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How to remove or control confounds in predictive models, with applications to brain biomarkers.

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Abstract:	<p>With increasing data sizes and more easily available computational methods, neurosciences rely more and more on predictive modeling with machine learning, eg to extract biomarkers of pathologies. Yet, a successful prediction may capture a confounding effect correlated with the outcome instead of brain features specific to the outcome of interest –eg the pathology. For instance, as patients tend to move more in the scanner than controls, imaging biomarkers of a pathology may mostly reflect head motion, leading to inefficient use of resources and wrong interpretation of the biomarkers. Here we study how to adapt methods that control for confounds in statistical analyses to predictive modeling settings. We review how to train predictors that are not driven by such spurious effects. We also show how to measure the unbiased predictive accuracy of these biomarkers, based on a confounded dataset. For this purpose, cross-validation must be modified to account for the nuisance effect. To guide understanding and practical recommendations, we apply various strategies to assess predictive models in the presence of confounds on simulated data and population brain imaging settings. Theoretical and empirical studies show that deconfounding should not be applied to the train and test data jointly: modeling the effect of confounds, on the train data only, should instead be decoupled from removing confounds. Cross-validation that isolates nuisance effects gives an additional piece of information: confound-free prediction accuracy.</p>	
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Response to Reviewers:	Dear editor and reviewers, Thank you for reviewing our manuscript, we have revised the manuscript according to	

your comments and suggestions. Please see the point by point responses to the comments as listed below.

Reviewer #1

I made minor notes on improving the writing on the manuscript. In a few places, where I could not figure out what you meant that has been noted.

- Thank you for providing comments in that form, this was really helpful to us. We have tried to clarify the formulation in places that you outlined.

Generally, the writing was imprecise. Your meaning of confounds was not clear to me. You say "MRI biomarkers of brain aging may be nothing more than expensive measurements of head motion." which can mean that using a feature for head motion will be utilizing a confound. Of course, the head motion will affect many measured image features also. I think you would mean both.

- Indeed. Part of the difficulty comes from the fact that an effect is a "confound" depends on the research question. The scenario that we think of is that people may use brain image features for prediction, ignoring that some these features have little or nothing to do with neural activity, but mostly reflect the motion effect, and yet are predictive of an outcome (say, age). People would then use the motion signal carried by brain data to predict age, which is suboptimal (if motion is the feature of interest, it can be captured more efficiently) and misleading (motion-related information being interpreted as neural feature).

It would be helpful to have a nice, clear description of what you intend to address. You say "procedures need to be adapted to predictive-modelling settings" when talking of confound removal in statistics. Yet, statistics underlies (in one way or another) arguably all learned predictive models. The sentence "For this we introduce confound-isolating cross-validation, sampling test sets in which the effect of interest is independent from the confounding effect." was not helpful. Later we can see you try to get a test set without a confound (assuming you realize there may be one).

- We thank the reviewer for raising this point.

Regarding considerations about statistics vs predictive models, we agree that predictive models are statistical models, but we want to emphasize the distinction between classical statistics, that perform in-sample inference, use extensively the maximum likelihood principle and thus rely heavily on distributional hypotheses, vs predictive models that are weakly parametric and are merely built to optimize accuracy scores. This distinction was probably best characterized by L. Breiman (see reference [18] in the new version).

We have added the following statement: "However, these procedures need to be adapted to high-dimensional predictive-modeling settings, where the focus is to achieve high-prediction accuracy based on imaging data.[...] It is not to identify treatment effects size nor to perform other types of causal inference."

On page 2, notation gets confusing. You do not tell us what p is. Usually, it would be features. Then z in R^n is trying to say a dimensionless confound is in all the examples, I think. Needs clarity. Trying to look at confounds in CV is interesting especially if it enables good performance on unseen sources (which in medical data can be very challenging unless you have a diverse training set, which I assume you well know but want to note a perspective).

- Point well-taken. We have added the following sentence: "Such prediction may be misleading or useless [...]. Moreover, this can be detrimental to accuracy: if a future dataset shows an altered relation between the confound and the features, the accuracy may be compromised."

