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CRITICAL APPRAISAL AND ISSUES REGARDING GENERALISABILITY OF COMPARATIVE EFFECTIVENESS STUDIES OF NOACS IN ATRIAL FIBRILLATION AND THEIR RELATION TO CLINCIAL TRIAL DATA - A SYSTEMATIC REVIEW

Journal:	BMJ Open
Manuscript ID	bmjopen-2020-042024
Article Type:	Original research
Date Submitted by the Author:	24-Jun-2020
Complete List of Authors:	Bunge, Eveline; Pallas health research and consultancy van Hout, Ben; University of Sheffield, HEDS, ScHARR Haas, Sylvia; Formerly Technical University Of Munich Spentzouris, Georgios; Daiichi Sankyo Europe GmbH Cohen, Alexander; Guys and St Thomas' NHS Foundation Trust, King's College
Keywords:	Adult cardiology < CARDIOLOGY, STATISTICS & RESEARCH METHODS, Cardiology < INTERNAL MEDICINE
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CRITICAL APPRAISAL AND ISSUES REGARDING GENERALISABILITY OF COMPARATIVE

EFFECTIVENESS STUDIES OF NOACS IN ATRIAL FIBRILLATION AND THEIR RELATION TO CLINCIAL

TRIAL DATA - A SYSTEMATIC REVIEW

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Word count: 4196

KEY WORDS

atrial fibrillation; non-vitamin K antagonist oral anticoagulants (NOAC); comparative effectiveness research; systematic review

ABSTRACT

Objective: To critically appraise the published comparative effectiveness studies on non-vitamin K antagonist oral anticoagulants (NOAC) in nonvalvular atrial fibrillation (NVAF).

Materials and Methods: We performed a systematic literature review in Medline and EMbase to investigate the way comparisons were made. Results were also compared with expectations formulated on the basis of trial results with specific attention to the patient years in each study.

Results: We included 39 studies in which direct comparison between at least two NOACs were made. Almost all studies concerned patient registries, pharmacy or prescription databases and/or health insurance database studies using a cohort design. Corrections for differences in patient characteristics was applied in all but two studies. Eighteen studies matched using propensity scores, eight studies weighted patients based on the inverse probability of treatment, one study used propensity score stratification and ten studies applied a proportional hazards model. These studies have some important limitations, even though the larger part of the studies were well conducted technically. On the basis of trial results, expected differences are small and a naïve analysis suggests trials with between 7,700 and 59,500 patients are needed to confirm the observed differences in bleedings and between 51,800 and 7,994,300 to confirm differences in efficacy.

Conclusion: Meaningful comparisons between NOACs on the basis of observational data, even after correction for baseline characteristics, may not be reliable due to unmeasured confounders, channelling bias and insufficient sample size. These limitations should be kept in mind when results of these studies are used to decide on NOAC treatment options.

STRENGHTS AND LIMITATIONS OF THIS STUDY

- To our knowledge, this is the first systematic review that critically appraised the quality and generalisability of the comparative effectiveness studies on NOACs in atrial fibrillation patients and to relate this to clinical trial data
- A naïve trial analysis was conducted to estimate the number of patients needed in a randomised clinical



INTRODUCTION

Guidelines state a preference for non-vitamin K antagonist oral anticoagulants (NOACs) above vitamin K antagonists (VKA) in patients with nonvalvular atrial fibrillation (NVAF) requiring prevention of stroke and systemic embolism.[1, 2] However, no recommendation for a specific NOAC is made in these guidelines, and in daily practice, physicians have to make a choice which of the four available NOACs (dabigatran, rivaroxaban, apixaban, edoxaban) they prescribe for a particular patient.[3-6]

In the absence of head-to-head trials, comparative effectiveness research (CER) has been conducted to compare the NOACs with regard to effectiveness and safety. This is also described as real-world evidence; i.e. the data will come from patients treated in daily practice. Comparisons on effectiveness and safety between NOACs are however not easy to make, as patients will not be prescribed one of the NOACs at random. The choice of a certain NOAC for a patient will at least partly be driven by patient characteristics, such as age, concomitant medications, and the risk of stroke and/or bleeding. This can lead to systematic differences between the treatment groups, which is known as channeling bias. [7] In order to make a valid comparison on effectiveness and safety between the NOACs, adjusting for these characteristics is necessary when these characteristics are also related to the outcome (confounding variables).

Several techniques exist to correct for imbalances in risks. Cox proportional hazards (Cox PH) regression model adjustment can be used but large sample sizes are needed when number of events is relatively low and the number of covariates is high (as a rule of thumb, about 10 events per predictor variable [8]) and these large sample sizes are not always available. Event rates are low, around 1 per 100 patient years for efficacy outcomes and to detect differences, even in a randomized clinical trial, one needs substantial numbers of patients. This number would only increase when the results are contaminated by a lack of balance between the patients groups. Another method to adjust for confounding is using propensity scores (PS) to create comparable patient groups before the analysis. A propensity score is the probability of an individual receiving a specific treatment given a specific set of patient characteristics (e.g. age, gender, comorbidities).[9] Variables related to the outcome should be included in the propensity score despite their strength of association on treatment (exposure) selection. This will increase the precision of the estimated exposure effect, while bias will not be increased. Variables that are related to the exposure but not the outcome will decrease the precision of the estimated exposure effect without decreasing bias.[10] Adjustment for confounding using PS can be done by matching the treatment groups on the PS, by weighing treatment groups based on the PS inverse probability of treatment weighting (IPTW), by PS stratification, or by covariate adjustment using the PS.[9, 11] Well conducted PS methods will lead to treatment groups that are very well comparable regarding important confounders, which increases the confidence in the results, however, there are also some disadvantages. For instance, in PS matching studies, patients who cannot

be matched to another patient will be excluded from the analyses, and in IPTW, when patients on one treatment have a low propensity score and patients treated with the other treatment have a high propensity score, extreme weights can occur which can bias the results.[12]

To gain more understanding in how the above described methodologies were applied in peer-reviewed CER on effectiveness and safety in NOACs in NVAF patients, we conducted a systematic literature review. Within this we compare the results with those from a naïve analysis of the results of the four major trial for rivaroxaban (Rivaroxaban Once Daily Oral Direct Factor Xa Inhibition Compared with Vitamin K Antagonism for Prevention of Stroke and Embolism Trial in Atrial Fibrillation (ROCKET AF)), apixaban (Apixaban for Reduction in Stroke and Other Thromboembolic Events in Atrial Fibrillation (ARISTOTLE)), dabigatran (Randomized Evaluation of Long-Term Anticoagulation Therapy (RE-LY)) and edoxaban (Effective Anticoagulation with Factor Xa Next Generation in Atrial Fibrillation—Thrombolysis in Myocardial Infarction 48 (ENGAGE AF)) and compare the results from the various analyses with those from the trials.

METHODS

Information sources, search strategy and eligibility criteria

We performed a systematic literature review to identify peer-reviewed comparative effectiveness research on NOACs in patients with atrial fibrillation. A search in Medline (access through PubMed) and EMbase was performed combining search strings on NOAC, VKA and atrial fibrillation (see appendix 1 for the search strings). The search was conducted on 23-04-2019 and we checked all articles published in English language. The title and abstract selection was done in duplicate by two independent researchers.

The following inclusion criteria were used:

- Population: patients with NVAF
- Intervention: NOAC (dabigatran, rivaroxaban, apixaban or edoxaban)
- Comparator: other NOAC(s) (dabigatran, rivaroxaban, apixaban and/or edoxaban)
- Outcomes: effectiveness and safety
- Study type: comparative effectiveness studies with a cohort design

The following exclusion criteria were applied:

- Studies on only one NOAC
- Studies in which VKA is the comparator for the NOACs, and NOACs are not compared against each other
- · Studies on cost-effectiveness and healthcare resources use
- Studies on adherence or persistence

Critical appraisal

We checked the setting, in- and exclusion criteria and the following baseline characteristics: age, proportion males, CHA2DS2Vasc score and comorbidity index.

We used the criteria suggested by ISPOR, Yao et al., and Austin et al. as a guidance to critically appraise the articles in which PS were used.[12-15] The criteria we checked concerned:

- The variables included in the propensity score model
- Explanation of the variable selection procedure for propensity score model
- Distribution of baseline characteristics for each group before propensity score analysis
- In case of PSM:
 - o matching ratio,
 - distance metric,

- with or without replacement,
- comparability of baseline characteristics in the matched groups,
- sample size before and after matching
- In case of IPTW:
 - comparability of baseline characteristics in the weighted groups
 - o extreme weights
- In case of PS stratification:
 - o number of strata, comparability of baseline characteristics

In case of analyses in which no PS was used in the main analyses, we evaluated whether the ratio number of covariates to the number of events seemed sufficient to produce valid results.[8]

Naïve trial analysis

Trials are quite often designed with a null hypothesis and associated with a power calculation while real world studies are often dictated by the number of observations available. To give the results from the real-world-evidence some perspective we undertook a naïve trial analysis in which the risk reductions from each trial with respect to efficacy and safety outcomes were applied to an average number of outcomes observed in the warfarin arms in each trial. This leads to an estimate of the relevant rates for each drug and the differences are illustrated by the number of patients (sample size) needed in a randomised clinical trial to confirm the estimated differences.

RESULTS

In total, we found 1302 unique articles in our search, of which 39 articles fulfilled the in- and exclusion criteria and were included for data extraction, see figure 1. In table 1 to 5, study characteristics are presented. The most important differences between the studies are outlined in table 6.



Table 1. Characteristics of the included articles that used propensity score matching (PSM) as primary analyses (n=18)

Author and country	Setting and study period	Study population	Patient characteristics before PSM (range between NOACs)	Primary outcome definition	PSM details	Sample size before matching	Sample size after matching	Results/conclusion as reported in the article
Abraham, 2017 USA	OptumLabs Data Warehouse Oct 1, 2010 through Feb 28, 2015	NVAF patients, 18 years of age or older, identified by their index prescription of a NOAC during study period (excluded if NOAC prescribed during 12 months before index date). No reporting on earlier VKA use	Male: 54.0-60.5% CHA2DS2-VASc: 3.2-4.0	Gastrointestinal bleeding: definition by Lewis et al. 2002 using inpatient hospital claims for relevant primary and secondary discharge diagnoses.	3 matched cohorts; 1:1 PSM without replacement and with a calliper of 0.01; After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=19,301 dabigatran: n=17,426 apixaban: n=6,576	rivaroxaban vs. dabigatran: n=31,574 apixaban vs. rivaroxaban: n=13,130 apixaban vs. dabigatran: n=13,084 (more than 90% of original smallest samples size)	Apixaban had the most favorable gastrointestinal bleeding profile and rivaroxaban had the least favorable safety profile. Apixaban had the most favorable gastrointestinal safety profile among all ag groups.
Amin, 2018 (J Manag Care Spec Pharm) USA	Medicare & Medicaid Services Jan 1, 2012, to Dec 31, 2014	NVAF patients of at least 65 years old, OAC treatment- naïve, ≥ 1 prescription claim for OAC during study period. Excluded if OAC pharmacy claim during the 12-month before study start.	Age: 77.2±7.0-78.4±7.4 Male: 47.4-50.6% CHA2DS2-VASc: 4.4-4.6 CCI: 2.5-2.7	Hospitalization for stroke, systemic embolism and major bleeding: ICD-9- code as primary discharge diagnosis	2 matched cohorts; 1:1 PSM with nearest neighbour without replacement and with a calliper of 0.01; After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=53,146 apixaban: n=20,853 dabigatran: n=16,743	rivaroxaban vs. apixaban: n=41,608 dabigatran vs. apixaban: n=30,836 (more than 90% of original smallest samples size)	Apixaban was associated with significantly lower risks of all-cause, stroke/SE-related, and MB-related hospitalizations compared with dabigatran, and rivaroxaban
Amin 2018 (J Med Econ) USA	OptumInsight research database Jan 1, 2012 – Sept 30, 2015	NVAF patients of at least 18 years old, OAC treatment- naïve, ≥ 1 prescription claim for OAC during study period. Excluded if OAC pharmacy claim during the 12-month before study start.	NR	Hospitalization for stroke, systemic embolism and major bleeding: ICD-9- code as primary discharge diagnosis	2 matched cohorts; 1:1 PSM with nearest neighbour without replacement and with a calliper of 0.01; After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=14,163 apixaban: n=8,652 dabigatran: n=3,684	apixaban vs. rivaroxaban: N=16,880 apixaban vs. dabigatran: N=7,114 (more than 90% of original smallest samples size)	Rivaroxaban patients were associated with a significantly higher risk of all-cause and major bleeding related hospitalisations and dabigatran patients were associated with a significantly higher risk of major bleeding hospitalisation compared with apixaban
Andersen, 2018 Denmark	National patient register, Register of Medicinal Product Statistics July 1, 2013 – March 31, 2016	NVAF patients who were new users of NOAC aged 45 years of age or older, with a recent diagnosis of NVAF (received no OAC treatment in the 12 months before inclusion; 'recent diagnosis' is not defined)	Online material not available	Stroke, systemic embolism and major bleeding (i,e, intracranial bleeding, gastro-intestinal bleeding (bleeding ulcer, hematemesis ormelena) or other serious bleeding (anemia caused by bleeding, bleeding of unknown origin, bleeding of the respiratory or urinary tract, peritoneal, retinal or orbital bleeding); hospital admission with a primary or secondary	3 matched cohorts; 1:1 PSM with nearest neighbour with a caliper of 0. (replacement yes or no not reported) All baseline characteristics were well balanced after matching, except for calendar year	apixaban: n=4,292 dabigatran: n=3,913 rivaroxaban: n=3,805	apixaban vs. rivaroxaban: N=7,352 apixaban vs. dabigatran: N=6,470 rivaroxaban vs. dabigatran: N=5,440	There were no statistically significant differences in risk of stroke or systemic embolism or major bleeding in propensity-matched comparisons between apixaban, dabigatran, and rivaroxaban used in standard doses.

Author and country	Setting and study period	Study population	Patient characteristics before PSM (range between NOACs)	Primary outcome definition	PSM details	Sample size before matching	Sample size after matching	Results/conclusion as reported in the article
Blin, 2019 France	French nationwide claims and hospitalization database, Système National des Données de Santé	NVAF patients of at least 18 years old, all new users of standard or reduced doses of NOAC in (received no OAC treatment in the three years before the index date)	Age: 65.3±10.2-69.0±11.1 Male: 62.7-68.3% Modified CHA2DS2-VASc ≥2: 57.1-67.4% Comorbidities: NR	Hospitalization with a main diagnosis of ischemic stroke or systemic embolism or major bleeding and all-cause death. (ICD-10) codes	1 matched cohort PSM method not reported. After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=18,829 dabigatran: n=10,847	dabigatran vs. rivaroxaban: n=16,580	Dabigatran had similar or better effectiveness than rivaroxaban but lower bleeding risk. Death rates were not different.
Briasoulis, 2018 USA	Medicare and Medicaid Services Jan 1, 2010 - Dec 31, 2013	NVAF patients newly diagnosed of ≥65 years old and initiated OAC treatment during study period	Age: 75.4±6–75.5±6 Male: 50-53% CHA2DS2-VASc: 4.1-4.1 Gagne: 2.7-2.7	All-cause mortality, stroke, including ischemic stroke or transient ischemic attack, gastrointestinal bleeding, any bleeding, non-gastrointestinal bleeding, acute myocardial infarction. ICD-9-CM reported in inpatient claims, whether primary and secondary codes were used is not described	1 matched cohort; 3-way propensity matching (VKA was one of the groups, but not further discussed here) After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=14,257 dabigatran: n=13,522	dabigatran vs. rivaroxaban: n =26,814	Rivaroxaban was associated with higher gastrointestinal bleeding rates than dabigatran
Deitelzweig, 2017 USA	Humana Research Database (Medicare coverage) Jan 2013 - 30 Sept 2015	NVAF patients age of ≥65 years, OAC treatment naïve (excluded if they had a pharmacy claim for OAC during the baseline period, which was 12 months before index date)	Age: 76.8±8.3-78.0±9.0 Male: 51.5-55.1% CHA2DS2-VASc: 4.3-4.6 CCI: 2.7-3.0	Hospitalisation claims of stroke, systemic embolism and major bleeding: ICD-9-code as primary discharge diagnosis	2 matched cohorts; 1:1 PSM with nearest neighbour (replacement yes or no and calliper not reported) balanced with key patient characteristics not statistically different (p>.05).	rivaroxaban: n=11,082 apixaban: n=8,250 dabigatran: n=2,474	apixaban vs. rivaroxaban: n=13,620 apixaban vs. dabigatran: n= 4,654	Apixaban is associated with significantly lower risk of stroke/systemic embolism and major bleeding than rivaroxaban, and a trend towards better outcomes vs. dabigatran.
Gupta, 2018 USA	Department of Defence data Jan 1, 2012, to Sept 30, 2015	NVAF patients, treatment-naïve (excluded if a pharmacy claim for an OAC during the baseline period)	NR	Inpatient claim of stroke, systemic embolism or major bleeding as primary or secondary diagnosis based on validated administrative claims-based algorithms	2 matched cohorts; 1:1 PSM with nearest neighbour without replacement with a calliper of 0.01 After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=15,680 apixaban: n=11,754 dabigatran: n=4,312	rivaroxaban vs. apixaban: n=22,568 dabigatran vs. apixaban: n=8,258	Rivaroxaban was associated with a significantly higher risk of stroke/systemic embolism and major bleeding compared with apixaban. Dabigatran use was associated with a numerically higher risk of stroke/systemic embolism and a significantly higher risk of major bleeding compared with apixaban

