

# MeToo Movement and Perceptions of Sexual Assault: Annotated analysis code

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# Introduction

## R Version:

x86\_64-redhat-linux-gnu, x86\_64, linux-gnu, x86\_64, linux-gnu, , 3, 4.1, 2017, 06, 30, 72865, R, R version 3.4.1 (2017-06-30), Single Candle

Run date: 2019-11-17 00:19:24

## Initial Seed

```
set.seed(3107941)
```

What follows is annotated code for *The #MeToo Movement and Perceptions of Sexual Assault: College Students' Recognition of Sexual Assault Experiences Over Time*. Its purpose is to allow readers to reproduce our analyses as exactly as possible.<sup>1</sup> You can read more about .Rmd files at [rmarkdown.rstudio.com](http://rmarkdown.rstudio.com).

This file contains all of the data import and model estimation code and output.

Note, our code was originally written in modules to keep each .Rmd document a manageable size. However, for portability, all of those documents have been knitted together in one large document here. To produce this document, the original .Rmd files (list below), were all run in the following order:

Section 1: Importation and model estimation

1. data\_wrangling.Rmd
2. analysis\_functions.Rmd
3. multiple\_imputation.Rmd
4. models\_and\_priors.Rmd
5. model\_comparison.Rmd
6. mcmc\_chain\_convergence.Rmd

Section 2: Processing and summarizing results

1. aws\_results\_summary.Rmd

---

<sup>1</sup>Note there are always limits to reproducibility. For example, Monte Carlo simulation is used in some of our analyses. Depending on the random seed and randomization strategy you use, you may get different results for those analyses. However, the differences should be small.

The output (text, code, and results) of these files is given in their corresponding sections below. None of the code blocks used to produce our results in this document are hidden. The code you see is all of the code that was run, from start to finish.

## Data Wrangling

### Import participant data

Data were originally stored in SPSS format, so were imported with the `memisc` package. The resulting dataframe included  $n = 2,616$  cases.

```
library(tidyverse)

spss_dataset <- memisc::spss.system.file('data/MeTooData_updatedAug2019.sav')

participant_df <- spss_dataset %>%
  memisc::as.data.set() %>%
  as.data.frame()

nrow(participant_df)

## [1] 2616
```

### Coding date-related variables

We continued by scoring date-related variables. Note, there are odd parameters for the `as.Date()` function that have been commented out (lines 37-38). These were intentional, as there was a previous problem converting SPSS dates to POSIXct format. They appear to have been fixed by the new version of `memisc`, so we commented those lines out and simply use the `ymd_hsm()` function from `lubridate`. However, we leave the old lines available in the code block in case attempts to reproduce our work require them for some unforeseen reason.

We also denoted a dichotomous pre-post MeToo timing variable (i.e., before/after 2017-10-15).

**NOTE:** In order to get the numerical optimization algorithms to behave appropriately, we produced a `week_num` variable by taking the number of days between a participant's study day and the minimum study day, then dividing by 7. This in turn reduces the magnitude of the resulting `week_num_SQ` variable. Our previous testing has shown that without this modification, MCMCpack is not able to compute marginal likelihoods effectively. Note that because we are dividing days by 7, a participants `week_num` variable can be a decimal (e.g., `week_num` =  $13/7 = 1.85$ ).

```
library(lubridate)

participant_df <- participant_df %>%
  mutate(
    # start_date = as.Date(startdate/86400, origin = '1582-10-14'),
    # end_date = as.Date(enddate/86400, origin = '1582-10-14'),
    start_date = ymd_hms(startdate),
    end_date = ymd_hms(enddate),
    hours_to_complete = difftime(end_date, start_date, units = 'hours'),
    start_day = as.Date(floor_date(start_date, unit = 'days')),
```

```

study_week = as.Date(floor_date(start_day, unit = 'week')),
# day_num = as.numeric(start_day - min(start_day)) # submission 1
day_num = as.numeric(start_day - as.Date('2017-10-15')), # rri
week_num = day_num / 7,
post_metoo = ifelse(start_day >= '2017-10-15', 1, 0))

```

## Coding demographic variables

Once date-related variables are scored, we continued with demographics. Race is coded such that anyone who endorses being Latino, American Indian, Asian, Black, Pacific Islander, or “Other Race” is coded as “racial\_minority = 1”; people who endorse being non-Hispanic White are coded 0; and people who are missing on all race questions are coded NA.

LGBTQ status is coded 1 (yes) for anyone who endorsed same sex partners, same sex attraction, or non cis-gender identity. It is coded 0 for anyone who answered no to all of those questions. People with missing scores for all of those questions are coded NA.

Gender is coded 0 = female, 1 = male (as non-cis-gender participants are already accounted for in the LGBTQ group).

```

participant_df$latino = as.numeric(participant_df$latino == 'Yes')
race_vars <- c('white', 'latino', 'amin', 'asian', 'black', 'pacisl', 'othrace')
missing_all_race <- apply(
  participant_df[, race_vars], 1, function(row) all(is.na(row)))
is_racial_minority <- apply(
  participant_df[, race_vars[2:7]], 1, function(row) any(row == 1, na.rm = T))

participant_df <- participant_df %>%
  mutate(
    has_race = !missing_all_race,
    racial_minority = as.numeric(has_race & is_racial_minority),
    male = gender == 'Male',
    male_partners = preference %in% c('Men', 'Both'),
    female_partners = preference %in% c('Women', 'Both'),
    non_hetero_identification = orientation != 'Heterosexual / Straight',
    lgbtq = (male & male_partners) | (!male & female_partners) |
      non_hetero_identification | (!gender %in% c('Male', 'Female')))
```

## Coding sexual experience variables

Then, we scored substantive variables related to participants’ perception of their sexual experiences. Some of these variables were already coded in SPSS, but were renamed for convenience.

```

participant_df <- participant_df %>%
  mutate(
    vic_screener = msesyn,
    vic_worst_is_sa = as.numeric(mses_sa.1),
    vic_worst_prior_sex = msesperp1prior.1 %in% c('Yes', 'Unsure'),
    vic_worst_verbal = (as.numeric(mses_tactic1.1) + as.numeric(mses_tactic2.1) +
      as.numeric(mses_tactic3.1) + as.numeric(mses_tactic6.1))/4,
    vic_worst_incap = (as.numeric(mses_tactic4.1) + as.numeric(mses_tactic5.1))/2,
    vic_worst_force = (as.numeric(mses_tactic7.1) + as.numeric(mses_tactic8.1) +
```

```
    as.numeric(mses_tactic9.1))/3  
)
```

## Interactions and variable reduction

Finally, we computed quadratic and interaction terms, saving the result to a new dataframe with only the variables we need. Also, participants who took longer than 24 hours to complete their survey were excluded.

```
df <- participant_df %>%  
  filter(hours_to_complete <= 24) %>%  
  mutate(  
    week_num_SQ = week_num^2,  
    post_metoo_week = post_metoo*week_num,  
    post_metoo_week_SQ = post_metoo*week_num_SQ) %>%  
  dplyr::select(  
    day_num, post_metoo, study_week, week_num, week_num_SQ,  
    post_metoo_week, post_metoo_week_SQ,  
    male, lgbtq, racial_minority,  
    vic_screener, vic_worst_is_sa, vic_worst_prior_sex, vic_worst_verbal,  
    vic_worst_incap, vic_worst_force)  
  
nrow(df)
```

```
## [1] 2566
```

## Demographics

Total participants in the analysis

```
n <- participant_df %>%  
  filter(hours_to_complete <= 24) %>%  
  nrow()  
  
n
```

```
## [1] 2566
```

Gender distribution

```
participant_df %>%  
  filter(hours_to_complete <= 24) %>%  
  group_by(gender) %>%  
  summarise(n = n()) %>%  
  ungroup() %>%  
  mutate(p = round(n/sum(n), 3))  
  
## Warning: Factor `gender` contains implicit NA, consider using  
## `forcats::fct_explicit_na`
```

```

## # A tibble: 8 x 3
##   gender              n     p
##   <fct>            <int> <dbl>
## 1 Female             1845  0.719
## 2 Male               699   0.272
## 3 Transgender woman    3   0.001
## 4 Transgender man     3   0.001
## 5 Gender queer        8   0.003
## 6 Other - Please specify  2   0.001
## 7 Decline to state    3   0.001
## 8 <NA>                3   0.001

levels(participant_df$gender_txt) #determine text response for "other"

```

```

## [1] ""           "genderfluid"  "Velociraptor"

```

Racial and ethnic distribution

```

participant_df %>%
  filter(hours_to_complete <= 24) %>%
  dplyr::select(race_vars) %>%
  colSums(na.rm = T) %>%
  magrittr::divide_by(n) %>%
  round(3)

##   white  latino   amin   asian   black  pacisl othrace
##   0.825  0.093  0.019  0.099  0.054  0.003  0.040

```

```

table(participant_df$white)

```

```

##
##      1
## 2159

```

```

table(participant_df$latino)

```

```

##
##      0      1
## 2368  244

```

```

table(participant_df$amin)

```

```

##
##      1
## 49

```

```

table(participant_df$asian)

```

```

##
##      1
## 257

```

```
table(participant_df$black)
```

```
##  
##    1  
## 141
```

```
table(participant_df$pacisl)
```

```
##  
##    1  
##  8
```

```
table(participant_df$othrace)
```

```
##  
##    1  
## 107
```

Racial/ethnic minority variable from above (FALSE = non-Hispanic White)

```
participant_df %>%  
  filter(hours_to_complete <= 24) %>%  
  group_by(racial_minority) %>%  
  summarise(n = n()) %>%  
  ungroup() %>%  
  mutate(p = round(n/sum(n), 3))
```

```
## # A tibble: 2 x 3  
##   racial_minority     n      p  
##             <dbl> <int> <dbl>  
## 1                 0  1900  0.74  
## 2                 1   666  0.26
```

LGBTQ distribution

```
participant_df %>%  
  filter(hours_to_complete <= 24) %>%  
  group_by(orientation) %>%  
  summarise(n = n()) %>%  
  ungroup() %>%  
  mutate(p = round(n/sum(n), 3))
```

```
## Warning: Factor `orientation` contains implicit NA, consider using  
## `forcats::fct_explicit_na`
```

```
## # A tibble: 6 x 3  
##   orientation           n      p  
##   <fct>              <int> <dbl>  
## 1 Heterosexual / Straight  2368  0.923
```

```

## 2 Lesbian / Gay          40 0.016
## 3 Bisexual               124 0.048
## 4 Something Else          12 0.005
## 5 Don't know              19 0.007
## 6 <NA>                     3 0.001

```

LGBTQ variable from above (FALSE = heterosexual *and* cisgender)

```

participant_df %>%
  filter(hours_to_complete <= 24) %>%
  group_by(lgbtq) %>%
  summarise(n = n()) %>%
  ungroup() %>%
  mutate(p = round(n/sum(n), 3))

```

```

## # A tibble: 2 x 3
##   lgbtq     n      p
##   <lgl> <int> <dbl>
## 1 FALSE   2331  0.908
## 2 TRUE    235   0.092

```

Age distribution

```

participant_df %>%
  filter(hours_to_complete <= 24) %>%
  summarise(
    mean = mean(age, na.rm = T),
    sd = sd(age, na.rm = T),
    min = min(age, na.rm = T),
    max = max(age, na.rm = T),
    n = n(),
    n_missing = sum(is.na(age))) %>%
  round(2)

##   mean   sd min max   n n_missing
## 1 20.47 2.5  19  55 2566         3

```

## Descriptives for study variables

Sexual assault screener

```

participant_df %>%
  filter(hours_to_complete <= 24) %>%
  filter(vic_screener == 'Yes' | vic_screener == 'No') %>% # only consider those with data on screener
  group_by(vic_screener) %>%
  summarise(n = n()) %>%
  ungroup() %>%
  mutate(p = round(n/sum(n), 3))

```

```

## # A tibble: 2 x 3
##   vic_screener     n      p
##   <fct>       <int> <dbl>
## 1 No            124  0.048
## 2 Yes           2331 0.908

```

```

##   <fct>      <int> <dbl>
## 1 No          1904  0.762
## 2 Yes         596   0.238

```

Sexual assault screener by gender

```

participant_df %>%
  filter(hours_to_complete <= 24) %>%
  filter(vic_screener == 'Yes' | vic_screener == 'No') %>% # only consider those with data on screener .
  group_by(gender, vic_screener) %>%
  summarise(n = n()) %>%
  ungroup() %>%
  group_by(gender) %>% #other gender is specified as "genderfluid" in gender_txt
  mutate(p = round(n/sum(n), 3))

```

```

## # A tibble: 12 x 4
## # Groups:   gender [7]
##   gender      vic_screener     n     p
##   <fct>      <fct>       <int> <dbl>
## 1 Female      No           1266  0.703
## 2 Female      Yes          534   0.297
## 3 Male        No           626   0.918
## 4 Male        Yes          56    0.082
## 5 Transgender woman  No           3   1
## 6 Transgender man   No          2   0.667
## 7 Transgender man   Yes          1   0.333
## 8 Gender queer    No           4   0.5
## 9 Gender queer    Yes          4   0.5
## 10 Other - Please specify No           1   0.5
## 11 Other - Please specify Yes          1   0.5
## 12 Decline to state  No           2   1

```

Sexual assault screener by LGBTQ

```

participant_df %>%
  filter(hours_to_complete <= 24) %>%
  filter(vic_screener == 'Yes' | vic_screener == 'No') %>% # only consider those with data on screener .
  group_by(lgbtq, vic_screener) %>%
  summarise(n = n()) %>%
  ungroup() %>%
  group_by(lgbtq) %>%
  mutate(p = round(n/sum(n), 3))

```

```

## # A tibble: 4 x 4
## # Groups:   lgbtq [2]
##   lgbtq vic_screener     n     p
##   <lgl> <fct>       <int> <dbl>
## 1 FALSE No           1757  0.773
## 2 FALSE Yes          517   0.227
## 3 TRUE  No            147   0.65
## 4 TRUE  Yes           79    0.35

```

Sexual assault screener by Race/Ethnicity

```
participant_df %>%
  filter(hours_to_complete <= 24) %>%
  filter(vic_screener == 'Yes' | vic_screener == 'No') %>% # only consider those with data on screener
  group_by(racial_minority, vic_screener) %>%
  summarise(n = n()) %>%
  ungroup() %>%
  group_by(racial_minority) %>%
  mutate(p = round(n/sum(n), 3))
```

```
## # A tibble: 4 x 4
## # Groups:   racial_minority [2]
##   racial_minority vic_screener     n      p
##             <dbl> <fct>     <int> <dbl>
## 1             0 No         1388 0.748
## 2             0 Yes        468 0.252
## 3             1 No         516 0.801
## 4             1 Yes        128 0.199
```

Sexual assault characteristics (only asked of participants with positive screens)

- Verbal coercion

```
participant_df %>%
  filter(hours_to_complete <= 24) %>%
  filter(vic_screener == 'Yes') %>% # only given to people with positive screens
  summarise(
    mean = mean(vic_worst_verbal, na.rm = T),
    sd = sd(vic_worst_verbal, na.rm = T),
    min = min(vic_worst_verbal, na.rm = T),
    max = max(vic_worst_verbal, na.rm = T),
    n = n(),
    n_missing = sum(is.na(vic_worst_verbal))) %>%
  round(2)

##   mean   sd min max   n n_missing
## 1 1.78 0.76   1   5 596        25
```

- Incapacitation due to substances

```
participant_df %>%
  filter(hours_to_complete <= 24) %>%
  filter(vic_screener == 'Yes') %>% # only given to people with positive screens
  summarise(
    mean = mean(vic_worst_incap, na.rm = T),
    sd = sd(vic_worst_incap, na.rm = T),
    min = min(vic_worst_incap, na.rm = T),
    max = max(vic_worst_incap, na.rm = T),
    n = n(),
    n_missing = sum(is.na(vic_worst_incap))) %>%
  round(2)
```

```
##   mean   sd min max   n n_missing
## 1 1.62 0.94   1    5 596          24
```

- Threatened or actual force

```
participant_df %>%
  filter(hours_to_complete <= 24) %>%
  filter(vic_screener == 'Yes') %>% # only given to people with positive screens
  summarise(
    mean = mean(vic_worst_force, na.rm = T),
    sd = sd(vic_worst_force, na.rm = T),
    min = min(vic_worst_force, na.rm = T),
    max = max(vic_worst_force, na.rm = T),
    n = n(),
    n_missing = sum(is.na(vic_worst_force))) %>%
  round(2)
```

```
##   mean   sd min max   n n_missing
## 1 1.33 0.67   1    5 596          24
```

- History of consent with perpetrator

```
participant_df %>%
  filter(hours_to_complete <= 24) %>%
  filter(vic_screener == 'Yes') %>% # only given to people with positive screens
  group_by(vic_worst_prior_sex) %>%
  summarise(n = n()) %>%
  ungroup() %>%
  mutate(p = round(n/sum(n), 3))
```

```
## # A tibble: 2 x 3
##   vic_worst_prior_sex     n      p
##   <lgl>             <int> <dbl>
## 1 FALSE              338  0.567
## 2 TRUE               258  0.433
```

- Sexual assault acknowledgment

```
participant_df %>%
  filter(hours_to_complete <= 24) %>%
  filter(vic_screener == 'Yes') %>% # only given to people with positive screens
  summarise(
    mean = mean(vic_worst_is_sa, na.rm = T),
    sd = sd(vic_worst_is_sa, na.rm = T),
    min = min(vic_worst_is_sa, na.rm = T),
    max = max(vic_worst_is_sa, na.rm = T),
    n = n(),
    n_missing = sum(is.na(vic_worst_is_sa))) %>%
  round(2)
```

```
##   mean   sd min max   n n_missing
## 1 3.8  2.2   1    7 596          19
```

How many people were there before and after #MeToo?

```
table(df$post_metoo)
```

```
##  
##      0      1  
## 1068 1498
```

```
1498/(1068+1498)
```

```
## [1] 0.583788
```

Semester-by-semester

- Sexual assault prevalence

```
participant_df %>%  
  filter(hours_to_complete <= 24) %>%  
  filter(vic_screener == 'Yes' | vic_screener == 'No') %>% # only consider those with data on screener  
  group_by(semester, vic_screener) %>%  
  summarise(n = n()) %>%  
  ungroup() %>%  
  group_by(semester) %>%  
  mutate(p = round(n/sum(n), 3))
```

```
## # A tibble: 12 x 4  
## # Groups:   semester [6]  
##   semester    vic_screener     n      p  
##   <fct>        <fct>     <int>  <dbl>  
##   1 Fall 2016    No         348  0.777  
##   2 Fall 2016    Yes        100  0.223  
##   3 Spring 2017  No         268  0.766  
##   4 Spring 2017  Yes        82  0.234  
##   5 Fall 2017    No         434  0.768  
##   6 Fall 2017    Yes        131  0.232  
##   7 Spring 2018  No         359  0.765  
##   8 Spring 2018  Yes        110  0.235  
##   9 Fall 2018    No         201  0.753  
##  10 Fall 2018   Yes        66  0.247  
##  11 Spring 2019  No         294  0.733  
##  12 Spring 2019  Yes        107  0.267
```

- Mean acknowledgment

```
participant_df %>%  
  group_by(semester) %>%  
  summarise(  
    n = n(),  
    mean_ack = mean(vic_worst_is_sa, na.rm = T))
```

```

## # A tibble: 6 x 3
##   semester      n  mean_ack
##   <fct>     <int>    <dbl>
## 1 Fall 2016    462     3.35
## 2 Spring 2017  372     3.92
## 3 Fall 2017    587     3.55
## 4 Spring 2018  483     3.95
## 5 Fall 2018    287     4.29
## 6 Spring 2019  425     4.05

```

## Analysis Convenience Functions

The following functions are used to streamline analysis of the Bayesian models in the accompanying paper.