When you talk of re-balancing the data, that is in the case that the confound is highly imbalanced data where predicting the majority class all the time is very accurate, but not useful?

- Not exactly. We consider the case of discrete confounders, where the distribution of the target is non-independent from that of the confounders. Resampling the distributions to reduce imbalance is an effective way of deconfounding, as explained in the Categorical confound section.

When you talk about target population at the bottom of page 2, you mean test

population? Seems obvious that if train/test are exactly the same then you have to remove confounds from both, but your comment is unclear when discussing [20].

- We have changed “target population” for “test population”. Note that in [20] (now [22]), the discussion is not about confounding but about covariate shift between train and test samples, thus assuming a simple validation procedure. We have tried to make it clear that we are not addressing the same question as [20]([22]): “However, [22] have shown that compensating for the confound does not improve prediction if the test population is not markedly different from the training population. Note that train and test samples are often drawn from the same population, either because only one cohort is available or if a proper stratification scheme is used. Our question is different: we are interested in knowing whether learning a biomarker on a confounded dataset leads to predictions that are fully driven by the confound.”

For Algorithm 1 and 2, the features (p) are missing. If not an oversight, needs explanation.

- This has been fixed, thanks

Why 4 subjects to discard vs. 2, 3, etc.?

- This is a choice tailored to the sample size considered here: 4 makes the algorithm faster than using 1, yet is low enough not to compromise mutual information minimization.

What are i and $i-1$ in the equation for z in column 2 on page 4? When you talk about multi-modal MRI data, I think you mean multi-sequence.

- These correspond to the time index of fMRI data, that are time series. We have improved the description here.

Your figures are hard to read. MAE for UKBB is low, but you say it is worse than chance. I think you mean something else. Figures 5 and 7 need better explanation in the text to convince us you have really achieved your goals.

- This is an important point. The “worse than chance” expression comes from a comparison with a permuted distribution. Note that MAE strongly depends on the age span of the observed cohort. This has been clarified in the main text.

Don't know what you mean by: Using deconfounding to condition on a putative confound z help isolating causal links between the data X and the prediction target y , when z is a common cause of X and y .

- Thanks for pointing this. Indeed this formulation was quite impressive and was requiring the reader to adopt a causal perspective, which is not our main aim here. We have rephrased this to better clarify the conceptual shift involved : “Using deconfounding to cancel the impact of a putative confound z removes any bias incurred by the spurious association between the data X and the prediction target y , when z is associated with both X and y .”

Overall, identifying confounds and removing them when doing cross validation can be useful. It is not clear to me what kinds of confounds you can find. Different machines for two classes? Motion (seems so as it is an example)? Sequence differences for classes? Etc. A good, clear takeaway of the impacts and limits would help (yes I know you have done something on this, but it left questions).

- Thanks for raising this, indeed it is worth illustrating the point: confounds can indeed relate to any aspect of the setup (acquisition device, data processing routine when it is not homogeneous across all dataset, measurement-related covariate such as motion, individual conditions, such as age, sex or genetics, that is correlated with the imaging variable and with the outcome. This has been added to the main text.

Reviewer #2

The manuscripts presents a test scheme for validating predictive models named confound-isolating cross-validation, which allows investigators to look at whether a model is fully driven by a confounder or has additional predictive power with respect to the target variable. Experiments include several large-scale datasets including both imaging and non-imaging data. The manuscript is prepared in a good manner. I believe the authors are addressing an important problem of handling confounders in predictive modeling, considering there is a surge of using machine learning to probe

neuroscientific research.

- Thanks for your comments.

Here are some of my suggestions:

The results seem to have valuable information in it, but it becomes very difficult for readers to harmonize the results with the conclusion/discussion. There seems to be a lot of messages being conveyed, but not in a coherent line. This is partly due to the results of different tasks (Figs. 5-7) being so variant. E.g., the text states "Prediction from confounds leads to good prediction of the target in all datasets.", but this is not the case in Fig 5, especially Fig 5c. The discussion states "Out-sample deconfounding give valid information", but how do we know this given that it gives worse-than-chance results in Fig. 5c. Basically the violin plots are so different across methods and tasks, which are not intuitively understandable.