Author and country	Setting and study period	Study population	Patient characteristics before PSM (range between NOACs)	Primary outcome definition	PSM details	Sample size before matching	Sample size after matching	Results/conclusion as reported in the article
Lai, 2017 Taiwan	National Health Insurance program 2011 to 2014	NVAF and flutter patients, ≥20 years, new-users (new users not further defined).	Age: 75.1±9.7-75.4±9.6 Male: 54-7-56.7% CHA2DS2-VASc: 3.3-3.3 Comorbidity index: NR	All-cause death	1 matched cohort; 1:1 PSM with calliper < 0.2 (neighbour and replacement not reported) Balance checked with p-values and standardized difference	dabigatran: n=10,625; rivaroxaban: n=4,609	dabigatran vs. rivaroxaban: N=9,200	Rivaroxaban therapy was associated with a statistically significant increase in all-cause death compared with dabigatran
Lin, 2017 USA	IMS Pharmetrics Plus database Jan 2013 – Sept 2015	NVAF patients of at least 18 years old who initiated OAC (received no OAC treatment received 12 months before the index date)	NR	Major bleeding first listed in ICD-9 diagnosis or procedure codes	2 cohorts; 1:1 PSM with nearest neighbour (replacement and calliper not reported) Patient key characteristic being similar with p>0.05	NR	apixaban vs. rivaroxaban: N=8,124 apixaban vs. dabigatran: N=5,368	Apixaban is associated with reduced risk of hospitalisation compared with dabigatran and rivaroxaban.
Lip, 2016 (Thromb Haemost) USA	Truven MarketScan® Commercial Claims and Encounter and Medicare Supplemental and Coordination of Benefits Databases Jan 2012 to Dec 2014	NVAF patients ≥18 years who newly initiated OACs (patients with a prescription claim for OAC prior to the index date were excluded)	Age: 66.5±12.4- 68.5±12.4 Male: 61.4-65.0% CHA2DS2-VASc: 2.6-2.8 CDI: 1.6-1.8	Major bleeding listed first primary ICD-9 code	3 cohorts; 1: 1 PSM with nearest neighbour without replacement with a maximum calliper of 0.01. After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=17,801 apixaban: n=7,438 dabigatran: n=4,661	apixaban vs dabigatran: n=14,798 rivaroxaban vs dabigatran: n=9,314 apixaban vs rivaroxaban: n=8,814	Compared to apixaban, rivaroxaban initiation was associated with significantly higher risk of major bleeding. The difference for dabigatran was not statistically significant
Lip, 2018 USA	Medicare and Medicaid Services Medicare; Truven MarketScan, IMS PharMetrics Plus Database, Optum Clinformatics Data Mart, and the Humana Research Database Jan 1, 2013, to Sept 30, 2015	NVAF patients newly prescribed OAC, (received no OAC treatment in the 12 months before the index date)	Age: 71.4±11.4-73.1±11.6 Male: 55.0-59.6% CHA2DS2-VASc: 3.3-3.6 CDI: 2.4-2.8	Hospitalizations with stroke, systemic embolism or major bleeding as the principal or first-listed diagnosis	3 cohorts 1:1 PSM with nearest neighbour without replacement with a maximum calliper of 0.01 After PSM, standardized differences of all baseline characteristics were <10%	rivaroxaban: n= 103,477 apixaban: n= 63,484 dabigatran: n= 27,571	apixaban-rivaroxaban: n=125,238 dabigatran-rivaroxaban: n=55,076 apixaban-dabigatran: n=54,192	Apixaban was associated with a lower rate of stroke/systemic embolism and major bleeding compared with dabigatran and rivaroxaban. Dabigatran was associated with a lower rate of major bleeding compared with rivaroxaban, with similar rates of stroke/systemic embolism.

Author and country	Setting and study period	Study population	Patient characteristics before PSM (range between NOACs)	Primary outcome definition	PSM details	Sample size before matching	Sample size after matching	Results/conclusion as reported in the article
Lutsey, 2018 USA	MarketScan Commercial Database Jan 1, 2010 through Sept 30, 2015	NVAF patients aged 45 and older with at least one prescription for OAC after their first AF claim (de novo patients or first initiation of treatment)	Age: 69.1±11.4-69.9 ± 11.7 Male: 59.4-63.7 CHA2DS2-VASC: 3.3-3.6 Comorbidity index: NR	venous thromboembolism: at least one inpatient ICD 9 claim (first listed or not is not specified)	3 cohorts 1:1 PSM with a maximum calliper of 0.25 (neighbour and replacement not reported) Balance not described	rivaroxaban: n=31,119 dabigatran: n=28,089 apixaban: n=17,112	rivaroxaban vs. apixaban: n=32,468 dabigatran vs. rivaroxaban: n=21,160 dabigatran vs. apixaban: n=6,200	Risk of VTE was lowest among those prescribed apixaban and dabigatran
Mentias, 2018 USA	Medicare & Medicaid Services Jan 1, 2010, to Dec 31, 2013	NVAF patients, newly diagnosed who initiated an OAC within 90 days of diagnosis	Age: 75.8±6.4-75.8±6.4 Male: 48.9-50.1% CHA2DS2-VASc: 4.3-4.3 Gagne: 3.0-3.0	inpatient admission for acute ischemic stroke or major bleeding as defined by Rothendler* and Suh based on the primary ICD-9-CM diagnosis on inpatient standard analytical files claims for acute care stays.	1 cohort 3-way PSM. (VKA was one of the groups, but not further discussed here) After PSM, standardized differences of all baseline characteristics were <10%	rivaroxaban: n=23,177 dabigatran: n=21,979	NR	Rivaroxaban users had significantly higher major bleeding risk compared with dabigatran users in the medium and high comorbidity groups
Norby, 2017 USA	Truven Health MarketScan Commercial Claims and Encounters Database and the Medicare Supplemental and Coordination of Benefits Database	NVAF patients with at least one prescription of NOAC after their first AF claim (first prescription of OAC)		ischemic stroke (primary discharge), intracranial bleeding (primary discharge), myocardial infarction (1st or 2nd position of an inpatient discharge diagnosis, and gastrointestinal bleeding (primary and secondary diagnoses, presence of transfusion codes, and presence/absence of trauma codes to exclude trauma-related bleeding based on ICD-9 codes	1 cohort; 1:1 PSM, greedy matching technique with a calliper of 0.25	NR NR	rivaroxaban vs dabigatran: n=16,957	Endpoint rates were similar when comparing anticoagulant-naïve rivaroxaban and dabigatran initiators, with the exception of higher gastrointestinal bleeding risk in rivaroxaban users
Noseworthy, 2016 USA	Optum Labs Data Warehouse Oct 1, 2010 - Feb 28, 2015	NVAF patients ≥ 18 years, who were OAC users during study period.	NR	inpatient admission for stroke or systemic embolism or major bleeding (ICD-9 codes in the primary or secondary diagnosis positions of inpatient claims)	3 cohorts; 1:1 PSM without replacement and with a caliper of 0.01. A standardized difference < 10% was considered acceptable	NR	rivaroxaban vs. dabigatran: n=31,574 apixaban vs. rivaroxaban: n=13,130 apixaban vs. dabigatran: n=13,084	Dabigatran, rivaroxaban, and apixaban appear to have similar effectiveness, although apixaban may be associated with a lower bleeding risk and rivaroxaban may be associated with an elevated bleeding risk
Shantha, 2017 USA	Medicare and Medicaid Nov 1, 2011 - Dec 31, 2013	Newly diagnosed NVAF patients and initiated OAC use.	Males: Age: 74.7±5.9-74.9±6. CHADS2-Vasc: 3.7-3.8 Gagne score: 2.9-2.9 Women: Age: 76.6±6.6-76.9±6.6 CHADS2-Vasc: 4.8-4.9	inpatient admissions for acute ischemic stroke or major bleeding (primary ICD-9-CM diagnosis on inpatient standard analytical files claims for acute care stays)	1 cohort; Three-way PSM (VKA was one of the groups, but not further discussed here)	rivaroxaban: n=23,177 dabigatran: n=21,979	dabigatran vs. rivaroxaban: n= 37,298	The reduced risk of ischemic stroke in patients taking rivaroxaban, compared with dabigatran, seems to be limited to men, whereas the higher risk of

Author and country	Setting and study period	Study population	Patient characteristics before PSM (range between NOACs)	Primary outcome definition	PSM details	Sample size before matching	Sample size after matching	Results/conclusion as reported in the article
			Gagne: 3.0-3.1		A standardized difference < 10% was considered acceptable			bleeding seems to be limited to women
Villines, 2019 USA	US Department of Defence Military Health System database 1 July 2010 to 30 June 2016 for the dabigatran vs. rivaroxaban cohort, and 28 Dec 2011 to 30 June 2016 for the dabigatran vs. apixaban cohort	NVAF patients ≥18 years newly initiated on standard-dose NOAC (first initiation of treatment, AF diagnosis in the 12 months before the index date or on the index date)	Age (mean): 70.9-71.3 Male: 60-62% CHA2DS2-VASc: 3.1-3.1 CCI score: 4.3-4.3	Stroke or major bleeding, ICD-9 or 10 codes, whether primary and secondary codes were used is not described	2 cohorts 1:1 PSM nearest neighbour with a calliper of 0.20 (replacement not reported). Balanced if the absolute value of the STD was ≤10%.	NR	dabigatran vs. rivaroxaban: n=25,526 dabigatran vs. apixaban: n=9,604	Dabigatran was associated with significantly lower major bleeding risk vs. rivaroxaban, and no significant difference in stroke risk. For dabigatran vs. apixaban, the reduced sample size limited the ability to draw definitive conclusions.

Age: mean, SD unless stated otherwise; CCI: Charlson comorbidity index; CDI: Charlson-Deyo index; Gagne: Gagne comorbidity score;

Table 2. Characteristics of the included articles that used inverse probability of treatment weighting as primary analyses (n=8)

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	IPWT details	Sample size	Result/conclusion as reported in the article
Adeboyeje, 2017 USA	HealthCore integrated research environment Nov 1, 2009 - Jan 31, 2016	NVAF patients newly prescribed OAC (no prescriptions for any anticoagulant in the 6-month period preceding their index dates).	Age (mean): 66-69 Male: 59.1-65.5% CHA2DS2-VASc: 2.7-3.2 Comorbidity index: NR	Hospitalization for major bleeding (ICD 9-CM codes; whether primary and secondary codes were used is not described	Extreme weights: not reported. Balanced if the absolute value of the STD was ≤10%.	dabigatran: n=8,539 rivaroxaban: n=8,398 apixaban: n=3,689	Apixaban and dabigatran were associated with lower major bleeding risk compared with rivaroxaban; however, apixaban had a lower risk of major gastrointestinal bleeding than dabigatran.
Chan, 2018 Taiwan	Taiwan National Health Insurance Research June 1, 2012 - Dec 31, 2016	NVAF patients with their first prescription of OAC	Age: 75±10-76±10 Male: 55-60% CHA2DS2-VASc: 3.7-3.9 Comorbidity index: NR	Hospitalization for ischemic stroke/systemic embolism, intracranial hemorrhage, major gastrointestinal bleeding, acute myocardial infarction, all major bleeding events, and all-cause mortality. ICD 9 and 10 codes, whether primary and secondary codes were used is not described	Extreme weights: not reported. Balanced if the absolute value of the STD was ≤10%.	rivaroxaban: n=27,777 dabigatran: n=20,079 apixaban: n=5,843	Three low-dose NOACs showed similar performance as without subgrouping
Charlton, 2018 USA	HealthCore Integrated Research Evironment database Nov 1, 2010 - March 31, 2014	NVAF patients hospitalized for bleeding after starting OAC (AF diagnosis 6 months before starting one of the index drugs).	Age: 68.0±12.5-69.6±12.6 Male: 61.8-62.9 CHA2DS2-VASc: 3.8 -3.8 CDI: 2.0-2.3	Total length of hospital stay, proportion of patients admitted to the ICU, mean length of ICU stay, and all-cause 30- and 90-day mortality, ICD 9 codes, whether primary and secondary codes were used is not described.	Extreme weights: not reported. Balance was tested using ANOVAs for significant differences	dabigatran: n=442 rivaroxaban n=256	There were no significant differences in relative risk of all-cause 30- or 90-day
Graham, 2016 USA	Medicare Nov 4, 2011 - June 30, 2014	NVAF patients, at least 65 years old, initiating OAC at standard doses (first treatment, received no NOAC treatment for other indications in the last 6 months before the index date)	Age: 65-74 y: 50-51% Age: 75-84: 40-40% Age ≥85: 9-10% Male: 53-53% CHADS2 ≥2: 66-67% Comorbidity index: NR	Thromboembolic stroke, ICH, major extracranial bleeding events and mortality (as the first study outcome or within 30 days after hospitalization for another primary outcome event), ICD 9 codes, whether primary and secondary codes were used is not described.	Extreme weights: not reported Balanced if the absolute value of the STD was ≤10%.	rivaroxaban: n=66,651 dabigatran: n=52,240 Weighted cohorts rivaroxaban: n=66,630 dabigatran: n=52,264	Treatment with rivaroxaban was associated with statistically significant increases in intracranial bleeding and major extracranial bleeding, including major gastrointestinal bleeding, compared with dabigatran

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	IPWT details	Sample size	Result/conclusion as reported in the article
Graham, 2019 USA	fee-for-service Medicare Part A (hospitalization), Part B (office-based care), and Part D (prescription drug coverage) Oct, 2010 - Sept, 2015	NVAF patients of ≥65 years old (first initiation of treatment)	Age (mean): 74.9-75.5 Male: 52.2-59.3% CHA2DS2-VASc ≥2: 96.6- 97.4% Comorbidity index: NR	Hospitalized due to thromboembolic stroke, intracranial haemorrhage, major extracranial bleeding, and all-cause mortality. ICD codes from the first hospital discharge diagnosis position	Not described how weighted cohort was composed. Balanced if the absolute value of the STD was ≤10%.	rivaroxaban: n=106,389 dabigatran: n=86,198 apixaban: n=73,039 Weighted cohort rivaroxaban: n=106,369 dabigatran: n=86,293 apixaban: n=72,921	Dabigatran and apixaban were associated with a more favourable benefit- harm profile than rivaroxaban.
Hernandez, 2017 USA	Medicare Nov 4, 2011 -Dec 31, 2013	NVAF patients (at any time before the index date; no NOAC treatment at least 3 months before the index date)	High dose: Age: <65: 5.0-6.3% Age: 65-74: 38.4-39.3% Age: ≥75: 55.3-55.7 Male: 45.9-49.5 CHADS2: 3.3-3.3 Comorbidity index: NR	ischemic stroke (inpatient, emergency room, or outpatient claim with primary or secondary, ICD-9 codes), other thromboembolic events, and all-cause mortality; ICD 9 codes, whether primary and secondary codes were used is not described. Any bleeding event and major bleeding; intracranial hemorrhage and gastrointestinal bleeding, not further described.	Extreme weights: not reported Balanced if the absolute value of the STD was ≤10%.	dabigatran n=9,138 rivaroxaban n=8,367	There was no difference in stroke prevention between rivaroxaban and dabigatran; however, rivaroxaban was associated with a higher risk of thromboembolic events other than stroke, death, and bleeding.
Larsen, 2016 Denmark	Danish national prescription registry, Danish national patient register, Danish civil registration system August, 2011 -Oct, 2015	NVAF patients who were naïve to oral anticoagulants (no use of oral anticoagulant within one year)	Age (median, IQR): 67.6 (62.0-72.4)-71.8 (65.7-78.9) Male: 56.9-66.1% CHA2DS2VASc: 2.2-2.8 Comorbidity index: NR	Ischaemic stroke or systemic embolism, ICD-10 codes whether primary and secondary codes were used is not described.	Extreme weights: not reported Balanced if the absolute value of the STD was ≤10%.	dabigatran: n=12,701 rivaroxaban: n=7,192 apixaban: n=6,349	Apixaban and dabigatran were associated with a significantly lower risk of death compared with rivaroxaban. Risk of any bleeding or major bleeding were significantly lower for apixaban and dabigatran than for rivaroxaban
Meng, 2019 Taiwan	National Health Insurance claims database June 1, 2012 - May 31, 2015	All NVAF patients aged ≥20 years who initiated NOACs during study period	Age <65: 11.8-13.5% Age 65-74: 29.7-32.7% Age ≥75: 53.8-58.4% Male: 54.6-56.2% CHA2DS2-VASc: 3.2-3.3 Comorbidity index: NR	all-cause death, ischemic stroke, intracranial hemorrhage, gastrointestinal hemorrhage needing transfusion, ICD-10 codes, whether primary and	Extreme weights: not reported Balanced if the absolute value of the STD was ≤10%.	dabigatran: n=13,505 rivaroxaban: n=6,551 Weighted pseudo- cohort	Rivaroxaban seemed to be associated with an increased risk of all- cause death compared with dabigatran

Author and country	Setting and study period	Patient characteristics (range between NOACs)	Primary outcome definition	IPWT details		Result/conclusion as reported in the article
			secondary codes were used is not described		dabigatran: n=13,508; rivaroxaban: n=6,547	



Table 3. Characteristics of the included articles that used adjusted Cox-proportional hazard models as primary analyses (n=10)

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	Sample size	Results/conclusion as reported in the article
Al-Khahili, 2016 Sweden	tertiary referral cardiology outpatient clinic (the Stockholm Heart Center) Dec, 2011 - May, 2014	NVAF patients from a single cardiology outpatient clinic incorporating the AF unit (initiate NOAC treatment)	Age: 72±8-73±8 Male: 50-51% CHA2DS2-VASc: 3-3 Comorbidity index: NR	Major bleeding was defined according to the criteria of the International Society of Thrombosis and Hemostasis	rivaroxaban: n=282; apixaban: n=251 dabigatran: n=233;	Rivaroxaban was associated with the highest bleeding rates owing mainly to the highest number of minor bleedings, and apixaban had the lowest bleeding rates and side effects
Alonso, 2017 USA	Truven Health MarketScan® Commercial Claims and Encounter Database and the Medicare Supplemental and Coordination of Benefits Database Jan 1, 2007 - Dec 31, 2014	NVAF patients with a first prescription of OAC after Nov 2, 2011.	Age: 67.2±12.4- 69.3±12.5 Male: 60.1-65.1% CHA2DS2-VASc: 2.9-3.6 Comorbidity index: NR	Hospitalization for liver injury potentially related to drug hepatotoxicity, ICD-9-CM codes in any position	rivaroxaban: n=30,347; dabigatran: n=17,286; apixaban: n=9,205	Risk of liver disease hospitalization was higher in rivaroxaban users compared to dabigatran and apixaban users
Chan, 2016 Taiwan	Taiwan National Health Insurance Research Database. Jan 1, 1996 - Dec 31, 2013	NVAF patients newly diagnosed	Age: 75±9-76±9 Male: 54-58 CHA2DS2-VASc: 4.1-4.1 Comorbidity index: NR	Ischemic stroke or systemic embolism, ICH, hospitalization for GI bleeding, acute myocardial infarction (AMI), all hospitalizations for bleeding, and all-cause mortality. All discharge diagnosis according to the ICD, whether primary and secondary codes were used is not described	dabigatran 110 mg: n= 5,921 rivaroxaban 10 mg: n=3,916	No differences were found between rivaroxaban and dabigatran in risk for thromboembolic events, intracranial haemorrhage, critical gastrointestinal bleeding, or all-cause mortality. However, rivaroxaban was associated with a higher risk for noncritical gastrointestinal bleeding than dabigatran

