Supplementary Material to the Article

Koko M, Motelow JE, Stanley KE, Bobbili DR, Dhindsa RS, May P; Canadian Epilepsy Network; Epi4K Consortium; Epilepsy Phenome/Genome Project, EpiPGX Consortium; EuroEPINOMICS-CoGIE Consortium. Association of ultra-rare coding variants with genetic generalized epilepsy: a casecontrol whole exome sequencing study. *Epilepsia.* 2022. https://doi.org/10.1111/epi.17166

Table of Contents

Supplementary Methods:

The first dataset:

Samples:

Participants whose sequencing data formed the first dataset included 1,214 GGE patients recruited by the Epi4K Consortium and Epilepsy Phenome/Genome Project as previously described,^{1,2} and sequenced at the Institute for Genomic Medicine (IGM) at Columbia University (New York, USA). The diagnosis of a GGE syndrome required the patients to have generalized epilepsy with absence, myoclonic or tonic-clonic seizures and generalized spike-and-wave discharge on electroencephalography (EEG). To qualify for the familial analysis, patients were required to have at least one relative (up to the third degree) who had been diagnosed with epilepsy. ¹ Ancestry matched controls (*n* = 14,100 before quality control) were selected from multiple collections of control cohorts at the IGM. 3

Sequence data generation:

Whole exome sequencing of DNA samples from participants forming the first dataset was performed at IGM using Illumina's HiSeq 2000, HiSeq 2500 or NovaSeq 6000 platforms (Illumina, San Diego, CA, USA) following enrichment with Agilent All Exon Enrichment kits (Agilent Technologies, Santa Clara, CA, USA), NimbleGen SeqCap EZ Exome Enrichment kit (Roche NimbleGen, Madison, WI, USA), Twist Human Core Exome (Twist Bioscience, San Francisco, CA, USA) or IDT xGen Exome Research Panel (Integrated DNA Technologies, Coralville, IA, USA). The sequence data from all cases and controls were processed according to the IGM bioinformatics pipeline.^{1,4} Sequencing reads were aligned to the human reference genome build 37 (GRCh37) using Illumina's Dynamic Read Analysis for GENomics (DRAGEN) Bio-IT Platform.^{5,6} Picard (https://broadinstitute.github.io/picard/) and the Genome Analysis Tool Kit v3.6 (GATK) were used to perform duplicate read marking, base quality scores recalibration, indel realignment and haplotype calling. The samples were processed individually at different time points and the variants obtained from single sample calling were imported and integrated in the Analysis Tool for Annotated Variants (ATAV) Database. 4

Sample quality control:

For the purpose of this study, samples with possible contamination (heterozygosity exceeding 2%) determined using VerifyBamID, ⁷ with discordance between self-declared and sequence-derived sex, or with low coverage (less than 85% of the consensus coding sequence release 20 (CCDS20) targets covered at a minimum of 10x) were removed. Related individuals were identified using Kinship-based Inference for GWAS⁸ (KING). One of each pair that had an inferred relationship of third-degree or

closer was dropped, preferentially retaining affected over control individuals and samples with higher coverage for pairs with similar disease status. EIGENSTRAT⁹ was then used to remove ethnicity outliers to minimize the effects of residual population stratification (Fig. S2).

Variant quality control and call rate harmonization:

The following variant-level parameters (hard filters) were enforced: Quality/Depth (QD) > 5 , Quality (QUAL) > 50, Mapping Quality (MQ) > 40, SOR < 3 (SNVs) or < 10 (indels), Fisher's Strand bias score (FS) ≤ 60 (SNVs) or ≤ 200 (indels), Read Position Rank Sum score (RPRS) ≤ -3 , and Mapping Quality Rank Sum score (MQRS) < -10. Variants were required to pass GATK Variant Quality Score Recalibration (VQSR) filter. Known artifacts and variants failing quality filters in population databases (Exome Variant Server, ExAC, gnomAD) were excluded. Low quality genotype calls with total allelic depth (DP) \leq 10 or genotype quality (GQ) \leq 20 were filtered. Heterozygous calls had a minimum alternate allele fraction (AD/DP) of 0.3. As previously described, a coverage harmonization procedure was employed to remove the variants that are differentially covered across the cases and controls.¹ Briefly, this was based on plotted the cumulative difference in site coverage between cases and controls to identify a filtering cut-off that will minimize this difference while allowing the largest possible number of variants to be retained.

