

Supplementary Material

1 NEUTRAL MODEL

Here, we review the neutral BDI model in which there is no heterogeneity in either proliferation or immigration rates, $\pi(\alpha, r) = \delta(\alpha - \bar{\alpha})\delta(r - \bar{r})$. Upon inserting this expression for $\pi(\alpha, r)$ in Eq. 8, we find that the clone abundance c_k follows a negative binomial distribution (?):

$$c_k = Q \left(1 - \frac{\bar{r}}{\mu^*}\right)^{\bar{\alpha}/\bar{r}} \left(\frac{\bar{r}}{\mu^*}\right)^k \frac{1}{k!} \prod_{\ell=0}^{k-1} \left(\frac{\bar{\alpha}}{\bar{r}} + \ell\right). \quad (\text{S1})$$

We can also express c_k/C , the clone abundance distribution normalized by the mean richness C in the body, as

$$\frac{c_k}{C} = \frac{c_k}{\sum_{\ell \geq 1} c_\ell} \quad (\text{S2})$$

where $C = \sum_{\ell=1} c_\ell = Q(1 - (1 - \bar{r}/\mu^*)^{\bar{\alpha}/\bar{r}})$ is C^S in Eq. 12 with $\eta = 1$. Using $\bar{\alpha} \approx 1.6 \times 10^{-8}/\text{day}$, $\bar{r} \sim 5 \times 10^{-4}/\text{day}$, and $\mu^* \approx 6.4 \times 10^{-4}$, we find $\bar{\alpha}/\bar{r} \ll \bar{r}/\mu^* \sim O(1)$. The $\bar{\alpha}/\bar{r} \ll 1$ regime allows us to approximate c_k/C as a log-series distribution with parameter \bar{r}/μ^* . To mathematically show this, consider a random variable X that follows a negative binomial distribution of parameters $\bar{\alpha}/\bar{r}$ and \bar{r}/μ^*

$$\mathbb{P}[X = k] = \left(1 - \frac{\bar{r}}{\mu^*}\right)^{\bar{\alpha}/\bar{r}} \left(\frac{\bar{r}}{\mu^*}\right)^k \frac{1}{k!} \prod_{\ell=0}^{k-1} \left(\frac{\bar{\alpha}}{\bar{r}} + \ell\right). \quad (\text{S3})$$

Note that the probability mass function of X above is also given by c_k/Q as can be seen from Eq. S1, the clone abundance distribution for all possible Q clones, which includes c_0 , the number of all clones that are not represented in the organism. To find the clone abundance distribution c_k/C , for all the C clones present in the organism, we must exclude the case $k = 0$ by marginalizing the distribution of X over all $X > 0$:

$$\mathbb{P}[X = k | X > 0] = \frac{\mathbb{P}[X = k]}{\sum_{\ell \geq 1} \mathbb{P}[X = \ell]} = \frac{c_k/Q}{\sum_{\ell \geq 1} c_\ell/Q} = \frac{c_k}{C}. \quad (\text{S4})$$

What remains is to show that the distribution in Eq. S4 converges to a log-series distribution of parameter \bar{r}/μ^* when $\bar{\alpha}/\bar{r} \rightarrow 0$. Consider the moment generating function of $X | X > 0$ given by

$$\mathbb{E}\left[e^{\xi X} | X > 0\right] = \frac{\mathbb{E}\left[e^{\xi X}\right] - \mathbb{E}\left[e^{\xi X} | X = 0\right] \mathbb{P}[X = 0]}{\mathbb{P}[X > 0]}. \quad (\text{S5})$$

Since the moment generating function of a negative binomial distribution $\mathbb{E}\left[e^{\xi X}\right]$ is known, and since $\mathbb{P}[X > 0] = 1 - \mathbb{P}[X = 0]$ (see Eq. S3), we can write

$$\mathbb{E} \left[e^{\xi X} | X > 0 \right] = \frac{\left(\frac{1 - \bar{r}/\mu^*}{1 - e^{\xi \bar{r}/\mu^*}} \right)^{\bar{\alpha}/\bar{r}} - \left(1 - \frac{\bar{r}}{\mu^*} \right)^{\bar{\alpha}/\bar{r}}}{1 - \left(1 - \frac{\bar{r}}{\mu^*} \right)^{\bar{\alpha}/\bar{r}}}. \quad (\text{S6})$$