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	Sample size	Results/conclusion as reported in the article
Hemandez, 2017 USA	Medicare database Jan 1, 2013 - Dec 31, 2014	NVAF patients newly diagnosed	Age: 74.9±8.7-77.4±8.6 Male: 42.5-47.0% CHA2DS2-VASc: 4.3-4.7 Comorbidity index: NR	Ischemic stroke, death, bleeding events, gastrointestinal bleeding, treatment persistence. ICD-9 codes, whether primary and secondary codes were used is not described	rivaroxaban: n=5,139; apixaban: n=2,358; dabigatran: 1,415;	Apixaban had the most favourable effectiveness and safety profile
Lamberts, 2017 Denmark	Danish national patient registry, Danish national prescription registry, Danish civil personal registry up to December 31, 2015	NVAF patients ≥18 years, with newly prescribed OAC (no prescription at least 6 months before inclusion)	Age: 71.5±11.0-75.4±11.10 Male: 50.8-56.7% CHA2DS2-VASc: 2.7-3.2 Comorbidity index: NR	major bleeding events requiring hospitalisation, ICD-10 codes, whether primary and secondary codes were used is not described	dabigatran: n=15,413; apixaban: n=7,963; rivaroxaban: n=6,715;	Apixaban had a lower adjusted major bleeding risk compared with rivaroxaban and dabigatran
Lip, 2016 (Int J Clin Pract) USA	Truven MarketScan® Commercial & Medicare supplemental US database Jan 1, 2013 - Dec 31, 2013	NVAF patients ≥18 years with newly prescribed OAC (no OACs received at least 1 year before the start of the OAC treatment)	Age: 66.8±12.2-69.3±12.3 Male: 63.1-65.8% CHA2DS2–VASc: 2.6-2.8 CCI: 1.7-1.9	Major bleeding was identified using hospital claims, which had a bleeding diagnosis code as the first listed primary ICD-9 diagnosis code	rivaroxaban: n=10,050 dabigatran: n=4,173 apixaban: n=2,402	Initiation with rivaroxaban was associated with a significantly greater risk of major bleeding compared with initiation on apixaban. There was no significant difference in the risk of major bleeding among patients newly initiated on dabigatran compared with apixaban.
Mueller, 2019 Scotland	Prescribing Information System, the Scottish Morbidity Records/ Hospital Inpatients and Outpatient attendance datasets; National Records of Scotland Drug's approval date – Dec 2015	NVAF patients who initiated NOAC treatment	Age: 71.1±12.0-74.8±11.0 Male: 53.5-73.1% CHA2DS2-VASc: 2.5-3.0 CCI: 1.1-1.4	strokes, systemic embolism, death due to cardiovascular, pulmonary embolism, bleeding events, clinical endpoints, according to ICD-10 codes whether primary and secondary codes were used is not described	rivaroxaban: n=7,265 apixaban: n=6,200; dabigatran: n=1,112;	All NOACs were similarly effective in preventing strokes and systemic embolisms, while patients being treated with rivaroxaban exhibited the highest bleeding risks.
Staerk, 2018 Denmark	Danish national patient registry, Danish national prescription registry,	NVAF patients, first-time OAC users (no previous OAC use), between 30 and 100 years old	Standard dose: Age (median, IQR): 67(61, 71)-71(65, 78) Male: 55.4-63.7% CHA2DS2-VASc (median); 2-3	stroke/thromboembolism (TE), ischaemic stroke, major bleeding, intracranial bleeding and gastrointestinal bleeding, ICD-10 codes whether primary and	dabigatran: n=11,492 apixaban: n=11,064 rivaroxaban: n=8966	Rivaroxaban was associated with higher bleeding risk

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	Sample size	Results/conclusion as reported in the article
	Danish civil registration system March 1, 2012 - Dec 31, 2016		Comorbidity index: NR	secondary codes were used is not described		compared with dabigatran and apixaban and dabigatran was associated with lower intracranial bleeding risk compare with rivaroxaban and apixaban.
Tepper, 2018 USA	Truven MarketScan Commercial Claims and Encounter and Medicare Supplemental & Coordination of Benefits Early View Database Jan 1, 2013 - Oct 31, 2014	NVAF patients aged ≥18 years with new initiators of NOACs or switched from warfarin to a NOAC	Age: 68±12-70±12 Male: 65.3-62.7 CHA2DS2-VASc: 2.4-2.5 CCI: 1.6-1.8	Bleeding, ICD-9-CM codes, whether primary and secondary codes were used is not described	rivaroxaban: n=30,529 dabigatran: n=20,963 apixaban: n=8,785;;	Rivaroxaban appeared to have an increased risk of any bleeding, clinically relevant nonmajor bleeding, and major inpatient bleeding, compared to apixaban patients. There was no significant difference in any bleeding, clinically relevant non-major bleeding, or inpatient major bleeding risks between patients treated with dabigatran and apixaban.
Vinogradova 2018 UK	UK general practices contributing to QResearch or Clinical Practice Research Datalink 2011 - 2016	NVAF patients, new NOAC (received no OAC treatment in at least the last 12 months)	QResearch: Age: 74.7±10.7-76.5±10.9 Male: 51.8-58.0% CHA2DS2-VASc: NR Comorbidity index: NR	Major bleeding after entry to the study which led to a hospital admission or death, based on linked hospital or mortality records.	rivaroxaban: n= 16,547 apixaban: n= 10,601 dabigatran: n=5,537	Apixaban was associated with a lowe risk of major bleed than rivaroxaban. Rivaroxaban was associated with a higher risk of intracranial bleed compared to apixaban rivaroxaban was associated with higher risks compared with apixaban for haematuria, all gastrointestinal bleed and upper gastrointestinal bleed.

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	Sample size	Results/conclusion as reported in the article
						The risk of primary ischaemic stroke did not differ between any of the anticoagulants

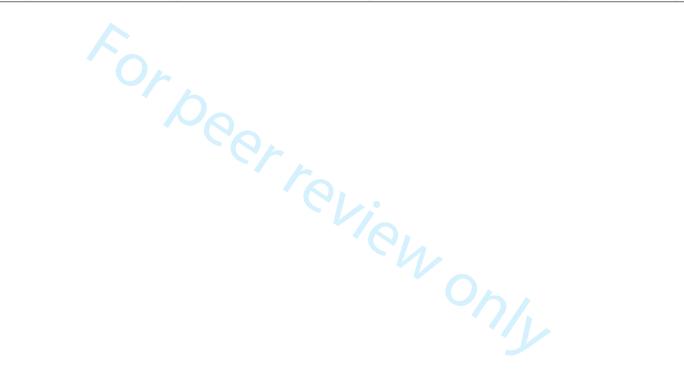


Table 4. Characteristics of the included articles that used unadjusted primary analysis (n=2)

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	Primary analysis	Sample size	Results/conclusio n as reported in the article
Cerda, 2019 Spain	Oral Anticoagulant Treatment Unit of the Hemostasis and Thrombosis Department of the University Hospital Vall d'Hebron from Barcelona (Spain) Jan, 2015 - Sept, 2017	NVAF patients with nonvalvular AF, with or without prior stroke, that had started treatment with any NOAC for the prevention of stroke	Age: 73.1±15.2-78.9±8.7 Male: 45.1-63.4% CHA2DS2-VASCc: 3.9-4.4 Comorbidity index: NR	Major bleeding according to ISTH 2005	log-rank test	rivaroxaban: n=663; dabigatran: n=352 apixaban: n=325 edoxaban: n=103	Rates of ischemic stroke and intracranial hemorrhage were similar among different NOACs, but rates of major bleeding were higher with dabigatran and apixaban and lower with rivaroxaban.
Li, 2017 China	Queen Mary Hospital, Hong Kong Jan, 2008 - Dec, 2014	NVAF patients diagnosed during study period.	Age: 71.9±11.1-73.3±12.1 Male: 53.1-59.8% CHA2DS2-VASc: 3.6-3.7 Comorbidity index: NR	The primary outcome was a composite of hospital admission with ischemic stroke or ICH, or death during the follow-up period. ICD-10 codes in medical records, and discharge summaries, whether primary and secondary codes were used is not described	Cox proportional hazard model (likely unadjusted, but this is not clearly described in the article)	rivaroxaban: n=669; dabigatran: n=467	Dabigatran had a lower ischemic stroke risk compared with patients on rivaroxaban. There was no significant difference in ischemic stroke risk between those on rivaroxaban and dabigatran.

Table 5. Characteristics of the included articles that used propensity score stratification as primary analyses (n=1)

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	PS details	Sample size	Results/conclusion as reported in the article	
Gorst- Rasmussen, 2016 Denmark	Danish national prescription registry, Danish national patient register, Danish civil registration system Feb. 1, 2012 - July 31, 2014	NVAF patients who were new-users of OAC (no OAC treatment in at least the last two years)	CHA2DS2-Vasc: 2.1-3.0 Comorbidity index: NR	ischemic stroke/systemic embolism/transient ischemic attack, any bleeding and all-cause death. ICD-10 codes, whether primary and secondary codes were used is not described	Asymmetric trimming of the propensity score. Trimmed propensity score was used in 10 deciles as strata Balanced if the absolute value of the STD was ≤10%.	dabigatran: n=8,908 rivaroxaban: n=1,405;	Rivaroxaban and dabigatran had similar stroke rates. Bleeding and mortality rates were higher in rivaroxaban versus dabigatran.	
STD was ≤10%.								

Table 6. Main differences between the included studies (n=39)

Study item	Range, total number of studies, or description				
Country	USA: Denmark: Taiwan: China: France: Scotland: Sweden: Spain: UK:	n=24 n=5 n=4 n=1 n=1 n=1 n=1 n=1			
NOAC included in included studies	Dabigatran: Rivaroxaban: Apixaban: Edoxaban:	n=39 n=39 n=26 n=1			
Most prescribed NOAC in included studies per country	Dabigatran: Rivaroxaban: Apixaban: Edoxaban: About equal*:	Denmark USA, UK, China, Scotland, and Taiwan In none of the included studies In none of the included studies France, Spain, Sweden			
Baseline characteristics	Mean age, years: % males: Mean CHA2DS2-Vasc:	65-84 39-73 2.1-4.9			
Primary study outcomes	Effectiveness outcomes:				
Statistical approaches	PS matching: IPTW: PS stratification: Cox PH regression model: Unadjusted analyses:	n=18 n=8 n=1 n=10 n=2			
Sample size	N=698 - N=265,583				
Study results	included:	xaban, rivaroxaban and dabigatran were urable compared to dabigatran and e NOAC: n=13			

^{*} about equal distribution between dabigatran, rivaroxaban and apixaban. Edoxaban is not included in these studies.

More than 50% of the studies were conducted in the USA (n=24),[16-39] five were conducted in Denmark,[40-44] four in Taiwan,[45-48] and one in France,[49] Sweden,[50] Scotland,[51] the UK,[52] Spain,[53] and China.[54] Dabigatran and rivaroxaban were included in all 39 studies, apixaban was included in 26 studies and edoxaban was included in 1 study. Next to these NOACs, VKA was included in 25 of these studies as one of the comparators. The results below focus on the NOAC to NOAC comparisons only.

In the studies that included apixaban, dabigatran and rivaroxaban, rivaroxaban was most dominantly used in the USA, UK, Scotland, and Taiwan, while dabigatran was the most prescribed NOAC in Denmark. In three other European studies the distribution was about equal between the three NOACs. In none of the included studies, apixaban was the most dominantly prescribed NOAC.

Setting

Most studies concerned patient registries, pharmacy or prescription databases and/or health insurance databases (n=39), while there were three clinical practice based studies.[50, 53, 54]

Study population

All studies included only NVAF patients. In seven studies, it was specifically described that patients were newly diagnosed with NVAF and initiated NOAC treatment during study period.[21, 27, 34, 37, 40, 45, 54] None of the other studies included prevalent users of (N)OAC, but included e.g. 'newly treated', 'initiating treatment', 'new users', 'first-time prescription' of NVAF patients who were prescribed (N)OAC. In some studies (N)OAC use in the past (between 3 months and 2 years before index date) was allowed, while this seemed not be allowed in some other studies, or it was not described.

Inclusion criteria

Five studies concerned elderly patients specifically (i.e. ≥65 years old),[19, 21, 23-25] two included adults ≥45 years old,[33, 40] and one study included patients between 30 and 100 years of age.[44] The other studies included all adults with atrial fibrillation (it was assumed that if no further age specification was provided, 'adults' meant that all >18 years old were included). In one study only patients who were hospitalised for bleeding after start with OAC treatment were included.[22] No other focus on a specific group of AF patients was found.

Exclusion criteria

NOAC use that could be related to other disorders, such as transient AF, major knee or hip surgery, venous thromboembolism or pulmonary embolism, were specifically described as exclusion criteria in most studies,

except in ten studies.[16, 27, 28, 33-35, 50, 52-54] In one study patients with liver injury before their first oral anticoagulant (OAC) prescription were specifically excluded.[18]

Baseline characteristics

Baseline characteristics of the NVAF patients differed between studies. Mean age ranged from 65-84 years between the studies. The percentage of males ranged from 39-73%, and the mean CHA2DS2-Vasc Score ranged from 2.1-4.9. Excluding the five studies that specifically focussed on an elderly population of ≥65 years old and the two additional studies that used the Medicare database (only patients of 65 years or older are in Medicare), the mean age ranged from 65-78 years old. Different measures were used to assess the comorbidity index: Charlson comorbidity index, Charlson-Deyo index and Gagne comorbidity score, while in 30 of the 43 studies no comorbidity index was presented.

Selection of covariates

Most studies (n=34) did not provide a rationale for the selection of covariates that were included in the PS model or in adjusted analysis. However, in one of the articles an extensive rationale and selection procedure of covariates that were included in the analysis was provided.[33] In three other studies, the authors selected covariates based on medical knowledge on risk factors with reference to earlier published studies.[31, 39, 52] In one other study it was reported that sociodemographic and clinical characteristics that were associated with treatment initiation and the risk of major bleeding were included in the model to adjust for differences across cohorts, without further explanation or reference.[30]

Definition primary study outcomes

Primary outcomes differed between the studies. Effectiveness outcomes included in the studies included stroke, systemic embolism, (or composite of stroke/systemic embolism), all-cause death, myocardial infarction, venous thromboembolism and safety outcomes included major bleeding, or a specific type of bleeding (e.g. intracranial haemorrhage, gastrointestinal bleeding etc.) and liver injury. In most studies, ICD-9 or ICD-10 codes were used, but whether this concerned a primary diagnosis only or whether it could be either a primary or a second diagnosis differed between the studies. In some studies it was not described whether the ICD codes referred to primary diagnosis only or to a primary or secondary diagnosis.

Statistical approaches to adjust for confounding (primary analysis)

In 18 studies, PS matching was done.[16, 19-21, 23, 26, 29, 30, 32-37, 39, 40, 47, 49] IPTW was used in eight studies.[17, 22, 24, 25, 28, 43, 46, 48] PS stratified analyses was done in one study.[41] In twelve studies, the primary analyses utilised a Cox PH regression model in which adjustment for confounding was done.[18, 27, 31,

38, 42, 44, 45, 50-52] Finally, in two studies no adjustment for differences in baseline characteristics was performed.[53, 54]

PS matching

Co-variates

Creatinine clearance was not included as a covariate in any of the 18 studies. All 18 studies took the following covariates into account: age, sex, CHA2DS2-VASc score and/or the individual comorbidities included in this score, HAS-BLED score and/or the individual conditions included in this score (except alcohol use in Lai et al. [47]), renal disease, and co-medication use such as antiplatelets. Some included other comorbidities, such as cancer, rheumatic disease, specific heart diseases, COPD, HIV, dementia, depression, neurological disorders, and/or a various list of co-medications as well.

Matching method

In one study the matching method was not described.[49] In two studies, the calliper used was not described.[23, 29] In seven studies 1:1 PS matching without replacement was used and a calliper of 0.01 was applied.[16, 19, 20, 26, 30, 32, 36] Five other studies also matched 1:1 without replacement but used another calliper: in three studies a calliper of 0.2 was used,[39, 40, 47] while two others used a calliper of <0.25.[33, 35] In three studies, three-way matching was used.[21, 34, 37]

Balance co-variates

In two studies it was not described how the balance between covariates was evaluated.[33, 35] In two studies the balance was evaluated using p<0.05 (of which one also used standardized difference of <10%),[23, 47] and in another study it was stated that the groups were comparable even though a p value of >0.05 was found.[29] Balance was checked with an absolute standardized difference of <10% in 13 studies.[16, 19-21, 26, 30, 32, 34, 36, 37, 39, 40, 47, 49] Balance was reached in all studies after matching.

Sample size

In four studies the sample size before matching was not reported[29, 35, 36, 39] and in one study the sample size after matching was not reported.[34] At study start (before PSM), sample size between the NOACs differed greatly, except in three studies.[21, 37, 40]

IPTW

In one study, balance was tested using ANOVAs for significant differences.[22] Balance was checked with an absolute standardized difference of <10% in the other nine studies.[17, 24, 25, 28, 43, 46, 48] Balance was reached in all studies after IPTW.

There was no reporting on extreme weights in the eight included studies.[17, 22, 24, 25, 28, 43, 46, 48]

PS stratification

In one study, asymmetric trimming of the PS was done, which resulted in a small part of both treatment groups being removed in order to gain in comparability. Balance in co-variates was reached with standardized difference of <10%. In a Cox model this trimmed PS was used in 10 deciles as strata.[41]

Cox HP regression models

In ten studies, Cox HP regression models were applied with adjustment for a number of confounders.[18, 27, 31, 38, 42, 44, 45, 50-52] In one of these studies, the number of events per variable was not sufficient for such an analyses.[50] The ratio was acceptable in the other studies for at least some of the outcomes.[18, 28, 31, 38, 42, 44, 45, 51, 52]

Unadjusted analysis

In two studies no adjustment for confounding factors seemed to have been done, even though significant differences between treatment groups existed at baseline. Cerda et al. presented events per 100 patient-years and used a log-rank test to determine whether outcomes differed between the NOACs.[53] Li et al. conducted a Cox proportional hazard model, likely unadjusted, but this was not clearly described in the article.[54]

Study results

Which NOAC performed best differed between the included studies. We found only one study that included all four NOACs, in which no preference for one specific NOAC was found, except that rates of major bleeding were lower with rivaroxaban.[53] Of the 26 studies in which apixaban, rivaroxaban and dabigatran were included, apixaban was favourable compared to dabigatran and rivaroxaban in 13 studies, of which 10 were from the USA, two from Europe and one from Asia,[16, 17, 19, 20, 23, 26, 28, 29, 32, 36, 42, 50, 52] while dabigatran and rivaroxaban were not found to be the single most favourable NOAC in any of the remaining 13 studies. Results for these 13 studies were mixed, with either no favourable NOAC at all, or one NOAC was selected as the least favourable, while the other two NOACs did not differ.

