Supplementary information

RNA timestamps identify the age of single molecules in RNA sequencing

In the format provided by the authors and unedited

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Supplementary Figure 1: (A) Five editing templates that were tested are shown. Of these, only templates A and B showed robust temporal editing that seemed appropriate for the construction of the timestamp system. Notes on the templates are in Supplementary Table 3. (**B**) The mean number of edits per RNA for several different timepoints is shown for three different ADAR variants, and for templates A and B. The protocol used here is identical to that in Fig. 1E. Some combinations, such as dmE488Q with template A, may show greater temporal resolution at short 9 timescales. Data represented as mean \pm std (n = 3 biological replicates). (C) Example editing histograms are shown for three different timepoints, for each combination of the three enzymes 11 and two templates in (B).

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- **Supplementary Figure 2:** The qPCR for the iRFP transcript, normalized to GAPDH, is shown as a function of time during the experiment in Fig. 1E. Values are normalized to the pre-
- as a function of time during the experiment in Fig. 1E. Values are normalized to the pre-
- 17 doxycycline timepoint. Error bars show mean \pm std (N= 3 biological samples).
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Supplementary Figure 3. We designed a statistical model to predict the RNA age distribution as a function of time since doxycycline induction. If the adenosines on the timestamp template are edited independently and uniformly in time, then for each adenosine on the timestamp, the fraction of RNAs with adenosines at that site should decrease exponentially with the time since transcription, with a site-specific rate constant that depends on the local sequence context. (**A**) For each adenosine on the timestamp, we fitted an exponential cumulative distribution function (CDF) to the editing fraction over time at that base. The fraction of A to I edits as a function of time is shown for three different bases on the timestamp, data from one replicate of 1E. Best exponential fits are shown. The black dotted line indicates the addition of actinomycin D. (**B**) We found 24 bases which fit well to the model (i.e., for which the value of R2 was greater than 30 0.9 across all replicates). For the same replicate as in (A), the R^2 value of the exponential fit is 31 shown for each base on the transcript. The black dotted line indicates the $R^2 > 0.9$ cutoff used for the exponential model. (**C**) Analyzing only those bases, the distribution of edits per RNAs was well-approximated by a Poisson binomial distribution with a single parameter, t, which represents time since doxycycline was added to the medium (see Methods), with the weights in the Poisson binomial distribution given by the exponential CDFs. The masked editing histograms 36 for four timepoints from the same replicate are shown (only the bases with $R^2 > 0.9$ are included). In green, the Poisson binomial distribution for each timepoint including all the bases 38 with $R^2 > 0.9$ (see Methods). (**D**) We used this Poisson binomial distribution to infer the times of cells induced at 2.5 and 4.5 hours prior to lysis, timepoints that had not been included in the

- dataset used to fit the exponential CDFs. By minimizing the Kullback-Leibler divergence (which
- is equivalent to maximizing the likelihood) between the test distributions and the Poisson
- 42 binomial distribution over t, we inferred that timing of those events to be $2.35\text{hr} \pm 0.09\text{hr}$ and
- 4.45hr \pm 0.03hr (mean \pm s.d., N=3 biological samples), respectively. In orange, the masked (R² >
- 0.9 in all 3 replicates from 1E, see Methods) editing histogram for a single 2.5 hour replicate
- along with Poisson binomial distribution for 2.5 hours (red line), and the Poisson binomial
- distribution with least KL divergence from the empirical distribution (blue line). The time
- estimate is mean ± s.d. (N=3 biological samples). (**E**) As in (D), but for the 4.5 hour timepoint.
- 48 (F) The mean absolute error is shown for the (D) and (E). Error bars show mean \pm S.D. (N = 3
- biological samples).
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Supplementary Figure 4. The fact that timestamps work with multiple promoters raises the possibility of recording the activity of multiple promoters simultaneously in a single cell population, and we validated that this is possible using barcoded timestamps responsive to the

Tet and Vivid promoters. All editing histograms are normalized to sum to 1. (**A**) For cells

transfected with a barcoded TRE-responsive timestamp construct, a barcoded Vivid-responsive

timestamp construct, or both, the number of reads for the TRE-responsive timestamp, Vivid-

- responsive timestamp, or both are shown. When only one timestamp is transfected, only one
- barcode is detected in significant numbers, confirming that there is minimal crossover between
- timestamp barcodes. Note that the third column is not the sum of the first and second columns,
- because it includes barcodes that did not perfectly align to either the Tet or Vivid timestamp
- barcodes. (**B**) To further confirm the possibility of multiplexing using barcoded timestamps, we
- analyzed the editing histograms for cells that were transfected with a barcoded TRE-responsive
- timestamp construct, a barcoded Vivid-responsive timestamp construct, or both. The editing
- histograms for the Vivid-responsive and TRE-responsive timestamps do not seem to change
- when the other timestamp is also present, again suggesting that there is minimal cross-talk
- between barcoded timestamp constructs. All editing histograms are normalized to sum to 1.
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Supp Fig. 5: Sequencing-Based Activity Measurement in Neurons using c-Fos Timestamps.

