## Appendix A Data Description and sources

In this appendix we describe in more detail the variables and their sources. Table [6](#page-5-0) provides summary statistics. The first subset of explanatory variables relates to *demographic* characteristics. They include population density in each province, average age, the average size of families, the share of students, the share of secondary school acquisition among 19+ years old residents, the share of postgraduate degree acquisition, the share of families with only one component, and the share of families with five or more components. The Italian national statistical agency ISTAT provides these measures either in 2019 or in 2011, the last year of the full Census. We also create a variable that weighs the number of students with the percent of remote-teaching conducted in each province on 15th November  $2020^{8}$ . The second subset of explanatory variables relates to economic characteristics. They include average income per capita (source: Eurostat, 2017), the share of employed workers in the population, share of the agricultural sector, the share of the industrial sector, the share of the service sector, and share of retail and accommodation activities (source: all ISTAT 2019). We also create a variable that weighs the share of retail and accommodation with the percent of businesses that remained open in each province (Ministry of Health) during fall 2020. The third subset of explanatory variables relates to commuting activities. We build two measures based on on the total commuting by public transport with trips longer than 15 minutes for i) work, and ii) study reasons. Using the detail of the hour at which commuters leave home and by what transportation mean, we build a measure of iii) concentration of long  $(215$ minutes) trips on public transport, weighted by the covid concentration in the province of destination. Finally, we build four measures of exposure through outgoing (OUT) or incoming (IN) commuters to covid. The variables are calculates as

$$
X_{ij} = \frac{\sum\limits_{ab \neq ij} C_{ab} \text{flow}_{(ij)(ab)}^D}{\sum\limits_{ab \neq ij} \text{flow}_{(ij)(ab)}^D}
$$
(6)

Where ab is any other province different from  $ij$ ,  $C_{ab}$  is the covid incidence per capita in province ab and flow $\bigcup_{(ij)(ab)}^D$  is the flow from either ij to ab if  $D = \text{OUT}$  or from ab to ij if  $D = \text{IN}$ . In practice, these variables are the average of neighbours' covid incidence, weighted by the commuting flows. These aim to capture whether commuting is a relevant predictor of local covid incidence as a function of whether local commuters work in provinces with high incidence (OUT) or local workers come from provinces with high incidence  $(IN)$ . We build four variables of this kind: iv) commuting covid  $IN, v$  commuting covid OUT, vi) commuting covid IN (using public transport flows only), and vii) commuting covid OUT (using public transport flows only). The original commuting data are from ISTAT, 2011 Census; we use the official cases in the whole second wave  $(1/09/2020-23/12/2020)$  to construct covid exposure. The fourth subset of variables relates to the *health and public health system*. They include mortality rate for cancer in the period 2012-2016, the mortality rate for heart attack in the period 2012-2016, increased life expectancy in the period 2002-2017, asthma incidence, measured as pro-capita consumption of medicine for asthma and Chronic Obstructive Pulmonary Disease (COPD), diabetes incidence, measured as pro-capita consumption of medicine for diabetes, hypertension incidence, measured as pro-capita consumption of medicine for hypertension, the average number of general practitioner doctor per capita, average number of hospital beds per capita. These data are retrieved from the Health index survey from il Sole  $24$  ore. The fifth subset of variables includes a geographical characteristic: the temperature registered in the period 2007-2016. (source: ISTAT). Finally, we include a measure of covid-19 incidence pre-September 2020, which captures the first wave's strength across provinces.

#### Hence, the dataset of explanatory variables is composed of 35 variable.

In addition to these, we collect data on the covid-19 incidence between 1/09/2020-3/11/2020, 4/11/2020- 23/12/2020, 25/11/2020-23/12/2020, 1/09/2020-26/01/2021, and 26/02/2020-26/01/2020. We do not include these variables in the LASSO selection procedure, as we use them as dependent variables.



#### Table 4: Data

Note: The health data from il Sole 24 ore can be retrevied here: [https://lab24.ilsole24ore.com/](https://lab24.ilsole24ore.com/indice-della-salute/indexT.php) [indice-della-salute/indexT.php](https://lab24.ilsole24ore.com/indice-della-salute/indexT.php)

# Appendix B Pre- and Post-Policy incidence

## B.1 Pre-Policy incidence

Figure 3: Weekly Cases per 100k people: Pre-Policy, 1/09/2020 - 3/11/2020



# B.2 Post-Policy incidence

Figure 4: Weekly Cases per 100k people: Post-Policy, 25/11/2020 - 23/12/2020



# Appendix C Robustness tables

### C.1 OLS estimates - Red and Yellow Tiers

### Table 5: Yellow and Red Tiers OLS Results



Note: Significance levels:  $* = 0.10; ** = 0.05; ** = 0.01.$  "Sh." stands for "Share. In the interaction terms, "Y" stand for "Yellow Tier" and "R" for "Red Tier". Number is parenthesis report the p-value of the t-test. All models are based on Equation [5.](#page--1-1) Specifications (1) and (3) test the model  $\tilde{\beta}X_{ij} + \gamma^Y S_i^{\text{Yellow}} X_{ij}$  with null hypothesis  $H_0: \tilde{\beta} + \gamma^Y = 0$ . Specifications (2) and (4) test the model  $\tilde{\beta}X_{ij} + \gamma^R S_i^{\text{Red}} X_{ij}$  against the null hypothesis  $H_0: \tilde{\beta} + \gamma^R = 0$ .