The second dataset:

Samples:

The individuals with generalized epilepsy analyzed here were selected from 2,524 individuals recruited by the EuroEPINOMICS-CoGIE Consortia, EpiPGX Consortium, and CENet as described previously.¹⁰ For the purpose of this work, we used the sequence data of 989 individuals ascertained to have classical GGE phenotypes (childhood or juvenile absence epilepsy, juvenile myoclonic epilepsy, or epilepsy with generalized tonic-clonic seizures alone), early-onset absence epilepsy (age of onset < 3 years), epilepsy with myoclonic absences, or unclassified GGE. Familial cases had one or more, selfreported, first- or second-degree, affected relative. The controls for this dataset $(n = 4.904)$ before quality control) were obtained from the database of Genotypes and Phenotypes¹¹ (dbGAP studies: MIGen Ottawa Heart Study controls, Rotterdam study controls and Alzheimer Disease Genetics Study controls) or from the Epi25 Collaborative.¹²

Sequence data generation:

Whole-exome sequencing of EuroEPINOMICS-CoGIE cases was done on Illumina's HiSeq 2000 platform using NimbleGen SeqCap EZ Human Exome Library (NimbleGen, Madison, WI, USA) at Cologne Center for Genomics (Cologne, Germany). Whole-exome sequencing for the EpiPGX cohort was done at deCODE genetics (Reykjavik, Iceland) on Illumina's HiSeq 2500 platform with Nextera Rapid Capture Expanded Exome kit (Illumina, San Diego, CA, USA). Whole exome sequencing of individuals recruited by the Canadian Epilepsy Network (CENet) was performed by the McGill University and Génome Québec Innovation Center (MAGQUIC, Québec, Canada) on Illumina's HiSeq sequencing platforms using TruSeq or Roche Nimblegen EZ libraries. Controls from the Epi25 Collaborative were sequenced at the Broad Institute of Harvard and the Massachusetts Institute of Technology on Illumina's HiSeq platform using Illumina's Nextera Rapid Capture or TruSeq Rapid Exome enrichment kits. The Rotterdam Study controls were sequenced on Illumina's HiSeq 2000 platform using EZ Human Exome Library. The Alzheimer study controls were sequenced over multiple time points at the Broad Institute using different capture kits. Fastq files were aligned to GRCh37 as previously described^{10,13} and jointly called using DRAGEN Bio-IT Platform.^{5,6}

Sample quality control:

The sample-level call rate, autosomal and chrX inbreeding coefficients were collected using Plink¹⁴ v1.9. and Picard (from $GATK¹⁵$ v4.1.4.1). Samples with phenotypes other than GGEs or without appropriate permissions for inclusion and samples with extremely low variant counts (< 10,000 nonmissing calls) were removed. Samples with genotyping rates lower than 80%, outlier samples on autosomal heterozygosity (> 4 median absolute deviations on autosomal inbreeding coefficient estimates), and samples with discordant or ambiguous sequencing sex based on chromosome X inbreeding co-efficient estimates ($F < 0.3$ for female and $F > 0.7$ for male predicted sequencing sex) were excluded. The remaining samples were scanned for relatedness (third degree) using KING. ⁸ For duplicates and pairs with matching phenotypes, the sample with the higher genotyping rate was retained. Otherwise, cases were preferentially retained. Next, multi-dimensional scaling (MDS) was used to project the major continental ancestry of the study samples on the MDS space of 1000 Genomes data (2,504 samples) using KING. The top principal components were visualized and used to classify the ancestry with a support vector machine using R package *e1071*. ¹⁶ Samples with predicted European ancestry were retained. Also, samples clustering with Finnish 1000 Genomes samples on PC1/2 were filtered. Following the baseline variant filtering steps outlined below, the variant calling metrics were re-examined to exclude any additional sample outliers. Here, all samples with SNV counts < 15,000 were filtered (this removed all Rotterdam Study controls and most Alzheimer Study controls). Outliers beyond 3 standard deviations per cohort on key variant calling metrics (Heterozygous-Homozygous calls ratio, Transitions-Transversions ratio, and Insertions-Deletions ratio) were filtered. To ensure adequate case control matching and the removal of ancestry outliers, PCA analysis using EIGENSTRAT⁹ was employed (Fig. S2).

