and *Cry1/2* KO mice. Target genes have been identified by published chromatin immunoprecipitation sequencing (ChIP-Seq) data in mouse liver.

C. Representative examples of PPAR α target genes that show a *Cry1/2* KO specific advance in phase in gene expression.

D. Hepatic *Ppara* expression exhibit lower mean levels in *Bmal1* KO animals.

E. *Cry1/2* KO mice show a specific phase advance in nuclear PPAR α protein levels (data from ref. (28)).

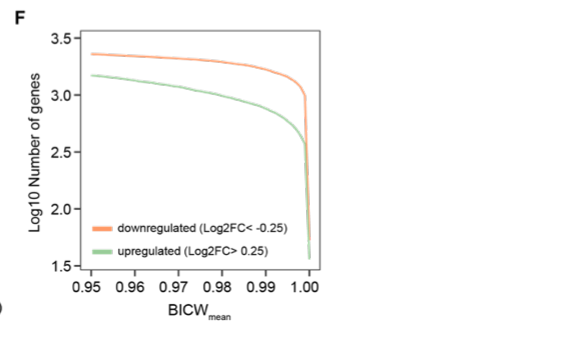
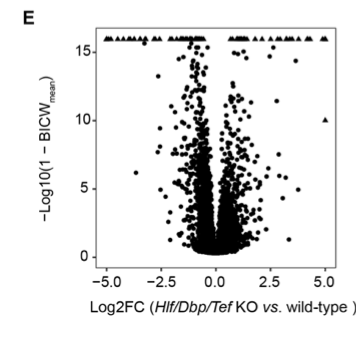
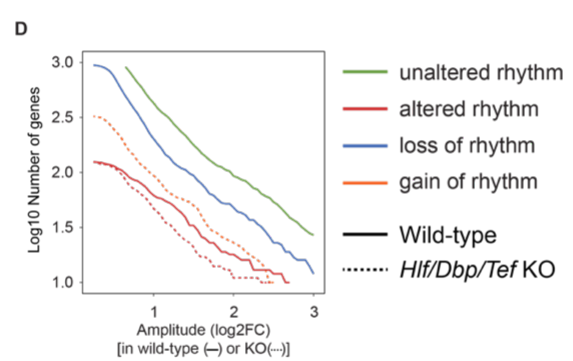
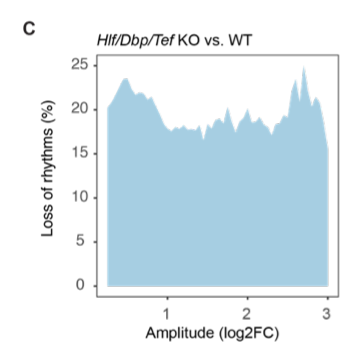
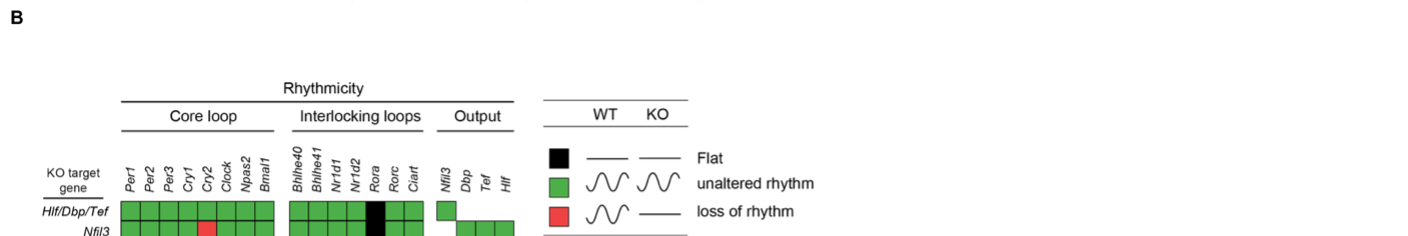
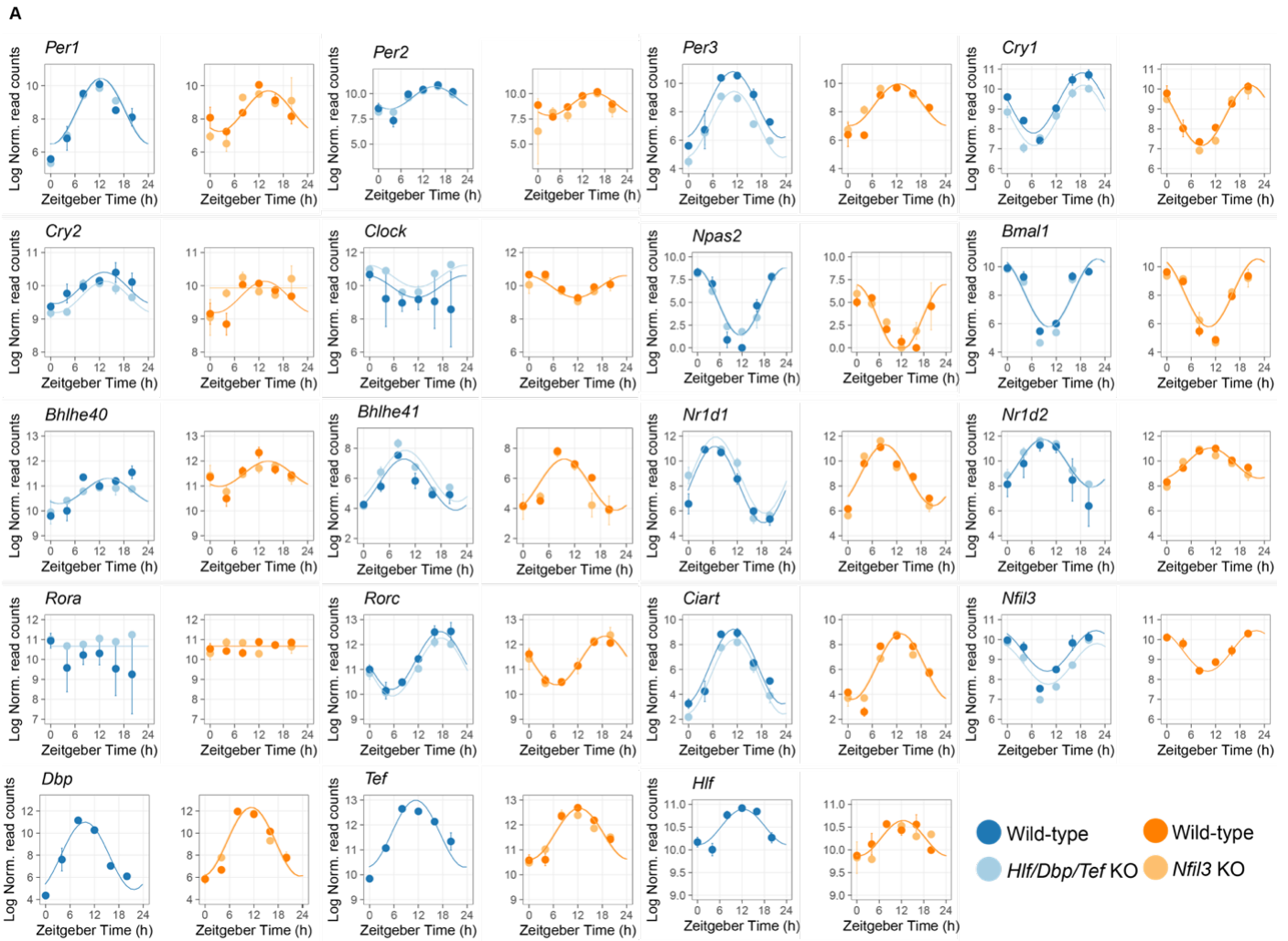