### <span id="page-5-0"></span>C.2 Robustness Checks



#### Table 6: Robustness checks

Note: Significance levels:  $* = 0.10$ ;  $** = 0.05$ ;  $*** = 0.01$ . All specifications use Conley Spatial Standard Errors with a cutoff of 150km. P-values of coefficients in parenthesis. "Public Tran. Trips Conc." stands for "Public Transport Trips Concentration". All regressions are controlled for region fixed effects. Therefore, the  $\beta$ coefficient on each variable can be interpreted as contributing to increasing (decreasing) Covid-19 cases per capita beyond (below) the regional mean. Specification (1) removers Sardegna due to its isolated status. Specification (2) removes also Campania and Sicilia, as they introduced some limited city-wide red tiers before the regional policies. Specification (3) extends the sample to 26th January 2021. Specification (4) considers the whole pandemic period.

## <span id="page-6-0"></span>Appendix D Robust Inference and Model Selection

The reader may be worried that the model selection through LASSO may change the inference approach that one should take in assessing the significance of the results. That is: can we really reject the null hypothesis that there are local-level effects in the pre-policy period, since we have selected the regressors in order to maximize R2 adjusted?

The worry here is that under small sample, the pre-selection over a large number of regressors may lead to overfitting and the selection of covariates uncorrelated to the dependent variable in the true data generating process, but correlated in the data due to small sample bias.

In this section, we show that simulating synthetic data allows us to produce an empirical distribution of post-selection OLS F-statistics under the null hypothesis. Using this distribution, we can build confidence intervals and rejection regions that account for the model selection algorithm. In particular, we generate 1000 draws of sets of 38 normally iid distributed regressors (random iid data, henceforth). We subtract regionals means in order to be centered within region. Then, we apply to each of them our model selection procedure and store the F-test p-value of the subsequent OLS regression (we take as reference specification 4, Table 1), assigning a value of one when no variable is selected ( $\approx 15\%$ of the cases). Then, we check the 5th percentile of the distribution of p-values so obtained, which represents the critical value representing the OLS F-test p-value such that less than 5% of draws under the null hypothesis of no correlation between covariates and dependent variable sit at lower p-values. Finally, we compare this critical value with the p-value obtained in the real data. We repeat this exercise by drawing 1000 sets of 38 jointly normally distributed regressors, with covariance matrix replicating the one of our true dataset (random correlated data, henceforth). This allows to account for the preference of LASSO of selecting predictors with low correlation, selecting less variables than in the case of uncorrelated sets of regressors.

Our results are confirmed by this empirical, stricter rejection criteria, built to account jointly for the selection and post-selection steps. Table [7](#page-7-0) shows how only 0.1% of the simulations in the iid data and 0% of the simulations in the correlated data have an F-test pvalue smaller than the one built using the real data. This is true whether we apply (right column) or do not apply (left column) the refinement process to maximize R2-adjusted after the LASSO. This means that the post-selection OLS p-value of the true data is much smaller than the one of most random data, with 99.9% of all simulations achieving a larger p-value. This means that our results are indeed significant at the 5% level and thus unlikely to be produced by covariates uncorrelated to the dependent variable.

In Table [8](#page-7-1) we show similar results for the R2 adjusted: it is highly unlikely for randomly generated covariates to generate an amount of R2-adjusted similar to the one of the true data.

<span id="page-7-0"></span>Table 7: Random Generated Samples and Statistical significance, with and without refinement. Share of simulations

|                        | $p-value(Fstat_z) < p-value(Fstat Data)$ |                 |  |
|------------------------|------------------------------------------|-----------------|--|
|                        | Without refinement                       | With refinement |  |
| Random iid data        | $0.1\%$                                  | $0.1\%$         |  |
| Random correlated data | $0.0\%$                                  | $0.0\%$         |  |

Note: this table displays the share of simulations (out of 1000), in percent, for which the p-value of the F-statistics (null hypothesis:  $H_0$ :  $\beta = 0$ , in model [2\)](#page--1-2) is less than the one found in the data. The first row displays the results when the regressors are assumed to be iid. The second row displays the results when the regressors are assumed to have the same covariance matrix as the regressors in the data. The first column presents the results without the refinement, while the second column presents the results with the refinement.