Variant quality control and call rate harmonization:

The variants were filtered for those located in the CCDS exonic coding regions (padded on each side to accommodate canonical splice sites and masked for low-complexity regions) using bcftools¹⁷ v1.9. The variants were decomposed, normalized and sorted using beftools and vt^{18} v0.5. Low quality genotypes were filtered by setting calls with total allelic depth ≤ 10 or genotype quality ≤ 20 to missing. Heterozygous calls had a minimum alternate allele depth fraction (AD/DP) of 0.25. This genotype filtering was performed using bcftools. A combination of hard filtering and filtering based on recalibrated variant quality scores was employed to remove low quality variants. Variant calls with low quality were filtered (SNVs: QUAL < 10, QD < 2, MQ < 30, FS > 60, MQRankSum < -12.5, RPRS < -8; Indels: QUAL < 10, QD < 2, RPRS < -20, FS > 200). Variant Quality Score Recalibration (VQSR) was performed on the normalized and genotype-filtered call set using GATK based on these annotations: QD, FS, SOR, MQRankSum, and RPRS. SNVs and Indels failing VQSR Tranche 99.0 filter were removed. Since the datasets were sequenced using different capture kits, we performed additional harmonization steps to limit our analysis to the coding regions covered in all kits & to minimize the spurious effects caused differences in capture kits. Variants were retained only if they had genotyping rates $\geq 90\%$ both in EpiPGX cases (largest case dataset; representing Illumina capture targets) and MIGen Ottawa controls (largest controls dataset; representing Agilent capture targets). After removal of sample outliers (see above), a final round of call rate harmonization was then performed where the variant call rate was calculated among the remaining cases and controls and variants with call rates below 95% in cases or controls were filtered. Also, the cumulative difference of call rate between cases and controls was plotted and 9.4% of the variants were removed to minimize this difference while retaining the largest possible number of variants.

Duplicates and ancestry harmonization across cohorts:

Ancestry matching between the case cohorts:

To maximize the ancestry matching between the two analyzed patient cohorts, the ancestry prediction among the cases was harmonized in our two cases datasets by using the same ancestry prediction model to ensure homogeneity in ancestry assignment. Principal components analysis was performed on genotypes of previously defined well covered exonic autosomal polymorphic markers. ¹ A neural network model that uses the first five principal component axes as the independent variables, trained on more than two thousand individuals with pre-evaluated genetic ancestry from six ethnic groups (European, Middle Eastern, Hispanic, East Asian, South Asian, and African), was then used to predict the probability of a European ancestry. Cases with < 95% probability were excluded.

Duplicates between the two case cohorts:

To exclude likely duplicates between the two case cohorts, a genotype hashing approach adopted from the Gencrypt method¹⁹ was used to avoid the need for genotype sharing across the two study sites. A group of variants with minor allele frequency > 0.1 and genotyping rate $> 98\%$ in both cohorts was identified. From this pool, 200 sets were created, each consisting of randomly selected non-overlapping 150 SNPs. For each sample, the genotypes over each set were concatenated keeping their order and converted to *sha256* cryptographic hashes. The hashes were exchanged and compared between cohorts. In total, 57 cases shared one or more hashes (according, likely to have identical genotypes in > 150 randomly selected polymorphic markers) were considered possible duplicates. These were retained only in the first dataset and were removed from the second set.