### Missing data imputation

A function for imputing missing data to pass to parallel processing functions.

```

impute_mice <- function(seed, data, m = 25){
  mice(
    data = data,
    m = m,
    print = F,
    seed = seed)
}

```

Unpack mids (multiply imputed complete datasets) objects and reformat as a list that can be passed to later model functions.

```

unpack_parallel_mids <- function(par_mids_list){
  par_mids_list %>%
    map(complete, action = 'long') %>%
    bind_rows(.id = 'core') %>%
    mutate(
      run = group_indices(., core, .imp), # Compute unique imputation datasets
      vic_screener = as.numeric(vic_screener) - 1) %>% # MCMCpack wants numerics
    split(.run)
}

```

### Priors

Calculate the precision parameter (tau - the inverse of the variance) for the normally distributed priors.

```

calc_tau <- function(prior_sd){
  1/(prior_sd**2)
}

```

Count number of priors needed for a given model formula (2 + the number of +'s)

```

n_priors <- function(model_formula){
  pluses <- max(str_count(model_formula, '\\+'))

  if(pluses == 0){
    return(1)
  } else {
    return(pluses + 2)
  }
}

```

## MCMC logistic regression functions

Run a Bayesian logistic regression, given data, models, and a prior.

```

run_logistic_mcmc <- function(df, model, prior_sd, mcmc = 1e4, seed, ...){
  if(missing(seed)){
    seed <- as.integer(1e7*runif(1))
  }
  MCMClogit(
    formula = model,
    data = df,
    b0 = rep(0, n_priors(model)),
    B0 = calc_tau(prior_sd[1:n_priors(model)])*diag(n_priors(model)),
    mcmc = mcmc,
    marginal.likelihood = 'Laplace',
    seed = seed,
    ...)
}

```

Parallelize Bayesian logistic regressions across imputed data sets, then draw random samples from their posteriors.

```

library(MCMCpack)

## Loading required package: coda

## Loading required package: MASS

##
## Attaching package: 'MASS'

## The following object is masked from 'package:dplyr':
##       select

## ##
## ## Markov Chain Monte Carlo Package (MCMCpack)

## ## Copyright (C) 2003-2019 Andrew D. Martin, Kevin M. Quinn, and Jong Hee Park

```

```

## ##
## ## Support provided by the U.S. National Science Foundation

## ## (Grants SES-0350646 and SES-0350613)
## ##

library(snow)

parallel_logistic_mcmc <- function(df_list, model, prior_sd, mcmc = 1e4, ...){
  parallel::mclapply(
    X = df_list,
    FUN = run_logistic_mcmc,
    model = model,
    prior_sd = prior_sd,
    mcmc = mcmc,
    ...) %>%
  map(as.data.frame) %>%
  bind_rows() %>%
  sample_n(size = mcmc, replace = T) %>%
  as.mcmc()
}

parallel_logistic_mcmc2 <- function(df_list, model, prior_sd, mcmc = 1e4, ...){

  cl <- makeCluster(spec = parallel::detectCores(), type='SOCK')

  clusterEvalQ(cl, library(MCMCpack))
  clusterExport(cl, c('n_priors', 'str_count', 'calc_tau', 'as.mcmc'))

  result <- snow::clusterApply(
    cl = cl,
    x = df_list,
    fun = run_logistic_mcmc,
    model = model,
    prior_sd = prior_sd,
    mcmc = mcmc,
    ...) %>%
  map(as.data.frame) %>%
  bind_rows() %>%
  sample_n(size = mcmc, replace = T) %>%
  as.mcmc()

  stopCluster(cl)

  return(result)
}

```

## MCMC linear regression functions

Run a Bayesian linear regression, given data, models, and a prior.

```

run_linear_mcmc <- function(df, model, prior_sd, mcmc = 1e4, seed, ...){
  if(missing(seed)){
    seed <- as.integer(1e7*runif(1))
  }

  MCMCregress(
    formula = model,
    data = df,
    b0 = rep(0, n_priors(model)),
    B0 = calc_tau(prior_sd[1:n_priors(model)])*diag(n_priors(model)),
    mcmc = mcmc,
    marginal.likelihood = 'Chib95',
    seed = seed,
    ...)
}

}

```

Parallelize Bayesian linear regressions across imputed data sets, then draw random samples from their posteriors.

```

parallel_linear_mcmc <- function(df_list, model, prior_sd, mcmc = 1e4, ...){
  parallel::mclapply(
    X = df_list,
    FUN = run_linear_mcmc,
    model = model,
    prior_sd = prior_sd,
    mcmc = mcmc,
    ...) %>%
  map(as.data.frame) %>%
  bind_rows() %>%
  sample_n(size = mcmc, replace = T) %>%
  as.mcmc()
}

parallel_linear_mcmc2 <- function(df_list, model, prior_sd, mcmc = 1e4, ...){

  cl <- makeCluster(spec = parallel::detectCores(), type='SOCK')

  clusterEvalQ(cl, library(MCMCpack))
  clusterExport(cl, c('n_priors', 'str_count', 'calc_tau', 'as.mcmc'))

  result <- snow::clusterApply(
    cl = cl,
    x = df_list,
    fun = run_linear_mcmc,
    model = model,
    prior_sd = prior_sd,
    mcmc = mcmc,
    ...) %>%
  map(as.data.frame) %>%
  bind_rows() %>%
  sample_n(size = mcmc, replace = T) %>%
  as.mcmc()
}

```

```

    stopCluster(cl)

    return(result)
}

```

## Multiple Imputation

Missing data were handled using multiple imputation, implemented by the `mice` package. To ensure `mice` handled the dataset correctly, we first recast all categorical variables as factors.

Note, two imputed datasets were created: one for all participants and one specifically for participants who screened positive for history of sexual assault. This was done because some analyses only involved the subset of participants who screened positive for a history of sexual assault.

```

library(mice)

categorical_vars <- c( 'post_metoo', 'male', 'lgbtq', 'racial_minority',
                      'vic_screener', 'vic_worst_prior_sex')

df[, categorical_vars] <- purrr::map(df[, categorical_vars], as.factor)

imp_seeds <- map_int(1:parallel::detectCores(), ~ as.integer(runif(1)*1e8))
sets_per_core <- ceiling(100/parallel::detectCores())

full_imp <- parallel::mclapply(
  X = imp_seeds,
  FUN = impute_mice,
  data = df,
  m = sets_per_core) %>%
  unpack_parallel_mids()

vic_imp <- parallel::mclapply(
  X = imp_seeds,
  FUN = impute_mice,
  data = filter(df, vic_screener == 'Yes'),
  m = sets_per_core) %>%
  unpack_parallel_mids()

paste(sets_per_core, 'imputed datasets per core')

## [1] "7 imputed datasets per core"

paste(parallel::detectCores(), 'cores detected')

## [1] "16 cores detected"

```

## Models and Priors

These are the models and priors used throughout the rest of the analysis. Note that in the paper, we clarify that the priors are all normally distributed, with mean = 0 and SD = 10. The attentive reader will thus

observe that we didn't need to code all of our priors individually, if they are all the same. However, we did this so that if the code needed to be changed later, it would be easy to modify.

## Constant prevalence hypothesis

**H1:** There are no changes in behavioral reports of sexual assault over time.

We begin by establishing the models to test.

```
h1_models <- list(
  intercept_only = formula(vic_screener ~ 1),
  demos = formula(vic_screener ~ male + lgbtq + racial_minority),
  demos_and_time = formula(vic_screener ~ male + lgbtq + racial_minority + week_num),
  demos_and_time_SQ = formula(
    vic_screener ~ male + lgbtq + racial_minority + week_num + week_num_SQ),
  post_metoo = formula(
    vic_screener ~ male + lgbtq + racial_minority + week_num + week_num_SQ +
    post_metoo),
  post_metoo_time = formula(
    vic_screener ~ male + lgbtq + racial_minority + week_num +
    week_num_SQ + post_metoo + post_metoo_week),
  post_metoo_time_SQ = formula(
    vic_screener ~ male + lgbtq + racial_minority + week_num + week_num_SQ +
    post_metoo + post_metoo_week + post_metoo_week_SQ))
```

Then identify uninformative priors, with SDs = 10.

```
h1_priors <- c(
  intercept = 10,
  male = 10,
  lgbtq = 10,
  racial_minority = 10,
  week_num = 10,
  week_num_SQ = 10,
  post_metoo = 10,
  post_metoo_week = 10,
  post_metoo_week_SQ = 10)
```

## Increasing acknowledgment hypothesis

**H2:** A reliable increase will be observed in victims' acknowledgment of their experiences as 'sexual assault.'

```
h2_models <- list(
  intercept_only = formula(vic_worst_is_sa ~ 1),
  demos = formula(vic_worst_is_sa ~ male + lgbtq + racial_minority),
  demos_and_context = formula(vic_worst_is_sa ~ male + lgbtq + racial_minority +
    vic_worst_prior_sex + vic_worst_verbal + vic_worst_incap +
    vic_worst_force),
  time = formula(vic_worst_is_sa ~ male + lgbtq + racial_minority +
    vic_worst_prior_sex + vic_worst_verbal + vic_worst_incap +
    vic_worst_force +
    week_num),
```

```

time_SQ = formula(vic_worst_is_sa ~ male + lgbtq + racial_minority +
  vic_worst_prior_sex + vic_worst_verbal + vic_worst_incap +
  vic_worst_force +
  week_num + week_num_SQ),
post_metoo = formula(vic_worst_is_sa ~ male + lgbtq + racial_minority +
  vic_worst_prior_sex + vic_worst_verbal + vic_worst_incap +
  vic_worst_force +
  week_num + week_num_SQ +
  post_metoo),
post_metoo_time = formula(vic_worst_is_sa ~ male + lgbtq + racial_minority +
  vic_worst_prior_sex + vic_worst_verbal + vic_worst_incap +
  vic_worst_force +
  week_num + week_num_SQ +
  post_metoo + post_metoo_week),
post_metoo_time_SQ = formula(vic_worst_is_sa ~ male + lgbtq + racial_minority +
  vic_worst_prior_sex + vic_worst_verbal + vic_worst_incap +
  vic_worst_force +
  week_num + week_num_SQ +
  post_metoo + post_metoo_week + post_metoo_week_SQ))

```

We use the same priors as above.

```

h2_priors <- c(
  intercept = 10,
  male = 10,
  lgbtq = 10,
  racial_minority = 10,
  vic_worst_prior_sex = 10,
  vic_worst_verbal = 10,
  vic_worst_incap = 10,
  vic_worst_force = 10,
  week_num = 10,
  week_num_SQ = 10,
  post_metoo = 10,
  post_metoo_week = 10,
  post_metoo_week_SQ = 10)

```

## Model Estimation and Comparison

### Hypothesis 1 - Constant Prevalence

Bayes factors are calculated with complete data because MCMCpack requires this.

```

library(MCMCpack)

h1_comp_post <- map(
  .x = h1_models,
  .f = ~ run_logistic_mcmc(
    df = df %>%
      mutate(vic_screener = as.numeric(vic_screener) - 1),

```

```
model = .x,
prior_sd = h1_priors))
```

Estimate the posterior distribution of complete data (`comp_post`), then compute Bayes factors comparing each subsequent model.

```
h1_bayes_factors <- h1_comp_post %>%
  map_dbl(attr, which = 'logmarglike') %>%
  data.frame(model = names(.), logmarglike = ., row.names = 1:length(.)) %>%
  mutate(
    model = as.character(model),
    BF = exp(logmarglike - lag(logmarglike))) %>%
  map_if(is.numeric, formatC, digits = 5, format = 'f', big.mark = ',',') %>%
  as.data.frame()

h1_bayes_factors
```

model	logmarglike	BF
## 1	intercept_only	-1,378.43983
## 2	demos	-1,306.38387
## 3	demos_and_time	-1,313.92861
## 4	demos_and_time_SQ	-1,326.41575
## 5	post_metoo	-1,328.89719
## 6	post_metoo_time	-1,336.21497
## 7	post_metoo_time_SQ	-1,346.57343
## 1		NA
## 2	19,656,419,195,564,889,022,319,527,198,720.00000	
## 3		0.00053
## 4		0.00000
## 5		0.08362
## 6		0.00066
## 7		0.00003

Estimate the posterior distributions implied by each imputed dataset, then splice them. Perform this process for each model.

```
library(tidyverse)
library(MCMCpack)

h1_posteriors <- map(
  .x = h1_models,
  .f = ~ parallel_logistic_mcmc(
    df_list = full_imp,
    model = .x,
    prior_sd = h1_priors,
    mcmc = 6*1e4,
    thin = 2))

# Save to CSV
h1_posteriors %>%
  map(as.data.frame) %>%
```

```

bind_rows(.id = 'model') %>%
gather(param, estimate, -model) %>%
filter(complete.cases(.)) %>%
write.csv(file = paste0('h1_posteriors_', Sys.Date(), '.csv'), row.names = F)

```

Summaries of the posteriors

```
map(h1_posteriors, summary)
```

```

## $intercept_only
##
## Iterations = 1:60000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 60000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##           Mean        SD   Naive SE Time-series SE
## -1.172e+00 6.064e-03 2.476e-05 2.476e-05
##
## 2. Quantiles for each variable:
##
##    2.5%    25%    50%    75%  97.5%
## -1.185 -1.176 -1.174 -1.170 -1.159
##
##
## $demos
##
## Iterations = 1:60000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 60000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##           Mean        SD   Naive SE Time-series SE
## (Intercept) -0.8698 0.05949 0.0002429 0.0002429
## maleTRUE     -1.5570 0.15006 0.0006126 0.0006126
## lgbtqTRUE     0.6382 0.15487 0.0006323 0.0006323
## racial_minority1 -0.2963 0.11602 0.0004737 0.0004737
##
## 2. Quantiles for each variable:
##
##           2.5%    25%    50%    75%  97.5%
## (Intercept) -0.9873 -0.9096 -0.8694 -0.8298 -0.75330
## maleTRUE     -1.8573 -1.6571 -1.5553 -1.4537 -1.26933
## lgbtqTRUE      0.3309  0.5355  0.6389  0.7425  0.94003
## racial_minority1 -0.5258 -0.3744 -0.2954 -0.2181 -0.07174
##
## 
## 

```

```

## $demos_and_time
##
## Iterations = 1:60000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 60000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##              Mean        SD  Naive SE Time-series SE
## (Intercept) -0.886736 0.060148 2.456e-04      2.441e-04
## maleTRUE     -1.561754 0.148697 6.071e-04      6.071e-04
## lgbtqTRUE     0.636865 0.154759 6.318e-04      6.318e-04
## racial_minority1 -0.304527 0.115571 4.718e-04      4.718e-04
## week_num      0.001993 0.001149 4.692e-06      4.763e-06
##
## 2. Quantiles for each variable:
##
##              2.5%       25%       50%       75%       97.5%
## (Intercept) -1.0054211 -0.927050 -0.886553 -0.845901 -0.769915
## maleTRUE     -1.8600301 -1.660887 -1.558806 -1.460145 -1.276943
## lgbtqTRUE     0.3297768  0.532910  0.637834  0.741181  0.936027
## racial_minority1 -0.5341765 -0.382223 -0.303737 -0.226100 -0.080856
## week_num      -0.0002546  0.001216  0.001997  0.002765  0.004245
##
## $demos_and_time_SQ
##
## Iterations = 1:60000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 60000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##              Mean        SD  Naive SE Time-series SE
## (Intercept) -9.019e-01 7.743e-02 3.161e-04      3.183e-04
## maleTRUE     -1.564e+00 1.496e-01 6.107e-04      6.107e-04
## lgbtqTRUE     6.378e-01 1.548e-01 6.320e-04      6.320e-04
## racial_minority1 -3.048e-01 1.160e-01 4.736e-04      4.736e-04
## week_num      1.758e-03 1.357e-03 5.541e-06      5.541e-06
## week_num_SQ    9.147e-06 2.987e-05 1.219e-07      1.219e-07
##
## 2. Quantiles for each variable:
##
##              2.5%       25%       50%       75%       97.5%
## (Intercept) -1.054e+00 -9.541e-01 -9.017e-01 -8.495e-01 -7.519e-01
## maleTRUE     -1.865e+00 -1.663e+00 -1.561e+00 -1.461e+00 -1.276e+00
## lgbtqTRUE     3.312e-01 5.338e-01 6.380e-01 7.424e-01 9.396e-01
## racial_minority1 -5.335e-01 -3.832e-01 -3.044e-01 -2.263e-01 -7.855e-02
## week_num     -8.882e-04 8.456e-04 1.754e-03 2.670e-03 4.439e-03
## week_num_SQ   -4.919e-05 -1.102e-05 9.297e-06 2.931e-05 6.735e-05

```

```

##  

##  

## $post_metoo  

##  

## Iterations = 1:60000  

## Thinning interval = 1  

## Number of chains = 1  

## Sample size per chain = 60000  

##  

## 1. Empirical mean and standard deviation for each variable,  

##     plus standard error of the mean:  

##  

##  

##               Mean        SD  Naive SE Time-series SE  

## (Intercept) -0.7025251 1.380e-01 5.632e-04    5.571e-04  

## maleTRUE     -1.5589399 1.494e-01 6.098e-04    6.098e-04  

## lgbtqTRUE     0.6367577 1.547e-01 6.315e-04    6.315e-04  

## racial_minority1 -0.3067300 1.164e-01 4.754e-04    4.754e-04  

## week_num      0.0055338 2.562e-03 1.046e-05   1.039e-05  

## week_num_SQ    -0.0000139 3.286e-05 1.342e-07   1.340e-07  

## post_metoo1   -0.3329885 1.910e-01 7.799e-04    7.584e-04  

##  

## 2. Quantiles for each variable:  

##  

##               2.5%       25%       50%       75%      97.5%  

## (Intercept) -9.758e-01 -7.949e-01 -7.023e-01 -6.086e-01 -4.336e-01  

## maleTRUE     -1.856e+00 -1.659e+00 -1.556e+00 -1.457e+00 -1.273e+00  

## lgbtqTRUE     3.294e-01  5.332e-01  6.376e-01  7.417e-01  9.371e-01  

## racial_minority1 -5.375e-01 -3.848e-01 -3.061e-01 -2.274e-01 -8.095e-02  

## week_num      5.208e-04  3.791e-03  5.530e-03  7.253e-03  1.057e-02  

## week_num_SQ    -7.858e-05 -3.599e-05 -1.407e-05  8.474e-06  5.021e-05  

## post_metoo1   -7.078e-01 -4.624e-01 -3.317e-01 -2.044e-01  4.069e-02  

##  

##  

## $post_metoo_time  

##  

## Iterations = 1:60000  

## Thinning interval = 1  

## Number of chains = 1  

## Sample size per chain = 60000  

##  

## 1. Empirical mean and standard deviation for each variable,  

##     plus standard error of the mean:  

##  

##               Mean        SD  Naive SE Time-series SE  

## (Intercept) -6.922e-01 0.1569546 6.408e-04    6.408e-04  

## maleTRUE     -1.562e+00 0.1502724 6.135e-04    6.135e-04  

## lgbtqTRUE     6.363e-01 0.1547698 6.318e-04    6.318e-04  

## racial_minority1 -3.073e-01 0.1161208 4.741e-04    4.741e-04  

## week_num      6.416e-03 0.0076420 3.120e-05   3.120e-05  

## week_num_SQ    -2.053e-07 0.0001171 4.781e-07   4.781e-07  

## post_metoo1   -3.299e-01 0.1943278 7.933e-04    7.933e-04  

## post_metoo_week -2.032e-03 0.0167645 6.844e-05    6.844e-05  

##  

## 2. Quantiles for each variable:  