- Indeed the description of the results was sometimes too sketchy. We have striven to make it more accurate. By "valid" we mean that deconfounding is conservative in the worst case, which is certainly suboptimal but not misleading. We have clarified in the main text.

"In the worst case, these approaches can be conservative, but they don't yield spurious associations."

Please note two additional points:

* the quality of methods can only be asserted based on simulated data, where the ground truth is known.

* the variability of the results can be related to the complex relationship between confounder, data and target: the causal link are not necessarily from confounder to data and target. For this reason, as explained in the main text, deconfounding can be beneficial or detrimental, and it is hard to know in advance what will occur.

We have tried to convey these more clearly in the results section.

It is not clear how the proposed approach can make an impact in the clinical setting. The learned model in the context of CICV is still confounded; i.e., it can not be used to perform diagnosis, pinpoint biomarkers (e.g., find which functional connections predict age), nor hypothesis testing. Even if CICV shows that a model is fully confounded (no additional power), it does not mean there is no relationship between X and y just because the model is not learned correctly. The only benefit I see here is that it can validate whether the model has additional power beyond capturing confounding effect and therefore can compare models, but this is only useful in the machine learning context, not in neuroscience applications.

- The reviewer raises a very important point. Indeed, our only claim is that it is possible to learn a confounded model yet evaluate it in an unbiased fashion. What matters in this logic is that the predictive accuracy after CICV remains better than chance, which amounts to performing an omnibus test of the variables of the model. Note that we explicitly recommend to use deconfounding if the goal is to obtain a model free of confounds.

The case where CICV would yield a null or weak result certainly means that one should be cautious in claiming an association between X and y, as slight variation in the confounding model may render the association significant or not: indeed the apparent association between X and y is dominated by z and is thus spurious. This has been added to the main text.

We think that even in a neuroscience context, practitioners should be made aware that the claimed association between covariates and target is dominated by a confounding effect, and in that sense, spurious. In other words this has an impact on the practical significance of claimed associations.

The proposed concept of using confounder-invariant test set is closely related to concepts of demographic parity in fair machine learning and has been explored in Zhao et al. Training confounder-free models for medical applications.

- Thanks for the reference to this nice piece of work, which we have added to the main text.

Some technical concerns:

the discussion on categorical variables seems to be flawed. Avoiding samples from the same site both in the train and the test sets is the OPPOSITE of having independent site and target. Instead, one should have equal (or proportional) number of control and

diseased subjects from each site (i.e. a non-significant chi-square test). In an extremely scenario, if all diseased subjects are from one site, the proposed construction would be fully confounded.

- Thanks for raising this point, which is very important. Actually there are two alternatives strategies here: stratification versus generalization across confounder values. While the first one corresponds to classical deconfounding in statistics, the second one is inspired by the machine learning point of view of generalization across contexts. Generalizing across discrete confounder values is indeed a more stringent test than stratification. And it is useful, because standard stratification may leave some latent association in the data, which is impossible in the strict generalization approach. For this reason, we recommend it. This has been clarified in the main text: "We note that this procedure is different from the stratification strategy used in classical statistics, but it clearly avoids any bias due to imperfectly corrected association between z and the other variables."

A confounder, by definition, impacts both X and y, i.e., a 3-way interaction, so removing confounding effects solely using z and X is theoretically questionable (Eqs 1-3).

- Indeed our operational definition of confounders is narrower than the general definition that would allow more complex interactions. We have clarified the point in the main text: "Note that in all this work we assume that the confounder is associated with X and y without creating three ways interactions between X, y and z."

The generative process in the direct-link scenario in the Simulation studies seems wrong. x and y are still independent, but x_obs and y are dependent. One should instead generate dependent x and y, and generate x_obs directly from x.