Overlap with the Epi25 Collaborative datasets:

There is marginal overlap between the samples of individuals with GGE and the published analyses of the Epi25 Collaborative (before QC: less than 100 samples from the EuroEUROEPINOMICS-CoGIE and EpiPGX dataset). The controls used for the first dataset were similar to those used in the Epi25 Collaborative Y1-3 analysis (internal control datasets of the Institute for Genomics Medicine, NY). The controls used for the second dataset had substantial overlap with the Epi25 Collaborative Y1-2 analysis (978 individuals from MIGen Ottawa study and 332 individuals from Epi25 controls).

Qualifying Variants' distribution plots:

To ensure that we achieved an adequate case control matching and coverage/call rate harmonization in each dataset, we examined the distribution plots of qualifying variants tallies. Variant tallies were examined separately for each study dataset and collectively for the final merged dataset. The significance of the differences in the distribution density of ultra-rare synonymous variants was examined using Wilcoxon Rank Sum test with continuity correction as implemented in R (Fig. S3 – S5).

Quantile-Quantile (QQ) plots:

To obtain cohort-level QQ plots, the case-control labels were shuffled before running 1,000 iterations of gene collapsing analyses using Fisher's exact test. *P* values were ranked from these permutations and the negative log_{10} of the mean *p* values at each rank was plotted against the negative log_{10} of observed *p* values. Confidence intervals were obtained by plotting negative log_{10} of the expected *p* values corresponding the 2.5th and 97th percentiles from the permutations. The genomic inflation factors (λ) were calculated according to the previously described regression method implemented in the function *estlambda2()* from R package *QQperm*. 20

The QQ plots for the combined analysis (Cochran-Mantel-Haenszel exact test) show the *p* values from those genes with at least one qualifying variant in the joint case cohort and the expected *p* values from a uniform distribution. The negative log_{10} of the observed *p* values was plotted against the negative log_{10} of an equal number of uniformly distributed *p* values $(-log_{10}((k-0.5)/n))$, where *k* is the gene rank and *n* the total genes). The confidence intervals for the expected *p* values were based on values drawn from a beta distribution $(-\log_{10}(\phi \text{beta}(\alpha/2, k, n-k \text{ and } -\log_{10}(\phi \text{beta}(\alpha/1-\alpha/2, k, n-k), \text{ where } \alpha = 0.05 \text{ for a})$ 95% confidence interval) using the *stats* package²¹ in R 3.3. The genomic inflation factors (λ) was calculated using *QQperm*.

Gene sets for gene set association analysis:

The list of genes composing the two gene sets of inhibitory signaling tested in this study were obtained from previously published work.^{10,12}

Genes encoding GABAA receptors: *GABRA1, GABRA2, GABRA3, GABRA4, GABRA5, GABRA6, GABRB1, GABRB2, GABRB3, GABRD, GABRE, GABRG1, GABRG2, GABRG3, GABRP, GABRQ, GABRR1, GABRR2, GABRR3*.

GABAergic pathway genes: *ABAT, ADCY1, ADCY2, ADCY3, ADCY4, ADCY5, ADCY6, ADCY7, ADCY8, ADCY9, ANK2, ANK3, ARHGEF9, DISC1, DLC1, DNAI1, FGF13, GABARAP, GABARAPL1, GABARAPL2, GABBR1, GABBR2, GABRA1, GABRA2, GABRA3, GABRA4, GABRA5, GABRA6, GABRB1, GABRB2, GABRB3, GABRD, GABRE, GABRG1, GABRG2, GABRG3, GABRP, GABRQ, GABRR1, GABRR2, GABRR3, GAD1, GAD2, GLS, GLS2, GLUL, GNAI1, GNAI2, GNAI3, GNAO1, GNB1, GNB2, GNB3, GNB4, GNB5, GNG10, GNG11, GNG12, GNG13, GNG2, GNG3, GNG4, GNG5, GNG7, GNG8, GNGT1, GNGT2, GPHN, HAP1, KCNB2, KCNC1, KCNC2, KCNC3, KCNJ6, KIF5A, KIF5B, KIF5C, MAGI1, MKLN1, MTOR, MYO5A, NLGN2, NRXN1, NSF, PFN1, PLCL1, PRKACA, PRKACB, PRKACG, PRKCA, PRKCB, PRKCG, RDX, SCN1A, SCN1B, SCN2B, SCN3A, SCN8A, SEMA4D, SLC12A2, SLC12A5, SLC32A1, SLC38A1, SLC38A2, SLC38A3, SLC38A5, SLC6A1, SLC6A11, SLC6A13, SRC, STARD13, TRAK1, TRAK2*.