```

```

##          2.5%      25%      50%      75%    97.5%
## (Intercept) -1.0040027 -7.973e-01 -6.918e-01 -5.864e-01 -0.38485
## maleTRUE     -1.8635482 -1.662e+00 -1.561e+00 -1.459e+00 -1.27532
## lgbtqTRUE     0.3316921  5.327e-01  6.357e-01  7.418e-01  0.93651
## racial_minority1 -0.5337385 -3.853e-01 -3.069e-01 -2.284e-01 -0.08207
## week_num     -0.0085259  1.257e-03  6.462e-03  1.156e-02  0.02148
## week_num_SQ   -0.0002298 -7.869e-05  2.343e-07  7.826e-05  0.00023
## post_metoo1   -0.7119947 -4.597e-01 -3.293e-01 -1.989e-01  0.05012
## post_metoo_week -0.0350430 -1.322e-02 -2.114e-03  9.196e-03  0.03077
##
##
## $post_metoo_time_SQ
##
## Iterations = 1:60000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 60000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##          Mean        SD  Naive SE Time-series SE
## (Intercept) -0.7274405 0.1860037 7.594e-04 7.594e-04
## maleTRUE     -1.5615981 0.1500449 6.126e-04 6.088e-04
## lgbtqTRUE     0.6350938 0.1553620 6.343e-04 6.343e-04
## racial_minority1 -0.3079653 0.1167400 4.766e-04 4.766e-04
## week_num      0.0016845 0.0151256 6.175e-05 6.175e-05
## week_num_SQ   -0.0000857 0.0002623 1.071e-06 1.071e-06
## post_metoo1   -0.2738007 0.2498288 1.020e-03 1.012e-03
## post_metoo_week 0.0009877 0.0187768 7.666e-05 7.666e-05
## post_metoo_week_SQ 0.0001061 0.0002935 1.198e-06 1.198e-06
##
## 2. Quantiles for each variable:
##
##          2.5%      25%      50%      75%    97.5%
## (Intercept) -1.0982963 -8.523e-01 -0.7258182 -6.008e-01 -0.3681649
## maleTRUE     -1.8615089 -1.662e+00 -1.5597026 -1.458e+00 -1.2751076
## lgbtqTRUE     0.3298350  5.309e-01  0.6358265  7.395e-01  0.9398068
## racial_minority1 -0.5377022 -3.869e-01 -0.3076681 -2.286e-01 -0.0802777
## week_num     -0.0283175 -8.451e-03  0.0017351  1.191e-02  0.0309030
## week_num_SQ   -0.0006032 -2.609e-04 -0.0000840  9.032e-05  0.0004255
## post_metoo1   -0.7624513 -4.424e-01 -0.2733351 -1.067e-01  0.2144674
## post_metoo_week -0.0352580 -1.176e-02  0.0009742  1.350e-02  0.0382627
## post_metoo_week_SQ -0.0004649 -9.285e-05  0.0001060  3.027e-04  0.0006873

h1_param_means <- map(h1_comp_post, colMeans) %>%
  map(data.frame) %>%
  map(~ set_names(.x, 'mean_est')) %>%
  map(~ mutate(.data = .x, param = rownames(.x))) %>%
  bind_rows(.id = 'model') %>%
  mutate(original_order = row_number())

h1_hpdi_97 <- map(h1_comp_post, HPDinterval, prob = .97) %>%

```

```

map(data.frame) %>%
  map(~ mutate(.data = .x, param = rownames(.x))) %>%
  bind_rows(.id = 'model') %>%
  dplyr::rename(lower_97 = lower, upper_97 = upper)

h1_hpdi_95 <- map(h1_comp_post, HPDinterval, prob = .95) %>%
  map(data.frame) %>%
  map(~ mutate(.data = .x, param = rownames(.x))) %>%
  bind_rows(.id = 'model') %>%
  dplyr::rename(lower_95 = lower, upper_95 = upper)

h1_params <- h1_param_means %>%
  left_join(h1_hpdi_97, by = c('model', 'param')) %>%
  left_join(h1_hpdi_95, by = c('model', 'param'))

h1_params

```

	model	mean_est	param	original_order
## 1	intercept_only	-1.161472e+00	(Intercept)	1
## 2	demos	-8.605255e-01	(Intercept)	2
## 3	demos	-1.557825e+00	maleTRUE	3
## 4	demos	6.510994e-01	lgbtqTRUE	4
## 5	demos	-2.889889e-01	racial_minority1	5
## 6	demos_and_time	-8.751223e-01	(Intercept)	6
## 7	demos_and_time	-1.570328e+00	maleTRUE	7
## 8	demos_and_time	6.545828e-01	lgbtqTRUE	8
## 9	demos_and_time	-3.080189e-01	racial_minority1	9
## 10	demos_and_time	2.063230e-03	week_num	10
## 11	demos_and_time_SQ	-8.913287e-01	(Intercept)	11
## 12	demos_and_time_SQ	-1.565685e+00	maleTRUE	12
## 13	demos_and_time_SQ	6.421851e-01	lgbtqTRUE	13
## 14	demos_and_time_SQ	-3.005697e-01	racial_minority1	14
## 15	demos_and_time_SQ	1.799801e-03	week_num	15
## 16	demos_and_time_SQ	9.387570e-06	week_num_SQ	16
## 17	post_metoo	-6.850868e-01	(Intercept)	17
## 18	post_metoo	-1.555198e+00	maleTRUE	18
## 19	post_metoo	6.430341e-01	lgbtqTRUE	19
## 20	post_metoo	-3.067355e-01	racial_minority1	20
## 21	post_metoo	5.582406e-03	week_num	21
## 22	post_metoo	-1.400851e-05	week_num_SQ	22
## 23	post_metoo	-3.422851e-01	post_metoo1	23
## 24	post_metoo_time	-6.703347e-01	(Intercept)	24
## 25	post_metoo_time	-1.566750e+00	maleTRUE	25
## 26	post_metoo_time	6.389077e-01	lgbtqTRUE	26
## 27	post_metoo_time	-3.135895e-01	racial_minority1	27
## 28	post_metoo_time	6.746423e-03	week_num	28
## 29	post_metoo_time	3.708741e-09	week_num_SQ	29
## 30	post_metoo_time	-3.383958e-01	post_metoo1	30
## 31	post_metoo_time	-2.442531e-03	post_metoo_week	31
## 32	post_metoo_time_SQ	-7.212757e-01	(Intercept)	32
## 33	post_metoo_time_SQ	-1.563486e+00	maleTRUE	33
## 34	post_metoo_time_SQ	6.589805e-01	lgbtqTRUE	34
## 35	post_metoo_time_SQ	-3.113713e-01	racial_minority1	35

```

## 36 post_metoo_time_SQ 1.697278e-03 week_num 36
## 37 post_metoo_time_SQ -8.519496e-05 week_num_SQ 37
## 38 post_metoo_time_SQ -2.696796e-01 post_metoo1 38
## 39 post_metoo_time_SQ 1.040243e-03 post_metoo_week 39
## 40 post_metoo_time_SQ 1.029846e-04 post_metoo_week_SQ 40
## lower_97 upper_97 lower_95 upper_95
## 1 -1.161472e+00 -1.161472e+00 -1.161472e+00 -1.161472e+00
## 2 -9.823657e-01 -7.240646e-01 -9.729835e-01 -7.419398e-01
## 3 -1.887389e+00 -1.226621e+00 -1.858027e+00 -1.257161e+00
## 4 3.099024e-01 9.758841e-01 3.672192e-01 9.621820e-01
## 5 -5.367332e-01 -1.967686e-02 -5.315866e-01 -6.160924e-02
## 6 -1.008214e+00 -7.518152e-01 -9.918802e-01 -7.608503e-01
## 7 -1.894559e+00 -1.268462e+00 -1.842552e+00 -1.283482e+00
## 8 3.403698e-01 1.005535e+00 3.403698e-01 9.351123e-01
## 9 -5.663265e-01 -6.099349e-02 -5.362572e-01 -8.379391e-02
## 10 -1.897394e-04 4.707469e-03 -4.215562e-06 4.320673e-03
## 11 -1.064814e+00 -7.269613e-01 -1.044271e+00 -7.365631e-01
## 12 -1.902394e+00 -1.266427e+00 -1.867637e+00 -1.300179e+00
## 13 3.391574e-01 1.002241e+00 3.391574e-01 9.406148e-01
## 14 -5.650169e-01 -6.584165e-02 -5.247370e-01 -8.040466e-02
## 15 -9.939577e-04 4.811566e-03 -9.919030e-04 4.253359e-03
## 16 -5.797015e-05 7.371878e-05 -4.608183e-05 7.200555e-05
## 17 -9.859085e-01 -3.888784e-01 -9.415601e-01 -4.008513e-01
## 18 -1.880643e+00 -1.248171e+00 -1.860183e+00 -1.291462e+00
## 19 2.949259e-01 9.655723e-01 3.212118e-01 9.231800e-01
## 20 -5.892825e-01 -6.431702e-02 -5.325179e-01 -6.639377e-02
## 21 9.456724e-05 1.069464e-02 7.844936e-04 1.060286e-02
## 22 -7.979668e-05 5.949311e-05 -7.403035e-05 5.136987e-05
## 23 -7.255369e-01 8.747438e-02 -7.048203e-01 2.619543e-02
## 24 -1.030885e+00 -3.094063e-01 -1.022198e+00 -3.631725e-01
## 25 -1.882985e+00 -1.221086e+00 -1.856603e+00 -1.264509e+00
## 26 2.951389e-01 9.614669e-01 3.469794e-01 9.600234e-01
## 27 -5.596408e-01 -5.596697e-02 -5.467531e-01 -9.114183e-02
## 28 -1.174790e-02 2.338161e-02 -1.112278e-02 2.095019e-02
## 29 -2.727531e-04 2.644408e-04 -2.546684e-04 2.265615e-04
## 30 -7.619915e-01 1.072312e-01 -7.508691e-01 3.697182e-02
## 31 -3.929453e-02 3.718791e-02 -3.241094e-02 3.718791e-02
## 32 -1.122691e+00 -2.876781e-01 -1.093877e+00 -3.407429e-01
## 33 -1.873238e+00 -1.213594e+00 -1.861249e+00 -1.278610e+00
## 34 3.326345e-01 9.909227e-01 3.319415e-01 9.371740e-01
## 35 -5.434988e-01 -3.242927e-02 -5.498439e-01 -8.975239e-02
## 36 -3.018353e-02 3.635220e-02 -2.710022e-02 3.424042e-02
## 37 -6.441850e-04 5.223892e-04 -6.352473e-04 4.470427e-04
## 38 -8.506460e-01 2.715678e-01 -8.577533e-01 1.765016e-01
## 39 -4.521653e-02 3.950017e-02 -4.002423e-02 3.516989e-02
## 40 -5.819643e-04 6.927084e-04 -4.501488e-04 6.833069e-04

```

```
write.csv(h1_params, paste0('h1_params_', Sys.Date(), '.csv'), row.names = F)
```

## Hypothesis 2 - Increasing Acknowledgment

Bayes factors are calculated with complete data because MCMCpack requires this.

```

h2_comp_post <- map(
  .x = h2_models,
  .f = ~ run_linear_mcmc(
    df = df,
    model = .x,
    prior_sd = h2_priors))

```

Estimate the posterior distribution of complete data (`comp_post`), then compute Bayes factors comparing each subsequent model.

```

h2_bayes_factors <- h2_comp_post %>%
  map_dbl(attr, which = 'logmarglike') %>%
  data.frame(model = names(.), logmarglike = ., row.names = 1:length(.)) %>%
  mutate(
    model = as.character(model),
    BF = exp(logmarglike - lag(logmarglike))) %>%
  map_if(is.numeric, formatC, digits = 5, format = 'f', big.mark = ',',) %>%
  as.data.frame()

h2_bayes_factors

##           model logmarglike
## 1   intercept_only -1,288.69005
## 2           demos -1,288.65833
## 3  demos_and_context -1,210.58796
## 4            time -1,215.35232
## 5        time_SQ -1,227.55852
## 6      post_metoo -1,231.02514
## 7  post_metoo_time -1,236.78897
## 8 post_metoo_time_SQ -1,245.38177
##                               BF
## 1                               NA
## 2                         1.03222
## 3 8,045,103,140,013,963,422,936,160,087,834,624.00000
## 4                           0.00853
## 5                           0.00000
## 6                           0.03122
## 7                           0.00314
## 8                           0.00019

```

Estimate the posterior distributions implied by each imputed dataset, then splice them. Perform this process for each model.

```

set.seed(3019301)

h2_posteriors <- map(
  .x = h2_models,
  .f = ~ parallel_linear_mcmc2(
    df_list = vic_imp,
    model = .x,
    prior_sd = h2_priors,
    mcmc = 1e4))

```

```

# Save to CSV
h2_posteriors %>%
  map(as.data.frame) %>%
  bind_rows(.id = 'model') %>%
  gather(param, estimate, -model) %>%
  filter(complete.cases(.)) %>%
  write.csv(file = paste0('h2_posteriors_', Sys.Date(), '.csv'), row.names = F)

```

Summaries of the posteriors

```
map(h2_posteriors, summary)
```

```

## $intercept_only
##
## Iterations = 1:10000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 10000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##           Mean      SD Naive SE Time-series SE
## (Intercept) 3.810 0.0920 0.000920      0.000920
## sigma2      4.884 0.2843 0.002843      0.002843
##
## 2. Quantiles for each variable:
##
##           2.5%   25%   50%   75% 97.5%
## (Intercept) 3.631 3.748 3.810 3.871 3.992
## sigma2      4.366 4.690 4.868 5.068 5.476
##
## $demos
##
## Iterations = 1:10000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 10000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##           Mean      SD Naive SE Time-series SE
## (Intercept) 3.73367 0.1119 0.001119      0.001119
## maleTRUE     -0.73898 0.3226 0.003226      0.003226
## lgbtqTRUE     1.11969 0.2672 0.002672      0.002672
## racial_minority1 -0.02818 0.2196 0.002196      0.002196
## sigma2       4.72537 0.2797 0.002797      0.002797
##
## 2. Quantiles for each variable:
##

```

```

##          2.5%    25%    50%    75%   97.5%
## (Intercept) 3.5141  3.6573  3.73347 3.8094  3.9518
## maleTRUE     -1.3698 -0.9570 -0.73847 -0.5239 -0.1186
## lgbtqTRUE    0.5993  0.9424  1.11589  1.3011  1.6389
## racial_minority1 -0.4521 -0.1811 -0.03054  0.1235  0.3973
## sigma2       4.2176  4.5291  4.71413  4.9111  5.2981
##
##
## $demos_and_context
##
## Iterations = 1:10000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 10000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##           Mean        SD  Naive SE Time-series SE
## (Intercept) 1.2182 0.29127 0.0029127      0.0029606
## maleTRUE    -0.6511 0.28096 0.0028096      0.0028096
## lgbtqTRUE    0.9348 0.23363 0.0023363      0.0023363
## racial_minority1 -0.1499 0.19329 0.0019329      0.0020017
## vic_worst_prior_sexTRUE -0.6714 0.16757 0.0016757      0.0016757
## vic_worst_verbal    0.2473 0.11560 0.0011560      0.0012255
## vic_worst_incap     0.5306 0.08886 0.0008886      0.0008886
## vic_worst_force     1.1621 0.12636 0.0012636      0.0013073
## sigma2      3.5703 0.21176 0.0021176      0.0021176
##
## 2. Quantiles for each variable:
##
##          2.5%    25%    50%    75%   97.5%
## (Intercept) 0.6502  1.0200  1.2140  1.41743  1.7885
## maleTRUE     -1.2015 -0.8417 -0.6469 -0.45927 -0.1047
## lgbtqTRUE    0.4770  0.7761  0.9345  1.08773  1.3956
## racial_minority1 -0.5302 -0.2799 -0.1498 -0.02163  0.2252
## vic_worst_prior_sexTRUE -1.0026 -0.7826 -0.6726 -0.55971 -0.3459
## vic_worst_verbal    0.0190  0.1720  0.2472  0.32416  0.4745
## vic_worst_incap     0.3577  0.4697  0.5312  0.59114  0.7034
## vic_worst_force     0.9185  1.0740  1.1618  1.24897  1.4111
## sigma2      3.1850  3.4248  3.5624  3.70528  4.0073
##
##
## $time
##
## Iterations = 1:10000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 10000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##           Mean        SD  Naive SE Time-series SE

```

```

## (Intercept) 1.185304 0.290484 2.905e-03 2.905e-03
## maleTRUE -0.668908 0.276581 2.766e-03 2.766e-03
## lgbtqTRUE 0.901318 0.233456 2.335e-03 2.335e-03
## racial_minority1 -0.178364 0.192854 1.929e-03 1.887e-03
## vic_worst_prior_sexTRUE -0.681519 0.165185 1.652e-03 1.652e-03
## vic_worst_verbal 0.234801 0.115111 1.151e-03 1.172e-03
## vic_worst_incap 0.535375 0.087179 8.718e-04 8.718e-04
## vic_worst_force 1.167349 0.127045 1.270e-03 1.270e-03
## week_num 0.005179 0.001853 1.853e-05 1.853e-05
## sigma2 3.528974 0.210354 2.104e-03 2.104e-03
##
## 2. Quantiles for each variable:
##
##          2.5%    25%    50%    75%   97.5%
## (Intercept) 0.621593 0.986956 1.185379 1.384512 1.748759
## maleTRUE -1.217587 -0.854935 -0.668950 -0.481599 -0.131860
## lgbtqTRUE 0.444952 0.740138 0.900976 1.060580 1.356137
## racial_minority1 -0.559032 -0.312014 -0.176412 -0.046850 0.199216
## vic_worst_prior_sexTRUE -1.008533 -0.792325 -0.682452 -0.570487 -0.362859
## vic_worst_verbal 0.008912 0.156759 0.236692 0.313168 0.458661
## vic_worst_incap 0.362322 0.476971 0.536518 0.592584 0.707294
## vic_worst_force 0.916142 1.081446 1.167293 1.253914 1.415760
## week_num 0.001560 0.003892 0.005172 0.006442 0.008831
## sigma2 3.142547 3.382056 3.519674 3.665954 3.961080
##
## 
## $time_SQ
##
## Iterations = 1:10000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 10000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##          Mean        SD  Naive SE Time-series SE
## (Intercept) 1.1737414 3.031e-01 3.031e-03 3.031e-03
## maleTRUE -0.6693122 2.767e-01 2.767e-03 2.767e-03
## lgbtqTRUE 0.9030929 2.345e-01 2.345e-03 2.345e-03
## racial_minority1 -0.1843849 1.937e-01 1.937e-03 1.937e-03
## vic_worst_prior_sexTRUE -0.6815818 1.677e-01 1.677e-03 1.677e-03
## vic_worst_verbal 0.2329343 1.152e-01 1.152e-03 1.152e-03
## vic_worst_incap 0.5342882 8.775e-02 8.775e-04 8.932e-04
## vic_worst_force 1.1669589 1.265e-01 1.265e-03 1.265e-03
## week_num 0.0048456 2.241e-03 2.241e-05 2.241e-05
## week_num_SQ 0.0000116 4.861e-05 4.861e-07 4.861e-07
## sigma2 3.5365455 2.091e-01 2.091e-03 2.091e-03
##
## 2. Quantiles for each variable:
##
##          2.5%    25%    50%    75%
## (Intercept) 5.806e-01 9.705e-01 1.179e+00 1.373e+00
## maleTRUE -1.206e+00 -8.573e-01 -6.696e-01 -4.834e-01

```

```

## lgbtqTRUE          4.478e-01  7.458e-01  9.016e-01  1.062e+00
## racial_minority1 -5.718e-01 -3.130e-01 -1.821e-01 -5.146e-02
## vic_worst_prior_sexTRUE -1.008e+00 -7.945e-01 -6.810e-01 -5.675e-01
## vic_worst_verbal    6.135e-03  1.572e-01  2.322e-01  3.098e-01
## vic_worst_incap     3.620e-01  4.757e-01  5.343e-01  5.930e-01
## vic_worst_force     9.184e-01  1.082e+00  1.167e+00  1.252e+00
## week_num            3.928e-04  3.326e-03  4.851e-03  6.364e-03
## week_num_SQ          -8.441e-05 -2.106e-05  1.117e-05  4.411e-05
## sigma2              3.150e+00  3.389e+00  3.531e+00  3.673e+00
##
##                               97.5%
## (Intercept)                1.7713119
## maleTRUE                  -0.1292070
## lgbtqTRUE                 1.3662450
## racial_minority1           0.1853507
## vic_worst_prior_sexTRUE   -0.3543508
## vic_worst_verbal           0.4582967
## vic_worst_incap             0.7063463
## vic_worst_force             1.4108434
## week_num                   0.0092250
## week_num_SQ                 0.0001075
## sigma2                      3.9683379
##
##
## $post_metoo
##
## Iterations = 1:10000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 10000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##                                Mean        SD  Naive SE Time-series SE
## (Intercept)                1.1436994 3.429e-01 3.429e-03 3.607e-03
## maleTRUE                  -0.6689025 2.794e-01 2.794e-03 2.794e-03
## lgbtqTRUE                 0.9025535 2.340e-01 2.340e-03 2.340e-03
## racial_minority1           -0.1816868 1.919e-01 1.919e-03 1.919e-03
## vic_worst_prior_sexTRUE   -0.6836100 1.666e-01 1.666e-03 1.666e-03
## vic_worst_verbal           0.2355147 1.149e-01 1.149e-03 1.149e-03
## vic_worst_incap             0.5356594 8.860e-02 8.860e-04 8.860e-04
## vic_worst_force             1.1668511 1.268e-01 1.268e-03 1.268e-03
## week_num                   0.0044443 4.189e-03 4.189e-05 4.122e-05
## week_num_SQ                 0.0000148 5.356e-05 5.356e-07 5.514e-07
## post_metoo1                 0.0386708 3.115e-01 3.115e-03 3.041e-03
## sigma2                      3.5414844 2.110e-01 2.110e-03 2.110e-03
##
## 2. Quantiles for each variable:
##
##                                2.5%       25%       50%       75%
## (Intercept)                4.665e-01  9.145e-01  1.141e+00  1.372e+00
## maleTRUE                  -1.204e+00 -8.566e-01 -6.731e-01 -4.827e-01
## lgbtqTRUE                 4.521e-01  7.426e-01  9.030e-01  1.061e+00
## racial_minority1           -5.595e-01 -3.117e-01 -1.813e-01 -5.130e-02

