

Appendix 1. Descriptions of consistency and inconsistency models for a simple network with a closed loop consisting of interventions A, B, and C.

The data consist of $\hat{\theta}_{ijk}$, the observed relative effect size of treatment k compared to treatment j from study i ($j = A, B, k = B, C, j \neq k$), and its variance, $\widehat{var}(\hat{\theta}_{ijk})$. The notation is identical to that used in the main body.

- Consistency model

$$\hat{\theta}_{ijk} \sim N\left(\theta_{ijk}, \widehat{var}(\hat{\theta}_{ijk})\right)$$

$$\theta_{ijk} \sim \text{Normal}(d_{jk}, \tau^2)$$

$$\theta_{iAB} \sim \text{Normal}(d_{AB}, \tau^2), \theta_{iAC} \sim \text{Normal}(d_{AC}, \tau^2), \theta_{iBC} \sim \text{Normal}(d_{AC} - d_{AB}, \tau^2)$$

$$d_{AB} \sim \text{Normal}(0, 10000), d_{AC} \sim \text{Normal}(0, 10000), \tau \sim \text{Uniform}(0, 5)$$

- Inconsistency model

$$\hat{\theta}_{ijk} \sim N\left(\theta_{ijk}, \widehat{var}(\hat{\theta}_{ijk})\right)$$

$$\theta_{ijk} \sim \text{Normal}(d_{jk}, \tau^2)$$

$$\theta_{iAB} \sim \text{Normal}(d_{AB}, \tau^2), \theta_{iAC} \sim \text{Normal}(d_{AC}, \tau^2), \theta_{iBC} \sim \text{Normal}(d_{BC}, \tau^2)$$

$$d_{AB} \sim \text{Normal}(0, 10000), d_{AC} \sim \text{Normal}(0, 10000), d_{BC} \sim \text{Normal}(0, 10000),$$

$$\tau \sim \text{Uniform}(0, 5)$$

Appendix 2. Details of how to derive the formula for calculating $\widehat{var}(\hat{\theta}_{kBC}^*)$, $k= 1, \dots, l$.

The imputed data consisted of a number of treatment effect sizes of treatment C compared to B and their variances, $\hat{\theta}_{kBC}^*$ and $\widehat{var}(\hat{\theta}_{kBC}^*)$, $k= 1, \dots, l$ (where l is the number of missing trials for comparison between B and C). The estimated variance of a pooled effect size from a pairwise random-effects meta-analysis of the data is $\frac{1}{\sum_{k=1}^l \frac{1}{\widehat{var}(\hat{\theta}_{kBC}^*) + \hat{\tau}_{BC}^2}}$, where $\hat{\tau}_{BC}^2$ is the extent of heterogeneity of the effect size [1].

Based on conditions II and III, we conservatively assumed that this variance would be the same as the variance of the effect size estimated from the original star-shaped network meta-analysis, $\widehat{var}(\hat{d}_{AC} - \hat{d}_{AB})$, and also assumed that $\hat{\tau}_{BC}^2$ would be equal to that of the overall heterogeneity across contrasts in the star-shaped network, $\hat{\tau}^2$. Accordingly, the following equation was established:

$$\widehat{var}(\hat{d}_{AC} - \hat{d}_{AB}) = \frac{1}{\sum_{k=1}^l \frac{1}{\widehat{var}(\hat{\theta}_{kBC}^*) + \hat{\tau}^2}} .$$

For simplicity, $\widehat{var}(\hat{\theta}_{1BC}^*), \dots, \widehat{var}(\hat{\theta}_{lBC}^*)$ were set up to be identical. The equation was thus expressed as:

$$\widehat{var}(\hat{d}_{AC} - \hat{d}_{AB}) = \frac{1}{\frac{l}{\widehat{var}(\hat{\theta}_{kBC}^*) + \hat{\tau}^2}} .$$

Finally, from the above equation, we obtained a formula, $\widehat{var}(\hat{\theta}_{kBC}^*) = l \cdot \widehat{var}(\hat{d}_{AC} - \hat{d}_{AB}) - \hat{\tau}^2$.

1. Whitehead A, Whitehead J: **A general parametric approach to the meta-analysis of randomized clinical trials.** *Statistics in medicine* 1991, **10**(11):1665-1677.