Naïve trial analysis

The primary efficacy endpoint (Strokes/SE) in the warfarin arms were estimated at 1.69% (RE-LY),[3] 2.2% (ROCKET),[6] 1.60% (ARISTOTLE),[5] and 1.50% (ENGAGE),[4] see table 7. From this range we chose a relatively arbitrary base rate of 1.6% and applied the observed risk reduction to estimate comparable base rates of 1.05% for dabigatran, 1.24% for rivaroxaban, 1.26% for edoxaban and 1.27% for apixaban. Using the sample size calculator[55] the biggest expected difference was between dabigatran and apixaban and it was estimated that a trial sample size with 51,847 patients would be needed to confirm this difference. The smallest difference was between edoxaban and apixaban and a trial of 7,994,340 patients required to confirm that difference.



Table 7. Primary efficacy and safety endpoints of the four pivotal trials.

	RE-LY [3]			ROCKET	AF [6] ARISTOTLE [5]		ENGAGE-AF [4]		4]	
	Dabigatr	Dabigatr	Warfari	Rivaroxab	Warfari	Apixaba	Warfari	Edoxab	Edoxab	Warfari
	an 150	an 110	n	an	n	n	n	an	an	n
	mg	mg						60 mg	30 mg	
	N=6076	N=6015	N=602	N=7131	N=713	N=9120	N=908	N=7035	N=7034	N=703
			2		3		1			6
Stroke/S	1.11	1.53	1.69	1.7	2.2	1.27	1.60	1.18	1.61	1.50
E										
(%/year)										
Major	3.11	2.71	3.36	3.6	3.4	2.13	3.09	2.75	1.61	3.43
bleeding										
(%/year)										



The primary safety endpoint was major bleeding for RE-LY, ARISTOTLE, and ENGAGE AF and major bleeding plus clinically relevant non-major bleeding for ROCKET AF, but data on major bleeds only for ROCKET-AF are available as well. Major bleeds in the warfarin arms were estimated at 3.36% (RE-LY),[3] 3.4% (ROCKET),[6] 3.09% (ARISTOTLE),[5] and 3.43% (ENGAGE).[4] From this range we choose a relatively arbitrary base rate of 3.2% and applied the observed risk reduction to estimate comparable base rates of 2.21% for apixaban 2.57% for edoxaban, 2.96% for dabigatran and 3.29% for rivaroxaban. Using the sample size calculator,[55] the biggest expected difference was between rivaroxaban and apixaban and it was estimated that a trial with 7,196 patients would be needed to confirm this difference. A much smaller difference is between edoxaban and apixaban which would require a trial of 56,512 patients to confirm that difference.

DISCUSSION

In total, we found 39 studies directly comparing the effectiveness and/or safety of at least two NOACs in NVAF patients. Three studies can be considered to be of low quality due to insufficiently described methods and/or small sample size[50, 53, 54].

Even though the remaining studies could be considered of sufficiently quality based on the technical aspects of the studies, there are some issues that can hamper the generalisability of the results. These issues concern remaining confounding, the use of a smaller or broader calliper, differences in baseline characteristics between studies, channelling bias and change in treatment paradigm.

Balance in baseline characteristics between NOACs was checked with p-values or a standardized difference of <10%. Balance was well at baseline in some studies, or was reached after PS matching or IWTP.[56] Even though some studies included over 40 covariates in their PS, and balance was reached for all of these variables, one should keep in mind that balance between unmeasured or unmeasurable factors cannot be assumed.[14] Therefore, there is always a possibility of residual confounding. This possibility was acknowledged in all included studies. Creatinine clearance for instance, seems to be an important covariate as subgroup analyses from the pivotal trials suggest that renal clearance might be an effect modifier.[5, 57] Only in one study however, the authors were able to take renal clearance into account in the adjusted analyses.[50] Especially when prescription of a certain NOAC in daily practice is driven by creatinine clearance, not adjusting for this variable may lead to biased results.

In general, a calliper of <0.2 of the standard deviation of the logit of the PS is considered to be 'optimal'.[58] About half of the included PS matching studies used a smaller calliper, namely of <0.1. This means that the matching is more precise in these studies, but the disadvantage is that possibly more patients cannot be matched to another patient due to this smaller allowed maximum differences, and thus will be excluded from the analysis. Excluding patients from the analysis will limit the generalisability of the results to the total patient population, especially when the excluded patients differ from the included patients, e.g. on the baseline risk for stroke.

All included studies focused on NVAF patients only. NOAC use that could have been related to other conditions was excluded specifically in 34 of the 36 included studies. In eight studies, inclusion criteria regarding age were applied. Three of these will likely still cover the largest part of NOAC users as they set relatively broad age ranges. The other five focussed on an elderly population of NVAF patients of ≥65 years old. Besides applying specific inclusion criteria regarding age in some studies, these differences also depended on the specific registry

or database that was used, e.g. Medicare is for people of 65 years old or older. Even though only five of the included studies focused on an elderly NVAF population, and the others applied broad age ranges, there were differences in mean age, proportion of males and mean CHA2DS2Vasc score between the studies, which can have an impact on the results and jeopardize the generalisability of the results.

Rivaroxaban was the most prescribed NOAC in almost all included studies from the USA. However, in the first quarter of 2017, apixaban was the most prescribed NOAC in NVAF in the USA (i.e. in 50% of new OAC prescriptions). Especially older patients, women, increased stroke or bleeding risk and having comorbidities was associated with prescription of apixaban versus other NOACs.[59] Rivaroxaban was also the most prescribed NOAC in the included studies from the UK and Scotland. Based on the CPRD, 56.5% of the OAC prescriptions concerned a NOAC, of which rivaroxaban was still described most often in 2015.[60] Dabigatran was described most often in the studies from Denmark. Haastrup et al. described that most AF patients that initiated NOAC received dabigatran between 2008 and 2016, but a trend was observed that per 1000 person-years the number of patients described dabigatran decreased and the number of patients receiving rivaroxaban and apixaban increased.[61] This shows that the treatment paradigm changed over time, and might still be changing, and this pattern differs between the USA, Europe and Asia. Channelling bias therefore likely occurs.

Our naïve analysis predicts that in terms of the primary efficacy outcome observational studies will need a relatively high number of patients to be able to demonstrate the differences between the NOACs and a small sample size will not allow robust comparison to be made.

The pattern of major bleeding events seen in the included observational studies, confirms the expectation from our naïve analysis of the pivotal clinical trials that rivaroxaban seems to have the least favourable safety profile among apixaban and dabigatran. The findings are not consistent to allow for a robust conclusion between apixaban and dabigatran which confirms the need for a high number of patients, although a trend for a slight better safety profile of apixaban can be observed.

The requirement for a high number of patients to compare NOACs both in terms of efficacy and safety as predicted by the pivotal trial results is confirmed by the findings of the observational studies. This finding may support the claim that the differences between the NOACs are relatively small.

In summary, even though the larger part of these studies are technically well conducted, these studies have some important limitations regarding the generalisability of the study results especially given the relatively high patient number required for a meaningful comparison between NOACs. Most studies included all NVAF patients on NOAC available in the registry/database during the study period and did not apply further specific in- and exclusion criteria, but differences between studies regarding baseline characteristics existed. Mean age at study

start and baseline risk for stroke (CHA2DS2-VASc score) differed between the studies. As channelling bias cannot be ruled out, the result of these studies might not be generalisable. Furthermore, results from the PS studies are only applicable to the patients that were kept in the analyses as patients excluded from the analysis likely differ from the ones that were included in the analysis. The 1:1 matched cohorts depended on the sample size of the NOAC with the least number of patients and as a result many patients from the larger of the two NOAC groups were excluded as they could not be matched. Besides these study specific limitations, differences in reimbursement and ICD-coding exist between the USA and Europe, and also within Europe. Differences in reimbursement may lead to differences in adherence and non-adherence could lead to worse outcomes. These limitations should be kept in mind when results of these studies are used to decide what NOAC should be prescribed for a certain patient. Finally, given the small differences between efficacy and safety outcomes between NOACs, the element of patient preference should be taken into consideration, [62] as tailoring anticoagulation treatment towards patient preferences can promote adherence to treatment.

ACKNOWLEDGEMENTS

The authors would like to thank Pearl Gumbs for the initiation of this project and her input with regard to study design and interpretation of the data. Pearl Gumbs was working at Daiichi Sankyo Europe at that time.



COMPETING INTEREST

EB reports grants from Daiichi Sankyo, during the conduct of the study; grants from Daiichi Sankyo, outside the submitted work. BvH reports grants from Daiichi Sankyo, during the conduct of the study. SH reports personal fees from Aspen, personal fees from Bayer, personal fees from BMS/Pfizer, personal fees from Daiichi-Sankyo, personal fees from Portola, outside the submitted work. GS reports personal fees from Daiichi Sankyo Europe Gmbh, outside the submitted work. AC reports personal fees from Daiichi Sankyo Europe, during the conduct of the study; grants and personal fees from Bayer AG, personal fees from Boehringer Ingelheim, grants and personal fees from Bristol-Myers Squibb, grants and personal fees from Pfizer Limited, personal fees from Portola Pharmaceuticals Inc, personal fees from Janssen, personal fees from ONO Pharmaceuticals, from AbbVie, outside the submitted work. OFK.

AUTHORS' CONTRIBUTION

EB: Conceptualization (support); Methodology (equal); Writing - Original Draft Preparation; Writing - Review & Editing (equal). BvH: Conceptualization (support); Methodology (equal); Writing - Review & Editing (equal). SH: Conceptualization (support); Writing - Review & Editing (equal). GS: Conceptualization (support), Supervision; Writing - Review & Editing (equal). ATC: Conceptualization (lead); Writing - Review & Editing (equal)



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DATA SHARING STATEMENT

No new data were generated or analysed in support of this research.

FUNDING

This work was supported by Daiichi Sankyo Europe.



PATIENT AND PUBLIC INVOLVEMENT

This research was done without patient involvement. Patients were not invited to comment on the study design and were not consulted to develop patient relevant outcomes or interpret the results. Patients were not invited to contribute to the writing or editing of this document for readability or accuracy.



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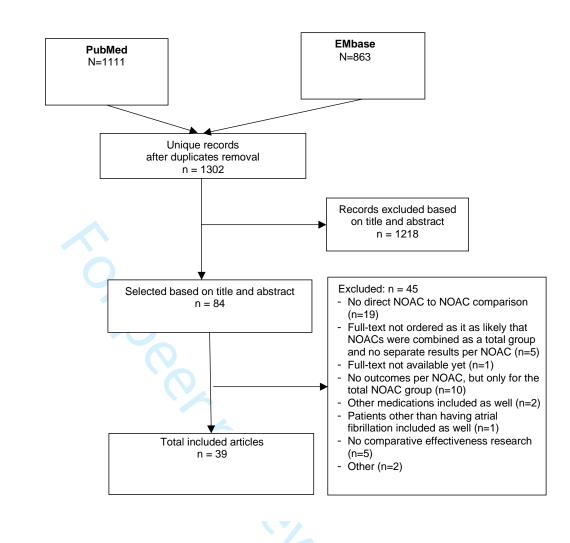
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Identification

Eligibility



The following search string was used for PubMed, and adapted for Cochrane and EMBase.

#1. NOAC

direct oral anticoagulant*[tiab] OR direct oral anti-coagulant*[tiab] OR direct oral anticoagulation[tiab] OR direct oral anticoagulation[tiab] OR direct oral anti-coagulation[tiab] OR direct-acting oral anti-coagulant*[tiab] OR direct-acting oral anti-coagulation[tiab] OR DOAC[tiab] OR novel oral anticoagulant*[tiab] OR novel oral anticoagulant*[tiab] OR Novel oral anticoagulation[tiab] OR Novel oral anti-coagulation[tiab] OR Novel oral anti-coagulation[tiab] OR Novel oral anticoagulation[tiab] OR Apixaban[tiab] OR Edoxaban[tiab] OR Dabigatran[tiab] OR "Non VKA Oral Anticoagulant"[tiab] OR "Non Vitamin K Antagonist Oral Anticoagulant"[tiab]

#2. Comparative effectiveness studies

comparative effectiveness research[mesh] OR comparative effectiveness[tiab] OR real-world[tiab] OR real-life[tiab] OR cohort studies[mesh] OR cohort[tiab]

#3. Atrial fibrillation

atrial fibrillation[tiab]

Limits:

Language: English



PRISMA 2009 Checklist

Section/topic	#	Checklist item	Reported on page #
TITLE			
Title	1	Identify the report as a systematic review, meta-analysis, or both.	1
ABSTRACT			
Structured summary 3	2	Provide a structured summary including, as applicable: background; objectives; data sources; study eligibility criteria, participants, and interventions; study appraisal and synthesis methods; results; limitations; conclusions and implications of key findings; systematic review registration number.	2
INTRODUCTION			
Rationale	3	Describe the rationale for the review in the context of what is already known.	3,4
8 Objectives	4	Provide an explicit statement of questions being addressed with reference to participants, interventions, comparisons, outcomes, and study design (PICOS).	4
METHODS			
2 Protocol and registration 3	5	Indicate if a review protocol exists, if and where it can be accessed (e.g., Web address), and, if available, provide registration information including registration number.	NA
Eligibility criteria	6	Specify study characteristics (e.g., PICOS, length of follow-up) and report characteristics (e.g., years considered, language, publication status) used as criteria for eligibility, giving rationale.	5,6
Information sources	7	Describe all information sources (e.g., databases with dates of coverage, contact with study authors to identify additional studies) in the search and date last searched.	5,6
Search	8	Present full electronic search strategy for at least one database, including any limits used, such that it could be repeated.	S1_Table
2 Study selection	9	State the process for selecting studies (i.e., screening, eligibility, included in systematic review, and, if applicable, included in the meta-analysis).	5
Data collection process	10	Describe method of data extraction from reports (e.g., piloted forms, independently, in duplicate) and any processes for obtaining and confirming data from investigators.	5
Data items	11	List and define all variables for which data were sought (e.g., PICOS, funding sources) and any assumptions and simplifications made.	5
⁹ Risk of bias in individual ¹⁰ studies	12	Describe methods used for assessing risk of bias of individual studies (including specification of whether this was done at the study or outcome level), and how this information is to be used in any data synthesis.	5,6
Summary measures	13	State the principal summary measures (e.g., risk ratio, difference in means).	NA
Synthesis of results	14	Describe the methods of handling data and combining results of studies, if done, including measures of consistency (e.g., I²) for each meta-analysis. For peer review only - http://bmjopen.bmj.com/site/about/guidelines.xhtml	NA

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PRISMA 2009 Checklist

		Page 1 of 2	
Section/topic	#	Checklist item	Reported on page #
Risk of bias across studies	15	Specify any assessment of risk of bias that may affect the cumulative evidence (e.g., publication bias, selective reporting within studies).	NA
Additional analyses	16	Describe methods of additional analyses (e.g., sensitivity or subgroup analyses, meta-regression), if done, indicating which were pre-specified.	NA
RESULTS			
Study selection	17	Give numbers of studies screened, assessed for eligibility, and included in the review, with reasons for exclusions at each stage, ideally with a flow diagram.	S2_Fig
Study characteristics	18	For each study, present characteristics for which data were extracted (e.g., study size, PICOS, follow-up period) and provide the citations.	8-21
Risk of bias within studies	19	Present data on risk of bias of each study and, if available, any outcome level assessment (see item 12).	8-27
Results of individual studies	20	For all outcomes considered (benefits or harms), present, for each study: (a) simple summary data for each intervention group (b) effect estimates and confidence intervals, ideally with a forest plot.	
Synthesis of results	21	Present results of each meta-analysis done, including confidence intervals and measures of consistency.	NA
Risk of bias across studies	22	Present results of any assessment of risk of bias across studies (see Item 15).	NA
Additional analysis	23	Give results of additional analyses, if done (e.g., sensitivity or subgroup analyses, meta-regression [see Item 16]).	NA
DISCUSSION			
Summary of evidence	24	Summarize the main findings including the strength of evidence for each main outcome; consider their relevance to key groups (e.g., healthcare providers, users, and policy makers).	30
Limitations	25	Discuss limitations at study and outcome level (e.g., risk of bias), and at review-level (e.g., incomplete retrieval of identified research, reporting bias).	30,31
Conclusions	26	Provide a general interpretation of the results in the context of other evidence, and implications for future research.	31,32
FUNDING			
Funding	27	Describe sources of funding for the systematic review and other support (e.g., supply of data); role of funders for the systematic review.	

41 From: Moher D, Liberati A, Tetzlaff J, Altman DG, The PRISMA Group (2009). Preferred Reporting Items for Systematic Reviews and Meta-Analyses: The PRISMA Statement. PLoS Med 6(7): e1000097. 42 doi:10.1371/journal.pmed1000097

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CRITICAL APPRAISAL AND ISSUES REGARDING GENERALISABILITY OF COMPARATIVE EFFECTIVENESS STUDIES OF NOACS IN ATRIAL FIBRILLATION AND THEIR RELATION TO CLINCIAL TRIAL DATA - A SYSTEMATIC REVIEW

Journal:	BMJ Open
Manuscript ID	bmjopen-2020-042024.R1
Article Type:	Original research
Date Submitted by the Author:	14-Jan-2021
Complete List of Authors:	Bunge, Eveline; Pallas Health Research and Consultancy van Hout, Ben; University of Sheffield, HEDS, ScHARR Haas, Sylvia; Formerly Technical University Of Munich Spentzouris, Georgios; Daiichi Sankyo Europe GmbH Cohen, Alexander; Guys and St Thomas' NHS Foundation Trust, King's College
Primary Subject Heading :	Cardiovascular medicine
Secondary Subject Heading:	Research methods
Keywords:	Adult cardiology < CARDIOLOGY, STATISTICS & RESEARCH METHODS, Cardiology < INTERNAL MEDICINE

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CRITICAL APPRAISAL AND ISSUES REGARDING GENERALISABILITY OF COMPARATIVE

EFFECTIVENESS STUDIES OF NOACS IN ATRIAL FIBRILLATION AND THEIR RELATION TO CLINCIAL

TRIAL DATA - A SYSTEMATIC REVIEW

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Word count: 4406

KEY WORDS

atrial fibrillation; non-vitamin K antagonist oral anticoagulants (NOAC); comparative effectiveness research; Tot beet crien only systematic review

ABSTRACT

Objective: To critically appraise the published comparative effectiveness studies on non-vitamin K antagonist oral anticoagulants (NOAC) in nonvalvular atrial fibrillation (NVAF). Results were compared with expectations formulated on the basis of trial results with specific attention to the patient years in each study.