- What the reviewer proposes here is indeed a canonical scheme in which the confounder would cause X and y. Since our work is not on causal inference per se, we aim at a statistical procedure that does not require a prescribed causal relationship between the variables (which is often unknown). This point has been made explicit in the main paper. Note that we have updated the notation to ease understanding.

In both old-fashioned statistical methods and modern machine learning, controlling confounders in neuroscience studies still primarily relies on matching confounders, i.e. constructing confounder-free training set instead of confounder-free test set. I wonder why this option is not discussed at all.

- Of course, when it is feasible to construct confounder-free training sets, this approach should be considered, and could actually rely on the procedure we use for the test set creation. However, in our experience this is not feasible, as soon as there are several confounders. We would also like to emphasize this does not circumvent the necessity to validate the model in an unbiased way using procedures like CICV. We have added in the introduction "We consider that practitioners should primarily avoid or reduce the impact of confounders on their model, but this is not always feasible or hard to check, hence, we choose to put the emphasis on the unbiased evaluation of models even in the presence of confounders."

Other confusion: in general I feel the figures contain many elements but the caption is concise, so it is hard to parse. The figures suggest CICV uses a model other than random forest, so what is it? What is the statistical test underlying the p-values?

- We have striven to improve the captions by making them more explicit. CICV uses kernel-density estimation.

Additional Information:	
Question	Response
Are you submitting this manuscript to a special series or article collection?	No
Experimental design and statistics	Yes
Full details of the experimental design and	

<p>statistical methods used should be given in the Methods section, as detailed in our Minimum Standards Reporting Checklist. Information essential to interpreting the data presented should be made available in the figure legends.</p> <p>Have you included all the information requested in your manuscript?</p>	
<p>Resources</p> <p>A description of all resources used, including antibodies, cell lines, animals and software tools, with enough information to allow them to be uniquely identified, should be included in the Methods section. Authors are strongly encouraged to cite Research Resource Identifiers (RRIDs) for antibodies, model organisms and tools, where possible.</p> <p>Have you included the information requested as detailed in our Minimum Standards Reporting Checklist?</p>	<p>Yes</p>
<p>Availability of data and materials</p> <p>All datasets and code on which the conclusions of the paper rely must be either included in your submission or deposited in publicly available repositories (where available and ethically appropriate), referencing such data using a unique identifier in the references and in the “Availability of Data and Materials” section of your manuscript.</p> <p>Have you have met the above requirement as detailed in our Minimum Standards Reporting Checklist?</p>	<p>Yes</p>



PAPER

How to remove or control confounds in predictive models, with applications to brain biomarkers

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Abstract

With increasing data sizes and more easily available computational methods, neurosciences rely more and more on predictive modeling with machine learning, *eg* to extract biomarkers of pathologies. Yet, a successful prediction may capture a confounding effect correlated with the outcome instead of brain features specific to the outcome of interest – *eg* the pathology. For instance, as patients tend to move more in the scanner than controls, imaging biomarkers of a pathology may mostly reflect head motion, leading to inefficient use of resources and wrong interpretation of the biomarkers. Here we study how to adapt methods that control for confounds in statistical analyses to predictive modeling settings. We review how to train predictors that are not driven by such spurious effects. We also show how to measure the unbiased predictive accuracy of these biomarkers, based on a confounded dataset. For this purpose, cross-validation must be modified to account for the nuisance effect. To guide understanding and practical recommendations, we apply various strategies to assess predictive models in the presence of confounds on simulated data and population brain imaging settings. Theoretical and empirical studies show that deconfounding should not be applied to the train and test data jointly: *modeling* the effect of confounds, on the train data only, should instead be decoupled from *removing* confounds. Cross-validation that isolates nuisance effects gives an additional piece of information: confound-free prediction accuracy.