Fig. S6. Differential gene expression in the liver of *Hlf/Dbp/Tef* KO animals.

A. Clock gene expression in the liver of *Hlf/Dbp/Tef* WT, *Hlf/Dbp/Tef* KO, *Nfil3* WT, and *Nfil3* KO mice.

B. Assignment of circadian clock genes to their corresponding temporal expression pattern (right) in the indicated mouse model (row).

C. Relative number of genes that lose rhythmicity in *Hlf/Dbp/Tef* KO compared to wild-type in function of minimal amplitude.

D. Cumulative number of genes classified in indicated rhythmic model in function of minimal amplitude in wild-type or *Hlf/Dbp/Tef* KO animals.

E. Volcano plot of mean differences between *Hlf/Dbp/Tef* KO and WT mice.

F. Cumulative number of genes that are considered up or downregulated with a BICW equal or larger than the value indicated on the x-axis.

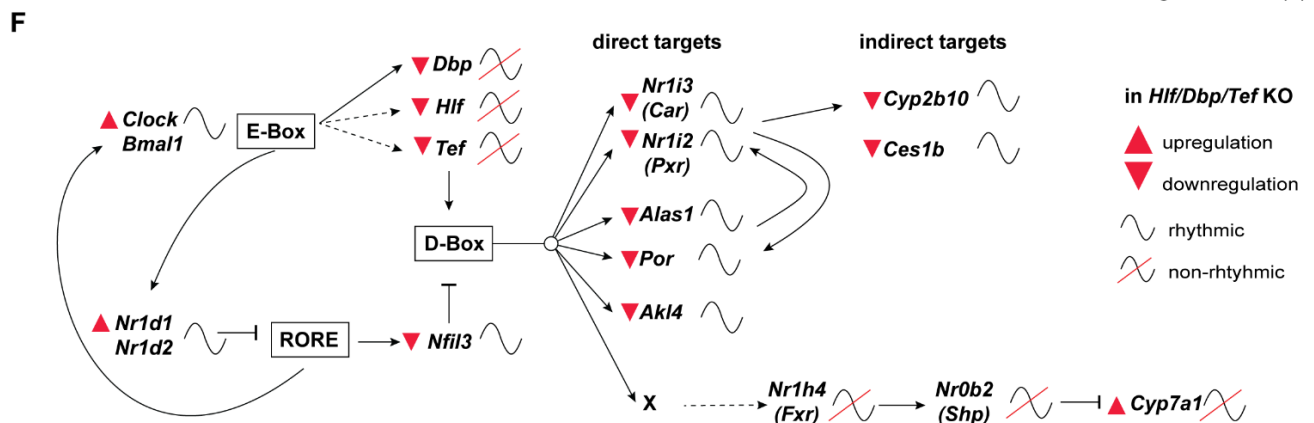
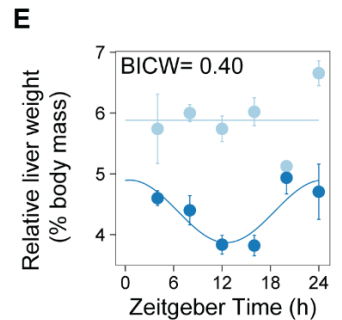
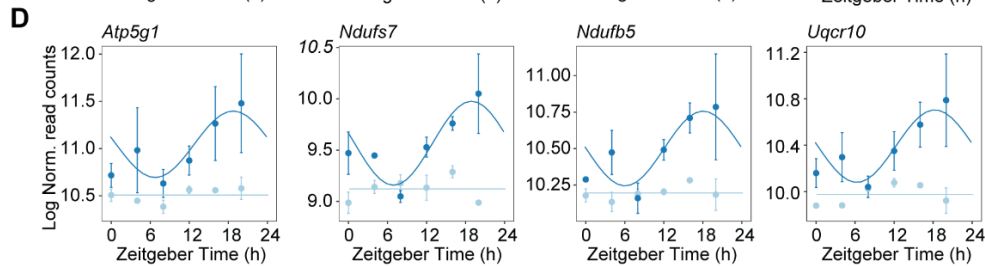
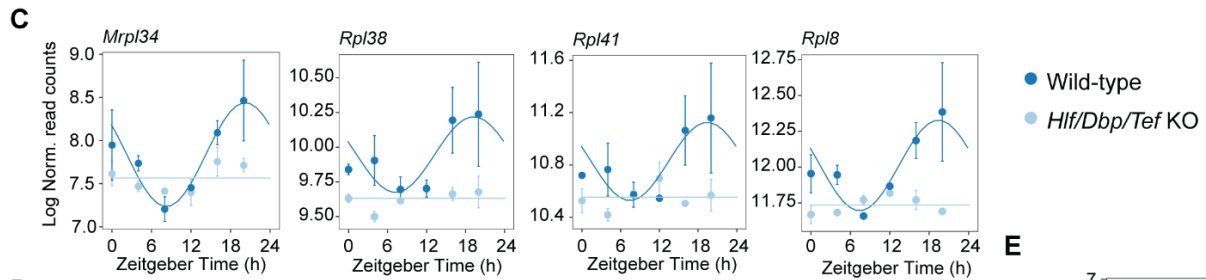
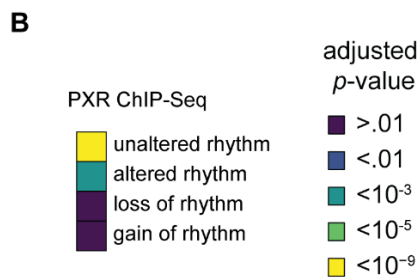
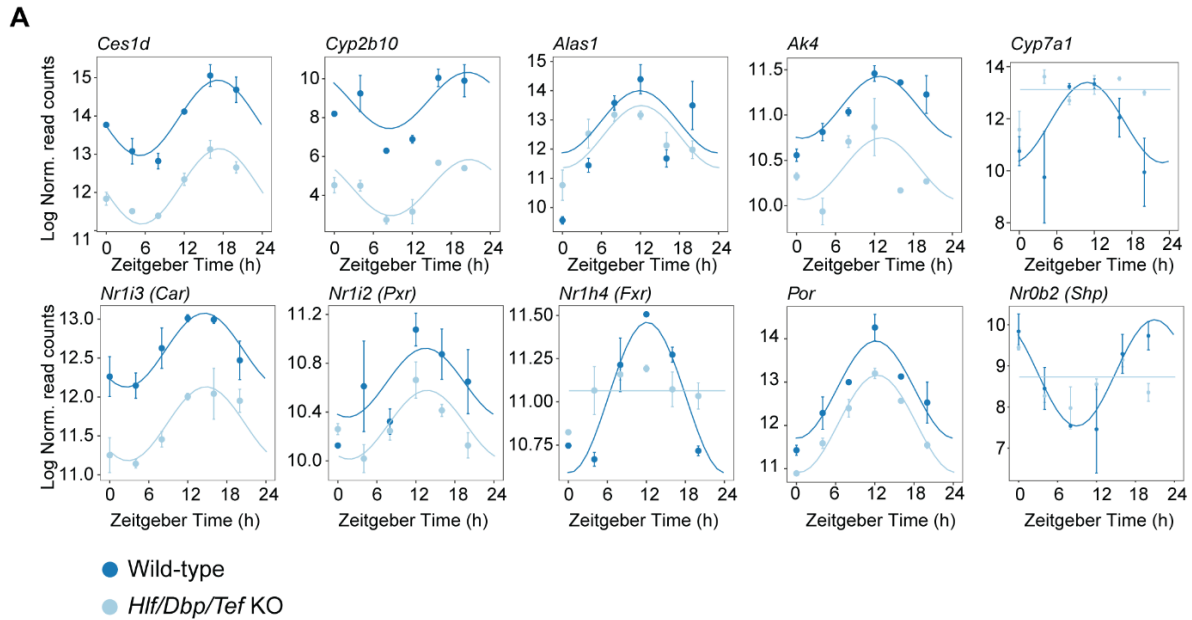