```

```

## vic_worst_prior_sexTRUE -1.004e+00 -7.973e-01 -6.856e-01 -5.752e-01
## vic_worst_verbal      7.489e-03  1.579e-01  2.356e-01  3.126e-01
## vic_worst_incap       3.617e-01  4.760e-01  5.358e-01  5.946e-01
## vic_worst_force       9.177e-01  1.080e+00  1.168e+00  1.253e+00
## week_num              -3.772e-03  1.694e-03  4.490e-03  7.250e-03
## week_num_SQ            -8.928e-05 -2.101e-05  1.474e-05  5.082e-05
## post_metoo1           -5.733e-01 -1.683e-01  3.853e-02  2.438e-01
## sigma2                3.148e+00  3.396e+00  3.532e+00  3.680e+00
##                                         97.5%
## (Intercept)                 1.8113707
## maleTRUE                  -0.1119072
## lgbtqTRUE                  1.3584174
## racial_minority1          0.1884960
## vic_worst_prior_sexTRUE   -0.3484704
## vic_worst_verbal          0.4648910
## vic_worst_incap           0.7098373
## vic_worst_force           1.4138863
## week_num                   0.0126118
## week_num_SQ                0.0001207
## post_metoo1               0.6541857
## sigma2                     3.9775325
##
##
## $post_metoo_time
##
## Iterations = 1:10000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 10000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##                               Mean        SD  Naive SE Time-series SE
## (Intercept)             1.090e+00 0.3624674 3.625e-03  3.468e-03
## maleTRUE                -6.525e-01 0.2837599 2.838e-03  2.838e-03
## lgbtqTRUE                8.956e-01 0.2340318 2.340e-03  2.340e-03
## racial_minority1         -1.750e-01 0.1911202 1.911e-03  1.838e-03
## vic_worst_prior_sexTRUE -6.789e-01 0.1655703 1.656e-03  1.671e-03
## vic_worst_verbal         2.346e-01 0.1148955 1.149e-03  1.149e-03
## vic_worst_incap          5.311e-01 0.0891224 8.912e-04  8.881e-04
## vic_worst_force          1.170e+00 0.1284959 1.285e-03  1.260e-03
## week_num                  -7.865e-04 0.0125081 1.251e-04  1.219e-04
## week_num_SQ                -6.695e-05 0.0001971 1.971e-06  1.971e-06
## post_metoo1               7.988e-03 0.3242601 3.243e-03  3.102e-03
## post_metoo_week           1.207e-02 0.0279267 2.793e-04  2.741e-04
## sigma2                     3.550e+00 0.2101298 2.101e-03  2.101e-03
##
## 2. Quantiles for each variable:
##
##                               2.5%        25%        50%        75%
## (Intercept)             0.3716147  0.8440158  1.0926286  1.335e+00
## maleTRUE                -1.2044304 -0.8432324 -0.6530988 -4.594e-01
## lgbtqTRUE                0.4408227  0.7402224  0.8951325  1.051e+00

```

```

## racial_minority1      -0.5517202 -0.3026533 -0.1735122 -4.649e-02
## vic_worst_prior_sexTRUE -1.0026116 -0.7897890 -0.6800948 -5.641e-01
## vic_worst_verbal      0.0122558  0.1556611  0.2340521  3.126e-01
## vic_worst_incap        0.3544190  0.4727184  0.5320265  5.911e-01
## vic_worst_force        0.9212393  1.0834088  1.1701317  1.256e+00
## week_num                -0.0254703 -0.0091709 -0.0006940  7.607e-03
## week_num_SQ              -0.0004579 -0.0001994 -0.0000649  6.555e-05
## post_metoo1             -0.6377534 -0.2075016  0.0104806  2.244e-01
## post_metoo_week         -0.0426597 -0.0066637  0.0117947  3.089e-02
## sigma2                  3.1547815  3.4052843  3.5421237  3.687e+00
##
##                               97.5%
## (Intercept)                 1.8009966
## maleTRUE                   -0.1003684
## lgbtqTRUE                   1.3692757
## racial_minority1            0.2020823
## vic_worst_prior_sexTRUE    -0.3594371
## vic_worst_verbal            0.4574037
## vic_worst_incap              0.7031699
## vic_worst_force              1.4229761
## week_num                     0.0237959
## week_num_SQ                  0.0003156
## post_metoo1                  0.6437823
## post_metoo_week                0.0669757
## sigma2                      3.9810691
##
##
## $post_metoo_time_SQ
##
## Iterations = 1:10000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 10000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##                               Mean        SD   Naive SE Time-series SE
## (Intercept)                 0.8292088 0.3958148 3.958e-03  3.958e-03
## maleTRUE                   -0.6605150 0.2833032 2.833e-03  2.833e-03
## lgbtqTRUE                   0.8946653 0.2348987 2.349e-03  2.349e-03
## racial_minority1            -0.1833450 0.1928774 1.929e-03  1.839e-03
## vic_worst_prior_sexTRUE    -0.6836718 0.1673858 1.674e-03  1.650e-03
## vic_worst_verbal            0.2440506 0.1142443 1.142e-03  1.101e-03
## vic_worst_incap              0.5333099 0.0883061 8.831e-04  8.831e-04
## vic_worst_force              1.1705309 0.1259183 1.259e-03  1.259e-03
## week_num                     -0.0362118 0.0250081 2.501e-04  2.553e-04
## week_num_SQ                  -0.0007244 0.0004464 4.464e-06  4.544e-06
## post_metoo1                  0.4034064 0.4007564 4.008e-03  4.090e-03
## post_metoo_week                0.0347544 0.0313294 3.133e-04  3.126e-04
## post_metoo_week_SQ            0.0008119 0.0004906 4.906e-06  4.997e-06
## sigma2                      3.5386304 0.2092159 2.092e-03  2.092e-03
##
## 2. Quantiles for each variable:
##

```

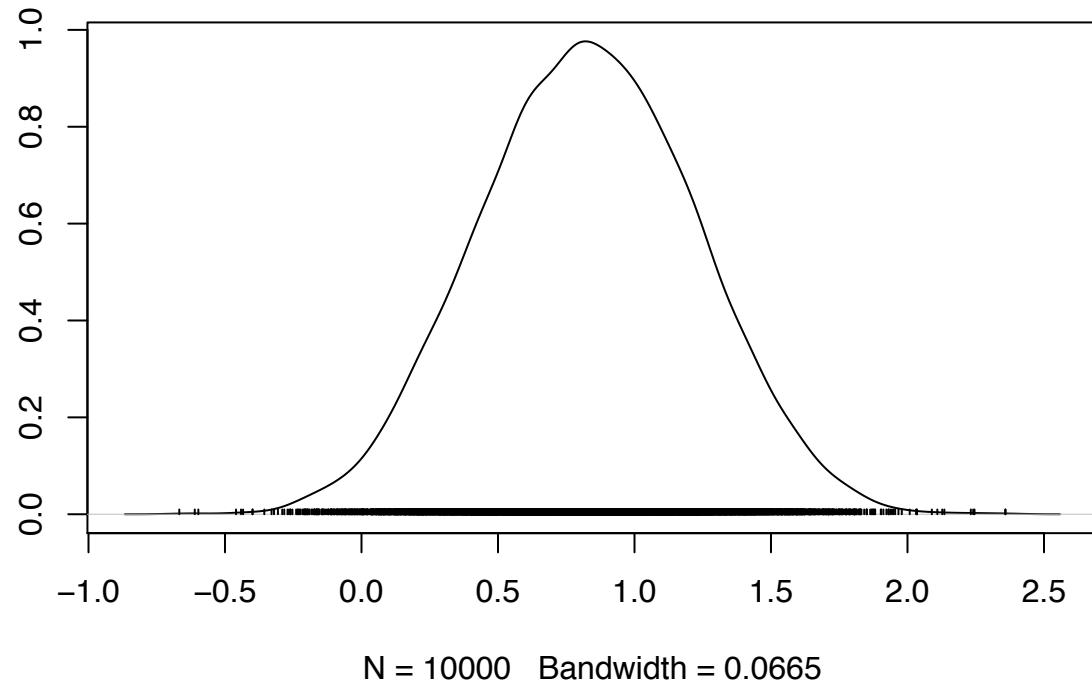
```

##          2.5%      25%      50%      75%
## (Intercept) 0.0677862 0.559701 0.8273073 1.1021928
## maleTRUE     -1.2105549 -0.849437 -0.6636819 -0.4686059
## lgbtqTRUE    0.4285330 0.737927 0.8936220 1.0536681
## racial_minority1 -0.5624145 -0.315432 -0.1833386 -0.0522772
## vic_worst_prior_sexTRUE -1.0109447 -0.796816 -0.6816811 -0.5712171
## vic_worst_verbal 0.0189556 0.168546 0.2434890 0.3204187
## vic_worst_incap 0.3598458 0.473716 0.5330752 0.5931340
## vic_worst_force 0.9301468 1.085198 1.1700653 1.2564103
## week_num      -0.0847719 -0.053118 -0.0364014 -0.0194960
## week_num_SQ   -0.0015824 -0.001023 -0.0007305 -0.0004334
## post_metoo1   -0.3900161 0.135788 0.4090188 0.6765989
## post_metoo_week -0.0267459 0.013879 0.0348062 0.0563137
## post_metoo_week_SQ -0.0001491 0.000486 0.0008107 0.0011393
## sigma2        3.1514004 3.396580 3.5300145 3.6730515
##          97.5%
## (Intercept) 1.5955248
## maleTRUE     -0.1011489
## lgbtqTRUE    1.3537085
## racial_minority1 0.1959153
## vic_worst_prior_sexTRUE -0.3596687
## vic_worst_verbal 0.4644550
## vic_worst_incap 0.7066714
## vic_worst_force 1.4191519
## week_num      0.0129870
## week_num_SQ   0.0001507
## post_metoo1   1.1789668
## post_metoo_week 0.0960883
## post_metoo_week_SQ 0.0017904
## sigma2        3.9739682

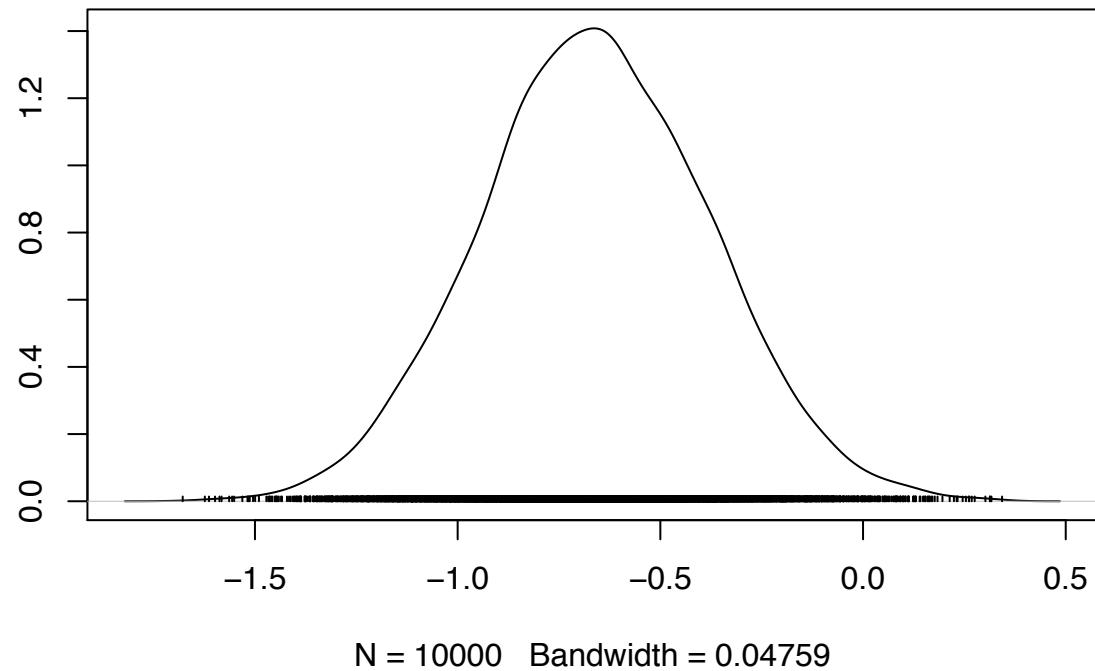
```

```
coda::densplot(h2_posteriors$post_metoo_time_SQ)
```

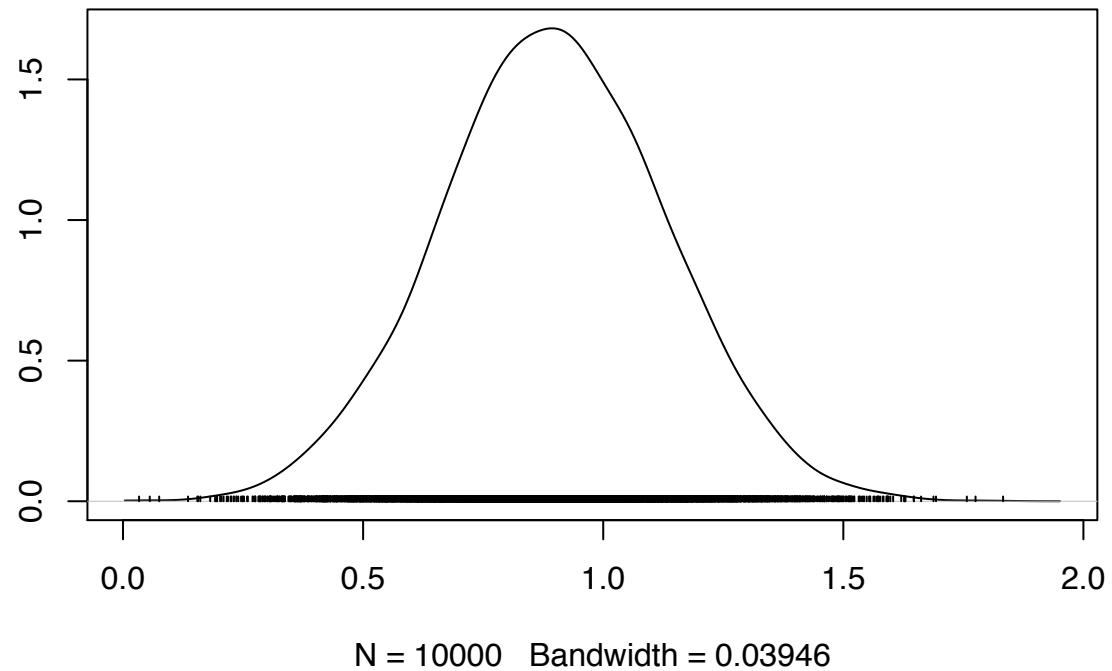
### Density of (Intercept)