Key words: confound, subsampling, phenotype, predictive models, biomarkers, statistical testing, deconfounding

Introduction

Predictive models, using machine learning, are becoming a standard tool for scientific inference. In cognitive neuroscience, they can be used for *decoding*, to conclude on mental processes given observed brain activity [1, 2, 3]. With the rise of large-scale brain-imaging cohorts, they can extract imaging biomarkers that predict across subjects phenotypes such as neuropsychiatric conditions [4, 5, 6] or individual traits [7, 8].

A crucial aspect of these biomarkers is their ability to *predict* the outcome of interest, *ie* generalize to new data [9]. How-

ever, these predictions can be driven by confounding effects. Such effects affect both the brain-imaging data and the prediction target but are considered as irrelevant. For instance, [10] showed that subjects' in-scanner motion severely affects the link between brain-imaging signals and their age: in-scanner motion varies with subjects' age and it creates systematic differences in brain signals. Given this confounding effect, MRI biomarkers of brain aging may be nothing more than expensive measurements of head motion. Other examples may be more subtle: brain imaging reflects age quite accurately [8, 11, 12], and age matters for diagnosing Alzheimer's disease, yet an im-

portant question is whether brain imaging yields an accurate diagnosis of Alzheimer disease beyond the mere effect of age.

More generally, the data at hand often capture effects not of direct interest to the investigation. In many situations, some confounds such as head motion cannot be fully avoided. To make matters worse, large cohorts developed in population imaging to answer epidemiological questions [as UK biobank, 13] are observational data: there is no controlled intervention or balanced case-control group; rather individuals are recruited from diverse populations with various sampling or selection biases. To conclude on the practical use of biomarkers, it is important to control that their predictions are not fully driven by such unwanted effects. This requires measuring model predictive accuracy after controlling for nuisance variables. Confounding effects can also make it hard to interpret brain-behavior relationships revealed by predictive models [14], as confounds can mediate the observed association or be a latent common cause to observations [15].

In experimental settings, *eg* as in a small cohort, it can be suppressed by balancing the acquisition for confounds, or using randomized control trials. However, constraints in the data acquisition, *eg* recruitment of a large cohort, often imply that confounds are present in the data, and a suitable analysis is needed to avoid reaching erroneous conclusions. The statistical literature on controlling confounding variables is well developed for classic statistical analysis, such as statistical testing in a linear model at the heart of the standard mass-univariate brain mapping [16, 17]. However, these procedures need to be adapted to high-dimensional predictive-modeling settings, where the focus is to achieve high-prediction accuracy based on imaging data. Indeed, predictive models do not rely on the same parametric assumptions, namely linearity of effects and Gaussian noise. Often, a predictive analysis does not build on a generative model of the signal but on optimizing discrimination [18]. In addition, predictive models draw their purpose and validity from out-of-sample prediction, rather than in-sample statistical testing [19]. The question tackled here is thus whether one can assess the predictive accuracy of brain measurements free of unwanted confounds. It is *not* to identify treatment effects size nor to perform other types of causal inference.

In this paper, we study statistical tools to control for confounding effects in predictive models. We consider that practitioners should primarily avoid or reduce the impact of confounds on their model, but this is not always feasible or hard to check, hence, we choose to put the emphasis on the unbiased evaluation of models even in the presence of confounds. A preliminary version of the work discussed here was presented at the PRNI conference [20]. While the core method is the same, it presents limited insights on the theoretical underpinnings and practical value of the method proposed. Experiments on simulated data are absent and experiments on neuroimaging data are limited to just one data set. In particular, statistical significance is not established thoroughly, and only one alternative approach is considered. In short the conference publication provides limited insights on the method, while the current work provides a complete description and points to the code for reuse.

We first review how the classic deconfounding procedures can be used in predictive-modeling settings, *i.e.* together with cross-validation. We then expose a complementary approach that is not based on removing confounding effects, but rather testing whether a given predictive model –*eg* a biomarker– predicts well when these confounds are not present. For this we introduce the *confound-isolating cross-validation* method, that consists in sampling test sets in which the effect of interest is independent from the confounding effect. The benefits of this approach are that it is non-parametric and that it directly

tests the quantity of interest in a predictive analysis. We then run an extensive empirical study on three population-imaging biomarker extraction problems as well as simulations. We draw practical recommendations to test predictive models in the presence of confounding effects.