Methods: All studies that compared the effectiveness or safety between at least two NOACs in patients with NVAF were eligible. We performed a systematic literature review in Medline and EMbase to investigate the way comparisons between NOACs were made, search date 23-04-2019. Critical appraisal of the studies was done using amongst other ISPOR checklists for comparative effectiveness research

Results: We included 39 studies in which direct comparison between at least two NOACs were made. Almost all studies concerned patient registries, pharmacy or prescription databases and/or health insurance database studies using a cohort design. Corrections for differences in patient characteristics was applied in all but two studies. Eighteen studies matched using propensity scores, eight studies weighted patients based on the inverse probability of treatment, one study used propensity score stratification and ten studies applied a proportional hazards model. These studies have some important limitations regarding unmeasured confounders and channelling bias, even though the larger part of the studies were well conducted technically. On the basis of trial results, expected differences are small and a naïve analysis suggests trials with between 7,700 and 59,500 patients are needed to confirm the observed differences in bleedings and between 51,800 and 7,994,300 to confirm differences in efficacy.

Discussion: Comparisons regarding effectiveness and safety between NOACs on the basis of observational data, even after correction for baseline characteristics, may not be reliable due to unmeasured confounders, channelling bias and insufficient sample size. These limitations should be kept in mind when results of these studies are used to decide on ranking NOAC treatment options.

STRENGHTS AND LIMITATIONS OF THIS STUDY

- To our knowledge, this is the first systematic review that critically appraised the quality and generalisability of the comparative effectiveness studies on NOACs in atrial fibrillation patients and to relate this to clinical trial data
- A naïve trial analysis was conducted to estimate the number of patients needed in a randomised clinical



INTRODUCTION

Guidelines state a preference for non-vitamin K antagonist oral anticoagulants (NOACs) above vitamin K antagonists (VKA) in patients with nonvalvular atrial fibrillation (NVAF) requiring prevention of stroke and systemic embolism.[1, 2] However, no recommendation for a specific NOAC is made in these guidelines, and in daily practice, physicians have to make a choice which of the four available NOACs (dabigatran, rivaroxaban, apixaban, edoxaban) they prescribe for a particular patient.[3-6]

In the absence of head-to-head trials, comparative effectiveness research (CER) has been conducted to compare the NOACs with regard to effectiveness and safety. This is also described as real-world evidence; i.e. the data will come from patients treated in daily practice. Comparisons on effectiveness and safety between NOACs are however not easy to make, as patients will not be prescribed one of the NOACs at random. The choice of a certain NOAC for a patient will at least partly be driven by patient characteristics, such as age, concomitant medications, and the risk of stroke and/or bleeding. This can lead to systematic differences between the treatment groups, which is known as channeling bias. [7] In order to make a valid comparison on effectiveness and safety between the NOACs, adjusting for these characteristics is necessary when these characteristics are also related to the outcome (confounding variables).

Several techniques exist to correct for imbalances in risks, but there is no gold standard and all methods have advantages and disadvantages. Cox proportional hazards (Cox PH) regression model adjustment can be used but large sample sizes are needed when number of events is relatively low and the number of covariates is high (as a rule of thumb, about 10 events per predictor variable [8]) and these large sample sizes are not always available. Event rates are low, around 1 per 100 patient years for efficacy outcomes and to detect differences, even in a randomized clinical trial, one needs substantial numbers of patients. This number would only increase when the results are contaminated by a lack of balance between the patients groups. Another method to adjust for confounding is using propensity scores (PS) to create comparable patient groups before the analysis. A propensity score is the probability of an individual receiving a specific treatment given a specific set of patient characteristics (e.g. age, gender, comorbidities). [9] Variables related to the outcome should be included in the propensity score despite their strength of association on treatment (exposure) selection. This will increase the precision of the estimated exposure effect, while bias will not be increased. Variables that are related to the exposure but not the outcome will decrease the precision of the estimated exposure effect without decreasing bias.[10] Adjustment for confounding using PS can be done by matching the treatment groups on the PS, by weighing treatment groups based on the PS inverse probability of treatment weighting (IPTW), by PS stratification, or by covariate adjustment using the PS.[9, 11] Well conducted PS methods will lead to treatment groups that are very well comparable regarding important confounders, which increases the confidence in the

results, however, there are also some disadvantages. For instance, in PS matching studies, patients who cannot be matched to another patient will be excluded from the analyses, and in IPTW, when patients on one treatment have a low propensity score and patients treated with the other treatment have a high propensity score, extreme weights can occur which can bias the results.[12]

To gain more understanding in how the above described methodologies were applied in peer-reviewed CER on effectiveness and safety in NOACs in NVAF patients, we conducted a systematic literature review. Within this we compare the results with those from a naïve analysis of the results of the four major trial for rivaroxaban, apixaban, dabigatran and edoxaban, and compare the results from the various analyses with those from the trials.

METHODS

Information sources, search strategy and eligibility criteria

We performed a systematic literature review to identify peer-reviewed comparative effectiveness research on NOACs in patients with atrial fibrillation. A search in Medline (access through PubMed) and EMbase was performed combining search strings on NOAC, VKA and atrial fibrillation (see appendix 1 for the search strings). The search was conducted on 23-04-2019 and we checked all articles published in English language. The title and abstract selection was done in duplicate by two independent researchers.

The following inclusion criteria were used:

- Population: patients with NVAF
- Intervention: NOAC (dabigatran, rivaroxaban, apixaban or edoxaban)
- Comparator: other NOAC(s) (dabigatran, rivaroxaban, apixaban and/or edoxaban)
- Outcomes: effectiveness and safety
- Study type: comparative effectiveness studies with a cohort design

The following exclusion criteria were applied:

- Studies on only one NOAC
- Studies in which VKA is the comparator for the NOACs, and NOACs are not compared against each other
- · Studies on cost-effectiveness and healthcare resources use
- Studies on adherence or persistence

Critical appraisal

We checked the setting, in- and exclusion criteria and the following baseline characteristics: age, proportion males, CHA2DS2Vasc score and comorbidity index.

We used the criteria suggested by ISPOR, Yao et al., and Austin et al. as a guidance to critically appraise the articles in which PS were used.[12-15] The criteria we checked concerned:

- The variables included in the propensity score model
- Explanation of the variable selection procedure for propensity score model
- Distribution of baseline characteristics for each group before propensity score analysis
- In case of PSM:
 - o matching ratio,
 - distance metric,

- with or without replacement,
- comparability of baseline characteristics in the matched groups,
- sample size before and after matching
- In case of IPTW:
 - comparability of baseline characteristics in the weighted groups
 - o extreme weights
- In case of PS stratification:
 - o number of strata, comparability of baseline characteristics
- In case of analyses in which no PS was used in the main analyses:
 - we evaluated whether the ratio number of covariates to the number of events seemed sufficient to produce valid results.[8]
- Sensitivity analyses to further explore the magnitude of residual confounding (i.e. case-crossover study designs; clinical details in a subsample; proxy measures; or instrumental variable (IV) techniques)

Naïve trial analysis

Trials are quite often designed with a null hypothesis and associated with a power calculation while real world studies are often dictated by the number of observations available. To give the results from the real-world-evidence some perspective we undertook a naïve trial analysis in which the risk reductions from each trial with respect to efficacy and safety outcomes were applied to an average number of outcomes observed in the warfarin arms in each trial. This leads to an estimate of the relevant rates for each drug and the differences are illustrated by the number of patients (sample size) needed in a randomised clinical trial to confirm the estimated differences.

RESULTS

In total, we found 1302 unique articles in our search, of which 39 articles fulfilled the in- and exclusion criteria and were included for data extraction, see figure 1. In table 1 to 5, study characteristics are presented. The most important differences between the studies are outlined in table 6.



Table 1. Characteristics of the included articles that used propensity score matching (PSM) as primary analyses (n=18)

Author and country	Setting and study period	Study population	Patient characteristics before PSM (range between NOACs)	Primary outcome definition	PSM details	Sample size before matching	Sample size after matching	Results/conclusion as reported in the article
Abraham, 2017 USA	OptumLabs Data Warehouse Oct 1, 2010 through Feb 28, 2015	NVAF patients, 18 years of age or older, identified by their index prescription of a NOAC during study period (excluded if NOAC prescribed during 12 months before index date). No reporting on earlier VKA use	Male: 54.0-60.5% CHA2DS2-VASc: 3.2-4.0	Gastrointestinal bleeding: definition by Lewis et al. 2002 using inpatient hospital claims for relevant primary and secondary discharge diagnoses.	3 matched cohorts; 1:1 PSM without replacement and with a calliper of 0.01; After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=19,301 dabigatran: n=17,426 apixaban: n=6,576	rivaroxaban vs. dabigatran: n=31,574 apixaban vs. rivaroxaban: n=13,130 apixaban vs. dabigatran: n=13,084 (more than 90% of original smallest samples size)	Apixaban had the most favorable gastrointestinal bleeding profile and rivaroxaban had the least favorable safety profile. Apixaban had the most favorable gastrointestinal safety profile among all age groups.
Amin, 2018 (J Manag Care Spec Pharm) USA	Medicare & Medicaid Services Jan 1, 2012, to Dec 31, 2014	NVAF patients of at least 65 years old, OAC treatment- naïve, ≥ 1 prescription claim for OAC during study period. Excluded if OAC pharmacy claim during the 12-month before study start.	Age: 77.2±7.0-78.4±7.4 Male: 47.4-50.6% CHA2DS2-VASc: 4.4-4.6 CCI: 2.5-2.7	Hospitalization for stroke, systemic embolism and major bleeding: ICD-9- code as primary discharge diagnosis	2 matched cohorts; 1:1 PSM with nearest neighbour without replacement and with a calliper of 0.01; After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=53,146 apixaban: n=20,853 dabigatran: n=16,743	rivaroxaban vs. apixaban: n=41,608 dabigatran vs. apixaban: n=30,836 (more than 90% of original smallest samples size)	Apixaban was associated with significantly lower risks of all-cause, stroke/SE-related, and MB-related hospitalizations compared with dabigatran, and rivaroxaban
Amin 2018 (J Med Econ) USA	OptumInsight research database Jan 1, 2012 – Sept 30, 2015	NVAF patients of at least 18 years old, OAC treatment- naïve, ≥ 1 prescription claim for OAC during study period. Excluded if OAC pharmacy claim during the 12-month before study start.	NR	Hospitalization for stroke, systemic embolism and major bleeding: ICD-9- code as primary discharge diagnosis	2 matched cohorts; 1:1 PSM with nearest neighbour without replacement and with a calliper of 0.01; After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=14,163 apixaban: n=8,652 dabigatran: n=3,684	apixaban vs. rivaroxaban: N=16,880 apixaban vs. dabigatran: N=7,114 (more than 90% of original smallest samples size)	Rivaroxaban patients were associated with a significantly higher risk of all-cause and major bleeding related hospitalisations and dabigatran patients were associated with a significantly higher risk of major bleeding hospitalisation compared with apixaban
Andersen, 2018 Denmark	National patient register, Register of Medicinal Product Statistics July 1, 2013 – March 31, 2016	NVAF patients who were new users of NOAC aged 45 years of age or older, with a recent diagnosis of NVAF (received no OAC treatment in the 12 months before inclusion; 'recent diagnosis' is not defined)	Online material not available	Stroke, systemic embolism and major bleeding (i.e., intracranial bleeding, gastro-intestinal bleeding (bleeding ulcer, hematemesis ormelena) or other serious bleeding (anemia caused by bleeding, bleeding of unknown origin, bleeding of the respiratory or urinary tract, peritoneal, retinal or orbital bleeding); hospital admission with a primary or secondary	3 matched cohorts; 1:1 PSM with nearest neighbour with a caliper of 0. (replacement yes or no not reported) All baseline characteristics were well balanced after matching, except for calendar year	apixaban: n=4,292 dabigatran: n=3,913 rivaroxaban: n=3,805	apixaban vs. rivaroxaban: N=7,352 apixaban vs. dabigatran: N=6,470 rivaroxaban vs. dabigatran: N=5,440	There were no statistically significant differences in risk of stroke or systemic embolism or major bleeding in propensity-matched comparisons between apixaban, dabigatran, and rivaroxaban used in standard doses.

Author and country	Setting and study period	Study population	Patient characteristics before PSM (range between NOACs)	Primary outcome definition	PSM details	Sample size before matching	Sample size after matching	Results/conclusion as reported in the article
Blin, 2019 France	French nationwide claims and hospitalization database, Système National des Données de Santé	NVAF patients of at least 18 years old, all new users of standard or reduced doses of NOAC in (received no OAC treatment in the three years before the index date)	Age: 65.3±10.2-69.0±11.1 Male: 62.7-68.3% Modified CHA2DS2-VASc ≥2: 57.1-67.4% Comorbidities: NR	Hospitalization with a main diagnosis of ischemic stroke or systemic embolism or major bleeding and all-cause death. (ICD-10) codes	1 matched cohort PSM method not reported. After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=18,829 dabigatran: n=10,847	dabigatran vs. rivaroxaban: n=16,580	Dabigatran had similar or better effectiveness than rivaroxaban but lower bleeding risk. Death rates were not different.
Briasoulis, 2018 USA	Medicare and Medicaid Services Jan 1, 2010 - Dec 31, 2013	NVAF patients newly diagnosed of ≥65 years old and initiated OAC treatment during study period	Age: 75.4±6–75.5±6 Male: 50-53% CHA2DS2-VASc: 4.1-4.1 Gagne: 2.7-2.7	All-cause mortality, stroke, including ischemic stroke or transient ischemic attack, gastrointestinal bleeding, any bleeding, non-gastrointestinal bleeding, acute myocardial infarction. ICD-9-CM reported in inpatient claims, whether primary and secondary codes were used is not described	1 matched cohort; 3-way propensity matching (VKA was one of the groups, but not further discussed here) After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=14,257 dabigatran: n=13,522	dabigatran vs. rivaroxaban: n =26,814	Rivaroxaban was associated with higher gastrointestinal bleeding rates than dabigatran
Deitelzweig, 2017 USA	Humana Research Database (Medicare coverage) Jan 2013 - 30 Sept 2015	NVAF patients age of ≥65 years, OAC treatment naïve (excluded if they had a pharmacy claim for OAC during the baseline period, which was 12 months before index date)	Age: 76.8±8.3-78.0±9.0 Male: 51.5-55.1% CHA2DS2-VASc: 4.3-4.6 CCI: 2.7-3.0	Hospitalisation claims of stroke, systemic embolism and major bleeding: ICD-9-code as primary discharge diagnosis	2 matched cohorts; 1:1 PSM with nearest neighbour (replacement yes or no and calliper not reported) balanced with key patient characteristics not statistically different (p>.05).	rivaroxaban: n=11,082 apixaban: n=8,250 dabigatran: n=2,474	apixaban vs. rivaroxaban: n=13,620 apixaban vs. dabigatran: n= 4,654	Apixaban is associated with significantly lower risk of stroke/systemic embolism and major bleeding than rivaroxaban, and a trend towards better outcomes vs. dabigatran.
Gupta, 2018 USA	Department of Defence data Jan 1, 2012, to Sept 30, 2015	NVAF patients, treatment-naïve (excluded if a pharmacy claim for an OAC during the baseline period)	NR	Inpatient claim of stroke, systemic embolism or major bleeding as primary or secondary diagnosis based on validated administrative claims-based algorithms	2 matched cohorts; 1:1 PSM with nearest neighbour without replacement with a calliper of 0.01 After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=15,680 apixaban: n=11,754 dabigatran: n=4,312	rivaroxaban vs. apixaban: n=22,568 dabigatran vs. apixaban: n=8,258	Rivaroxaban was associated with a significantly higher risk of stroke/systemic embolism and major bleeding compared with apixaban. Dabigatran use was associated with a numerically higher risk of stroke/systemic embolism and a significantly higher risk of major bleeding compared with apixaban

Author and country	Setting and study period	Study population	Patient characteristics before PSM (range between NOACs)	Primary outcome definition	PSM details	Sample size before matching	Sample size after matching	Results/conclusion as reported in the article
Lai, 2017 Taiwan	National Health Insurance program 2011 to 2014	NVAF and flutter patients, ≥20 years, new-users (new users not further defined).	Age: 75.1±9.7-75.4±9.6 Male: 54-7-56.7% CHA2DS2-VASc: 3.3-3.3 Comorbidity index: NR	All-cause death	1 matched cohort; 1:1 PSM with calliper < 0.2 (neighbour and replacement not reported) Balance checked with p-values and standardized difference	dabigatran: n=10,625; rivaroxaban: n=4,609	dabigatran vs. rivaroxaban: N=9,200	Rivaroxaban therapy was associated with a statistically significant increase in all-cause death compared with dabigatran
Lin, 2017 USA	IMS Pharmetrics Plus database Jan 2013 – Sept 2015	NVAF patients of at least 18 years old who initiated OAC (received no OAC treatment received 12 months before the index date)	NR	Major bleeding first listed in ICD-9 diagnosis or procedure codes	2 cohorts; 1:1 PSM with nearest neighbour (replacement and calliper not reported) Patient key characteristic being similar with p>0.05	NR	apixaban vs. rivaroxaban: N=8,124 apixaban vs. dabigatran: N=5,368	Apixaban is associated with reduced risk of hospitalisation compared with dabigatran and rivaroxaban.
Lip, 2016 (Thromb Haemost) USA	Truven MarketScan® Commercial Claims and Encounter and Medicare Supplemental and Coordination of Benefits Databases Jan 2012 to Dec 2014	NVAF patients ≥18 years who newly initiated OACs (patients with a prescription claim for OAC prior to the index date were excluded)	Age: 66.5±12.4- 68.5±12.4 Male: 61.4-65.0% CHA2DS2-VASc: 2.6-2.8 CDI: 1.6-1.8	Major bleeding listed first primary ICD-9 code	3 cohorts; 1: 1 PSM with nearest neighbour without replacement with a maximum calliper of 0.01. After PSM, standardized differences of all baseline characteristics were <10%.	rivaroxaban: n=17,801 apixaban: n=7,438 dabigatran: n=4,661	apixaban vs dabigatran: n=14,798 rivaroxaban vs dabigatran: n=9,314 apixaban vs rivaroxaban: n=8,814	Compared to apixaban, rivaroxaban initiation was associated with significantly higher risk of major bleeding. The difference for dabigatran was not statistically significant
Lip, 2018 USA	Medicare and Medicaid Services Medicare; Truven MarketScan, IMS PharMetrics Plus Database, Optum Clinformatics Data Mart, and the Humana Research Database Jan 1, 2013, to Sept 30, 2015	NVAF patients newly prescribed OAC, (received no OAC treatment in the 12 months before the index date)	Age: 71.4±11.4-73.1±11.6 Male: 55.0-59.6% CHA2DS2-VASc: 3.3-3.6 CDI: 2.4-2.8	Hospitalizations with stroke, systemic embolism or major bleeding as the principal or first-listed diagnosis	3 cohorts 1:1 PSM with nearest neighbour without replacement with a maximum calliper of 0.01 After PSM, standardized differences of all baseline characteristics were <10%	rivaroxaban: n= 103,477 apixaban: n= 63,484 dabigatran: n= 27,571	apixaban-rivaroxaban: n=125,238 dabigatran-rivaroxaban: n=55,076 apixaban-dabigatran: n=54,192	Apixaban was associated with a lower rate of stroke/systemic embolism and major bleeding compared with dabigatran and rivaroxaban. Dabigatran was associated with a lower rate of major bleeding compared with rivaroxaban, with similar rates of stroke/systemic embolism.