Fig. S7. Liver expression in *Hlf/Dbp/Tef* KO and WT mice.

A. Temporal profiles of genes associated with xenobiotic and bile acids metabolic network with altered expression in *Hlf/Dbp/Tef* KO mice.

B. Enrichment analysis of PXR targets for differentially rhythmically expressed genes in *Hlf/Dbp/Tef* KO and WT mice. Target genes have been identified by published chromatin immunoprecipitation sequencing (ChIP-Seq) data in mouse liver.

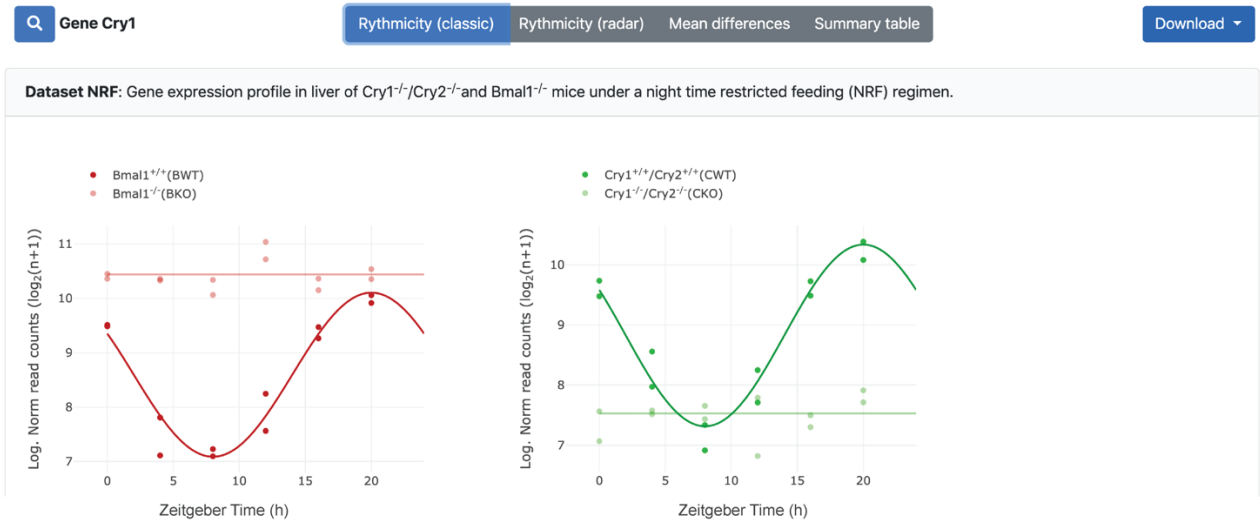
C-D. Gene expression of example genes associated with ribosomes (C) and oxidative phosphorylation (D).