Methods: controlling for confounds in predictive models

Formalizing the problem of prediction with a confound

Assessing predictive models

Predictive models are assessed by their prediction accuracy [19]. For this, cross-validation is the standard tool, typically k -fold cross-validation [21]. It consists in partitioning (potentially randomly) the original dataset into k equal size subsets or folds (each denoted by a color on Figure 1). One of these k sets is held out for testing, and the remaining $(k - 1)$ folds are used for training the model. This process is repeated k times, where each time a different group of observations compose the test set. Prediction accuracy is measured on the test set, then averaged across folds.

Confounding variables in a prediction task

To formalize prediction in presence of a confound, we consider a dataset of n observations –*eg* subjects or time-points– comprising p –dimensional brain signals $\mathbf{X} \in \mathbb{R}^{n \times p}$, an effect of interest¹ $\mathbf{y} \in \mathbb{R}^n$ –the biomarker target– and a confounding effect $\mathbf{z} \in \mathbb{R}^n$.

An imaging biomarker then predicts \mathbf{y} from \mathbf{X} . If \mathbf{X} and \mathbf{z} on the one hand, \mathbf{y} and \mathbf{z} on the other hand, are not independent, the prediction of the target \mathbf{y} might be affected or most accurately done by the confounding effect, \mathbf{z} . Such prediction may be misleading or useless. It can be misleading as it can be interpreted as a link between brain structures and \mathbf{y} –*eg* fluid intelligence– while such a link only reflects the effect of \mathbf{z} –*eg* age. It can be useless because brain imaging is likely much more costly to acquire than the phenotypic variable \mathbf{z} , hence it should be used only if it brings more diagnostic information. Moreover, this can be detrimental to accuracy: if a future dataset shows an altered relation between the confound and the features, the accuracy may be compromised.

A crucial problem for the validity of the biomarker is to measure whether it can predict \mathbf{y} from \mathbf{X} and not solely from \mathbf{z} . Prediction accuracy is ideally measured on an independent validation set, but most often, no large independent validation set is available and a cross-validation procedure, that iteratively separates train and test sets [21], is used. [22] discusses what cross-validation captures in the presence of a confounding variable. Though there can be many possible confounds in brain imaging (see section 8), we focus below on simple settings, assuming that such nuisance factors have been isolated in one confound variable.

There are two points-of-view to controlling confounds in predictive models. One is to try and *remove* the effect of the confounding variables from the data, by regressing them out (deconfounding) or resampling the data to cancel spurious correlations (re-balancing). The other is to test that the model's prediction captures more than the confound. Removing the confounding signal can test whether predictions are fully driven by the confound \mathbf{z} rather than the brain signal \mathbf{X} . However, it does not provide a good tool to measure the predictive power in

¹ In classification settings, \mathbf{y} does not take continuous values in \mathbb{R}^n , yet we use the most general notation to cover both classification and regression settings.

the presence of confounds: it can give a significance test, but not a measure of effect size.

Another point of view on confounding effects in predictive models consists in trying to learn a predictor from a biased population –with the confounding effect– that differs from the population of interest –without the confounding effect. The problem can then be tackled as a *domain adaptation* problem [23, 24]. However, [24] have shown that compensating for the confound does not improve prediction if the test population is not markedly different from the training population. Note that train and test samples are often drawn from the same population, either because only one cohort is available or because a proper stratification scheme is used. Our question is different: we are interested in assessing whether learning a biomarker on a confounded dataset leads to predictions that are fully driven by the confound.