Overrepresentation of gene sets among top ranked genes:

A hypergeometric test was employed to examine the probability that *n* genes from a gene set of *N* genes appeared by chance among the top-ranked *k* genes when examining a total of 18,834 protein coding genes. The enrichment was tested at each rank *k* occupied by a gene from the gene set using a using *phyper* function from R *stats* package as follows: *phyper*(*n*-1, *N*, 18834-*N*, *k*, lower.tail= FALSE). This was limited to those genes with nominal *p* values < 0.05. As the change in the direction of effect was the main outcome we intended to investigate, the outcomes from these secondary analyses were not corrected for multiple testing. The Online Mendelian Inheritance in Man (OMIM, https://www.omim.org/) database was used to obtain a list of genes associated with susceptibility to

GGE (IGE, CAE & JME; phenotypic series: PS600669, PS254770 and PS600131) or causing Developmental and Epileptic Encephalopathies (DEE; phenotypic series: PS308350).

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Supplementary Tables:

Table S1: Numbers of analyzed samples from the study cohorts.

Table S2: *GABRG2* **variants identified in individuals diagnosed with genetic generalized epilepsy.**

The variants were analyzed up to an external minor allele frequency (MAF) of 0.1%. All those detected in the cases had an internal leave-one-out MAF and external MAF = 0 (i.e., ultra-rare variants). No rare variants were seen in individuals with epilepsy and all ultra-rare variants were found only once. CAE: Childhood Absence Epilepsy. GEFS: Generalized Epilepsy Febrile Seizures. FS: Febrile Seizures. NAFE: Non-Acquired Focal Epilepsy. TM: Transmembrane segment. PPh2, REVEL and MTR scores considered damaging/deleterious/intolerant are underlined.

Table S3: *GABRG2* **variants in individuals without epilepsy (controls).**

The variants were analyzed up to an external minor allele frequency (MAF) of 0.1%. Ultra-rare variants had an internal leave-one-out MAF and external MAF = 0. Two variants (p.T58A, p.D231N) were seen in two control individuals. The remaining variants were found only once. GGE: Genetic Generalized Epilepsy. FS: Febrile Seizures. TM: Transmembrane segment. PPh2, REVEL and MTR scores considered damaging/deleterious/intolerant are underlined.

Table S4: Top-ranked genes in the secondary analyses of rare functional variation.

Odds Ratio (OR) and *P* values are given from a Cochran-Mantel-Haenszel exact test. The accompanying homogeneity *p* value indicates the lowest *p* value from Breslow-Day & Woolf tests for homogeneity of odds, where *p* values < 0.05 indicate significantly different odds between the two analysis datasets. CI: Confidence Interval. GGE: Genetic Generalized Epilepsy. HGNC: Human Gene Nomenclature Consortium genes names. OMIM: Online Mendelian Inheritance in Man database. URVs: Ultra-rare variants. OMIM phenotypes: *ZIC3*: VACTERL, *COPA*: Autoimmune interstitial lung, joint, and kidney disease.

Table S5: Top-ranked genes in the secondary analyses of predicted Loss of Function (pLoF) variants.