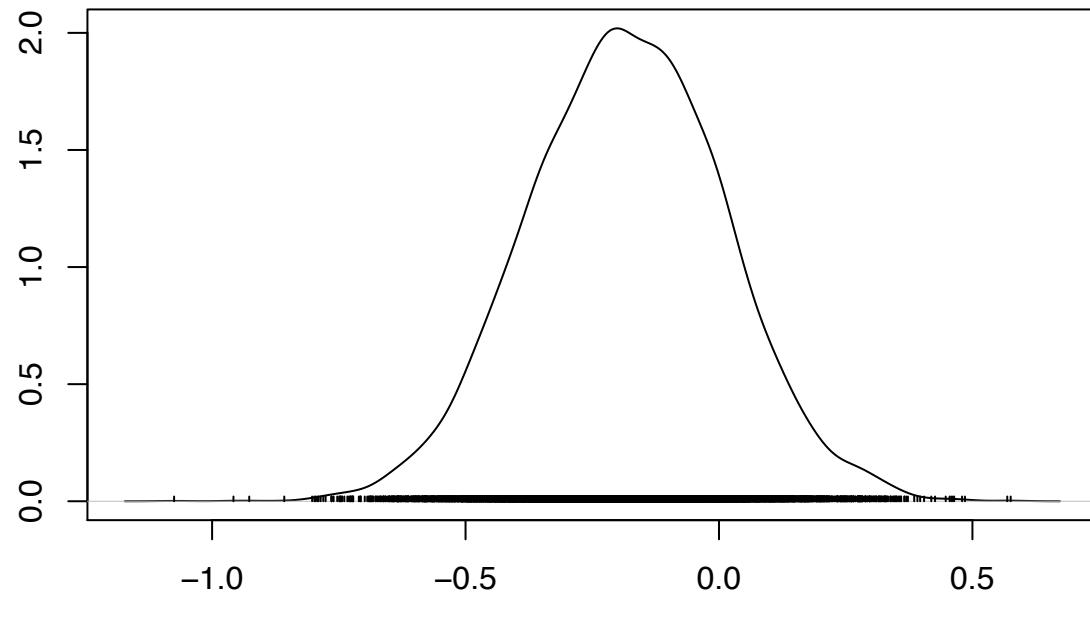
### Density of maleTRUE



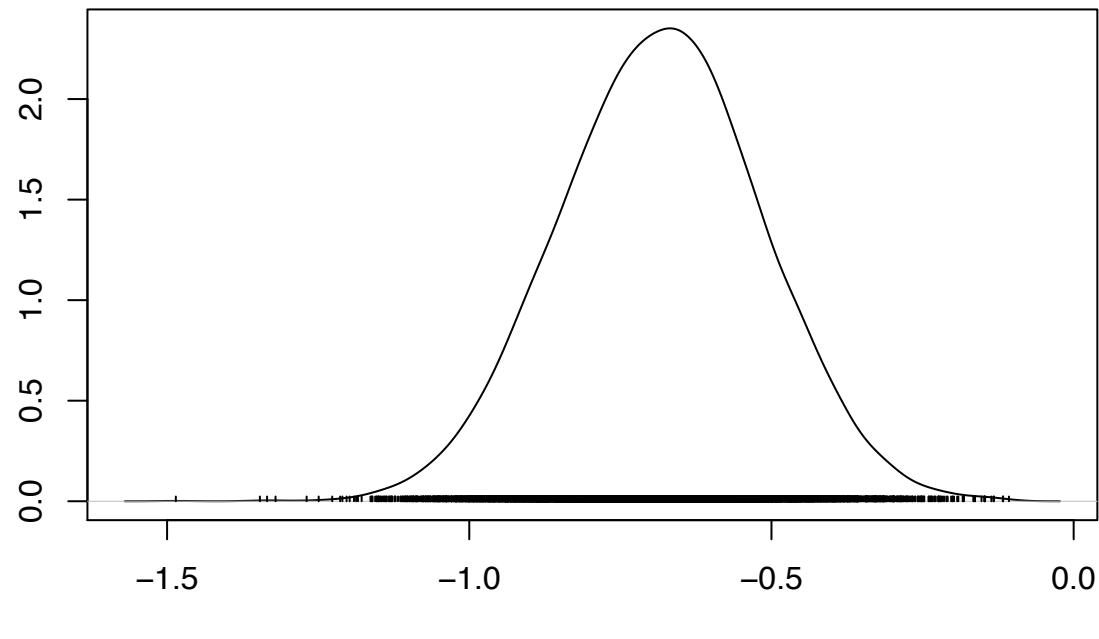
## Density of lgbtqTRUE



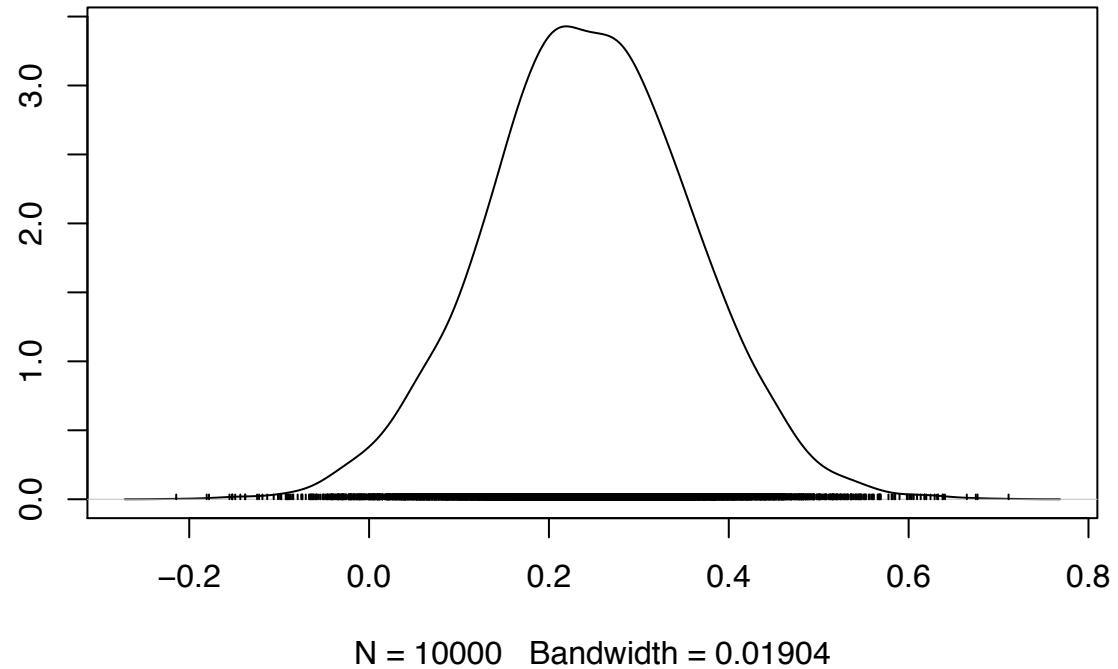
### **Density of racial\_minority1**



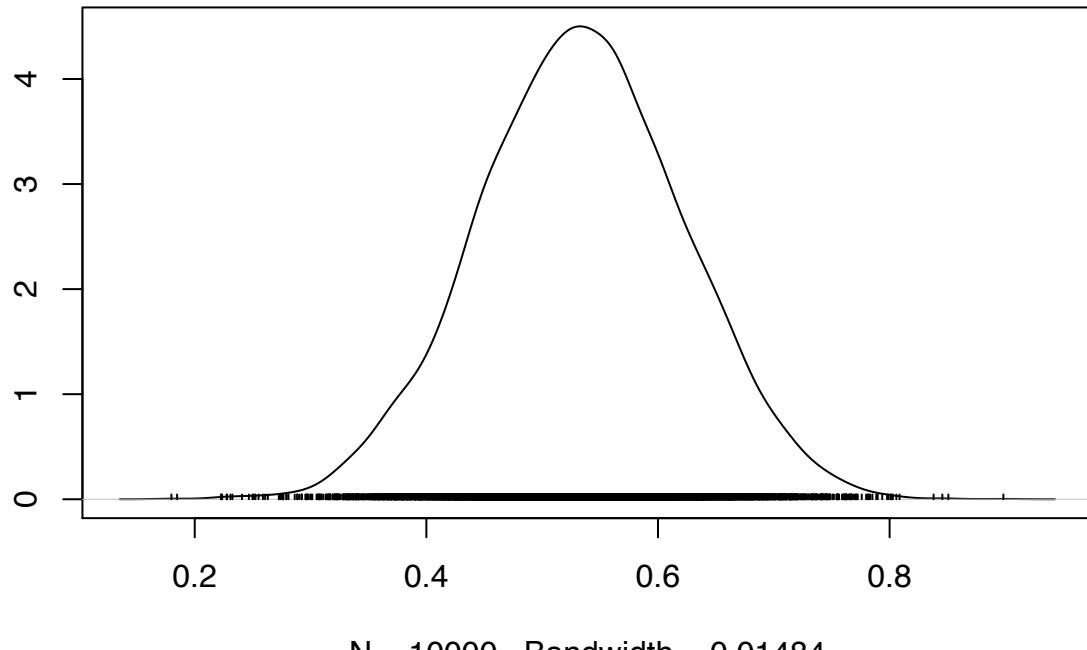
### **Density of vic\_worst\_prior\_sexTRUE**



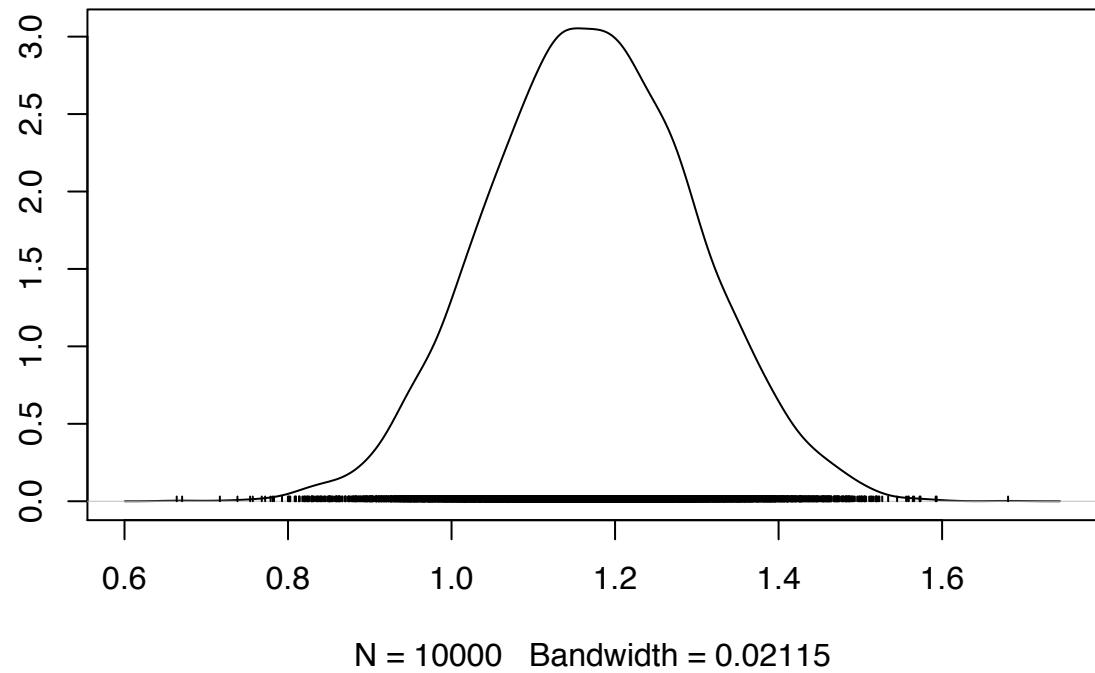
### **Density of vic\_worst\_verbal**



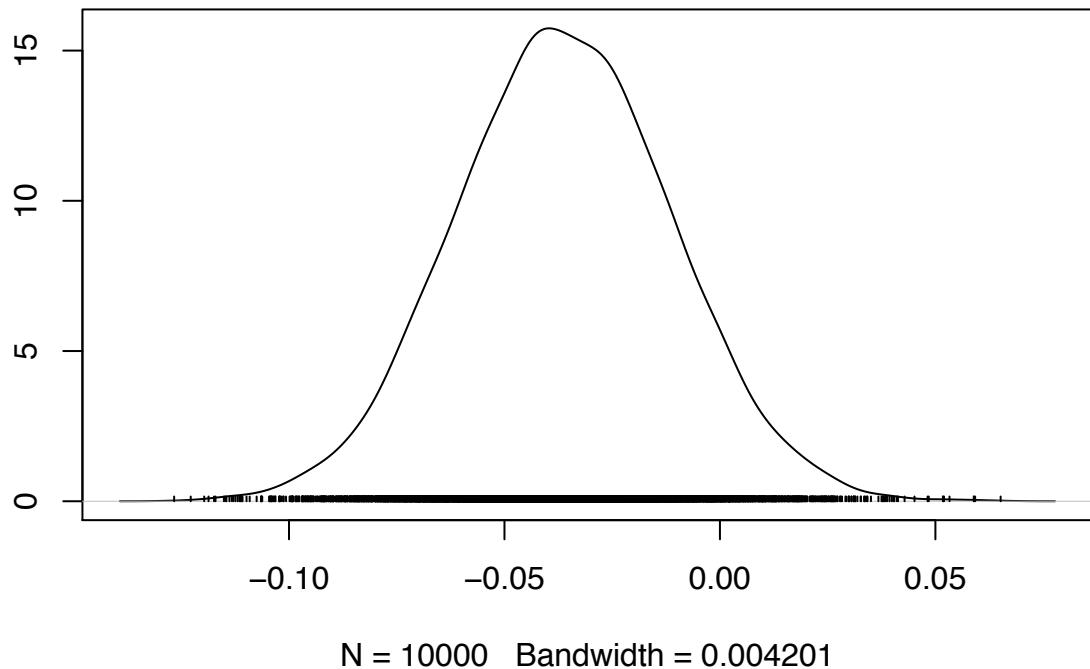
### **Density of vic\_worst\_incap**



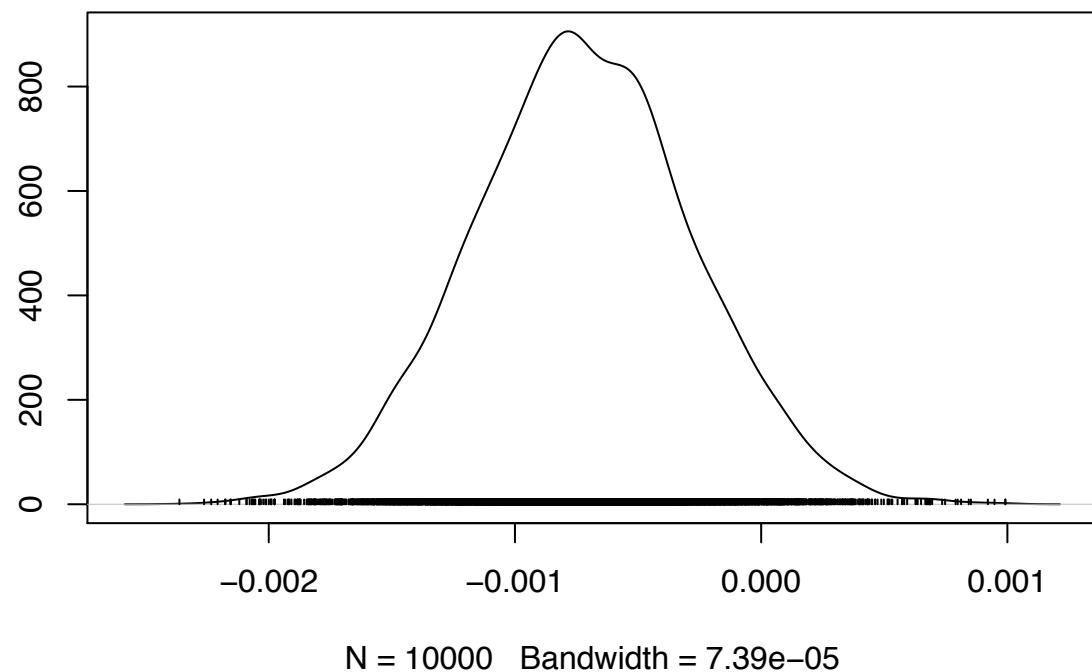
### **Density of vic\_worst\_force**



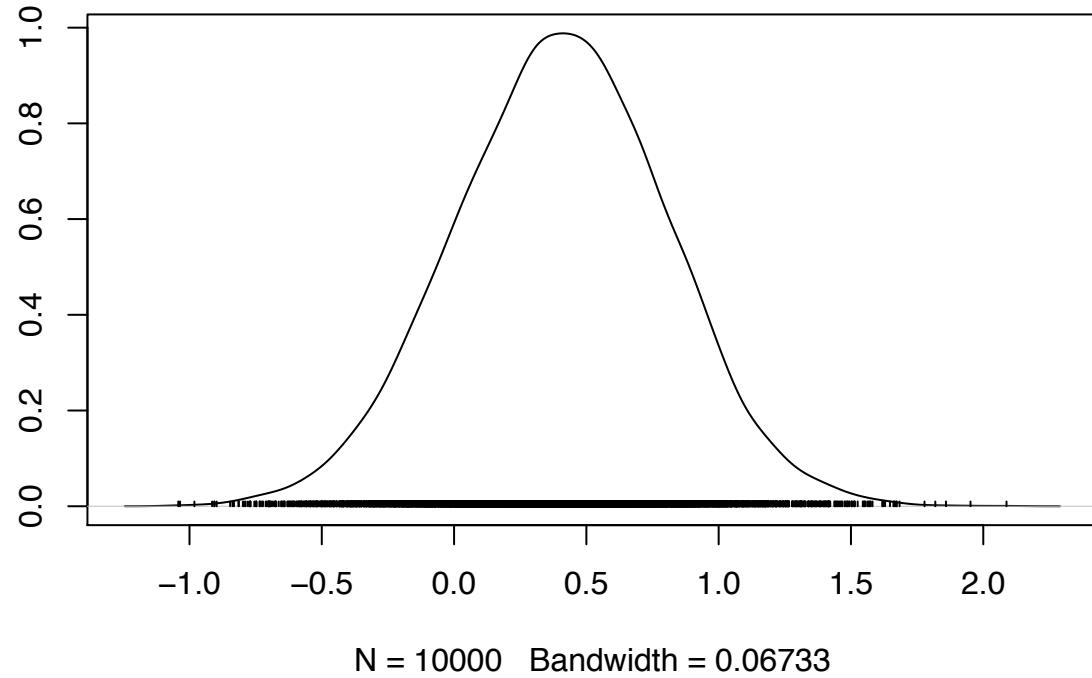
### **Density of week\_num**



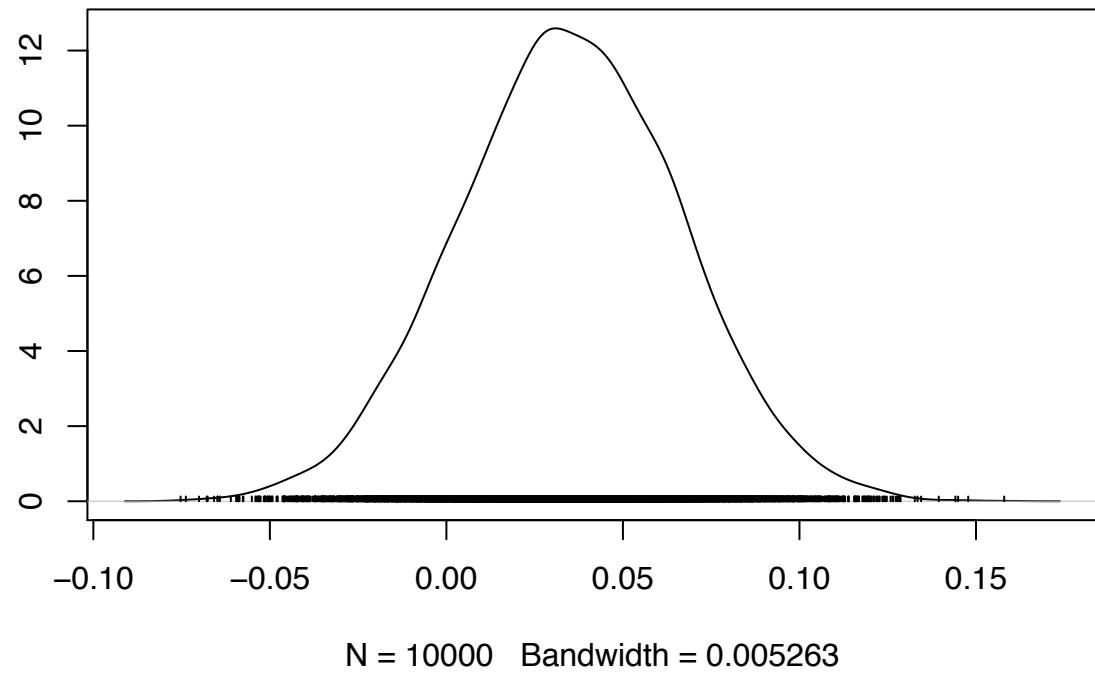
### **Density of week\_num\_SQ**



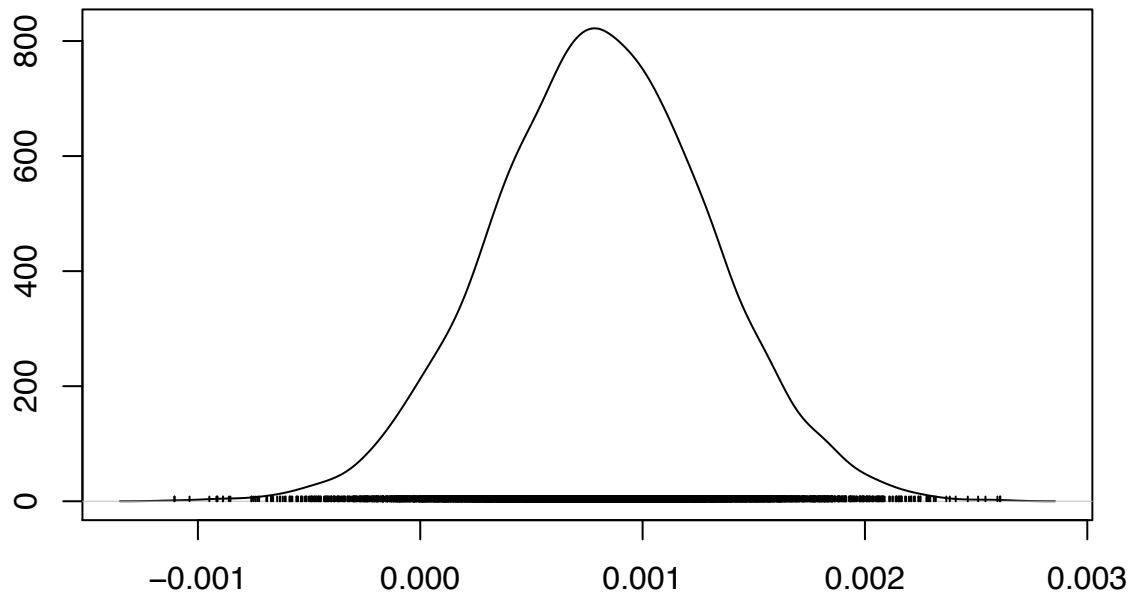
### **Density of post\_metoo1**



### **Density of post\_metoo\_week**

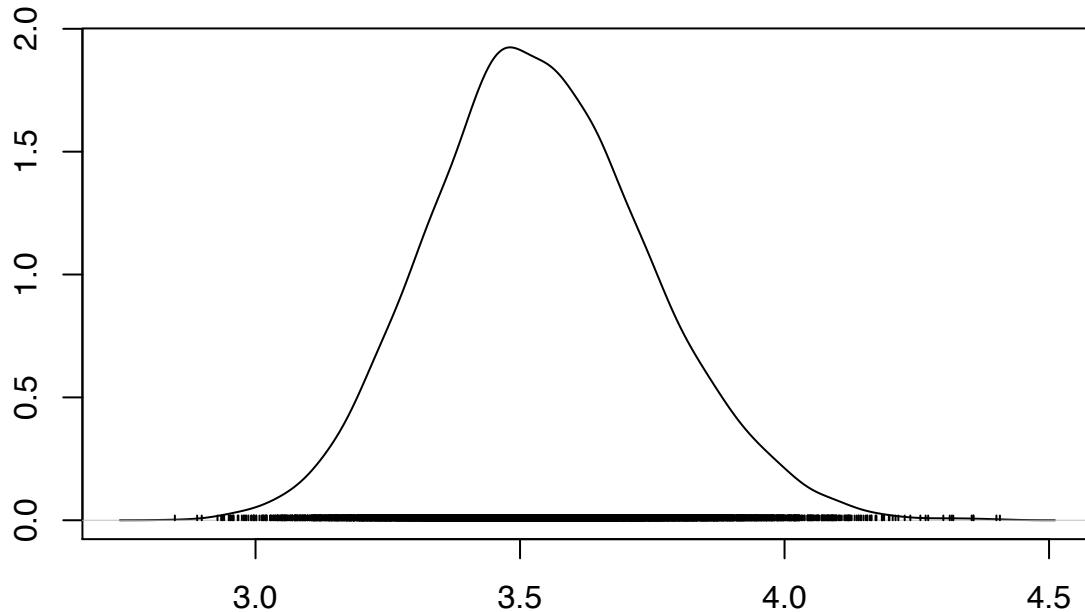


**Density of post\_metoo\_week\_SQ**



N = 10000 Bandwidth = 8.192e-05

## Density of sigma2



N = 10000 Bandwidth = 0.03466

```
h2_param_means <- map(h2_comp_post, colMeans) %>%
  map(data.frame) %>%
  map(~ set_names(.x, 'mean_est')) %>%
  map(~ mutate(.data = .x, param = rownames(.x))) %>%
  bind_rows(.id = 'model') %>%
  mutate(original_order = row_number())

h2_param_sd <- map(h2_comp_post, colMeans) %>%
  map(data.frame) %>%
  map(~ set_names(.x, 'mean_est')) %>%
  map(~ mutate(.data = .x, param = rownames(.x))) %>%
  bind_rows(.id = 'model') %>%
  mutate(original_order = row_number())

h2_hpdi_97 <- map(h2_comp_post, HPDinterval, prob = .97) %>%
  map(data.frame) %>%
  map(~ mutate(.data = .x, param = rownames(.x))) %>%
  bind_rows(.id = 'model') %>%
  dplyr::rename(lower_97 = lower, upper_97 = upper)

h2_hpdi_95 <- map(h2_comp_post, HPDinterval, prob = .95) %>%
  map(data.frame) %>%
  map(~ mutate(.data = .x, param = rownames(.x))) %>%
  bind_rows(.id = 'model') %>%
  dplyr::rename(lower_95 = lower, upper_95 = upper)
```

```

h2_params <- h2_param_means %>%
  left_join(h2_hpdi_97, by = c('model', 'param')) %>%
  left_join(h2_hpdi_95, by = c('model', 'param'))

h2_params