Author and country	Setting and study period	Study population	Patient characteristics before PSM (range between NOACs)	Primary outcome definition	PSM details	Sample size before matching	Sample size after matching	Results/conclusion as reported in the article
Lutsey, 2018 USA	MarketScan Commercial Database Jan 1, 2010 through Sept 30, 2015	NVAF patients aged 45 and older with at least one prescription for OAC after their first AF claim (de novo patients or first initiation of treatment)	Age: 69.1±11.4-69.9 ± 11.7 Male: 59.4-63.7 CHA2DS2-VASC: 3.3-3.6 Comorbidity index: NR	venous thromboembolism: at least one inpatient ICD 9 claim (first listed or not is not specified)	3 cohorts 1:1 PSM with a maximum calliper of 0.25 (neighbour and replacement not reported) Balance not described	rivaroxaban: n=31,119 dabigatran: n=28,089 apixaban: n=17,112	rivaroxaban vs. apixaban: n=32,468 dabigatran vs. rivaroxaban: n=21,160 dabigatran vs. apixaban: n=6,200	Risk of VTE was lowest among those prescribed apixaban and dabigatran
Mentias, 2018 USA	Medicare & Medicaid Services Jan 1, 2010, to Dec 31, 2013	NVAF patients, newly diagnosed who initiated an OAC within 90 days of diagnosis	Age: 75.8±6.4-75.8±6.4 Male: 48.9-50.1% CHA2DS2-VASc: 4.3-4.3 Gagne: 3.0-3.0	inpatient admission for acute ischemic stroke or major bleeding as defined by Rothendler* and Suh based on the primary ICD-9-CM diagnosis on inpatient standard analytical files claims for acute care stays.	1 cohort 3-way PSM. (VKA was one of the groups, but not further discussed here) After PSM, standardized differences of all baseline characteristics were <10%	rivaroxaban: n=23,177 dabigatran: n=21,979	NR	Rivaroxaban users had significantly higher major bleeding risk compared with dabigatran users in the medium and high comorbidity groups
Norby, 2017 USA	Truven Health MarketScan Commercial Claims and Encounters Database and the Medicare Supplemental and Coordination of Benefits Database	NVAF patients with at least one prescription of NOAC after their first AF claim (first prescription of OAC)		ischemic stroke (primary discharge), intracranial bleeding (primary discharge), myocardial infarction (1st or 2nd position of an inpatient discharge diagnosis, and gastrointestinal bleeding (primary and secondary diagnoses, presence of transfusion codes, and presence/absence of trauma codes to exclude trauma-related bleeding based on ICD-9 codes	1 cohort; 1:1 PSM, greedy matching technique with a calliper of 0.25	NR NR	rivaroxaban vs dabigatran: n=16,957	Endpoint rates were similar when comparing anticoagulant-naïve rivaroxaban and dabigatrar initiators, with the exceptior of higher gastrointestinal bleeding risk in rivaroxabar users
Noseworthy, 2016 USA	Optum Labs Data Warehouse Oct 1, 2010 - Feb 28, 2015	NVAF patients ≥ 18 years, who were OAC users during study period.	NR	inpatient admission for stroke or systemic embolism or major bleeding (ICD-9 codes in the primary or secondary diagnosis positions of inpatient claims)	3 cohorts; 1:1 PSM without replacement and with a caliper of 0.01. A standardized difference < 10% was considered acceptable	NR	rivaroxaban vs. dabigatran: n=31,574 apixaban vs. rivaroxaban: n=13,130 apixaban vs. dabigatran: n=13,084	Dabigatran, rivaroxaban, and apixaban appear to have similar effectiveness, although apixaban may be associated with a lower bleeding risk and rivaroxaban may be associated with an elevated bleeding risk
Shantha, 2017 USA	Medicare and Medicaid Nov 1, 2011 - Dec 31, 2013	Newly diagnosed NVAF patients and initiated OAC use.	Males: Age: 74.7±5.9-74.9±6. CHADS2-Vasc: 3.7-3.8 Gagne score: 2.9-2.9 Women: Age: 76.6±6.6-76.9±6.6 CHADS2-Vasc: 4.8-4.9	inpatient admissions for acute ischemic stroke or major bleeding (primary ICD-9-CM diagnosis on inpatient standard analytical files claims for acute care stays)	1 cohort; Three-way PSM (VKA was one of the groups, but not further discussed here)	rivaroxaban: n=23,177 dabigatran: n=21,979	dabigatran vs. rivaroxaban: n= 37,298	The reduced risk of ischemic stroke in patients taking rivaroxaban, compared with dabigatran, seems to be limited to men whereas the higher risk of

Author and country	Setting and study period	Study population	Patient characteristics before PSM (range between NOACs)	Primary outcome definition	PSM details	Sample size before matching	Sample size after matching	Results/conclusion as reported in the article
			Gagne: 3.0-3.1		A standardized difference < 10% was considered acceptable			bleeding seems to be limited to women
Villines, 2019 USA	US Department of Defence Military Health System database 1 July 2010 to 30 June 2016 for the dabigatran vs. rivaroxaban cohort, and 28 Dec 2011 to 30 June 2016 for the dabigatran vs. apixaban cohort	NVAF patients ≥18 years newly initiated on standard-dose NOAC (first initiation of treatment, AF diagnosis in the 12 months before the index date or on the index date)	Age (mean): 70.9-71.3 Male: 60-62% CHA2DS2-VASc: 3.1-3.1 CCI score: 4.3-4.3	Stroke or major bleeding, ICD-9 or 10 codes, whether primary and secondary codes were used is not described	2 cohorts 1:1 PSM nearest neighbour with a calliper of 0.20 (replacement not reported). Balanced if the absolute value of the STD was ≤10%.	NR	dabigatran vs. rivaroxaban: n=25,526 dabigatran vs. apixaban: n=9,604	Dabigatran was associated with significantly lower major bleeding risk vs. rivaroxaban, and no significant difference in stroke risk. For dabigatran vs. apixaban, the reduced sample size limited the ability to draw definitive conclusions.

Age: mean, SD unless stated otherwise; CCI: Charlson comorbidity index; CDI: Charlson-Deyo index; Gagne: Gagne comorbidity score;

Table 2. Characteristics of the included articles that used inverse probability of treatment weighting as primary analyses (n=8)

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	IPWT details	Sample size	Result/conclusion as reported in the article
Adeboyeje, 2017 USA	HealthCore integrated research environment Nov 1, 2009 - Jan 31, 2016	NVAF patients newly prescribed OAC (no prescriptions for any anticoagulant in the 6-month period preceding their index dates).	Age (mean): 66-69 Male: 59.1-65.5% CHA2DS2-VASc: 2.7-3.2 Comorbidity index: NR	Hospitalization for major bleeding (ICD 9-CM codes; whether primary and secondary codes were used is not described	Extreme weights: not reported. Balanced if the absolute value of the STD was ≤10%.	dabigatran: n=8,539 rivaroxaban: n=8,398 apixaban: n=3,689	Apixaban and dabigatran were associated with lower major bleeding risk compared with rivaroxaban; however, apixaban had a lower risk of major gastrointestinal bleeding than dabigatran.
Chan, 2018 Taiwan	Taiwan National Health Insurance Research June 1, 2012 - Dec 31, 2016	NVAF patients with their first prescription of OAC	Age: 75±10-76±10 Male: 55-60% CHA2DS2-VASc: 3.7-3.9 Comorbidity index: NR	Hospitalization for ischemic stroke/systemic embolism, intracranial hemorrhage, major gastrointestinal bleeding, acute myocardial infarction, all major bleeding events, and all-cause mortality. ICD 9 and 10 codes, whether primary and secondary codes were used is not described	Extreme weights: not reported. Balanced if the absolute value of the STD was ≤10%.	rivaroxaban: n=27,777 dabigatran: n=20,079 apixaban: n=5,843	Three low-dose NOACs showed similar performance as without subgrouping
Charlton, 2018 USA	HealthCore Integrated Research Evironment database Nov 1, 2010 - March 31, 2014	NVAF patients hospitalized for bleeding after starting OAC (AF diagnosis 6 months before starting one of the index drugs).	Age: 68.0±12.5-69.6±12.6 Male: 61.8-62.9 CHA2DS2-VASc: 3.8 -3.8 CDI: 2.0-2.3	Total length of hospital stay, proportion of patients admitted to the ICU, mean length of ICU stay, and all-cause 30- and 90-day mortality, ICD 9 codes, whether primary and secondary codes were used is not described.	Extreme weights: not reported. Balance was tested using ANOVAs for significant differences	dabigatran: n=442 rivaroxaban n=256	There were no significant differences in relative risk of all-cause 30- or 90-day
Graham, 2016 USA	Medicare Nov 4, 2011 - June 30, 2014	NVAF patients, at least 65 years old, initiating OAC at standard doses (first treatment, received no NOAC treatment for other indications in the last 6 months before the index date)	Age: 65-74 y: 50-51% Age: 75-84: 40-40% Age ≥85: 9-10% Male: 53-53% CHADS2 ≥2: 66-67% Comorbidity index: NR	Thromboembolic stroke, ICH, major extracranial bleeding events and mortality (as the first study outcome or within 30 days after hospitalization for another primary outcome event), ICD 9 codes, whether primary and secondary codes were used is not described.	Extreme weights: not reported Balanced if the absolute value of the STD was ≤10%.	rivaroxaban: n=66,651 dabigatran: n=52,240 Weighted cohorts rivaroxaban: n=66,630 dabigatran: n=52,264	Treatment with rivaroxaban was associated with statistically significant increases in intracranial bleeding and major extracranial bleeding, including major gastrointestinal bleeding, compared with dabigatran

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	IPWT details	Sample size	Result/conclusion as reported in the article
Graham, 2019 USA	fee-for-service Medicare Part A (hospitalization), Part B (office-based care), and Part D (prescription drug coverage) Oct, 2010 - Sept, 2015	NVAF patients of ≥65 years old (first initiation of treatment)	Age (mean): 74.9-75.5 Male: 52.2-59.3% CHA2DS2-VASc ≥2: 96.6- 97.4% Comorbidity index: NR	Hospitalized due to thromboembolic stroke, intracranial haemorrhage, major extracranial bleeding, and all-cause mortality. ICD codes from the first hospital discharge diagnosis position	Not described how weighted cohort was composed. Balanced if the absolute value of the STD was ≤10%.	rivaroxaban: n=106,389 dabigatran: n=86,198 apixaban: n=73,039 Weighted cohort rivaroxaban: n=106,369 dabigatran: n=86,293 apixaban: n=72,921	Dabigatran and apixaban were associated with a more favourable benefit- harm profile than rivaroxaban.
Hernandez, 2017 USA	Medicare Nov 4, 2011 -Dec 31, 2013	NVAF patients (at any time before the index date; no NOAC treatment at least 3 months before the index date)	High dose: Age: <65: 5.0-6.3% Age: 65-74: 38.4-39.3% Age: ≥75: 55.3-55.7 Male: 45.9-49.5 CHADS2: 3.3-3.3 Comorbidity index: NR	ischemic stroke (inpatient, emergency room, or outpatient claim with primary or secondary, ICD-9 codes), other thromboembolic events, and all-cause mortality; ICD 9 codes, whether primary and secondary codes were used is not described. Any bleeding event and major bleeding; intracranial hemorrhage and gastrointestinal bleeding, not further described.	Extreme weights: not reported Balanced if the absolute value of the STD was ≤10%.	dabigatran n=9,138 rivaroxaban n=8,367	There was no difference in stroke prevention between rivaroxaban and dabigatran; however, rivaroxaban was associated with a higher risk of thromboembolic events other than stroke, death, and bleeding.
Larsen, 2016 Denmark	Danish national prescription registry, Danish national patient register, Danish civil registration system August, 2011 -Oct, 2015	NVAF patients who were naïve to oral anticoagulants (no use of oral anticoagulant within one year)	Age (median, IQR): 67.6 (62.0-72.4)-71.8 (65.7-78.9) Male: 56.9-66.1% CHA2DS2VASc: 2.2-2.8 Comorbidity index: NR	Ischaemic stroke or systemic embolism, ICD-10 codes whether primary and secondary codes were used is not described.	Extreme weights: not reported Balanced if the absolute value of the STD was ≤10%.	dabigatran: n=12,701 rivaroxaban: n=7,192 apixaban: n=6,349	Apixaban and dabigatran were associated with a significantly lower risk of death compared with rivaroxaban. Risk of any bleeding or major bleeding were significantly lower for apixaban and dabigatran than for rivaroxaban
Meng, 2019 Taiwan	National Health Insurance claims database June 1, 2012 - May 31, 2015	All NVAF patients aged ≥20 years who initiated NOACs during study period	Age <65: 11.8-13.5% Age 65-74: 29.7-32.7% Age ≥75: 53.8-58.4% Male: 54.6-56.2% CHA2DS2-VASc: 3.2-3.3 Comorbidity index: NR	all-cause death, ischemic stroke, intracranial hemorrhage, gastrointestinal hemorrhage needing transfusion, ICD-10 codes, whether primary and	Extreme weights: not reported Balanced if the absolute value of the STD was ≤10%.	dabigatran: n=13,505 rivaroxaban: n=6,551 Weighted pseudo- cohort	Rivaroxaban seemed to be associated with an increased risk of all- cause death compared with dabigatran

Author and country	Setting and study period	Patient characteristics (range between NOACs)	Primary outcome definition	IPWT details		Result/conclusion as reported in the article
			secondary codes were used is not described		dabigatran: n=13,508; rivaroxaban: n=6,547	



Table 3. Characteristics of the included articles that used adjusted Cox-proportional hazard models as primary analyses (n=10)

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	Sample size	Results/conclusion as reported in the article
Al-Khahili, 2016 Sweden	tertiary referral cardiology outpatient clinic (the Stockholm Heart Center) Dec, 2011 - May, 2014	NVAF patients from a single cardiology outpatient clinic incorporating the AF unit (initiate NOAC treatment)	Age: 72±8-73±8 Male: 50-51% CHA2DS2-VASc: 3-3 Comorbidity index: NR	Major bleeding was defined according to the criteria of the International Society of Thrombosis and Hemostasis	rivaroxaban: n=282; apixaban: n=251 dabigatran: n=233;	Rivaroxaban was associated with the highest bleeding rates owing mainly to the highest number of minor bleedings, and apixaban had the lowest bleeding rates and side effects
Alonso, 2017 USA	Truven Health MarketScan® Commercial Claims and Encounter Database and the Medicare Supplemental and Coordination of Benefits Database Jan 1, 2007 - Dec 31, 2014	NVAF patients with a first prescription of OAC after Nov 2, 2011.	Age: 67.2±12.4- 69.3±12.5 Male: 60.1-65.1% CHA2DS2-VASc: 2.9-3.6 Comorbidity index: NR	Hospitalization for liver injury potentially related to drug hepatotoxicity, ICD-9-CM codes in any position	rivaroxaban: n=30,347; dabigatran: n=17,286; apixaban: n=9,205	Risk of liver disease hospitalization was higher in rivaroxaban users compared to dabigatran and apixaban users
Chan, 2016 Taiwan	Taiwan National Health Insurance Research Database. Jan 1, 1996 - Dec 31, 2013	NVAF patients newly diagnosed	Age: 75±9-76±9 Male: 54-58 CHA2DS2-VASc: 4.1-4.1 Comorbidity index: NR	Ischemic stroke or systemic embolism, ICH, hospitalization for GI bleeding, acute myocardial infarction (AMI), all hospitalizations for bleeding, and all-cause mortality. All discharge diagnosis according to the ICD, whether primary and secondary codes were used is not described	dabigatran 110 mg: n= 5,921 rivaroxaban 10 mg: n=3,916	No differences were found between rivaroxaban and dabigatran in risk for thromboembolic events, intracranial haemorrhage, critical gastrointestinal bleeding, or all-cause mortality. However, rivaroxaban was associated with a higher risk for noncritical gastrointestinal bleeding than dabigatran