For any $x > 0$, the limit $\bar{\alpha}/\bar{r} \rightarrow 0$ yields $x^{\bar{\alpha}/\bar{r}} = 1 + (\bar{\alpha}/\bar{r}) \log x + o(\bar{\alpha}/\bar{r})$. If we apply this result to Eq. S6 for $\mathbb{E} \left[e^{\xi X} | X > 0 \right]$, we find

$$\begin{aligned} \mathbb{E} \left[e^{\xi X} | X > 0 \right] &= \frac{1 + \frac{\bar{\alpha}}{\bar{r}} \log \left(\frac{\mu^* - \bar{r}}{\mu^* - e^{\xi \bar{r}}} \right) - \left(1 + \frac{\bar{\alpha}}{\bar{r}} \log \left(1 - \frac{\bar{r}}{\mu^*} \right) \right) + o\left(\frac{\bar{\alpha}}{\bar{r}}\right)}{-\frac{\bar{\alpha}}{\bar{r}} \log \left(1 - \frac{\bar{r}}{\mu^*} \right) + o\left(\frac{\bar{\alpha}}{\bar{r}}\right)} \\ &= \frac{\log \left(1 - e^{\xi \frac{\bar{r}}{\mu^*}} \right)}{\log \left(1 - \frac{\bar{r}}{\mu^*} \right)} + o(1), \end{aligned}$$

which we recognize as the moment generating function of a log series distribution of parameter \bar{r}/μ^* . Thus, we finally have

$$\lim_{\bar{\alpha}/\bar{r} \rightarrow 0} \frac{c_k}{C} = \frac{(\bar{r}/\mu^*)^k}{k \log \left(\frac{1}{1 - \bar{r}/\mu^*} \right)}. \quad (\text{S7})$$

2 EXPLICIT FORMS USING DIFFERENT IMMIGRATION AND PROLIFERATION RATE DISTRIBUTIONS

In the following, we propose four simplifying expressions for the heterogeneity-averaged clone counts $c_k^s(\bar{\alpha}, \mu^*, w, \eta)$ derived from Eq. 18.

Clone-independent Neutral model: $\pi(\alpha, r) = \delta(\alpha - \bar{\alpha})\delta(r - \bar{r})$

First, consider the simplest case where all naive T cells carry the same immigration and proliferation rates $\bar{\alpha}$ and \bar{r} , respectively, and define $\pi(\alpha, r) = \delta(\alpha - \bar{\alpha})\delta(r - \bar{r})$. This case corresponds to $w \rightarrow 0$ and $r \rightarrow \bar{r} = 1/2$ in the $\pi_r(r|w)$ box distribution in Eq. 13. The self-consistent condition for μ^* and $\bar{\alpha}/\bar{r}$ become

$$\frac{\bar{r}}{\mu^*} \rightarrow \frac{\lambda}{\lambda + 2\bar{\alpha}}, \quad \frac{\bar{\alpha}}{\bar{r}} \rightarrow 2\bar{\alpha}, \quad (\text{S8})$$

and the clone count given in Eq. 11 can be explicitly simplified to

$$c_k^s(\bar{\alpha}, \lambda, \eta) \equiv \frac{Q}{k!} \left(\frac{\eta\lambda}{\eta\lambda + 2\bar{\alpha}} \right)^k \left(\frac{2\bar{\alpha}}{\eta\lambda + 2\bar{\alpha}} \right)^{2\bar{\alpha}} \prod_{\ell=0}^{k-1} (2\bar{\alpha} + \ell). \quad (\text{S9})$$