```

	model	mean_est	param	original_order
## 1	intercept_only	3.793366e+00	(Intercept)	1
## 2	intercept_only	4.877615e+00	sigma2	2
## 3	demos	3.724704e+00	(Intercept)	3
## 4	demos	-7.753399e-01	maleTRUE	4
## 5	demos	1.118866e+00	lgbtqTRUE	5
## 6	demos	-3.221098e-02	racial_minority1	6
## 7	demos	4.720490e+00	sigma2	7
## 8	demos_and_context	1.197940e+00	(Intercept)	8
## 9	demos_and_context	-5.783731e-01	maleTRUE	9
## 10	demos_and_context	9.080412e-01	lgbtqTRUE	10
## 11	demos_and_context	-1.556264e-01	racial_minority1	11
## 12	demos_and_context	-6.402269e-01	vic_worst_prior_sexTRUE	12
## 13	demos_and_context	2.442476e-01	vic_worst_verbal	13
## 14	demos_and_context	5.322545e-01	vic_worst_incap	14
## 15	demos_and_context	1.167414e+00	vic_worst_force	15
## 16	demos_and_context	3.578458e+00	sigma2	16
## 17	time	1.163183e+00	(Intercept)	17
## 18	time	-5.934384e-01	maleTRUE	18
## 19	time	8.718091e-01	lgbtqTRUE	19
## 20	time	-1.817410e-01	racial_minority1	20
## 21	time	-6.495053e-01	vic_worst_prior_sexTRUE	21
## 22	time	2.325767e-01	vic_worst_verbal	22
## 23	time	5.330431e-01	vic_worst_incap	23
## 24	time	1.174795e+00	vic_worst_force	24
## 25	time	5.172195e-03	week_num	25
## 26	time	3.541068e+00	sigma2	26
## 27	time_SQ	1.145318e+00	(Intercept)	27
## 28	time_SQ	-5.961344e-01	maleTRUE	28
## 29	time_SQ	8.665201e-01	lgbtqTRUE	29
## 30	time_SQ	-1.808430e-01	racial_minority1	30
## 31	time_SQ	-6.492758e-01	vic_worst_prior_sexTRUE	31
## 32	time_SQ	2.329441e-01	vic_worst_verbal	32
## 33	time_SQ	5.358449e-01	vic_worst_incap	33
## 34	time_SQ	1.173081e+00	vic_worst_force	34
## 35	time_SQ	4.934240e-03	week_num	35
## 36	time_SQ	9.530423e-06	week_num_SQ	36
## 37	time_SQ	3.545892e+00	sigma2	37
## 38	post_metoo	1.147504e+00	(Intercept)	38
## 39	post_metoo	-5.919566e-01	maleTRUE	39
## 40	post_metoo	8.686691e-01	lgbtqTRUE	40
## 41	post_metoo	-1.824841e-01	racial_minority1	41
## 42	post_metoo	-6.466003e-01	vic_worst_prior_sexTRUE	42
## 43	post_metoo	2.342714e-01	vic_worst_verbal	43
## 44	post_metoo	5.347992e-01	vic_worst_incap	44
## 45	post_metoo	1.171170e+00	vic_worst_force	45
## 46	post_metoo	4.948438e-03	week_num	46

```

## 47      post_metoo 8.790726e-06      week_num_SQ 47
## 48      post_metoo -1.883065e-03     post_metoo1 48
## 49      post_metoo 3.549096e+00      sigma2 49
## 50      post_metoo_time 1.098153e+00 (Intercept) 50
## 51      post_metoo_time -5.746679e-01 maleTRUE 51
## 52      post_metoo_time 8.637244e-01 lgbtqTRUE 52
## 53      post_metoo_time -1.756329e-01 racial_minority1 53
## 54      post_metoo_time -6.463240e-01 vic_worst_prior_sexTRUE 54
## 55      post_metoo_time 2.318144e-01 vic_worst_verbal 55
## 56      post_metoo_time 5.302168e-01 vic_worst_incap 56
## 57      post_metoo_time 1.174840e+00 vic_worst_force 57
## 58      post_metoo_time -3.516417e-04 week_num 58
## 59      post_metoo_time -7.506711e-05 week_num_SQ 59
## 60      post_metoo_time -4.449188e-02 post_metoo1 60
## 61      post_metoo_time 1.256406e-02 post_metoo_week 61
## 62      post_metoo_time 3.553901e+00 sigma2 62
## 63 post_metoo_time_SQ 8.321770e-01 (Intercept) 63
## 64 post_metoo_time_SQ -5.749980e-01 maleTRUE 64
## 65 post_metoo_time_SQ 8.618804e-01 lgbtqTRUE 65
## 66 post_metoo_time_SQ -1.815823e-01 racial_minority1 66
## 67 post_metoo_time_SQ -6.461953e-01 vic_worst_prior_sexTRUE 67
## 68 post_metoo_time_SQ 2.420592e-01 vic_worst_verbal 68
## 69 post_metoo_time_SQ 5.342429e-01 vic_worst_incap 69
## 70 post_metoo_time_SQ 1.177701e+00 vic_worst_force 70
## 71 post_metoo_time_SQ -3.549490e-02 week_num 71
## 72 post_metoo_time_SQ -7.276398e-04 week_num_SQ 72
## 73 post_metoo_time_SQ 3.389893e-01 post_metoo1 73
## 74 post_metoo_time_SQ 3.548293e-02 post_metoo_week 74
## 75 post_metoo_time_SQ 8.005857e-04 post_metoo_week_SQ 75
## 76 post_metoo_time_SQ 3.540552e+00 sigma2 76
##           lower_97      upper_97      lower_95      upper_95
## 1  3.595397e+00  3.9923419929  3.609373e+00  3.9692691116
## 2  4.283655e+00  5.5250454563  4.320346e+00  5.4363445390
## 3  3.490107e+00  3.9694850016  3.508518e+00  3.9421184058
## 4 -1.469636e+00 -0.0815043853 -1.401183e+00 -0.1585566287
## 5  5.449931e-01  1.7031045017  5.795262e-01  1.6309328932
## 6 -5.150259e-01  0.4380258363 -4.559557e-01  0.4035402401
## 7  4.121814e+00  5.3353132132  4.174995e+00  5.2614568712
## 8  5.664221e-01  1.8015709091  6.228238e-01  1.7551876696
## 9 -1.228159e+00  0.0273697908 -1.157091e+00 -0.0288148735
## 10 4.015918e-01  1.4336852031  4.523162e-01  1.3756080356
## 11 -5.577394e-01  0.2680996567 -5.270874e-01  0.2204142483
## 12 -9.910790e-01 -0.2650852725 -9.743571e-01 -0.3230476440
## 13 2.507211e-03  0.4913711906  3.075708e-02  0.4750007915
## 14 3.461168e-01  0.7271641225  3.529341e-01  0.6978362449
## 15 8.947502e-01  1.4451671377  9.150060e-01  1.4116946985
## 16 3.114359e+00  4.0320463082  3.181807e+00  4.0106290433
## 17 5.265396e-01  1.8045676997  5.961834e-01  1.7464867162
## 18 -1.208659e+00  0.0173793405 -1.145612e+00 -0.0400976448
## 19 3.664116e-01  1.3922758599  4.033380e-01  1.3272063541
## 20 -5.897806e-01  0.2314314566 -5.470690e-01  0.1925110990
## 21 -9.900894e-01 -0.2672117237 -9.675444e-01 -0.3199687114
## 22 -2.480403e-02  0.4735678972  4.226998e-03  0.4534086952
## 23 3.406722e-01  0.7305106329  3.514120e-01  0.7023472250

```

```

## 24 9.047468e-01 1.4486851762 9.242989e-01 1.4180741842
## 25 1.128145e-03 0.0092431131 1.448875e-03 0.0088051190
## 26 3.095546e+00 4.0122961542 3.131023e+00 3.9628731471
## 27 4.577362e-01 1.7602012678 5.659763e-01 1.7404562907
## 28 -1.189769e+00 0.0443355477 -1.118275e+00 -0.0059373588
## 29 3.416386e-01 1.3510593312 4.413339e-01 1.3423427598
## 30 -6.099564e-01 0.2128147095 -5.361516e-01 0.2095838615
## 31 -1.015465e+00 -0.2915509857 -9.678972e-01 -0.3123163078
## 32 -1.413289e-02 0.4858169895 1.438567e-02 0.4580901631
## 33 3.504726e-01 0.7351151273 3.6633336e-01 0.7126602814
## 34 9.015375e-01 1.4505413457 9.246369e-01 1.4264650477
## 35 1.526988e-04 0.0098714358 6.477746e-04 0.0094347038
## 36 -9.741973e-05 0.0001158848 -8.796369e-05 0.0001041760
## 37 3.115256e+00 4.0482772082 3.123060e+00 3.9661856640
## 38 4.272145e-01 1.8889352147 4.787724e-01 1.7847078655
## 39 -1.203420e+00 0.0358071909 -1.123182e+00 -0.0060633287
## 40 3.461999e-01 1.3837353700 4.028743e-01 1.3263394417
## 41 -6.110770e-01 0.2354334799 -5.704930e-01 0.1873784586
## 42 -9.939175e-01 -0.2799219787 -9.792464e-01 -0.3342702559
## 43 -2.192261e-02 0.4815704406 -5.121179e-03 0.4515528519
## 44 3.453593e-01 0.7256130489 3.679052e-01 0.7131005301
## 45 9.062909e-01 1.4534120925 9.286032e-01 1.4212133539
## 46 -3.952112e-03 0.0140854048 -3.450538e-03 0.0129706425
## 47 -1.058441e-04 0.0001234489 -9.822715e-05 0.0001113750
## 48 -6.774438e-01 0.6674761808 -5.980444e-01 0.6108886145
## 49 3.105310e+00 4.0091568066 3.159567e+00 3.9754926900
## 50 2.812141e-01 1.8825840000 3.903436e-01 1.8256135675
## 51 -1.197199e+00 0.0580890929 -1.154370e+00 -0.0069186357
## 52 3.747055e-01 1.3900243447 3.834387e-01 1.3010919996
## 53 -5.875943e-01 0.2454246983 -5.583857e-01 0.2036097077
## 54 -1.021670e+00 -0.2897862180 -9.717779e-01 -0.3117497348
## 55 -2.005138e-02 0.4878068406 1.103638e-02 0.4686111053
## 56 3.450681e-01 0.7285504677 3.578463e-01 0.7093604811
## 57 8.964305e-01 1.4418058112 9.302642e-01 1.4264280205
## 58 -2.739358e-02 0.0264897289 -2.397863e-02 0.0249543729
## 59 -5.021014e-04 0.0003500008 -4.577510e-04 0.0003110959
## 60 -7.404560e-01 0.6408607868 -6.598600e-01 0.5865775845
## 61 -4.618784e-02 0.0746373337 -4.256796e-02 0.0671202577
## 62 3.102398e+00 4.0485957993 3.142366e+00 3.9849524977
## 63 -1.106045e-02 1.7042608063 3.601376e-02 1.5968292791
## 64 -1.185816e+00 0.0588546682 -1.127500e+00 0.0029290324
## 65 3.362588e-01 1.3668408351 3.936576e-01 1.3238052455
## 66 -5.915063e-01 0.2426707897 -5.471268e-01 0.2086451714
## 67 -1.011331e+00 -0.2936467910 -9.641094e-01 -0.3192015770
## 68 -3.8455585e-03 0.4935195324 2.412471e-02 0.4717972068
## 69 3.444411e-01 0.7243488050 3.691934e-01 0.7137884304
## 70 9.086130e-01 1.4598977149 9.220604e-01 1.4242891129
## 71 -8.790990e-02 0.0194771468 -8.659981e-02 0.0108990988
## 72 -1.658780e-03 0.0002612401 -1.605542e-03 0.0001342228
## 73 -5.197759e-01 1.1952091994 -4.494992e-01 1.1171694700
## 74 -3.429720e-02 0.1015448446 -2.545126e-02 0.0977908598
## 75 -2.352638e-04 0.0019041890 -1.588900e-04 0.0017658348
## 76 3.114227e+00 4.0303965893 3.136825e+00 3.9698779246

```

```
write.csv(h2_params, paste0('h2_params_', Sys.Date(), '.csv'), row.names = F)
```

Lastly, we write the table to a .csv file for later.

```
h2_model_table <- h2_bayes_factors
```

```
h2_model_table
```

```
##                  model logmarglike
## 1      intercept_only -1,288.69005
## 2              demos -1,288.65833
## 3  demos_and_context -1,210.58796
## 4              time -1,215.35232
## 5          time_SQ -1,227.55852
## 6      post_metoo -1,231.02514
## 7  post_metoo_time -1,236.78897
## 8 post_metoo_time_SQ -1,245.38177
##                                     BF
## 1                               NA
## 2                         1.03222
## 3 8,045,103,140,013,963,422,936,160,087,834,624.00000
## 4                           0.00853
## 5                           0.00000
## 6                           0.03122
## 7                           0.00314
## 8                           0.00019
```

```
write.csv(
  x = h2_model_table,
  file = paste0('h2_model_table', Sys.Date(), '.csv'),
  row.names = F)
```

## Model MCMC Chain Convergence

Model convergence is assessed using the first multiply imputed dataset for each hypothesis, using standard techniques from the `coda` package.

### H1 Convergence

```

Sys.time()

## [1] "2019-11-17 01:49:17 UTC"

library(MCMCpack)

test_df <- full_imp[[1]]

test_chains <- map(
  .x = h1_models,
  .f = ~ rerun(4, run_logistic_mcmc(
    df = test_df,
    model = .x,
    prior_sd = h1_priors,
    mcmc = 6*1e4,
    thin = 2)) %>%
    as.mcmc.list()) %>%
  set_names(names(h1_models))

Sys.time()

## [1] "2019-11-17 01:55:29 UTC"

# coda::traceplot(test_chains$post_metoo_time_SQ)

1 - coda::rejectionRate(test_chains$post_metoo_time_SQ)

##          (Intercept)      maleTRUE      lgbtqTRUE
## 0.2466332        0.2466332        0.2466332
##   racial_minority1      week_num      week_num_SQ
## 0.2466332        0.2466332        0.2466332
##   post_metoo1    post_metoo_week post_metoo_week_SQ
## 0.2466332        0.2466332        0.2466332

coda::gelman.diag(test_chains$post_metoo_time_SQ)

## Potential scale reduction factors:
##                               Point est. Upper C.I.
## (Intercept)                  1     1.00
## maleTRUE                     1     1.01
## lgbtqTRUE                     1     1.00
## racial_minority1              1     1.00
## week_num                      1     1.00
## week_num_SQ                   1     1.00
## post_metoo1                  1     1.01
## post_metoo_week                1     1.00
## post_metoo_week_SQ             1     1.00
##
## Multivariate psrf
##
## 1

```

## H2 Convergence

```
test_df <- vic_imp[[1]]  
  
test_chains <- map(  
  .x = h2_models,  
  .f = ~ rerun(4, run_linear_mcmc(  
    df = test_df,  
    model = .x,  
    prior_sd = h2_priors,  
    mcmc = 1e5,  
    thin = 1)) %>%  
  as.mcmc.list()) %>%  
  set_names(names(h2_models))  
  
# coda::traceplot(test_chains$post_metoo_time_SQ)
```

Note, `coda::rejectionRate` is likely producing errors here. See helpfile, where it describes doing a naive test of two successive iterations of the sampler (using `==`). This likely produces rounding errors, causing the rejection rate to look like 0. This happened on simulated data too, so it is not just our data that is producing this odd output. However, all other diagnostics for our chains look healthy, so we proceed to estimating the rest of these models.

```
map(test_chains[2:length(test_chains)], ~ 1 - coda::rejectionRate(.x))
```

```
## $demos  
##      (Intercept) maleTRUE lgbtqTRUE racial_minority1  
##            1          1          1          1  
##      sigma2  
##            1  
##  
## $demos_and_context  
##      (Intercept) maleTRUE lgbtqTRUE  
##            1          1          1  
##      racial_minority1 vic_worst_prior_sexTRUE vic_worst_verbal  
##            1          1          1  
##      vic_worst_incap    vic_worst_force sigma2  
##            1                  1          1  
##  
## $time  
##      (Intercept) maleTRUE lgbtqTRUE  
##            1          1          1  
##      racial_minority1 vic_worst_prior_sexTRUE vic_worst_verbal  
##            1          1          1  
##      vic_worst_incap    vic_worst_force week_num  
##            1                  1          1  
##      sigma2  
##            1  
##  
## $time_SQ  
##      (Intercept) maleTRUE lgbtqTRUE  
##            1          1          1
```

```

##      racial_minority1 vic_worst_prior_sexTRUE          vic_worst_verbal
##                           1                               1
##      vic_worst_incap           vic_worst_force          week_num
##                           1                               1
##      week_num_SQ                sigma2
##                           1                               1
##
## $post_metoo
##      (Intercept)           maleTRUE          lgbtqTRUE
##                           1                               1
##      racial_minority1 vic_worst_prior_sexTRUE          vic_worst_verbal
##                           1                               1
##      vic_worst_incap           vic_worst_force          week_num
##                           1                               1
##      week_num_SQ                post_metoo1          sigma2
##                           1                               1
##
## $post_metoo_time
##      (Intercept)           maleTRUE          lgbtqTRUE
##                           1                               1
##      racial_minority1 vic_worst_prior_sexTRUE          vic_worst_verbal
##                           1                               1
##      vic_worst_incap           vic_worst_force          week_num
##                           1                               1
##      week_num_SQ                post_metoo1          post_metoo_week
##                           1                               1
##      sigma2
##                           1
##
## $post_metoo_time_SQ
##      (Intercept)           maleTRUE          lgbtqTRUE
##                           1                               1
##      racial_minority1 vic_worst_prior_sexTRUE          vic_worst_verbal
##                           1                               1
##      vic_worst_incap           vic_worst_force          week_num
##                           1                               1
##      week_num_SQ                post_metoo1          post_metoo_week
##                           1                               1
##      post_metoo_week_SQ          sigma2
##                           1                               1

map(test_chains[2:length(test_chains)], coda::gelman.diag)

## $demos
## Potential scale reduction factors:
##
##      Point est. Upper C.I.
##      (Intercept)           1       1
##      maleTRUE              1       1
##      lgbtqTRUE              1       1
##      racial_minority1       1       1
##      sigma2                 1       1
##
## Multivariate psrf

