Dynamic Conditional Correlation Estimation

The Dynamic Conditional Correlation (DCC) method [\(Engle,](#page-3-0) [2002\)](#page-3-0) is an example of a model-based multivariate volatility method, popular in finance, and recently introduced to neuroimaging (Lindquist et al., 2014). It provides a way to estimate conditional correlations between time series.

We begin by introducing generalized autoregressive conditional heteroscedastic (GARCH) models (Engle, 1982; Bollerslev, 1986), often used to model the volatility in univariate time series. These models express the conditional variance of a time series at time *t* as a linear combination of past values of the conditional variance and of the squared process itself. To illustrate, assume that we are observing a univariate process

$$
y_t = \sigma_t \varepsilon_t
$$

where ε _{*t*} is a N (0, 1) random variable and σ _{*t*} represents the time-varying variance term we want to model. A GARCH(1,1) process models the conditional variance as follows:

$$
\sigma_t^2{=}\omega{+}\alpha y_{t-1}^2{+}\beta \sigma_{t-1}^2
$$

where $ω > 0$, $α, β ≥ 0$ and $α+β < 1$. Here $α$ controls the influence of past values of the time series on the variance and *β* the influence of past values of the conditional variance on its present value.

In the DCC approach, we assume $y_t = \varepsilon_t$ is a bivariate mean zero time series with conditional covariance matrix Σ*t*. The first order form of DCC can be expressed as follows:

$$
\sigma_{i,t}^2 = \omega_i + \alpha_i y_{i,t-1}^2 + \beta_i \sigma_{i,t-1}^2 \text{ for } i = 1, 2
$$
\n(1)

$$
\mathbf{D}_t = diag\{\sigma_{1,t}, \sigma_{2,t}\}\tag{2}
$$

$$
\varepsilon_t = \mathbf{D}_t^{-1} \mathbf{e}_t \tag{3}
$$

$$
\mathbf{Q}_t = (1 - \theta_1 - \theta_2) \overline{\mathbf{Q}} + \theta_1 \varepsilon_{t-1} \varepsilon_{t-1}^{\prime} + \theta_2 \mathbf{Q}_{t-1} \tag{4}
$$

$$
\mathbf{R}_t = diag\{\mathbf{Q}_t\}^{-1/2} \mathbf{Q}_t diag\{\mathbf{Q}_t\}^{-1/2}
$$
\n(5)

$$
\sum_{t} = \mathbf{D}_{t} \mathbf{R}_{t} \mathbf{D}_{t}
$$
\n(6)

In Eqs. 1-3, univariate GARCH(1,1) models are fit (Eq. 1) to each of the two univariate time series that make up **y***^t* , and used to compute standardized residuals (Eq. 3). In Eq. 4, an exponentially weighted moving average is applied to the standardized residuals and used to compute a non-normalized version of the time-varying correlation matrix. Here **Q̄** represents the unconditional covariance matrix of ε_t and (θ_1, θ_2) are non-negative scalars satisfying $0 < \theta_1 + \theta_2 < 1$. Eq. 5 is a rescaling that ensures a proper correlation matrix is created, while Eq. 6 computes the time-varying covariance matrix. Model parameters $(\omega_1,$ *α*1*, β*1*, ω*2*, α*2*, β*2*, θ*1*, θ*2) can be estimated using a two-step approach. This approach has been shown (Engle and Sheppard, 2001; Engle, 2002) to provide estimates that are consistent and asymptotically normal with a variance that can be computed using the generalized method of moments approach.

Figure S1. Network heritability estimates in each session for the DCC variance, DCC mean, and static connectivity measures. $VN =$ visual network, $SMN =$ sensorimotor network, $VAN =$ ventral attention network, $DAN =$ dorsal attention network, $DMN =$ default mode network, FPN = frontoparietal network

Table S1. Network heritability estimates in each session for the DCC variance, DCC mean, and static connectivity measures computed at different ICA dimensionalities. VN $=$ visual network, SMN $=$ sensorimotor network, VAN $=$ ventral attention network, DAN $=$ dorsal attention network, DMN = default mode network, FPN = frontoparietal network

Table S2. Individual connection heritability estimates in each session for the DCC variance, DCC mean, and static connectivity measures computed at the ICA dimensionality of 300 components.

References

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