The total sampled clone count is then

$$C^S(\bar{\alpha}, \lambda, \eta) = \sum_{k=1}^{\infty} c_k^S(\bar{\alpha}, \lambda, \eta) = Q \left[1 - \left(\frac{2\bar{\alpha}}{\eta\lambda + 2\bar{\alpha}} \right)^{2\bar{\alpha}} \right]. \quad (\text{S10})$$

Fixed immigration rate, distributed proliferation: $\pi(\alpha, r) = \delta(\alpha - \bar{\alpha})\pi_r(r)$

Next, consider a common immigration rate $\bar{\alpha}$ for all T cell clones and a box distribution $\pi_r(r|w)$ of full width $w = 1$. Eq. 14 yields $\mu^* = (1 - e^{-\lambda/\bar{\alpha}})^{-1}$, so that the averaged clone counts from Eq. 11 are now explicitly

$$c_k^S(\bar{\alpha}, \lambda, \eta) \equiv \frac{Q}{k!} \int_0^1 dr \left(\frac{\eta r / \mu^*}{1 - (1 - \eta)r / \mu^*} \right)^k \left(\frac{1 - r / \mu^*}{1 - (1 - \eta)r / \mu^*} \right)^{\frac{\bar{\alpha}}{r}} \prod_{j=0}^{k-1} \left(\frac{\bar{\alpha}}{r} + j \right). \quad (\text{S11})$$

The total sampled clone count can also be explicitly expressed as the integral over $C^S(\bar{\alpha}, r, \lambda|\eta)$ from Eq. 12:

$$C^S(\bar{\alpha}, \lambda, \eta) = Q \int_0^1 dr \left[1 - \left(\frac{1 - r / \mu^*}{1 - (1 - \eta)r / \mu^*} \right)^{\bar{\alpha}/r} \right]. \quad (\text{S12})$$

Clone-specific immigration, fixed proliferation rate: $\pi(\alpha, r) = \pi_\alpha(\alpha|\bar{\alpha})\delta(r - \bar{r})$

Finally, we consider the case whereby all proliferation occurs at a fixed rate \bar{r} and α is distributed according to Eq. 17, as determined from our OLGA sequence-drawing analysis. Using the same rate dimensionalization as before (Eqs. S8), we find explicitly

$$c_k^S(\bar{\alpha}, \lambda, \eta) = \frac{Q}{k!} \left(\frac{\eta\lambda}{\eta\lambda + 2\bar{\alpha}} \right)^k \sum_{j=1}^J \frac{b_j}{C_\star} \left(\frac{2\bar{\alpha}}{\eta\lambda + 2\bar{\alpha}} \right)^{2\alpha_j} \prod_{\ell=0}^{k-1} (2\alpha_j + \ell), \quad (\text{S13})$$

where α_j depends implicitly on $\bar{\alpha}$ through Eq. 16. Similarly, the total sampled clone count can be explicitly expressed as

$$C^S(\bar{\alpha}, \lambda, \eta) = Q \sum_{j=1}^J \frac{b_j}{C_\star} \left[1 - \left(\frac{2\bar{\alpha}}{\eta\lambda + 2\bar{\alpha}} \right)^{2\alpha_j} \right]. \quad (\text{S14})$$

3 SMALL AVERAGE IMMIGRATION RATE

Here, we show that if the support of $\pi_\alpha(\alpha)$ is sufficiently small, the exponential term in Eq. 11 $(\cdot)^{\alpha/r} \sim 1$, and the product term $\sim (\alpha/r)(k-1)!$. While α is summed or integrated over, for reasonable distributions $\pi_\alpha(\alpha)$, the lowest few rates contribute the most and the average of a function over $\pi_\alpha(\alpha)$ can be replaced by its value evaluated at the small average value $\bar{\alpha}$. Even though for r is integrated over $(0, 1)$ for $w = 1$, and the region near 0^+ would lead to a large α/r , the contribution from $c_k^S(\alpha, r, \lambda, \eta)$ is also small near $r = 0$. We have numerically checked that for all cases of $\bar{\alpha} \ll 1/2$, c_k^S can be approximated by