Odds Ratio (OR) and *P* values are given from a Cochran-Mantel-Haenszel exact test. No gene reached the study wide significant *p* value of 2.9 x 10⁻⁷. The accompanying homogeneity *p* value indicates the lowest *p* value from Breslow-Day & Woolf tests for homogeneity of odds, where *p* values < 0.05 indicate significantly different odds between the two analysis datasets. CI: Confidence Interval. GGE: Genetic Generalized Epilepsy. HGNC: Human Gene Nomenclature Consortium genes names. OMIM: Online Mendelian Inheritance in Man database. URVs: Ultra-rare variants.

Table S6: Comparisons of top-ranked genes with three previous large-scale rare variant association studies of genetic generalized epilepsy.

B. **Association of top his from the current analysis in recent studies**

Table S7: Association of genes encoding GABAA receptors.

Results from the primary limited to genes with *p* values < 0.05.

Table S8: Association of OMIM genes implicated in susceptibility to generalized epilepsy.

Results from the primary analyses limited to genes with *P* values < 0.05.

Table S9: Association of OMIM genes implicated autosomal dominant developmental and epileptic encephalopathies.

Results from the primary analyses limited to genes with *P* values < 0.05.

Table S10: PPh2, REVEL and MTR scores of epilepsy related *GABRG2* **missense variants.**

A. Functionally characterized variants

PPh2/REVEL/MTR scores considered damaging/deleterious/intolerant in this study are highlighted. CAE: Childhood Absence Epilepsy. DEE: Developmental and Epileptic Encephalopathy. DS: Dravet Syndrome. FS/+: Febrile Seizures/Plus. GEFS+: Generalized Epilepsy with Febrile Seizures Plus. GGE: Genetic Generalized Epilepsy. GGE-GTCS: Genetic Epilepsy with Generalized Tonic Clonic Seizures Only. (G-)LOF: Predominant loss-of-function with gain-of-function features. LGS: Lennox-Gastaut syndrome. LOF: Loss-of-function. MAE: Myoclonic Atonic/Astatic Epilepsy. NAFE: Non-acquired Focal Epilepsy.

B. Additional reported variants (not functionally characterized).

GEFS+: Generalized Epilepsy with Febrile Seizures Plus. GGE: Genetic Generalized Epilepsy. EMAS: Epilepsy with Myoclonic Atonic Seizures. NAFE: Non-acquired Focal Epilepsy. NDD-E: Neurodevelopmental Disorder with Epilepsy.

C. Additional (likely) pathogenic variants in ClinVar (not functionally characterized).

First Cohort (IGM)

Cases & controls from multiple studies Sequencing on Illumina platforms at **IGM**

Alignment & Calling using DRAGEN/GATK Imported to ATAV Database

Cases $&$ controls from multiple studies sequenced on Illumina platforms at different sites

Aligned sequencing data transferred to ULHPC Raw sequence data aligned at ULHPC Join calling using DRAGEN/GATK

Duplicate cases identified without genotype sharing $\&$ removed from the second dataset Ancestry prediction homogenized between cases using a random forest classifier

Ouality control on GGE samples & appropriate controls $(ATAV)$

Done separately for all, familial and sporadic GGEs vs. controls

Sample OC Samples with excess heterozygosity, ambiguous sequencing sex, low coverage removed One pair from duplicates & related individuals (KING) removed. Ethnicity outliers (EIGENSTRAT) removed.

> Variant OC Variants failing Hard/VQSR filters, with low GO, DP or AD/DP removed

Coverage harmonization Sites with extreme coverage differences across cohorts removed

Quality control on GGE samples & appropriate controls (bcftools, GATK, Plink)

Done collectively for all, familial and sporadic GGEs vs. controls

Sample OC Samples with excess heterozygosity, ambiguous sequencing sex, low coverage removed One pair from duplicates $&$ related individuals (KING) removed. Ethnicity outliers (EIGENSTRAT) removed.

> Variant OC Variants failing Hard/VQSR filters, with low GQ, DP or AD/DP removed.