(**A**) Schematic of timestamps constructs and experimental timeline for neuronal recording. (**B**)

Editing histograms are shown for neurons prior (blue) and one hour following (orange) KCl

induction. The lower overall editing rate for the +KCl case indicates the generation of new

repRNAs by the c-fos promoter. Editing histograms are normalized so the sum of all values is 1. (**C**) The mean editing rate is shown as a function of time following KCl induction. (**D**) The

76 predicted and actual time estimates are shown for all timepoints. Dotted line is a guide for $Y=X$.

There are no estimates for the 1 hour and 7 hour timepoints due to mean interpolation. (**E**) The

mean absolute error in the predictions from (C) is shown as a function of time since induction.

79 All error bars (C-E) show mean \pm SD. (For C-E N = biological replicates as follows, 0hr N = 2,

80 1hr N=3, 2hr N=4, 3hr N=5, 3.5hr N=3, 4hr N=5, 5hr N= 4, 5.5hr N = 3, 6hr N = 2, 7hr N= 2).

85 produced by the gradient descent decoder applied to the ramp profile in (A) with different

numbers of RNAs (rows). Columns show technical replicates, each corresponding to a different

random set of RNAs. Replicates for display were selected at random. (**C**) The fraction of weight

incorrectly assigned. "Random" was calculated on distributions chosen from a Dirichlet

89 distribution. Error bars show mean \pm std (N = 10 replicates).

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the analysis pipeline is shown. See Methods. (**C**) For one replicate from the experiment in Fig. 1E, a histogram of the number of reads with a given percentage of As with Q score >27 is shown. This includes all sites that are As on the timestamp template, i.e., it also counts Gs that are read at positions that are A on the template. The black line indicates the 90% cutoff, which was applied to all analysis. (**D**) For one replicate from the experiment in Fig. 1E, the percentage of reads having no edits in either R1 or R2 is shown as a function of time. These reads were

- excluded from analysis, except where otherwise stated in Fig. S2.
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 $\begin{array}{c} 100 \\ 101 \end{array}$ **Supplementary Fig. 8: Schematic for decoder of timestamps using gradient descent.** Given

an observed distribution of edits per RNA, we want to reconstruct the set of 1 hour weights

which generate this observed distribution. To infer the transcriptional program underlying this

observed editing distribution, we perform gradient descent to minimize the L2 norm between the observed distribution and a convex (a nonnegative linear combination with weights summing to

1) combination of the 12 basis distributions. These basis distributions are obtained from the he

average distribution of edits obtained from a transcriptional pulse lasting one hour that began

between 1 and 12 hours in the past in the calibration experiment performed in HEK293 cells.

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110 **Supplementary Figure 9:** For 1000 randomly generated weight vectors ("simulated vectors"), chosen according to a Dirichlet distribution with uniform weights, we used gradient descent to find the approximation ("approximated vectors") that minimized the L2 norm ("inner product") between the RNA editing distribution corresponding to the simulated vectors ("simulated distributions") and the RNA editing distribution corresponding to the approximated vectors ("approximated distributions"). We refer to the L2 norm between the distributions as the inner product to distinguish it from the L2 norm between the vectors, which we refer to as the mean squared error (MSE). (**A**) The inner product between simulated distributions and approximated distributions is shown in blue. By contrast, the inner product between simulated distributions and other random distributions is shown in orange. (**B**) The mean squared error between the simulated vectors and approximated vectors is shown in blue. By contrast, the inner product between the simulated distributions and other random distributions is shown in orange. Note that a substantial number of random weight vectors have lower mean squared error than the approximated vectors. This is possible because the noise in the basis distribution set used to generate the approximated distributions from the approximated vectors is different from the noise in the basis distribution set used to generate the simulated distributions from the simulated vectors, so the minimum of inner product between the simulated and approximated distributions is not always the same as the minimum of the MSE between the simulated and approximated vectors. (**C**) Another visualization of (B). For each simulated vector, we calculated both an approximated vector and a random vector. The difference in MSE between the approximated and random vectors is shown. Negative values correspond to test vectors for which the associated random vector was a better approximation to the simulated vector than the approximated vector. (**D**) Blue and orange bars are the same as in (B). Yellow bars correspond to the minimum MSE among all of the solutions found by gradient descent for a given test vector, indicating that the inner product minima found by the gradient descent are not in general minima of the MSE. (**E**) The difference in the inner product between the solutions with the minimum MSE found by gradient descent, and the solutions with the minimum inner product, as a fraction of the minimum inner product. The solutions with the minimum MSE discovered by gradient descent often have inner products several fold higher than the solution with the minimum inner product.

140 **Supplementary Table 1:** List of plasmids used in this study. This list excludes pCMV Tet3G,

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Supplementary Table 2: List of oligos used in this study 145
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149 **Supplementary Table 3:** List of RNA editing templates used in this study.

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151 The following sequences are the sequences that were analyzed for RNA editing. Notes are supplied as a courtesy to follow-on studies, and no representations are made as to their accuracy or reproducibility. courtesy to follow-on studies, and no representations are made as to their accuracy or reproducibility.