Deconfounding

Deconfounding in standard analysis

In inferential statistics –as opposed to predictive modeling– proper modeling of confounds is important to control the *interpretation* of model parameters, ensuring that they are not driven by the confounding effects. Classical statistic analysis in brain imaging is based on the general linear model (GLM) [25, 16], in which confounding effects are controlled by additional regressors to capture the corresponding variance. Such an approach shows limitations in predictive-modeling settings. First, it is based on maximum-likelihood estimates of linear models, while in general, predictive models are not explicitly based on a likelihood and are often not linear. Second, it is designed to control *in-sample* properties, while predictive models are designed for *out-of-sample* prediction. The two-step approach based on applying a classical GLM to remove the confounding effect, then a predictive model, may lead to pessimistic results, *eg* below-chance prediction [8, 26].

In the context of the GLM, an alternative implementation relies on removing the effect of variables that are correlated. [25]. Note that in all this work we assume that the confounder is associated with \mathbf{X} and \mathbf{y} without creating three ways interactions between \mathbf{X} , \mathbf{y} and \mathbf{z} . Given a sample $\mathbf{X} \in \mathbb{R}^{n \times p}$ of n observations (subjects) with p brain imaging features (*eg* connectivity matrices), $\mathbf{X}_i = (\mathbf{X}_{i1}, \mathbf{X}_{i2}, \dots, \mathbf{X}_{ip})$ and confounds $\mathbf{z} \in \mathbb{R}^n$, the model is:

$$\mathbf{X} = \mathbf{z}^T \mathbf{w} + \mathbf{e}, \quad (1)$$

where \mathbf{w} is a vector of weights (one per voxel, $\mathbf{w} \in \mathbb{R}^p$). The coefficients $\hat{\mathbf{w}}$ can be estimated from the data by solving the regression model:

$$\hat{\mathbf{w}} = (\mathbf{z}^T \mathbf{z})^{-1} \mathbf{z}^T \mathbf{X} \quad (2)$$

Given these equations, a linear model can be used prior to the predictive model to remove the effect of the confounds \mathbf{z} on the brain signals \mathbf{X} . It must be adapted to out-of-sample testing. One solution is to apply deconfounding jointly on the train and the test set, but it breaks the statistical validity of cross-validation because it couples the train and the test set [21]. Hence it can give biased results.

Out-of-sample deconfounding

To adapt the above *deconfounding* approach to the two phases of training and testing a predictive model, a useful view is to consider the deconfounding model as a predictive encoding model, predicting a fraction of the signal \mathbf{X} from \mathbf{z} . Deconfounding is then performed by removing the part of the signal captured by

\mathbf{z} from \mathbf{X} :

$$\hat{\mathbf{X}}_{\text{clean}} = \mathbf{X} - \mathbf{z} \hat{\mathbf{w}} \quad (3)$$

Where $\hat{\mathbf{w}}$ are the coefficients of the linear deconfounding model (Equation 1), estimated on the train data with Equation 2 and then applied to the test [26]. The full out-of-sample deconfounding procedure is listed in algorithm 1.

A drawback of such deconfounding is that it is strongly parametric, i.e. it relies on the model of confounds used. Equation 2 stands for the classic linear model, assuming linearity between the confounding variable \mathbf{z} and its effect on the brain signal \mathbf{X} . The linear model only takes into account second-order statistics (covariance or correlations) and ignores more complex dependencies.

Model-agnostic out-of-sample deconfounding

A common solution to go beyond linear effects of confounds is to use a polynomial expansion of the confounds \mathbf{z} in the linear deconfounding model. Another option is to use a more powerful predictive model in the confound removal. A predictive model –including a mere linear model– regressing \mathbf{X} on \mathbf{z} can be seen as estimating a function f so that $f(\mathbf{z}) = \mathbb{E}[\mathbf{X}|\mathbf{z}]$. There are many possibilities such as random forests or Gaussian processes. The procedure used for out-of-sample deconfounding can then be adapted as in Algorithm 2. While this approach is very powerful, the danger is to remove also part of the signal of interest. Indeed, using a more powerful predictive model, for instance a higher-order polynomial, leads to explaining in \mathbf{X} more data as a function of \mathbf{z} ; however too powerful models *overfit*, which means that they explain variance in \mathbf{X} by chance. In such a situation, the deconfounding procedure may remove signal of interest, unrelated to the confound.