```

```

##  

## 1  

##  

## $demos_and_context  

## Potential scale reduction factors:  

##  

##  

##          Point est. Upper C.I.  

## (Intercept)      1      1  

## maleTRUE        1      1  

## lgbtqTRUE       1      1  

## racial_minority1 1      1  

## vic_worst_prior_sexTRUE 1      1  

## vic_worst_verbal    1      1  

## vic_worst_incap     1      1  

## vic_worst_force     1      1  

## sigma2            1      1  

##  

## Multivariate psrf  

##  

##  

## 1  

##  

##  

## $time  

## Potential scale reduction factors:  

##  

##          Point est. Upper C.I.  

## (Intercept)      1      1  

## maleTRUE        1      1  

## lgbtqTRUE       1      1  

## racial_minority1 1      1  

## vic_worst_prior_sexTRUE 1      1  

## vic_worst_verbal    1      1  

## vic_worst_incap     1      1  

## vic_worst_force     1      1  

## week_num          1      1  

## sigma2            1      1  

##  

## Multivariate psrf  

##  

##  

## 1  

##  

##  

## $time_SQ  

## Potential scale reduction factors:  

##  

##          Point est. Upper C.I.  

## (Intercept)      1      1  

## maleTRUE        1      1  

## lgbtqTRUE       1      1  

## racial_minority1 1      1  

## vic_worst_prior_sexTRUE 1      1  

## vic_worst_verbal    1      1  

## vic_worst_incap     1      1  

## vic_worst_force     1      1  

## week_num          1      1  

## week_num_SQ        1      1

```

```

## sigma2          1          1
##
## Multivariate psrf
##
## 1
##
## $post_metoo
## Potential scale reduction factors:
##
##                               Point est. Upper C.I.
## (Intercept)                  1          1
## maleTRUE                     1          1
## lgbtqTRUE                    1          1
## racial_minority1              1          1
## vic_worst_prior_sexTRUE      1          1
## vic_worst_verbal             1          1
## vic_worst_incap               1          1
## vic_worst_force               1          1
## week_num                      1          1
## week_num_SQ                   1          1
## post_metoo1                  1          1
## sigma2                        1          1
##
## Multivariate psrf
##
## 1
##
## $post_metoo_time
## Potential scale reduction factors:
##
##                               Point est. Upper C.I.
## (Intercept)                  1          1
## maleTRUE                     1          1
## lgbtqTRUE                    1          1
## racial_minority1              1          1
## vic_worst_prior_sexTRUE      1          1
## vic_worst_verbal             1          1
## vic_worst_incap               1          1
## vic_worst_force               1          1
## week_num                      1          1
## week_num_SQ                   1          1
## post_metoo1                  1          1
## post_metoo_week               1          1
## sigma2                        1          1
##
## Multivariate psrf
##
## 1
##
## $post_metoo_time_SQ
## Potential scale reduction factors:
##
##                               Point est. Upper C.I.
## (Intercept)                  1          1

```

```
## maleTRUE           1           1
## lgbtqTRUE          1           1
## racial_minority1   1           1
## vic_worst_prior_sexTRUE 1           1
## vic_worst_verbal    1           1
## vic_worst_incap     1           1
## vic_worst_force     1           1
## week_num            1           1
## week_num_SQ          1           1
## post_metoo1          1           1
## post_metoo_week      1           1
## post_metoo_week_SQ   1           1
## sigma2              1           1
##
## Multivariate psrf
##
## 1
```

# MeToo Movement and Perceptions of Sexual Assault:

## Section 2

### Contents

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### Introduction

#### R Version:

x86\_64-pc-linux-gnu, x86\_64, linux-gnu, x86\_64, linux-gnu, , 3, 6.3, 2020, 02, 29, 77875, R, R version 3.6.3 (2020-02-29), Holding the Windsock

**Run date:** 2020-05-20 10:15:17

What follows is the second section of annotated code for *The #MeToo Movement and Perceptions of Sexual Assault: College Students' Recognition of Sexual Assault Experiences Over Time*. This includes all processing and summarization of results, *after* models had been estimated.

#### Setup - load participant dataframe

```
knitr::knit('data_wrangling.Rmd', output = tempfile())  
  
##  
##  
## processing file: data_wrangling.Rmd  
  
## -- Attaching packages ----- tidyverse 1.3.0 --  
  
## v ggplot2 3.2.1      v purrr    0.3.4  
## v tibble   3.0.1      v dplyr    0.8.5  
## v tidyr    1.0.3      v stringr  1.4.0  
## v readr    1.3.1      vforcats  0.4.0  
  
## -- Conflicts ----- tidyverse_conflicts() --  
## x dplyr::filter() masks stats::filter()  
## x dplyr::lag()   masks stats::lag()  
  
##  
## Attaching package: 'lubridate'
```

```

## The following object is masked from 'package:base':
##
##      date

## output file: /tmp/RtmpaWsX4S/file30e97968af34

```

## Figure - change over time

Create a dataframe for plotting participant probability of screening positive (*p\_vic*) and mean sexual assault acknowledgment (*mean\_sa*).

```

library(tidyverse)

fig_df <- df %>%
  group_by(study_week) %>%
  summarise(
    post_metoo = unique(post_metoo),
    n = n(),
    p_vic = sum(vic_screener == 'Yes', na.rm = T)/n,
    mean_sa = mean(vic_worst_is_sa, na.rm = T)) %>%
  mutate(post_metoo = ifelse(study_week >= '2017-10-16', 'Yes', 'No'))

fig_df

```

```

## # A tibble: 90 x 5
##   study_week post_metoo     n  p_vic mean_sa
##   <date>     <chr>     <int>  <dbl>   <dbl>
## 1 2016-09-04 No         32 0.0312     6
## 2 2016-09-11 No         18 0.333      4
## 3 2016-09-18 No         14 0.214      5.33
## 4 2016-09-25 No         33 0.242      2.88
## 5 2016-10-02 No         37 0.108      1.75
## 6 2016-10-09 No         23 0.174      2
## 7 2016-10-16 No         14 0.214      3.67
## 8 2016-10-23 No         17 0.294      5.8
## 9 2016-10-30 No         19 0.368      3.71
## 10 2016-11-06 No        26 0.269      4.33
## # ... with 80 more rows

```

Plot screener and acknowledgment over time.

```

post_metoo_plot_df <- participant_df %>%
  group_by(post_metoo) %>%
  summarise(
    begin = ymd(as.Date(min(study_week))),
    end = ymd(as.Date(max(study_week))),
    mid = begin + (end - begin)/2,
    p_vic = mean(vic_screener == 'Yes', na.rm = T),
    mean_sa = mean(vic_worst_is_sa, na.rm = T),
    n = n()) %>%
  gather(point_type, p, -post_metoo, -begin, -end, -mid, -n) %>%
  mutate(

```

```

post_metoo = ifelse(post_metoo == 0, 'No', 'Yes'),
point_type = factor(
  x = point_type,
  levels = c('p_vic', 'mean_sa'),
  labels = c(
    'Sexual Assault Prevalence', 'Mean Acknowledgment')))

post_metoo_plot_df

## # A tibble: 4 x 7
##   post_metoo begin      end      mid      n point_type      p
##   <chr>     <date>    <date>    <date>    <int> <fct>       <dbl>
## 1 No        2016-09-04 2017-10-08 2017-03-23 1086 Sexual Assault P~ 0.241
## 2 Yes       2017-10-15 2019-04-21 2018-07-18 1530 Sexual Assault P~ 0.239
## 3 No        2016-09-04 2017-10-08 2017-03-23 1086 Mean Acknowledgm~ 3.55
## 4 Yes       2017-10-15 2019-04-21 2018-07-18 1530 Mean Acknowledgm~ 4.00

p <- fig_df %>%
  filter(is.na(n) | n >= 5) %>%
  gather(point_type, p, -study_week, -n, -post_metoo) %>%
  mutate(
    n = ifelse(point_type == 'sa', 10, n),
    point_type = factor(
      x = point_type,
      levels = c('p_vic', 'mean_sa'),
      labels = c(
        'Sexual Assault Prevalence', 'Mean Acknowledgment')) %>%
  ggplot(aes(study_week, p)) +
  theme_light() +
  scale_color_manual(values = c('grey60', 'grey20')) +
  facet_grid(point_type ~ ., scales = 'free') +
  geom_vline(xintercept = ymd('2017-10-15'), color = 'grey80', size = 2, linetype = 1) +
  # geom_line() +
#  geom_line(data = indiv_fig_df, color = 'grey20', linetype = 1, size = 1.5) +
  geom_point(aes(color = factor(post_metoo), size = n), alpha = .50) +
  # geom_smooth(span = .54, color = 'black', weights = n) +
  geom_segment(
    data = post_metoo_plot_df,
    mapping = aes(
      x = begin,
      xend = end,
      y = p,
      yend = p,
      color = factor(post_metoo)),
    linetype = 2
  ) +
#  geom_segment(
#    data = sem_plot_df,
#    mapping = aes(
#      x = begin,
#      xend = end,
#      y = p,
#      yend = p)) +

```

```

# geom_point(data = sem_plot_df, mapping = aes(x = mid), size = 3) +
# geom_line(data = sem_plot_df, mapping = aes(x = mid)) +
scale_x_date(
  date_breaks = '3 months', date_labels = '%b %y', date_minor_breaks = '1 month') +
labs(
  y = 'Outcome',
  x = 'Study Timeline') +
guides(
  color = guide_legend(title = 'Post #MeToo'),
  size = guide_legend(title = 'Total Participants')) +
theme(
  strip.text = element_text(color = 'black'),
  panel.border = element_rect(color = 'grey'),
  strip.background = element_rect(color = 'grey'),
  panel.grid = element_blank())

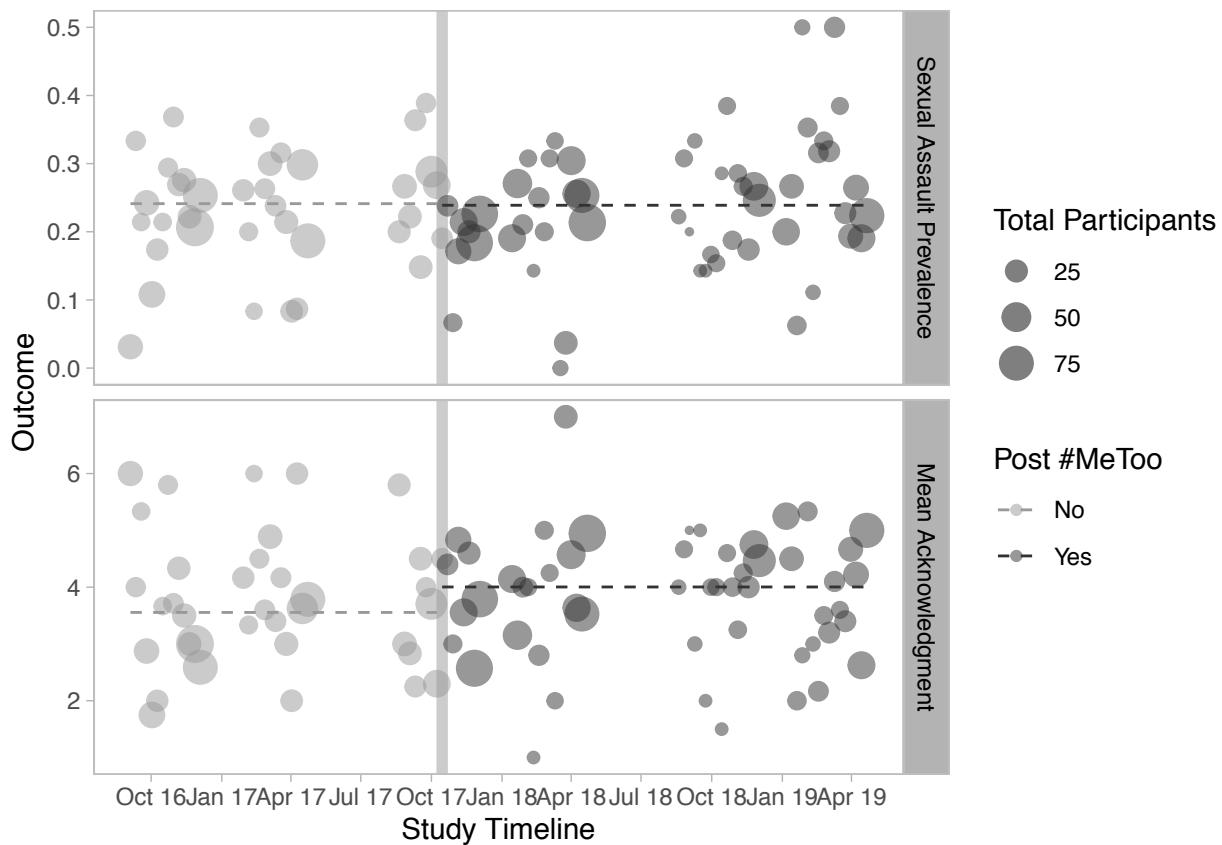
ggsave(
  filename = paste0('sa_endorsement_plot_', Sys.Date(), '.png'),
  plot = p,
  height = 6,
  width = 8,
  dpi = 800)

```

## Warning: Removed 1 rows containing missing values (geom\_point).

p

## Warning: Removed 1 rows containing missing values (geom\_point).



## Dummy coding table

Produce a dummy-coding table so readers can see how time and the MeToo interaction was coded.

```
coding_example_table <- df %>%
  group_by(study_week) %>%
  summarise(
    week_num = min(week_num),
    week_num_SQ = min(week_num_SQ),
    post_metoo = unique(post_metoo),
    post_metoo_week = min(post_metoo_week),
    post_metoo_week_SQ = min(post_metoo_week_SQ)) %>%
  filter(week_num >= 55 & week_num <= 60)

coding_example_table

## # A tibble: 5 x 6
##   study_week week_num week_num_SQ post_metoo post_metoo_week
##   <date>      <dbl>     <dbl>      <dbl>        <dbl>
## 1 2018-11-04     55     3025         1          55
## 2 2018-11-11     56     3136         1          56
## 3 2018-11-18     57     3249         1          57
## 4 2018-11-25     58     3364         1          58
## 5 2018-12-02     59     3481         1          59
## # ... with 1 more variable: post_metoo_week_SQ <dbl>
```

```
write.csv(coding_example_table, 'coding_example_table.csv', row.names = F)
```

## Tables for posterior distribution model parameters

Import posterior distributions from model fitting in Part 1.

```
latest_h1_posterior <- max(list.files()[grep('h1_posteriors', list.files())])
latest_h2_posterior <- max(list.files()[grep('h2_posteriors', list.files())])

h1_post <- readr::read_csv(latest_h1_posterior)

## Parsed with column specification:
## cols(
##   model = col_character(),
##   param = col_character(),
##   estimate = col_double()
## )

h2_post <- readr::read_csv(latest_h2_posterior)

## Parsed with column specification:
## cols(
##   model = col_character(),
##   param = col_character(),
##   estimate = col_double()
## )
```

Create tables of parameter estimates and uncertainty from posterior distributions for H1 and H2

### H1 parameter table

```
library(tidyverse)
library(MCMCpack)

## Loading required package: coda

## Loading required package: MASS

##
## Attaching package: 'MASS'

## The following object is masked from 'package:dplyr':
## 
##     select

## ##
## ## Markov Chain Monte Carlo Package (MCMCpack)
```

```

## ## Copyright (C) 2003-2020 Andrew D. Martin, Kevin M. Quinn, and Jong Hee Park
## ##
## ## Support provided by the U.S. National Science Foundation

## ## (Grants SES-0350646 and SES-0350613)
## ##

h1_param_table <- h1_post %>%
  filter(model == 'post_metoo_time_SQ') %>%
  group_by(param) %>%
  summarise(
    mean = mean(estimate),
    sd = sd(estimate),
    lower_hpdi_95 = HPDinterval(as.mcmc(estimate))[[1]],
    upper_hpdi_95 = HPDinterval(as.mcmc(estimate))[[2]],
    mean_OR = exp(mean(estimate))) %>%
  map_if(is.numeric, round, digits = 3) %>%
  as.data.frame()

h1_param_table

##           param   mean     sd lower_hpdi_95 upper_hpdi_95 mean_OR
## 1 (Intercept) -0.727 0.186      -1.097      -0.367  0.483
## 2 lgbtqTRUE   0.635 0.155       0.329      0.938  1.887
## 3 maleTRUE   -1.562 0.150      -1.857      -1.272  0.210
## 4 post_metoo_week  0.001 0.019      -0.035      0.039  1.001
## 5 post_metoo_week_SQ  0.000 0.000       0.000      0.001  1.000
## 6 post_metoo1  -0.274 0.250      -0.771      0.205  0.760
## 7 racial_minority1 -0.308 0.117      -0.535      -0.078  0.735
## 8 week_num    0.002 0.015      -0.027      0.032  1.002
## 9 week_num_SQ  0.000 0.000      -0.001      0.000  1.000

write.csv(
  x = h1_param_table,
  file = paste0('h1_param_table_', Sys.Date(), '.csv'),
  row.names = F)

```

## H2 parameter table

```

h2_param_table <- h2_post %>%
  filter(model == 'post_metoo_time_SQ') %>%
  group_by(param) %>%
  summarise(
    mean = mean(estimate),
    sd = sd(estimate),
    lower_hpdi_95 = HPDinterval(as.mcmc(estimate))[[1]],
    upper_hpdi_95 = HPDinterval(as.mcmc(estimate))[[2]]) %>%
  map_if(is.numeric, round, digits = 3) %>%
  as.data.frame()

h2_param_table

```

```

##                  param   mean     sd lower_hpdi_95 upper_hpdi_95
## 1          (Intercept) 0.829 0.396        0.070      1.597
## 2           lgbtqTRUE 0.895 0.235        0.411      1.333
## 3            maleTRUE -0.661 0.283      -1.222     -0.115
## 4    post_memeo_week  0.035 0.031      -0.028      0.095
## 5 post_memeo_week_SQ  0.001 0.000       0.000      0.002
## 6           post_memeo1 0.403 0.401      -0.393      1.175
## 7      racial_minority1 -0.183 0.193      -0.570      0.185
## 8             sigma2  3.539 0.209       3.138      3.957
## 9      vic_worst_force  1.171 0.126       0.925      1.413
## 10     vic_worst_incap  0.533 0.088       0.361      0.707
## 11 vic_worst_prior_sexTRUE -0.684 0.167      -1.005     -0.356
## 12     vic_worst_verbal  0.244 0.114       0.018      0.464
## 13           week_num -0.036 0.025      -0.085      0.013
## 14      week_num_SQ -0.001 0.000      -0.002      0.000

write.csv(
  x = h2_param_table,
  file = paste0('h2_param_table_', Sys.Date(), '.csv'),
  row.names = F)

```

# MeToo Movement and Perceptions of Sexual Assault:

## Section 3

Below are the additional analyses that occurred after primary model fitting. Among them are estimates of sample size and the computation of the #MeToo effect (post hoc interaction analyses).

### Setup - load participant dataframe

For cleanliness, we reload the data and model formulas from earlier in the analysis.

```
options(knitr.duplicate.label = 'allow')

knitr::knit('data_wrangling.Rmd', output = tempfile())
knitr::knit('models_and_priors.Rmd', output = tempfile())
```

### Effective sample size (ESS)

To evaluate the strength of our priors, we used Effective Sample Size methods from Morita et al. (2008), implemented in the BayesESS package.

For the constant prevalence hypothesis, the ESS was...

```
library(tidyverse)

set.seed(5647891)

reg_priors <- rerun(9, c('norm', 0, 10))

BayesESS::ess(
  model = 'logistic',
  prior = reg_priors,
  ncov = 10,
  m = 5,
  nsim = 1000,
  fast = T,
  svec1 = rep(1, 12),
  svec2 = rep(1, 12))

## 
## ESS was calculated for a logistic regression model
##
## ESSoverall: Overall ESS for the whole vector (theta1,theta2,theta3,theta4,theta5,theta6,theta7,theta8)
## ESSsubvector1: ESS for the first sub-vector (theta1,theta2,theta3,theta4,theta5,theta6,theta7,theta8)
## ESSsubvector2: ESS for the second sub-vector (theta1,theta2,theta3,theta4,theta5,theta6,theta7,theta8)

## $ESSoverall
## [1] 0.8213573
##
## $ESSsubvec1
```

```

## [1] 0.8213573
##
## $ESSsubvec2
## [1] 0.8213573

```

For the increasing acknowledgment hypothesis, the ESS was...

```

set.seed(5647891)

reg_priors <- rerun(11, c('norm', 0, 10))
reg_priors[[length(reg_priors) + 1]] <- c('gamma', 1, 1)

BayesESS::ess(
  model = 'linreg',
  prior = reg_priors,
  ncov = 10,
  m = 5,
  nsim = 1000,
  fast = T,
  svec1 = rep(1, 12),
  svec2 = rep(1, 12))

##
## ESS was calculated for a linear regression model
##
## ESSoverall: Overall ESS for the whole vector (theta1,theta2,theta3,theta4,theta5,theta6,theta7,theta8)
## ESSsubvector1: ESS for the first sub-vector (theta1,theta2,theta3,theta4,theta5,theta6,theta7,theta8)
## ESSsubvector2: ESS for the second sub-vector (theta1,theta2,theta3,theta4,theta5,theta6,theta7,theta8)

## $ESSoverall
## [1] 0.4292039
##
## $ESSsubvec1
## [1] 0.4292039
##
## $ESSsubvec2
## [1] 0.4292039

```

## Region of practical equivalence (ROPE)

For the H1 ROPE, we established a practical metric to evaluate changes in prevalence: the advantage or disadvantage of being in the sample's state, then moving to any other state at random. We calculated this by taking the average absolute difference between the study sample state and every other state.