Coverage harmonization Variants with extreme differences in call rates or $< 95\%$ call rate in cases & controls removed

Use identical annotations, CCDS boundaries, and variant models

Collapsing analysis in ATAV

Collapsing analysis in R

Exchange of summary statistics and qualifying variants counts

Joint analysis (CMH test) in R

Joint analysis (CMH test) in R

Outcomes compared to ensure matching results

Fig. S1: Flow chart summarizing the analysis strategy used in this study. IGM: Institute of Genomic Medicine, New York, USA. LCSB: Luxembourg Centre for Systems Bioscience, Esch-sur-Alzette, Luxembourg. DRAGEN: Dynamic Read Analysis for Genomic platform. GATK: Genome Analysis Toolkit. ATAV: Analysis Tool for Annotated Variants. ULHPC: University of Luxembourg High Performance Computing Cluster. GGE: Genetic Generalized Epilepsy. QC: Quality control. GQ: Genotype Quality. AD: Allele Depth. DP: Depth. VQSR: Variant Quality Score Recalibration. CCDS: Consensus Coding Sequence. CMH: Cochran Mantel Haenszel test. Details on ATAV: <https://github.com/igm-team/atav>. Details on ULHPC: [https://hpc.uni.lu/.](https://hpc.uni.lu/)

Fig. S2: Principal Component Analysis for ancestry matching. The plot shows eigenvectors on the first and second principal components from 1055 individuals with GGE vs. 6814 controls from the first dataset and 829 individuals with GGE vs. 1764 controls from the second dataset.

Fig. S3: Balance of ultra-rare synonymous qualifying variants tallies between cases and controls in the first dataset. There was no significant difference in the distribution of ultra-rare synonymous variants between the cases and controls. A single control sample in the first dataset with QVs tally exceeding 40 was not plotted but was included in significance testing. The QQ plots show the negative log₁₀ of observed *p* values *vs.* the expected *p* values from 1000 permutations (mean, 2.5th and 97.5th centiles). P values were obtained from a two-sided Fisher's exact test of the association of ultra-rare synonymous qualifying variants

Fig. S4: Balance of ultra-rare synonymous qualifying variants tallies between cases and controls in the second dataset. There was no significant difference in the distribution of ultra-rare synonymous variants between the cases and controls. A single control sample in the first dataset with QVs tally exceeding 40 was not plotted but was included in significance testing. The QQ plots show the negative log₁₀ of observed *p* values *vs.* the expected *p* values from 1000 permutations (mean, 2.5th and 97.5th centiles). P values were obtained from a two-sided Fisher's exact test of the association of ultra-rare synonymous qualifying variants

Fig. S5: Balance of ultra-rare synonymous qualifying variants tallies between cases and controls in the total dataset. The QQ plots show the negative log_{10} of observed *p* values *vs.* the expected *p* values from a uniform distribution. P values were obtained from a two-sided Cochran Mantel Haenszel exact test of the association of ultra-rare synonymous qualifying variants. The 95% confidence intervals are shown as grey solid lines. The slope of the solid black line indicates the genomic inflation factor, whereas the slope of the dotted line equals 1.

Fig. S6: Association of ultra-rare deleterious variants with genetic generalized epilepsy. The quantile-quantile plots compare observed *p* values (Cochran-Mantel-Haenszel exact test) and expected *p* values (drawn from a uniform distribution) in analyses of 1,928 individuals with genetic generalized epilepsy (GGEs) vs. 8,578 controls and subsets of familial GGEs (945 cases vs. 8,626 controls) or sporadic GGEs (1,005 cases vs. 8,621 controls). The 95% confidence intervals are shown as grey solid lines. The slope of the solid black line indicates the genomic inflation factor, whereas the slope of the dotted line equals 1. Labels: genes that are enriched in cases in both datasets among the five top-raking genes. Exome-wide significance after Bonferroni correction (dark red line) was 29 defined by a *p* value < 2.9 x 10⁻⁷.