E. The rhythmic fluctuation in relative liver weight (% body mass) is blunted in *Hlf/Dbp/Tef* KO mice.

F. Schematic representation of gene expression of xenobiotic metabolic gene network in the liver of *Hlf/Dbp/Tef* KO mice. Direction of triangles indicate if a gene is up (▲) or downregulated (▼) in *Hlf/Dbp/Tef* KOs.

A

B



C

Symbol	NRF phase				Model	BICW	Class	NRF mean				Model	BICW	Class	PAR bZip KO phase			
	BWT	CWT	BKO	CKO				BWT	CWT	BKO	CKO				WT	KO	Model	BICW
<input checked="" type="checkbox"/> Cry1					6	0.88	Clock driven					14	0.53	H			4	0.74

Fig. S8. The web app “ClockProfile” to assess temporal gene profiles in *Bmal1* KO, *Cry1/2* KO and PARbZip deficient mice. Screenshots from ClockProfile (<https://clockprofile.epfl.ch/>) illustrating features such as visualizing gene expression in WT and KO models, assessing statistical results from *dryR*, and browsing gene set enrichment analysis.

A. Homepage of ClockProfile with two options: search for 1. gene sets (search for gene set) or 2. individual genes (search by genes).

B. Gene expression visualization of *Cry1* as an example.

C. Summary table of *Cry1* as an example.

Dataset S1 (separate file). RNA-Seq analysis results of animal models with a genetically disrupted circadian clock under *ad libitum* or night restricted feeding regimen.

Sheet 1. Detailed table legend.

Sheet 2. Results of the rhythmicity analysis using multiple generalized linear regression and model selection in liver of WT and *Bmal1* KO series.

Sheet 3. Results of the rhythmicity analysis using multiple generalized linear regression and model selection in liver of WT and *Cry1/2* KO.

Dataset S2 (separate file). RNA-Seq analysis results of animal models with a genetically disrupted circadian clock under night restricted feeding regimen.

Sheet 1. Detailed table legend.

Sheet 2. Results of the rhythmicity analysis using multiple generalized linear regression and model selection in liver of WT, *Bmal1* KO and *Cry1/2* KO series.

Sheet 3. Functional enrichment analysis in system-driven genes.

Sheet 4. Functional enrichment analysis in clock-driven genes.

Sheet 5. Functional enrichment analysis in clock-modulated genes.

Sheet 6. Functional enrichment analysis of rhythmic genes that lose rhythmicity specifically in *Bmal1* KO.

Sheet 7. Functional enrichment analysis of rhythmic genes that lose rhythmicity specifically in *Cry1/2* KO.

Dataset S3 (separate file). RNA-Seq analysis results of *Hlf/Dbp/Tef* KO and WT under *ad libitum* feeding regimen.

Sheet 1. Detailed table legend.

Sheet 2. Results of the rhythmicity analysis using multiple generalized linear regression and model selection in liver of WT and *Hlf/Dbp/Tef* KO series.

Sheet 3. Functional enrichment analysis of genes that show an unaltered rhythmicity in *Hlf/Dbp/Tef* KO.

Sheet 4. Functional enrichment analysis of genes that lose rhythmicity in *Hlf/Dbp/Tef* KO.

Sheet 5. Functional enrichment analysis of genes that gain rhythmicity in *Hlf/Dbp/Tef* KO.

Sheet 6. Functional enrichment analysis of rhythmic genes with altered phase and/or amplitude in *Hlf/Dbp/Tef* KO.

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