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	Sample size	Results/conclusion as reported in the article
Hernandez, 2017 USA	Medicare database Jan 1, 2013 - Dec 31, 2014	NVAF patients newly diagnosed	Age: 74.9±8.7-77.4±8.6 Male: 42.5-47.0% CHA2DS2-VASc: 4.3-4.7 Comorbidity index: NR	Ischemic stroke, death, bleeding events, gastrointestinal bleeding, treatment persistence. ICD-9 codes, whether primary and secondary codes were used is not described	rivaroxaban: n=5,139; apixaban: n=2,358; dabigatran: 1,415;	Apixaban had the most favourable effectiveness and safety profile
Lamberts, 2017 Denmark	Danish national patient registry, Danish national prescription registry, Danish civil personal registry up to December 31, 2015	NVAF patients ≥18 years, with newly prescribed OAC (no prescription at least 6 months before inclusion)	Age: 71.5±11.0-75.4±11.10 Male: 50.8-56.7% CHA2DS2-VASc: 2.7-3.2 Comorbidity index: NR	major bleeding events requiring hospitalisation, ICD-10 codes, whether primary and secondary codes were used is not described	dabigatran: n=15,413; apixaban: n=7,963; rivaroxaban: n=6,715;	Apixaban had a lower adjusted major bleeding risk compared with rivaroxaban and dabigatran
Lip, 2016 (Int J Clin Pract) USA	Truven MarketScan® Commercial & Medicare supplemental US database Jan 1, 2013 - Dec 31, 2013	NVAF patients ≥18 years with newly prescribed OAC (no OACs received at least 1 year before the start of the OAC treatment)	Age: 66.8±12.2-69.3±12.3 Male: 63.1-65.8% CHA2DS2–VASc: 2.6-2.8 CCl: 1.7-1.9	Major bleeding was identified using hospital claims, which had a bleeding diagnosis code as the first listed primary ICD-9 diagnosis code	rivaroxaban: n=10,050 dabigatran: n=4,173 apixaban: n=2,402	Initiation with rivaroxaban was associated with a significantly greater risk of major bleeding compared with initiation on apixaban. There was no significant difference in the risk of major bleeding among patients newly initiated on dabigatran compared with apixaban.
Mueller, 2019 Scotland	Prescribing Information System, the Scottish Morbidity Records/ Hospital Inpatients and Outpatient attendance datasets; National Records of Scotland Drug's approval date – Dec 2015	NVAF patients who initiated NOAC treatment	Age: 71.1±12.0-74.8±11.0 Male: 53.5-73.1% CHA2DS2-VASc: 2.5-3.0 CCI: 1.1-1.4	strokes, systemic embolism, death due to cardiovascular, pulmonary embolism, bleeding events, clinical endpoints, according to ICD-10 codes whether primary and secondary codes were used is not described	rivaroxaban: n=7,265 apixaban: n=6,200; dabigatran: n=1,112;	All NOACs were similarly effective in preventing strokes and systemic embolisms, while patients being treated with rivaroxaban exhibited the highest bleeding risks.
Staerk, 2018 Denmark	Danish national patient registry, Danish national prescription registry,	NVAF patients, first-time OAC users (no previous OAC use), between 30 and 100 years old	Standard dose: Age (median, IQR): 67(61, 71)-71(65, 78) Male: 55.4-63.7% CHA2DS2-VASc (median); 2-3	stroke/thromboembolism (TE), ischaemic stroke, major bleeding, intracranial bleeding and gastrointestinal bleeding, ICD-10 codes whether primary and	dabigatran: n=11,492 apixaban: n=11,064 rivaroxaban: n=8966	Rivaroxaban was associated with higher bleeding risk

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	Sample size	Results/conclusion as reported in the article
	Danish civil registration system March 1, 2012 - Dec 31, 2016		Comorbidity index: NR	secondary codes were used is not described		compared with dabigatran and apixaban and dabigatran was associated with lower intracranial bleeding risk compared with rivaroxaban and apixaban.
Tepper, 2018 USA	Truven MarketScan Commercial Claims and Encounter and Medicare Supplemental & Coordination of Benefits Early View Database Jan 1, 2013 - Oct 31, 2014	NVAF patients aged ≥18 years with new initiators of NOACs or switched from warfarin to a NOAC	Age: 68±12-70±12 Male: 65.3-62.7 CHA2DS2-VASc: 2.4-2.5 CCI: 1.6-1.8	Bleeding, ICD-9-CM codes, whether primary and secondary codes were used is not described	rivaroxaban: n=30,529 dabigatran: n=20,963 apixaban: n=8,785;;	Rivaroxaban appeared to have an increased risk of any bleeding, clinically relevant nonmajor bleeding, and major inpatient bleeding, compared to apixaban patients. There was no significant difference in any bleeding, clinically relevant non-major bleeding, or inpatient major bleeding risks between patients treated with dabigatran and apixaban.
Vinogradova 2018 UK	UK general practices contributing to QResearch or Clinical Practice Research Datalink 2011 - 2016	NVAF patients, new NOAC (received no OAC treatment in at least the last 12 months)	QResearch: Age: 74.7±10.7-76.5±10.9 Male: 51.8-58.0% CHA2DS2-VASc: NR Comorbidity index: NR	Major bleeding after entry to the study which led to a hospital admission or death, based on linked hospital or mortality records.	rivaroxaban: n= 16,547 apixaban: n= 10,601 dabigatran: n=5,537	Apixaban was associated with a lowe risk of major bleed thar rivaroxaban. Rivaroxaban was associated with a higher risk of intracranial bleed compared to apixaban rivaroxaban was associated with higher risks compared with apixaban for haematuria, all gastrointestinal bleed and upper gastrointestinal bleed.

Author and country	Setting and study period	Patient characteristics (range between NOACs)	Primary outcome definition	 Results/conclusion as reported in the article
				The risk of primary ischaemic stroke did not differ between any of the anticoagulants

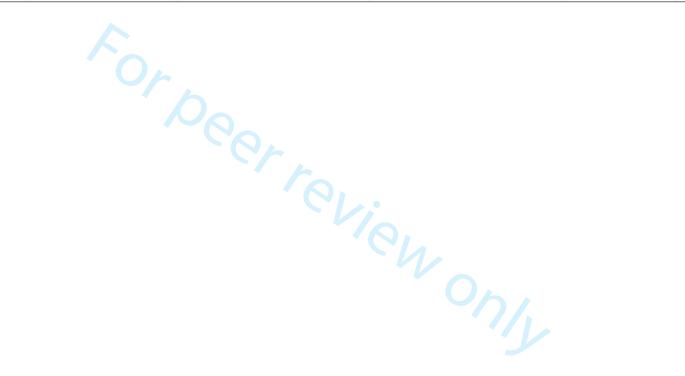


Table 4. Characteristics of the included articles that used unadjusted primary analysis (n=2)

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	Primary analysis	Sample size	Results/conclusio n as reported in the article
Cerda, 2019 Spain	Oral Anticoagulant Treatment Unit of the Hemostasis and Thrombosis Department of the University Hospital Vall d'Hebron from Barcelona (Spain) Jan, 2015 - Sept, 2017	NVAF patients with nonvalvular AF, with or without prior stroke, that had started treatment with any NOAC for the prevention of stroke	Age: 73.1±15.2-78.9±8.7 Male: 45.1-63.4% CHA2DS2-VASCc: 3.9-4.4 Comorbidity index: NR	Major bleeding according to ISTH 2005	log-rank test	rivaroxaban: n=663; dabigatran: n=352 apixaban: n=325 edoxaban: n=103	Rates of ischemic stroke and intracranial hemorrhage were similar among different NOACs, but rates of major bleeding were higher with dabigatran and apixaban and lower with rivaroxaban.
Li, 2017 China	Queen Mary Hospital, Hong Kong Jan, 2008 - Dec, 2014	NVAF patients diagnosed during study period.	Age: 71.9±11.1-73.3±12.1 Male: 53.1-59.8% CHA2DS2-VASc: 3.6-3.7 Comorbidity index: NR	The primary outcome was a composite of hospital admission with ischemic stroke or ICH, or death during the follow-up period. ICD-10 codes in medical records, and discharge summaries, whether primary and secondary codes were used is not described	Cox proportional hazard model (likely unadjusted, but this is not clearly described in the article)	rivaroxaban: n=669; dabigatran: n=467	Dabigatran had a lower ischemic stroke risk compared with patients on rivaroxaban. There was no significant difference in ischemic stroke risk between those on rivaroxaban and dabigatran.

Table 5. Characteristics of the included articles that used propensity score stratification as primary analyses (n=1)

Author and country	Setting and study period	Study population	Patient characteristics (range between NOACs)	Primary outcome definition	PS details	Sample size	Results/conclusion as reported in the article	
Gorst- Rasmussen, 2016 Denmark	Danish national prescription registry, Danish national patient register, Danish civil registration system Feb. 1, 2012 - July 31, 2014	NVAF patients who were new-users of OAC (no OAC treatment in at least the last two years)	CHA2DS2-Vasc: 2.1-3.0 Comorbidity index: NR	ischemic stroke/systemic embolism/transient ischemic attack, any bleeding and all-cause death. ICD-10 codes, whether primary and secondary codes were used is not described	Asymmetric trimming of the propensity score. Trimmed propensity score was used in 10 deciles as strata Balanced if the absolute value of the STD was ≤10%.		Rivaroxaban and dabigatran had similar stroke rates. Bleeding and mortality rates were higher in rivaroxaban versus dabigatran.	

Table 6. Main differences between the included studies (n=39)

Study item	Range, total number of stud	dies, or description
Country	USA: Denmark: Taiwan: China: France: Scotland: Sweden: Spain: UK:	n=24 n=5 n=4 n=1 n=1 n=1 n=1 n=1
NOAC included in included studies	Dabigatran: Rivaroxaban: Apixaban: Edoxaban:	n=39 n=39 n=26 n=1
Most prescribed NOAC in included studies per country	Dabigatran: Rivaroxaban: Apixaban: Edoxaban: About equal*:	Denmark USA, UK, China, Scotland, and Taiwan In none of the included studies In none of the included studies France, Spain, Sweden
Baseline characteristics	Mean age, years: % males: Mean CHA2DS2-Vasc:	65-84 39-73 2.1-4.9
Primary study outcomes	- all-cause death, - myocardial infarction - venous thromboen Safety outcomes: - major bleeding,	nbolism. leeding (e.g. intracranial haemorrhage,
Statistical approaches	PS matching: IPTW: PS stratification: Cox PH regression model: Unadjusted analyses:	n=18 n=8 n=1 n=10 n=2
Sample size	N=698 - N=265,583	
Study results	included:	xaban, rivaroxaban and dabigatran were urable compared to dabigatran and le NOAC: n=13

^{*} about equal distribution between dabigatran, rivaroxaban and apixaban. Edoxaban is not included in these studies.

More than 50% of the studies were conducted in the USA (n=24),[16-39] five were conducted in Denmark,[40-44] four in Taiwan,[45-48] and one in France,[49] Sweden,[50] Scotland,[51] the UK,[52] Spain,[53] and China.[54] Dabigatran and rivaroxaban were included in all 39 studies, apixaban was included in 26 studies and edoxaban was included in 1 study. Next to these NOACs, VKA was included in 25 of these studies as one of the comparators. The results below focus on the NOAC to NOAC comparisons only.

In the studies that included apixaban, dabigatran and rivaroxaban, rivaroxaban was most dominantly used in the USA, UK, Scotland, and Taiwan, while dabigatran was the most prescribed NOAC in Denmark. In three other European studies the distribution was about equal between the three NOACs. In none of the included studies, apixaban was the most dominantly prescribed NOAC.

Setting

Most studies concerned patient registries, pharmacy or prescription databases and/or health insurance databases (n=39), while there were three clinical practice based studies.[50, 53, 54]

Study population

All studies included only NVAF patients. In seven studies, it was specifically described that patients were newly diagnosed with NVAF and initiated NOAC treatment during study period.[21, 27, 34, 37, 40, 45, 54] None of the other studies included prevalent users of (N)OAC, but included e.g. 'newly treated', 'initiating treatment', 'new users', 'first-time prescription' of NVAF patients who were prescribed (N)OAC. In some studies (N)OAC use in the past (between 3 months and 2 years before index date) was allowed, while this seemed not be allowed in some other studies, or it was not described.

Inclusion criteria

Five studies concerned elderly patients specifically (i.e. ≥65 years old),[19, 21, 23-25] two included adults ≥45 years old,[33, 40] and one study included patients between 30 and 100 years of age.[44] The other studies included all adults with atrial fibrillation (it was assumed that if no further age specification was provided, 'adults' meant that all >18 years old were included). In one study only patients who were hospitalised for bleeding after start with OAC treatment were included.[22] No other focus on a specific group of AF patients was found.

Exclusion criteria

NOAC use that could be related to other disorders, such as transient AF, major knee or hip surgery, venous thromboembolism or pulmonary embolism, were specifically described as exclusion criteria in most studies,

except in ten studies.[16, 27, 28, 33-35, 50, 52-54] In one study patients with liver injury before their first oral anticoagulant (OAC) prescription were specifically excluded.[18]

Baseline characteristics

Baseline characteristics of the NVAF patients differed between studies. Mean age ranged from 65-84 years between the studies. The percentage of males ranged from 39-73%, and the mean CHA2DS2-Vasc Score ranged from 2.1-4.9. Excluding the five studies that specifically focussed on an elderly population of ≥65 years old and the two additional studies that used the Medicare database (only patients of 65 years or older are in Medicare), the mean age ranged from 65-78 years old. Different measures were used to assess the comorbidity index: Charlson comorbidity index, Charlson-Deyo index and Gagne comorbidity score, while in 30 of the 43 studies no comorbidity index was presented.

Selection of covariates

Most studies (n=34) did not provide a rationale for the selection of covariates that were included in the PS model or in adjusted analysis. However, in one of the articles an extensive rationale and selection procedure of covariates that were included in the analysis was provided.[33] In three other studies, the authors selected covariates based on medical knowledge on risk factors with reference to earlier published studies.[31, 39, 52] In one other study it was reported that sociodemographic and clinical characteristics that were associated with treatment initiation and the risk of major bleeding were included in the model to adjust for differences across cohorts, without further explanation or reference.[30]

Definition primary study outcomes

Primary outcomes differed between the studies. Effectiveness outcomes included in the studies included stroke, systemic embolism, (or composite of stroke/systemic embolism), all-cause death, myocardial infarction, venous thromboembolism and safety outcomes included major bleeding, or a specific type of bleeding (e.g. intracranial haemorrhage, gastrointestinal bleeding etc.) and liver injury. In most studies, ICD-9 or ICD-10 codes were used, but whether this concerned a primary diagnosis only or whether it could be either a primary or a second diagnosis differed between the studies. In some studies it was not described whether the ICD codes referred to primary diagnosis only or to a primary or secondary diagnosis.

Statistical approaches to adjust for confounding (primary analysis)

In 18 studies, PS matching was done.[16, 19-21, 23, 26, 29, 30, 32-37, 39, 40, 47, 49] IPTW was used in eight studies.[17, 22, 24, 25, 28, 43, 46, 48] PS stratified analyses was done in one study.[41] In twelve studies, the primary analyses utilised a Cox PH regression model in which adjustment for confounding was done.[18, 27, 31,

38, 42, 44, 45, 50-52] Finally, in two studies no adjustment for differences in baseline characteristics was performed.[53, 54]

PS matching

Co-variates

Creatinine clearance was not included as a covariate in any of the 18 studies. All 18 studies took the following covariates into account: age, sex, CHA2DS2-VASc score and/or the individual comorbidities included in this score, HAS-BLED score and/or the individual conditions included in this score (except alcohol use in Lai et al. [47]), renal disease, and co-medication use such as antiplatelets. Some included other comorbidities, such as cancer, rheumatic disease, specific heart diseases, COPD, HIV, dementia, depression, neurological disorders, and/or a various list of co-medications as well.

Matching method

In one study the matching method was not described.[49] In two studies, the calliper used was not described.[23, 29] In seven studies 1:1 PS matching without replacement was used and a calliper of 0.01 was applied.[16, 19, 20, 26, 30, 32, 36] Five other studies also matched 1:1 without replacement but used another calliper: in three studies a calliper of 0.2 was used,[39, 40, 47] while two others used a calliper of <0.25.[33, 35] In three studies, three-way matching was used.[21, 34, 37]

Balance co-variates

In two studies it was not described how the balance between covariates was evaluated.[33, 35] In two studies the balance was evaluated using p<0.05 (of which one also used standardized difference of <10%),[23, 47] and in another study it was stated that the groups were comparable even though a p value of >0.05 was found.[29] Balance was checked with an absolute standardized difference of <10% in 13 studies.[16, 19-21, 26, 30, 32, 34, 36, 37, 39, 40, 47, 49] Balance was reached in all studies after matching.

Sample size

In four studies the sample size before matching was not reported[29, 35, 36, 39] and in one study the sample size after matching was not reported.[34] At study start (before PSM), sample size between the NOACs differed greatly, except in three studies.[21, 37, 40]

IPTW

In one study, balance was tested using ANOVAs for significant differences.[22] Balance was checked with an absolute standardized difference of <10% in the other nine studies.[17, 24, 25, 28, 43, 46, 48] Balance was reached in all studies after IPTW.

There was no reporting on extreme weights in the eight included studies.[17, 22, 24, 25, 28, 43, 46, 48]

PS stratification

In one study, asymmetric trimming of the PS was done, which resulted in a small part of both treatment groups being removed in order to gain in comparability. Balance in co-variates was reached with standardized difference of <10%. In a Cox model this trimmed PS was used in 10 deciles as strata.[41]

Cox HP regression models

In ten studies, Cox HP regression models were applied with adjustment for a number of confounders.[18, 27, 31, 38, 42, 44, 45, 50-52] In one of these studies, the number of events per variable was not sufficient for such an analyses.[50] The ratio was acceptable in the other studies for at least some of the outcomes.[18, 28, 31, 38, 42, 44, 45, 51, 52]

Unadjusted analysis

In two studies no adjustment for confounding factors seemed to have been done, even though significant differences between treatment groups existed at baseline. Cerda et al. presented events per 100 patient-years and used a log-rank test to determine whether outcomes differed between the NOACs.[53] Li et al. conducted a Cox proportional hazard model, likely unadjusted, but this was not clearly described in the article.[54]

Sensitivity analyses

Although in some articles sensitivity analyses were done, none of the included studies further explored the magnitude of residual confounding in their sensitivity analyses using one of the approaches recommended by IPSOR (see methods section).