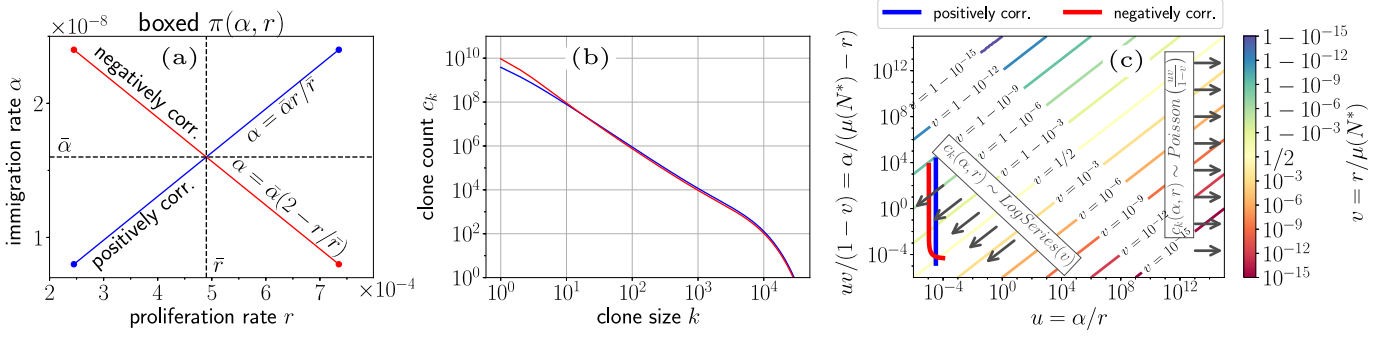


Figure S1. Positively and negatively correlated $\pi(\alpha, r)$. (a) For $\bar{r}/2 \leq r \leq 2\bar{r}$, we consider $\pi(\alpha, r)$ distributions with positively and negatively correlated α and r (Eqs. S18). (b) Mean sampled clone counts corresponding to positively and negatively correlated $\pi(\alpha, r)$ show negligible differences. (c) “Line integrals” of the positively and negatively correlated distributions $\pi(\alpha, r)$ in the $uv/(1-v)$ - u diagram. Clones counts predicted by such $\pi(\alpha, r)$ follow log-series distributions, similar to those of a neutral model.

$$c_k^s(\alpha, r, \mu^*, \eta) \approx \frac{\alpha Q}{rk} \left(\frac{\eta r / \mu^*}{1 - (1 - \eta)r / \mu^*} \right)^k. \quad (\text{S15})$$

Thus, for general w , f_k^s in Eq. 19 can be approximated by

$$f_k^s(\bar{\alpha}, \lambda, w, \eta) \equiv \frac{k c_k^s}{Q \eta \lambda} = \frac{\bar{\alpha}}{\eta \lambda w} \int_{\frac{1}{2} - \frac{w}{2}}^{\frac{1}{2} + \frac{w}{2}} \left(\frac{\eta r / \mu^*}{1 - (1 - \eta)r / \mu^*} \right)^k \frac{dr}{r}, \quad (\text{S16})$$

where $\lambda \equiv N^*/Q$ and μ^* is given by

$$\mu^* = \frac{\left(\frac{1}{2} + \frac{w}{2}\right) e^{\lambda w / \bar{\alpha}} - \left(\frac{1}{2} - \frac{w}{2}\right)}{e^{\lambda w / \bar{\alpha}} - 1}. \quad (\text{S17})$$

Since only $\bar{\alpha}$ appears in Eqs. S16 and S17, the irrelevance of the shape of $\pi_\alpha(\alpha)$ is apparent. We have explicitly shown that for small $\bar{\alpha} \ll 1/2$, the approximations in Eqs. S15 and S16 are quantitatively accurate. These simpler forms expedite our numerical analysis and fitting to data using Eq. 20.