<span id="page-7-1"></span>Table 8: Random Generated Samples and Explanatory power, with and without refinement: Additional  $R^2$  Adjusted

|                                   |                                                        | <b>Without Refinement</b> |                | With Refinement     |
|-----------------------------------|--------------------------------------------------------|---------------------------|----------------|---------------------|
|                                   | All Samples                                            | Significant Samples       | All Samples    | Significant Samples |
|                                   | Random iid data                                        |                           |                |                     |
| Average $R^{2,z}$                 | 0.09                                                   | 0.20                      | 0.10           | 0.22                |
| 95\% conf Interval                | $[0.0 - 0.21]$                                         | $[0.16 - 0.25]$           | $[0.0 - 0.22]$ | $[0.18 - 0.26]$     |
| Frequency: $R^{2,z} > R^{2,data}$ | $0.3\%$                                                | $6.0\%$                   | $0.9\%$        | 14\%                |
| Average $R^{2,z}$                 | Random correlated data<br>0.07<br>0.21<br>0.18<br>0.08 |                           |                |                     |
| 95\% conf Interval                | $[0.0 - 0.18]$                                         | $[0.13 - 0.22]$           | $[0.0 - 0.21]$ | $[0.17 - 0.26]$     |
| Frequency: $R^{2,z} > R^{2,data}$ | $0.0\%$                                                | $0.0\%$                   | $0.5\%$        | $6\%$               |

Note: this table displays the additional Adjusted  $R^2$  $R^2$  of model 2 with respect to model [1](#page--1-3) in the 1000 simulations. This statistic captures the additional explanatory power of the selected regressors in addition to the regional fixed effects. The top-panel displays the results when the regressors are assumed to be iid. The second panel displays the results when the regressors are assumed to have the same covariance matrix as the regressors in the data. The left panel presents the results without the refinement, while the right panel presents the results with the refinement. The first column presents the statistics for all the simulations (1000), while the second column presents the statistics for the 5% simulations with the lowest p-value of the F-statistics. The first line displays the average additional Adjusted  $R^2$ , across the simulations. The second line displays its 95 percent confidence interval. The third line displays the share of simulations, in percent, for which the Adjusted  $R^2$ with the synthetic data is larger than the one found in the data (equal to 0.2449 without the refinement and equal to 0.2524 with the refinement).

## Appendix E Post-LASSO Refinement Procedure

In this section, we discuss the role of the refinement to the LASSO selection discussed in the main text. The refinement works as follows: take all covariates selected by the LASSO procedure. Then, start iterating over the variables with the lowest p-value, perform an OLS regression and: (1) keep the variable if R2-adjusted does not increase, or (2) discard the variables if R2-adjusted increases. Under option (2), repeat the procedure until you find that R2-adjusted does not increase any further.

We have discussed in Appendix [D](#page-6-0) how this has little impact on the inference procedure and on the explained  $R^2$  adjusted of the selected model. In Table [9](#page-9-0) we present further evidence of how the variable selection in random data and in a bootstrap exercise is affected by this refinement.[9](#page-8-0) The refinement reduces the number of selected variables by 1.3 out of an average of 9.7 (when we use 38 random, uncorrelated regressors to simulate our procedure under the null hypothesis), and shrinks by 6 the upper bound of the 95% confidence interval of the distribution. When we simulate the procedure using correlated regressors with the same covariance matrix as the true data, the refinement shrinks the number of selected variables by 2 out of 8.2, and shrinks the upper bound of the confidence interval by 7 out of 29. Finally, when we bootstrap the error terms of the dependent variable, we find that the refinement shrinks the average selected covariates (from the true data) by 5.1 variables.

<span id="page-8-0"></span><sup>&</sup>lt;sup>9</sup>In the bootstrap exercise we resample [1](#page--1-3)000 times the residual obtained after estimating Model 1 to recreate 1000 dependent variables that have the same systematic component as the one estimated from the data but a different realization of the random component. We then repeat all the steps of our methodology (Lasso selection, and refinement) to each newly obtained dependent variables.

<span id="page-9-0"></span>

| Without Refinement                                   | With Refinement |  |  |
|------------------------------------------------------|-----------------|--|--|
| Random iid data                                      |                 |  |  |
| 13.9                                                 | 13.9            |  |  |
| 9.7                                                  | 8.4             |  |  |
| $\left[ \begin{array}{c} 0 - 27 \end{array} \right]$ | $[0-21]$        |  |  |
| Random correlated data                               |                 |  |  |
| 15.2                                                 | 15.2            |  |  |
| 8.2                                                  | 6.2             |  |  |
| $\left[ \begin{array}{c} 0 - 29 \end{array} \right]$ | $[0-22]$        |  |  |
|                                                      |                 |  |  |
| $\Omega$                                             | $\theta$        |  |  |
| 21.4                                                 | 16.3            |  |  |
| 11 - 33                                              | $[9-25]$        |  |  |
|                                                      | Bootstrap       |  |  |

Table 9: Regressor Selection: with and without refinement

Note: this table displays the share of simulations in which the selection procedure select zero regressors in percent, (first line); the average number of regressors selected (second line), and its 95% confidence interval (third line) obtained by using the Lasso procedure without (first column) and with (second column) our proposed refinement. The top and central panels display the results for the randomly generated data (iid and correlated, respectively). The bottom panel displays the results for the bootstrapping exercise.