Study results

Which NOAC performed best differed between the included studies. We found only one study that included all four NOACs, in which no preference for one specific NOAC was found, except that rates of major bleeding were lower with rivaroxaban.[53] Of the 26 studies in which apixaban, rivaroxaban and dabigatran were included, apixaban was favourable compared to dabigatran and rivaroxaban in 13 studies, of which 10 were from the USA, two from Europe and one from Asia,[16, 17, 19, 20, 23, 26, 28, 29, 32, 36, 42, 50, 52] while dabigatran and

rivaroxaban were not found to be the single most favourable NOAC in any of the remaining 13 studies. Results for these 13 studies were mixed, with either no favourable NOAC at all, or one NOAC was selected as the least favourable, while the other two NOACs did not differ.

Naïve trial analysis

The primary efficacy endpoint (Strokes/SE) in the warfarin arms were estimated at 1.69% (RE-LY),[3] 2.2% (ROCKET),[6] 1.60% (ARISTOTLE),[5] and 1.50% (ENGAGE),[4] see table 7. From this range we chose a relatively arbitrary base rate of 1.6% and applied the observed risk reduction to estimate comparable base rates of 1.05% for dabigatran, 1.24% for rivaroxaban, 1.26% for edoxaban and 1.27% for apixaban. Using the sample size calculator[55] the biggest expected difference was between dabigatran and apixaban and it was estimated that a trial sample size with 51,847 patients would be needed to confirm this difference. The smallest difference was between edoxaban and apixaban and a trial of 7,994,340 patients required to confirm that difference.

Table 7. Primary efficacy and safety endpoints of the four pivotal trials.

		RE-LY [3]		ROCKET-AF [6]		ARISTOTLE [5]		ENGAGE-AF [4]		
	Dabigatr	Dabigatr	Warfari	Rivaroxab	Warfari	Apixaba	Warfari	Edoxab	Edoxab	Warfari
	an 150 mg			an	n	n	n	an 60 mg	an 30 mg	n
	N=6076	N=6015	N=602 2	N=7131 N=713		N=9120 N=908 1		N=7035 N=7034		N=703 6
Stroke/S E (%/year)	1.11	1.53	1.69	1.7	2.2	1.27	1.60	1.18	1.61	1.50
Major bleeding (%/year)	3.11	2.71	3.36	3.6	3.4	2.13	3.09	2.75	1.61	3.43

The primary safety endpoint was major bleeding for RE-LY, ARISTOTLE, and ENGAGE AF and major bleeding plus clinically relevant non-major bleeding for ROCKET AF, but data on major bleeds only for ROCKET-AF are available as well. Major bleeds in the warfarin arms were estimated at 3.36% (RE-LY),[3] 3.4% (ROCKET),[6] 3.09% (ARISTOTLE),[5] and 3.43% (ENGAGE).[4] From this range we choose a relatively arbitrary base rate of 3.2% and applied the observed risk reduction to estimate comparable base rates of 2.21% for apixaban 2.57% for edoxaban, 2.96% for dabigatran and 3.29% for rivaroxaban. Using the sample size calculator,[55] the biggest expected difference was between rivaroxaban and apixaban and it was estimated that a trial with 7,196 patients would be needed to confirm this difference. A much smaller difference is between edoxaban and apixaban which would require a trial of 56,512 patients to confirm that difference.

DISCUSSION

In total, we found 39 studies directly comparing the effectiveness and/or safety of at least two NOACs in NVAF patients. Three studies can be considered to be of low quality due to insufficiently described methods and/or small sample size[50, 53, 54].

Even though the remaining studies could be considered of sufficiently quality based on the technical aspects of the studies, there are some issues that can hamper the generalisability of the results. These issues concern residual confounding, the use of a smaller or broader calliper, differences in baseline characteristics between studies, channelling bias and change in treatment paradigm, and the high number of patients needed.

Balance in baseline characteristics between NOACs was checked with p-values or a standardized difference of <10%. Balance was well at baseline in some studies, or was reached after PS matching or IWTP.[56] Even though some studies included over 40 covariates in their PS, in most studies it was not described how the covariates were selected. The ISPOR Good Research Practices for Retrospective Database Analysis recommends to include all factors that are theoretically related to outcome or treatment selection, even if the relation is weak or statistically non-significant. [14] Directed acyclic graphs might be helpful as well. [57] And even though balance was reached for all of these variables, one should keep in mind that balance between unmeasured or unmeasurable factors cannot be assumed.[14] Therefore, due to the lack of randomization, there is always a possibility of residual confounding. This possibility was acknowledged in all included studies, and all studies have largely the same missing covariates. Hardly any laboratory results and lifestyle information were included, such as body0mass index, smoking status and alcohol consumption, which are also risk factors for ischaemic stroke and bleeding events respectively. Creatinine clearance for instance, seems to be an important covariate as subgroup analyses from the pivotal trials suggest that renal clearance might be an effect modifier.[5, 58] Only in one study however, the authors were able to take renal clearance into account in the adjusted analyses.[50] Especially when prescription of a certain NOAC in daily practice is driven by creatinine clearance, not adjusting for this variable may lead to biased results. However, it is unknown what the magnitude and direction (i.e. will the differences in effectiveness and safety between NOACs be smaller or larger) of this potential bias due to lack of randomization would be. The magnitude of residual confounding was not further explored in the sensitivity of the included studies..

In general, a calliper of <0.2 of the standard deviation of the logit of the PS is considered to be 'optimal'.[59] About half of the included PS matching studies used a smaller calliper, namely of <0.1. This means that the matching is more precise in these studies, but the disadvantage is that possibly more patients cannot be matched to another patient due to this smaller allowed maximum differences, and thus will be excluded from the analysis. Excluding

patients from the analysis will limit the generalisability of the results to the total patient population, especially when the excluded patients differ from the included patients, e.g. on the baseline risk for stroke.

All included studies focused on NVAF patients only.. In eight studies, inclusion criteria regarding age were applied. Three of these will likely still cover the largest part of NOAC users as they set relatively broad age ranges. The other five focussed on an elderly population of NVAF patients of ≥65 years old. Besides applying specific inclusion criteria regarding age in some studies, these differences also depended on the specific registry or database that was used, e.g. Medicare is for people of 65 years old or older. Even though only five of the included studies focused on an elderly NVAF population, and the others applied broad age ranges, there were differences in mean age, proportion of males and mean CHA2DS2Vasc score between the studies, which can have an impact on the results and jeopardize the generalisability of the results.

Rivaroxaban was the most prescribed NOAC in almost all included studies from the USA. However, in the first quarter of 2017, apixaban was the most prescribed NOAC in NVAF in the USA (i.e. in 50% of new OAC prescriptions). Especially older patients, women, increased stroke or bleeding risk and having comorbidities was associated with prescription of apixaban versus other NOACs.[60] Rivaroxaban was also the most prescribed NOAC in the included studies from the UK and Scotland. Based on the CPRD, 56.5% of the OAC prescriptions concerned a NOAC, of which rivaroxaban was still described most often in 2015.[61] Dabigatran was described most often in the studies from Denmark. Haastrup et al. described that most AF patients that initiated NOAC received dabigatran between 2008 and 2016, but a trend was observed that per 1000 person-years the number of patients described dabigatran decreased and the number of patients receiving rivaroxaban and apixaban increased.[62] This shows that the treatment paradigm changed over time, and might still be changing, and this pattern differs between the USA, Europe and Asia. Channelling bias therefore likely occurs and might shift between the NOACs. Although in a few studies it was mentioned that selective prescriptions were noticed and that these might have changed over time, none of the included studies dealt with temporal trends in prescription patterns.

Our naïve analysis predicts that in terms of the primary efficacy outcome observational studies will need a relatively high number of patients to be able to demonstrate the differences between the NOACs and a small sample size will not allow robust comparison to be made.

The pattern of major bleeding events seen in the included observational studies, confirms the expectation from our naïve analysis of the pivotal clinical trials that rivaroxaban seems to have the least favourable safety profile among apixaban and dabigatran. The findings are not consistent to allow for a robust conclusion between

apixaban and dabigatran which confirms the need for a high number of patients, although a trend for a slight better safety profile of apixaban can be observed.

The requirement for a high number of patients to compare NOACs both in terms of efficacy and safety as predicted by the pivotal trial results is confirmed by the findings of the observational studies. This finding may support the claim that the differences between the NOACs are relatively small.

In the process of conducting systematic reviews it is inevitable that the review will never be completely up to date with the most recent published evidence. Even though our search ended in April 2019, recently published studies will have encountered the same issues as described above. Residual confounding and channelling bias cannot have been ruled out in newer publications. Ideally, head-to-head trials should be conducted to compare the efficacy/effectiveness and safety of the four NOACs to overcome the methodological issues in the comparative effectiveness studies. To our knowledge, one head to head trial including all four NOACs is currently running. This nationwide cluster randomized cross-over study aims to compare efficacy and safety of the four NOACs (clinicaltrials.gov; NCT03129490)

In conclusion, even though the larger part of these studies are conducted as well as possible considering what data are available, there are some important limitations regarding the generalisability of the study results especially given the relatively high patient number required for a meaningful comparison between NOACs. Most studies included all NVAF patients on NOAC available in the registry/database during the study period and did not apply further specific in- and exclusion criteria, but differences between studies regarding baseline characteristics existed. Mean age at study start and baseline risk for stroke (CHA2DS2-VASc score) differed between the studies. As channelling bias cannot be ruled out, the result of these studies might not be generalisable.

Furthermore, results from the PS studies are only applicable to the patients that were kept in the analyses as patients excluded from the analysis likely differ from the ones that were included in the analysis. The 1:1 matched cohorts depended on the sample size of the NOAC with the least number of patients and as a result many patients from the larger of the two NOAC groups were excluded as they could not be matched. In clinical practice, these limitations should be kept in mind when results of these studies are used to decide what NOAC should be prescribed for a certain patient. Given the small differences between efficacy and safety outcomes between NOACs, the element of patient preference should be taken into consideration, [63] as tailoring anti-coagulation treatment towards patient preferences can promote adherence to treatment.

FIGURE LEGEND

Figure 1: PRISMA flowchart



ACKNOWLEDGEMENTS

The authors would like to thank Pearl Gumbs for the initiation of this project and her input with regard to study design and interpretation of the data. Pearl Gumbs was working at Daiichi Sankyo Europe at that time.



COMPETING INTEREST

EB reports grants from Daiichi Sankyo, during the conduct of the study; grants from Daiichi Sankyo, outside the submitted work. BvH reports grants from Daiichi Sankyo, during the conduct of the study. SH reports personal fees from Aspen, personal fees from Bayer, personal fees from BMS/Pfizer, personal fees from Daiichi-Sankyo, personal fees from Portola, outside the submitted work. GS reports personal fees from Daiichi Sankyo Europe Gmbh, outside the submitted work. AC reports personal fees from Daiichi Sankyo Europe, during the conduct of the study; grants and personal fees from Bayer AG, personal fees from Boehringer Ingelheim, grants and personal fees from Bristol-Myers Squibb, grants and personal fees from Pfizer Limited, personal fees from Portola Pharmaceuticals Inc, personal fees from Janssen, personal fees from ONO Pharmaceuticals, from AbbVie, outside the submitted work. OFK.

AUTHORS' CONTRIBUTION

EB: Conceptualization (support); Methodology (equal); Writing - Original Draft Preparation; Writing - Review & Editing (equal). BvH: Conceptualization (support); Methodology (equal); Writing - Review & Editing (equal). SH: Conceptualization (support); Writing - Review & Editing (equal). GS: Conceptualization (support), Supervision; Writing - Review & Editing (equal). ATC: Conceptualization (lead); Writing - Review & Editing (equal)



DATA SHARING STATEMENT

No new data were generated or analysed in support of this research.



FUNDING

This work was supported by Daiichi Sankyo Europe. GRANT / AWARD NUMBER: N/A

PATIENT AND PUBLIC INVOLVEMENT

This research was done without patient involvement. Patients were not invited to comment on the study design and were not consulted to develop patient relevant outcomes or interpret the results. Patients were not invited to contribute to the writing or editing of this document for readability or accuracy.



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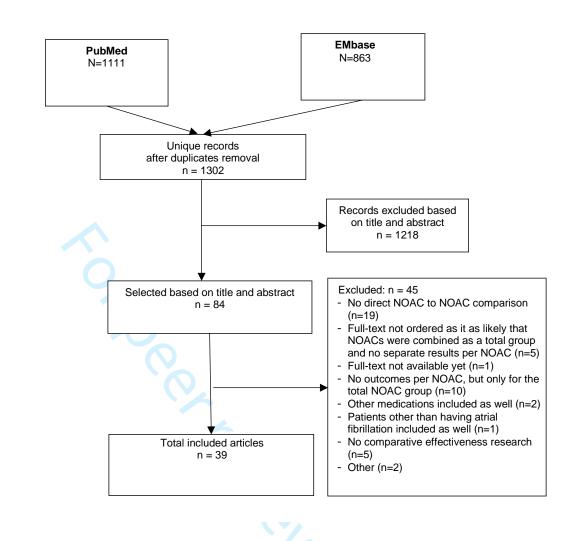
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Identification

Eligibility



The following search string was used for PubMed, and adapted for Cochrane and EMBase.

#1. NOAC

direct oral anticoagulant*[tiab] OR direct oral anti-coagulant*[tiab] OR direct oral anticoagulation[tiab] OR direct oral anticoagulation[tiab] OR direct oral anti-coagulation[tiab] OR direct-acting oral anti-coagulant*[tiab] OR direct-acting oral anti-coagulation[tiab] OR DOAC[tiab] OR novel oral anticoagulant*[tiab] OR novel oral anticoagulant*[tiab] OR Novel oral anticoagulation[tiab] OR Novel oral anti-coagulation[tiab] OR Novel oral anti-coagulation[tiab] OR Novel oral anticoagulation[tiab] OR Apixaban[tiab] OR Edoxaban[tiab] OR Dabigatran[tiab] OR "Non VKA Oral Anticoagulant"[tiab] OR "Non Vitamin K Antagonist Oral Anticoagulant"[tiab]

#2. Comparative effectiveness studies

comparative effectiveness research[mesh] OR comparative effectiveness[tiab] OR real-world[tiab] OR real-life[tiab] OR cohort studies[mesh] OR cohort[tiab]

#3. Atrial fibrillation

atrial fibrillation[tiab]

Limits:

Language: English



PRISMA 2009 Checklist

Section/topic	#	Checklist item	Reported on page #
TITLE			
Title	1	Identify the report as a systematic review, meta-analysis, or both.	1
ABSTRACT			
Structured summary	2	Provide a structured summary including, as applicable: background; objectives; data sources; study eligibility criteria, participants, and interventions; study appraisal and synthesis methods; results; limitations; conclusions and implications of key findings; systematic review registration number.	2
INTRODUCTION			
Rationale	3	Describe the rationale for the review in the context of what is already known.	3,4
Objectives	4	Provide an explicit statement of questions being addressed with reference to participants, interventions, comparisons, outcomes, and study design (PICOS).	4
METHODS			
Protocol and registration	5	Indicate if a review protocol exists, if and where it can be accessed (e.g., Web address), and, if available, provide registration information including registration number.	NA
Eligibility criteria	6	Specify study characteristics (e.g., PICOS, length of follow-up) and report characteristics (e.g., years considered, language, publication status) used as criteria for eligibility, giving rationale.	5,6
7 Information sources 8	7	Describe all information sources (e.g., databases with dates of coverage, contact with study authors to identify additional studies) in the search and date last searched.	5,6
Search	8	Present full electronic search strategy for at least one database, including any limits used, such that it could be repeated.	S1_Table
2 Study selection	9	State the process for selecting studies (i.e., screening, eligibility, included in systematic review, and, if applicable, included in the meta-analysis).	5
Data collection process	10	Describe method of data extraction from reports (e.g., piloted forms, independently, in duplicate) and any processes for obtaining and confirming data from investigators.	5
Data items	11	List and define all variables for which data were sought (e.g., PICOS, funding sources) and any assumptions and simplifications made.	5
Risk of bias in individual studies	12	Describe methods used for assessing risk of bias of individual studies (including specification of whether this was done at the study or outcome level), and how this information is to be used in any data synthesis.	5,6
Summary measures	13	State the principal summary measures (e.g., risk ratio, difference in means).	NA
Synthesis of results	14	Describe the methods of handling data and combining results of studies, if done, including measures of consistency (e.g., I²) for each meta-analysis. For peer review only - http://bmjopen.bmj.com/site/about/guidelines.xhtml	NA

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PRISMA 2009 Checklist

		Page 1 of 2	
Section/topic	#	Checklist item	Reported on page #
Risk of bias across studies	15	Specify any assessment of risk of bias that may affect the cumulative evidence (e.g., publication bias, selective reporting within studies).	NA
Additional analyses	16	Describe methods of additional analyses (e.g., sensitivity or subgroup analyses, meta-regression), if done, indicating which were pre-specified.	NA
RESULTS			
Study selection	17	Give numbers of studies screened, assessed for eligibility, and included in the review, with reasons for exclusions at each stage, ideally with a flow diagram.	S2_Fig
Study characteristics	18	For each study, present characteristics for which data were extracted (e.g., study size, PICOS, follow-up period) and provide the citations.	8-21
Risk of bias within studies	19	Present data on risk of bias of each study and, if available, any outcome level assessment (see item 12).	8-27
Results of individual studies	20	For all outcomes considered (benefits or harms), present, for each study: (a) simple summary data for each intervention group (b) effect estimates and confidence intervals, ideally with a forest plot.	NA
Synthesis of results	21	Present results of each meta-analysis done, including confidence intervals and measures of consistency.	NA
Risk of bias across studies	22	Present results of any assessment of risk of bias across studies (see Item 15).	NA
Additional analysis	23	Give results of additional analyses, if done (e.g., sensitivity or subgroup analyses, meta-regression [see Item 16]).	NA
DISCUSSION			
Summary of evidence	24	Summarize the main findings including the strength of evidence for each main outcome; consider their relevance to key groups (e.g., healthcare providers, users, and policy makers).	30
Limitations	25	Discuss limitations at study and outcome level (e.g., risk of bias), and at review-level (e.g., incomplete retrieval of identified research, reporting bias).	30,31
Conclusions	26	Provide a general interpretation of the results in the context of other evidence, and implications for future research.	31,32
FUNDING			
Funding	27	Describe sources of funding for the systematic review and other support (e.g., supply of data); role of funders for the systematic review.	

41 From: Moher D, Liberati A, Tetzlaff J, Altman DG, The PRISMA Group (2009). Preferred Reporting Items for Systematic Reviews and Meta-Analyses: The PRISMA Statement. PLoS Med 6(7): e1000097. 42 doi:10.1371/journal.pmed1000097

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