4 CORRELATED IMMIGRATION AND PROLIFERATION RATES

Hitherto, we have considered independent immigration and proliferation, and assumed a factorisable rate distribution $\pi(\alpha, r) = \pi_\alpha(\alpha)\pi_r(r)$. However, immigration and proliferation rates may be correlated for certain clones. For example, a frequent realization of V(D)J recombination may also result in a TCR that is more likely to be activated for proliferation. In this case, α would be positively correlated with r . In Fig. S1 we use dimensional rates and consider the effect of correlated $\pi(\alpha, r)$. For $\bar{r}/2 \leq r \leq 2\bar{r}$, we considered normalized, positively/negatively correlated box distributions as shown in Fig S1(a):

$$\begin{aligned} \text{Positively correlated : } \quad \pi(\alpha, r) &= \frac{1}{\bar{r}} \delta\left(\alpha - \frac{\bar{\alpha}}{\bar{r}} r\right), \\ \text{Negatively correlated : } \quad \pi(\alpha, r) &= \frac{1}{\bar{r}} \delta\left(\alpha - \bar{\alpha}\left(2 - \frac{r}{\bar{r}}\right)\right). \end{aligned} \quad (\text{S18})$$

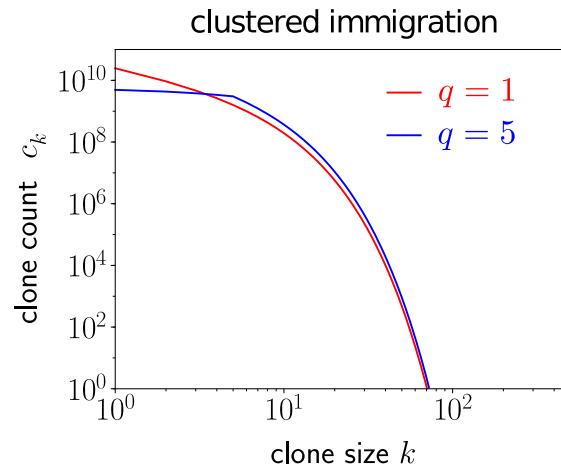


Figure S2. Clustered immigration in a neutral model. Comparison of clone abundances for a $q = 1$ and $q = 5$ models. The difference between the two predicted mean clone counts arise for $k \lesssim q$. Even after sampling, clone counts predicted under clustered immigration ($q > 1$) yields a more slowly decreasing c_k^S for small $k \lesssim q$.

Within our mean field model, these positively and negatively correlated distributions $\pi(\alpha, r)$ result in very similar expected clone abundance distributions c_k (Fig S1(b)). This insensitivity to correlations between immigration and proliferation can be qualitatively understood by considering the “line integral” over dominant paths of $\pi(\alpha, r)$ in the $uv/(1-v) = \alpha/(\mu^* - r)$ vs. $u = \alpha/r$ diagram, as shown in Fig. S1(c). Both line integrals remain in the log-series distribution regime, indicating that the clone abundance distributions are qualitatively similar to those predicted by a model with proliferation heterogeneity alone.

5 MEAN CLONE COUNTS FOR CLUSTERED IMMIGRATION

We explore how clustered emigration from the thymus affects the mean clone count c_k . Suppose that q cells of the same clone (TCR nucleotide or amino acid sequence) are simultaneously exported by the thymus. The equation for the mean clone count c_k becomes

$$\frac{dc_k}{dt} = \sum_q \alpha_q [c_{k-q} - c_k] + r [(k-1)c_{k-1} - kc_k] + \mu(N) [(k+1)c_{k+1} - kc_k]. \quad (\text{S19})$$

This equation does not admit a simple analytic solution so we numerically solved the equation assuming $\alpha_q = \alpha_5 \mathbb{1}(q, 5)$ and $Q = 10^{11}$. Fig. S2 compares the shapes of c_k for single cell immigration ($q = 1$) and simultaneous multicell immigration $q = 5$. In general, for $q > 1$, c_k , and ultimately c_k^S and f_k^S are flatter up to $k \approx q$, making the clone counts more sharply kink downwards near q . Thus, as can be seen from Fig. 9(a,b), we can reasonably conclude that some level of paired immigration would provide even better fits to the data at appropriately small values of λ , especially for the first few k -points.