State-level data were acquired from [Table 3.9 of this report](#) (<https://www.cdc.gov/violenceprevention/pdf/NISVS-StateReportBook.pdf>) and reflect the lifetime prevalence of contact sexual violence in **women**.

```

h1_rope <- readr::read_csv('rope_materials/NIP&SV_survey_2012.csv') %>%
  mutate(
    NE_change =
      abs(sex_vio_lifetime_prev -
          sex_vio_lifetime_prev[which(state.name == 'Nebraska')])) %>%

```

```

filter(State != 'Nebraska') %>%
dplyr::select(NE_change) %>%
unlist() %>%
mean() %>%
magrittr::divide_by(100) %>%
round(2)

h1_rope

## [1] 0.03

```

To clarify, this means any change - up or down - in the prevalence of reported sexual assault that is within  $\pm 3\%$  of the pre-#MeToo prevalence of sexual assault would be regarded as “negligibly different from no change.” This is because any change within that range is statistically equivalent to simply choosing another state in which to live at random. This is admittedly a rough metric and future research will be needed to come up with a more comprehensive set of benchmarks; however, it is used here for its simplicity and clarity, in lieu of another approach.

For the H2 ROPE, which relies on linear regression, consensus benchmarks for effect sizes are already well established. For that reason, we choose a ROPE of  $d = \pm .10$ , which would be a negligible effect size according to Cohen (1988).

```

d <- .10

h2_rope <- d*sd(df$vic_worst_is_sa, na.rm = T)
h2_rope

## [1] 0.2204741

```

## Quadratic effect probes

The effect of the #MeToo movement involves three terms in each model: a dichotomous term, a linear term, and a quadratic term. Because none of these terms represents the overall #MeToo effect in isolation, we have incorporated them all into a single metric that accounts for all of their influence simultaneously. To accomplish this, we followed Aiken and West (1991) and calculated the derivative of our regression equation (with respect to the #MeToo dummy variable).

This is simplest in the case of the linear regression, which requires only a basic application of the power rule. Specifically, if

$$y = b_0 + b_1(t) + b_2(\text{MeToo}) + b_3(\text{MeToo})(t) + b_4(t^2) + b_5(\text{MeToo})(t^2)$$

Where,  $t$  is *time*, then

$$\frac{dy}{d\text{MeToo}} = b_2 + b_3t + b_5t^2$$

This is slightly more complicated in the logistic regression from H1, where the log-link function must also be taken into account. In that case, a combination of the chain and power rule thus yields,

$$\frac{dy}{d\text{MeToo}} = p(1-p)(b_2 + b_3t + b_5t^2)$$

where  $p$  is any sensible value for baseline prevalence of sexual assault. In this case, we have chosen the mean sexual assault prevalence across all pre-#MeToo weeks in the dataset  $\approx 24\%$ .

```
pre_metoo_vic_prob <- df %>%
  filter(post_metoo == 0) %>%
  dplyr::select(vic_screener) %>%
  unlist(use.names = F) %>%
  map_lgl(~ .x == 'Yes') %>%
  mean(na.rm = T)
```

```
pre_metoo_vic_prob
```

```
## [1] 0.2382325
```

### Implementation of post-hoc #MeToo probes

To test this, we first import the posterior distributions from both of our models.

```
latest_h1_posterior <- max(list.files() [grep('h1_posteriors', list.files())])
latest_h2_posterior <- max(list.files() [grep('h2_posteriors', list.files())])
```

```
# h1_post <- readr::read_csv('h1_posteriors_2019-03-20.csv')
# h2_post <- readr::read_csv('h2_posteriors_2019-03-20.csv')
```

```
h1_post <- readr::read_csv(latest_h1_posterior)
```

```
## Parsed with column specification:
## cols(
##   model = col_character(),
##   param = col_character(),
##   estimate = col_double()
## )
```

```
h2_post <- readr::read_csv(latest_h2_posterior)
```

```
## Parsed with column specification:
## cols(
##   model = col_character(),
##   param = col_character(),
##   estimate = col_double()
## )
```

## H1

Beginning with H1, we probe the hypotheses at each of the 0-80 weeks after the MeToo movement starts by extracting the parameters of interest from our posterior distribution and multiplying by our probe points

```
h1_param_matrix <- h1_post %>%
  filter(model == 'post_metoo_time_SQ') %>%
  mutate(i = rep(1:(6*1e4), times = 9)) %>%
```

```

spread(param, estimate) %>%
dplyr::select(post_metoo1, post_metoo_week, post_metoo_week_SQ) %>%
as.matrix()

# study included 80 weeks after #metoo
probe_points <- cbind(1, 0:80, (0:80)^2)
h1_probes <- probe_points %*% t(h1_param_matrix)

# 60k samples for each of the 28 weeks
dim(h1_probes)

## [1] 81 60000

```

We then reformat the results into a dataframe.

```

h1_probes <- h1_probes %>%
  as.data.frame() %>%
  set_names(1:length(.)) %>%
  mutate(week_num = row_number()) %>%
  gather(mcmc_samp, probe_est, -week_num) %>%
  mutate(mcmc_samp = as.numeric(mcmc_samp)) %>%
  arrange(week_num, mcmc_samp)

h1_probe_results <- h1_probes %>%
  mutate(
    # See equations above for clarification
    prev_est = pre_metoo_vic_prob*(1 - pre_metoo_vic_prob)*probe_est) %>%
  group_by(week_num) %>%
  summarise(
    mean_prev_est = mean(prev_est),
    below_rope = mean(prev_est < -h1_rope),
    within_rope = mean(-h1_rope < prev_est & prev_est < h1_rope),
    above_rope = mean(h1_rope < prev_est)) %>%
  round(2) %>%
  ungroup() %>%
  filter(week_num %% 10 == 0)

h1_probe_results

## # A tibble: 8 x 5
##   week_num mean_prev_est below_rope within_rope above_rope
##       <dbl>        <dbl>      <dbl>      <dbl>      <dbl>
## 1       10       -0.05      0.61     0.290      0.1
## 2       20       -0.04      0.54     0.22       0.24
## 3       30       -0.03      0.5       0.16      0.34
## 4       40       -0.01      0.47     0.12       0.41
## 5       50        0.01      0.45     0.09      0.46
## 6       60        0.03      0.44     0.07      0.5
## 7       70        0.05      0.43     0.05      0.52
## 8       80        0.08      0.42     0.04      0.54

```

```
write.csv(h1_probe_results, paste0('h1_probe_results_', Sys.Date(), '.csv'), row.names = F)
```

Results show the posterior distribution is always overlapping - but not contained within - the ROPE. Thus, we can only conclude that we are uncertain about the population effect.

## H2

The same procedure is repeated for H2.

```
h2_post <- readr::read_csv(latest_h2_posterior)
```

```
## Parsed with column specification:  
## cols(  
##   model = col_character(),  
##   param = col_character(),  
##   estimate = col_double()  
## )  
  
h2_param_matrix <- h2_post %>%  
  filter(model == 'post_metoo_time_SQ') %>%  
  mutate(i = rep(1:1e4, times = 14)) %>%  
  spread(param, estimate) %>%  
  dplyr::select(post_metoo1, post_metoo_week, post_metoo_week_SQ) %>%  
  as.matrix()  
  
wks <- 1:80  
probe_points <- cbind(1, wks, wks^2)  
  
h2_probes <- probe_points %*% t(h2_param_matrix)
```

```
h2_probes <- h2_probes %>%  
  as.data.frame() %>%  
  set_names(1:length(.)) %>%  
  mutate(week_num = row_number()) %>%  
  gather(mcmc_samp, probe_est, -week_num) %>%  
  mutate(mcmc_samp = as.numeric(mcmc_samp)) %>%  
  arrange(week_num, mcmc_samp)
```

```
h2_probe_results <- h2_probes %>%  
  group_by(week_num) %>%  
  summarise(  
    below_rope = mean(probe_est < -h2_rope),  
    within_rope = mean(-h2_rope < probe_est & probe_est < h2_rope),  
    above_rope = mean(h2_rope < probe_est)) %>%  
  round(2) %>%  
  ungroup() %>%  
  filter(week_num %% 10 == 0)  
  
h2_probe_results
```

```

## # A tibble: 8 x 4
##   week_num below_rope within_rope above_rope
##   <dbl>      <dbl>      <dbl>      <dbl>
## 1      10     0.03     0.11     0.86
## 2      20     0.03     0.06     0.91
## 3      30     0.04     0.04     0.92
## 4      40     0.04     0.02     0.93
## 5      50     0.04     0.02     0.94
## 6      60     0.04     0.01     0.94
## 7      70     0.04     0.01     0.95
## 8      80     0.04     0.01     0.95

write.csv(h2_probe_results, paste0('h2_probe_results_', Sys.Date(), '.csv'), row.names = F)

```

## Power

The sample was collected as part of another study, so N was fixed before analyses could even begin. To that end, we conducted a sensitivity analysis to see what kinds of measurement precision and hypothesis testing power were available for a sample with this pattern of data and covariate relationships.

This process involved first taking the posterior distributions for both hypotheses and assuming that their means and covariance matrices were approximately “correct.” Then, we modified the means and standard deviations for the parameters involved in the MeToo effect displayed above.

```

knitr::knit('data_wrangling.Rmd', output = tempfile())
knitr::knit('analysis_functions.Rmd', output = tempfile())
knitr::knit('models_and_priors.Rmd', output = tempfile())

```

Note, to ensure power estimation accounted for missing data, we used a multiply-imputed dataset for our simulations. This will better mimic actual power in our analyses than would using the observed dataset with listwise deletion.

```

library(mice)

categorical_vars <- c( 'post_metoo', 'male', 'lgbtq', 'racial_minority',
                      'vic_screener', 'vic_worst_prior_sex')

df[, categorical_vars] <- purrr::map(df[, categorical_vars], as.factor)

imp_seeds <- map_int(1:parallel::detectCores(), ~ as.integer(runif(1)*1e8))
sets_per_core <- 2

full_imp <- parallel::mclapply(
  X = imp_seeds,
  FUN = impute_mice,
  data = df,
  m = sets_per_core) %>%
  unpack_parallel_mids()

vic_imp <- parallel::mclapply(
  X = imp_seeds,
  FUN = impute_mice,

```

```

    data = filter(df, vic_screener == 'Yes'),
    m = sets_per_core) %>%
  unpack_parallel_mids()

paste(sets_per_core, 'imputed datasets per core')

## [1] "2 imputed datasets per core"

paste(parallel::detectCores(), 'cores detected')

## [1] "4 cores detected"

```

## H1

For H1, we initially reload the posterior distributions for all model parameters from the main analysis.

```

h1_full_params <- h1_post %>%
  filter(model == 'post_metoo_time_SQ') %>%
  group_by(param) %>%
  mutate(i = row_number()) %>%
  # filter(i <= 1000) %>% XXX shrink posterior sample size when testing
  spread(param, estimate) %>%
  dplyr::select(-model, -i) %>%
  as.matrix()

head(h1_full_params)

##      (Intercept) lgbtqTRUE maleTRUE post_metoo_week post_metoo_week_SQ
## [1,] -0.8886429  0.7021671 -1.699400   0.026482121 -5.981125e-05
## [2,] -0.5937209  0.8038032 -1.717845  -0.026519924 -1.187580e-04
## [3,] -0.5015628  0.3757730 -1.635688   0.012636844  3.116562e-04
## [4,] -0.6471048  0.4610756 -1.664365   0.008383404  4.118130e-04
## [5,] -0.9455278  0.9011552 -1.560569   0.006029041  5.471285e-04
## [6,] -0.5702671  0.7200323 -1.691096   0.016829588  5.242529e-04
##      post_metoo1 racial_minority1 week_num week_num_SQ
## [1,] -0.5739410     -0.2485226 -0.003971182 -0.0001507965
## [2,] -0.2015509     -0.1793750  0.019833539  0.0002276976
## [3,] -0.5931556     -0.3809182 -0.003389214 -0.0003712934
## [4,] -0.1409018     -0.3297326 -0.008010630 -0.0004108269
## [5,]  0.1094466     -0.2327070 -0.014911860 -0.0004113637
## [6,] -0.3998894     -0.4711051 -0.011570891 -0.0005208271

```

We then isolated the variables involved in the power analysis and create a population-level effect for them, which will later be inserted into our simulated datasets for power analysis. In the case of H1, we created a population-level effect for the post-hoc effect of #MeToo that was centered at 0 and always contained within the ROPE, 6 months after #MeToo.

```

power_vars <- c('post_metoo1', 'post_metoo_week', 'post_metoo_week_SQ')
param_means <- colMeans(h1_full_params)
param_sds <- apply(h1_full_params, 2, sd)

```

```

param_cor <- cor(h1_full_params)

# This assumes no sudden shift, followed by no gradual, no accelerating growth
param_means[power_vars] <- c(0, 0, 0)

# values chosen by guessing and checking in next code block
# to produce 95% variance within H1 ROPE
param_sds[power_vars] <- c(.001, .0001, .00001)

param_cov <- diag(param_sds) %*% param_cor %*% diag(param_sds)

set.seed(76846235)

param_samples <- mvtnorm::rmvnorm(
  n = 10,
  mean = param_means,
  sigma = param_cov)

sim_metoo_effect <- param_samples[, power_vars] %*% c(1, 24, 24^2)
mean(pre_metoo_vic_prob*(1 - pre_metoo_vic_prob)*sim_metoo_effect)

## [1] -0.0004252763

1.96*sd(sim_metoo_effect)

## [1] 0.01211935

round(mean(sim_metoo_effect > - h1_rope & sim_metoo_effect < h1_rope), 3)

```

With the population-level effect distribution created, we inject samples of this effect into each of 100 datasets, then analyze them using our main analysis model.

```
design_names <- colnames(param_samples) %>%
  gsub('TRUE|1|\\\"|\\\'|', '', .)

design_matrix <- full_imp$`1` %>%
  sample_n(size = nrow(.) * 100, replace = T) %>%
  map_df(as.numeric) %>%
  cbind(Intercept = 1, .) %>%
  as.matrix()

sim_logit <- design_matrix[, design_names] %*% t(param_samples)
sim_p <- exp(sim_logit) / (1 + exp(sim_logit))
sim_y <- runif(length(sim_p)) <- sim_p

sim_y <- sim_y %>%
  as.data.frame() %>%
```

```

set_names(paste0(' ', 1:ncol(.))) %>%
  mutate(id = row_number()) %>%
  gather(samp, y, -id)

sim_df <- full_imp$`1` %>%
  mutate(id = row_number()) %>%
  right_join(sim_y, by = 'id') %>%
  mutate(vic_screener = y)

```

We then analyze them without the main analysis model...

```

library(HDInterval)

Sys.time()

## [1] "2020-05-23 10:41:03 EDT"

h1_power_models <- sim_df %>%
  split(.\$samp) %>%
  map(
    .f = ~ MCMClogit(
      formula = h1_models$post_metroo_time_SQ,
      b0 = 0,
      B0 = 1/10,
      data = .x))

Sys.time()

## [1] "2020-05-23 10:41:33 EDT"

```

... and compile the results from those analyses.

```

h1_power_posteriors <- h1_power_models %>%
  map(
    .f = ~ c(
      pre_metroo_vic_prob*(1-pre_metroo_vic_prob)*.x[, power_vars] %*%
        c(1, 24, 24^2)))

h1_power_results <- h1_power_posteriors %>%
  map(
    .f = ~ data.frame(
      mean_est = mean(.x),
      perc_within_rope = mean(h1_rope < .x & .x < h1_rope),
      perc_within_pm10 = mean(-.10 < .x & .x < .10),
      perc_within_pm20 = mean(-.20 < .x & .x < .20),
      perc_within_pm30 = mean(-.30 < .x & .x < .30),
      lhdi = hdi(.x)[['lower']],
      uhdi = hdi(.x)[['upper']]))) %>%
  bind_rows(.id = 'samp') %>%
  summarise(
    mean_est = mean(mean_est),

```

```

power_for_rope = mean(perc_within_rope >= .95),
power_for_pm10 = mean(perc_within_pm10 >= .95),
power_for_pm20 = mean(perc_within_pm20 >= .95),
power_for_pm30 = mean(perc_within_pm30 >= .95),
mean_hpdi_width = mean(uhdi - lhdi))

h1_power_results

##      mean_est power_for_rope power_for_pm10 power_for_pm20 power_for_pm30
## 1 0.01381373          0           0           0           0.2
##      mean_hpdi_width
## 1      0.5839684

```

## H2

For H2, we repeat an identical procedure as in H1.

```

h2_full_params <- h2_post %>%
  filter(model == 'post_metoo_time_SQ') %>%
  mutate(i = rep(1:1e4, times = 14)) %>%
  spread(param, estimate) %>%
  dplyr::select(-model, -i) %>%
  as.matrix()

power_vars <- c('post_metoo1', 'post_metoo_week', 'post_metoo_week_SQ')
param_means <- colMeans(h2_full_params)
param_sds <- apply(h2_full_params, 2, sd)
param_cor <- cor(h2_full_params)

# values that produce an effect of about .30 by week 24 post metoo
# This assumes no sudden shift, followed by gradual, accelerating growth
param_means[power_vars] <- c(0, .04, .0001)

# values chosen by guessing and checking in next code block
# to produce variance for post-metoo effect of about
# +1-point on the outcome scale
param_sds[power_vars] <- c(.01, .02, .0001)

param_cov <- diag(param_sds) %*% param_cor %*% diag(param_sds)

set.seed(156498416)

param_samples <- mvtnorm::rmvnorm(
  n = 100,
  mean = param_means,
  sigma = param_cov)

sim_metoo_effect <- param_samples[, power_vars] %*% c(1, 24, 24^2)
mean(sim_metoo_effect)

## [1] 0.9759749

```

```

sd(sim_metoo_effect)

## [1] 0.5252076

round(mean(sim_metoo_effect > .10), 3)

## [1] 0.96

sim_error <- param_samples[, 'sigma2']
param_samples <- param_samples[, colnames(param_samples) != 'sigma2']

design_names <- colnames(param_samples) %>%
  gsub('TRUE|1|\\)|\\\\(| ', '', .)

design_matrix <- vic_imp$`1` %>%
  map_df(as.numeric) %>%
  cbind(Intercept = 1, .) %>%
  as.matrix()

sim_y <- design_matrix[, design_names] %*% t(param_samples) + sim_error

sim_y <- sim_y %>%
  as.data.frame() %>%
  set_names(paste0('', 1:ncol(.))) %>%
  mutate(id = row_number()) %>%
  gather(samp, y, -id)

sim_df <- vic_imp$`1` %>%
  mutate(id = row_number()) %>%
  right_join(sim_y, by = 'id') %>%
  mutate(vic_worst_is_sa = y)

```

Results are again estimated using the analysis model and then compiled to produce a power estimate.

```

library(HDInterval)

h2_power_sims <- sim_df %>%
  split(.samp) %>%
  map(
    .f = ~ MCMCregress(
      formula = h2_models$post_memeo_time_SQ,
      b0 = 0,
      B0 = 1/10,
      data = .x)) %>%
  map(~ c(.x[, power_vars] %*% c(1, 24, 24^2))) %>%
  map(
    .f = ~ data.frame(
      mean_est = mean(.x),
      perc_above_rope = mean(.x > h2_rope),
      lhdi = hdi(.x)[['lower']],
      uhdi = hdi(.x)[['upper']])) %>%

```

```
bind_rows(.id = 'samp') %>%
summarise(
  power = mean(perc_above_rope > .95),
  mean_hdi_width = mean(uhdi - lhdi),
  sd_hd_width = sd(uhdi - lhdi))

h2_power_sims

##   power mean_hdi_width sd_hd_width
## 1  0.85      0.4896485 0.001734127
```

END OF DOCUMENT