Fig. S7: Association of ultra-rare deleterious and intolerant variants with genetic generalized epilepsy. The quantile-quantile plots compare observed *p* values (Cochran-Mantel-Haenszel exact test) and expected *p* values (drawn from a uniform distribution) in analyses of 1,928 individuals with genetic generalized epilepsy (GGEs) vs. 8,578 controls and subsets of familial GGEs (945 cases vs. 8,626 controls) or sporadic GGEs (1,005 cases vs. 8,621 controls). The 95% confidence intervals are shown as grey solid lines. The slope of the solid black line indicates the genomic inflation factor, whereas the slope of the dotted line equals 1. Labels: genes that are enriched in cases in both datasets among the five top-raking genes. Exome-wide significance after Bonferroni correction (dark red 30
line) was defined by a *p* value < 2.9 x 10⁻⁷.

Fig. S8: Association of rare deleterious variants with genetic generalized epilepsy. The quantile-quantile plots compare observed *p* values (Cochran-Mantel-Haenszel exact test) and expected *p* values (drawn from a uniform distribution) in analyses of 1,928 individuals with genetic generalized epilepsy (GGEs) vs. 8,578 controls and subsets of familial GGEs (945 cases vs. 8,626 controls) or sporadic GGEs (1,005 cases vs. 8,621 controls). The 95% confidence intervals are shown as grey solid lines. The slope of the solid black line indicates the genomic inflation factor, whereas the slope of the dotted line equals 1. Labels: genes that are enriched in cases in both datasets among the five top-raking genes. Exome-wide significance after Bonferroni correction (dark red line) was 31 defined by a *p* value < 2.9 x 10⁻⁷.

Fig. S9A: Association of rare predicted loss of function variants (including URVs) with genetic generalized epilepsy. The quantile-quantile plots compare observed *p* values (Cochran-Mantel-Haenszel exact test) and expected *p* values (drawn from a uniform distribution) in analyses of 1,928 individuals with genetic generalized epilepsy (GGEs) vs. 8,578 controls and subsets of familial GGEs (945 cases vs. 8,626 controls) or sporadic GGEs (1,005 cases vs. 8,621 controls). The 95% confidence intervals are shown as grey solid lines. The slope of the solid black line indicates the genomic inflation factor, whereas the slope of the dotted line equals 1. Labels: genes that are enriched in cases in both datasets among the five top-raking genes. Exome-wide significance after Bonferroni 32 correction (dark red line) was defined by a *p* value < 2.9 x 10⁻⁷.

Fig. S9B: Association of rare predicted loss of function variants (excluding URVs) with genetic generalized epilepsy. The quantile-quantile plots compare observed *p* values (Cochran-Mantel-Haenszel exact test) and expected *p* values (drawn from a uniform distribution) in analyses of 1,928 individuals with genetic generalized epilepsy (GGEs) vs. 8,578 controls and subsets of familial GGEs (945 cases vs. 8,626 controls) or sporadic GGEs (1,005 cases vs. 8,621 controls). The 95% confidence intervals are shown as grey solid lines. The slope of the solid black line indicates the genomic inflation factor, whereas the slope of the dotted line equals 1. Labels: genes that are enriched in cases in both datasets among the five top-raking genes. Exome-wide significance after Bonferroni 33 correction (dark red line) was defined by a *p* value < 2.9 x 10⁻⁷.

Fig. S10: Pedigrees of four families with genetic generalized epilepsy. Pedigree information was available for four carriers of ultra-rare variants in *GABRG2*. R177P, identified in a proband with Early Onset Absence Epilepsy (EOAE), was inherited from a parent with a similar phenotype (EOAE) and segregated in a sibling with Genetic Generalized Epilepsy (GGE) not further classified in a sub-syndrome. On the other hand, Y213* in a proband with GGE was inherited from a parent not diagnosed with epilepsy. Two other variants were identified in probands with Childhood Absence Epilepsy (CAE). These had siblings diagnosed with CAE, but segregation was not feasible at the time of this study. Pedigrees from six individuals carrying six other ultra-rare variants in *GABRG2* reported in this studt were not available.

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Additional Data

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