Comparing predictive power of confounds

A simple evaluation of the impact of \mathbf{z} on the prediction of \mathbf{y} is to use predictive models predicting \mathbf{y} from \mathbf{z} (*prediction from confound*) and compare the predictive accuracy to that obtained with biomarkers based on brain signals. This argument is used in [6] to control for the effect of movement on autism diagnostic.

Creating a test set to isolate the confounding effect

Rather than deconfounding, the investigator may ensure that the predictive model is useful by measuring its accuracy on a dataset where the confounding effect is absent. In a cross-validation setting, such a situation can be created by using as a test set a well-chosen subset of the data that isolates the confounding effect. See Figure 1 for a graphical illustration of the approach. Formally, it requires choosing a subset \mathcal{S} of the data such that $\mathbf{y}_{\mathcal{S}}$ and $\mathbf{z}_{\mathcal{S}}$ are independent. The remainder of the data is used as a training set, to learn to predict \mathbf{y} from \mathbf{X} . If the prediction generalizes to the test set \mathcal{S} , the learned

Algorithm 1: Out-of-sample deconfounding

Input: Brain signal $\mathbf{X} \in \mathbb{R}^{n \times p}$, confound $\mathbf{z} \in \mathbb{R}^n$, {train} and {test} indices

1 $\hat{\mathbf{w}}_{\text{confounds}} \leftarrow (\mathbf{z}_{\text{train}}^T \mathbf{z}_{\text{train}})^{-1} \mathbf{z}_{\text{train}}^T \mathbf{X}_{\text{train}}$
/* Regression of confounds on data */

2 $\hat{\mathbf{X}}_{\text{clean,test}} \leftarrow \mathbf{X}_{\text{test}} - \mathbf{z}_{\text{test}} \hat{\mathbf{w}}_{\text{confounds}}$
/* Remove confounds in the test set */

Output: Brain signal without confounds $\hat{\mathbf{X}}_{\text{clean,test}}$

Algorithm 2: Model-agnostic deconfounding

Input: Brain signal $\mathbf{X} \in \mathbb{R}^{n \times p}$, confound $\mathbf{z} \in \mathbb{R}^n$, {train} and {test} indices, machine-learning algorithm g

- 1 $f \leftarrow g(\mathbf{z}_{\text{train}}, \mathbf{X}_{\text{train}})$
/* Fit confound model capturing $\mathbb{E}[\mathbf{X}|\mathbf{z}]$ */
- 2 $\hat{\mathbf{X}}_{\text{clean,test}} \leftarrow \mathbf{X}_{\text{test}} - f(\mathbf{z}_{\text{test}})$
/* Remove confounds in the test set */

Output: Brain signal without confounds $\hat{\mathbf{X}}_{\text{clean,test}}$

Figure 1. Classic and confound-isolating cross-validation. a) k -fold cross-validation is the common procedure to evaluate predictive models. It consists in splitting the data into k equal groups. $k-1$ folds are used to fit a model and 1 fold is used to validate the model. This process is repeated k times so that each sample is taken once in the test set. b) In *confound-isolating cross-validation* sampling we divide the data in train and test sets, but in a different way. First, using subsampling, we create a test set on which \mathbf{y} and \mathbf{z} are independent. The train test is constructed from the rest of the samples that are not included in the test set. In this way, the method creates a test set that contains unrelated target and confound.

relationship between \mathbf{X} and \mathbf{y} is not entirely mediated by \mathbf{z} . In particular, the prediction accuracy then measures the gain in prediction brought by \mathbf{X} .

Categorical confound

The confounding effect can be “categorical”, for instance the site effect when learning predictive biomarkers on multi-site acquisitions as in [6]. In such settings, to test that the model can indeed predict independently from site effects, a simple solution is to resort to a cross-validation that avoids to have samples from the same site both in the train and the test sets. This may imply resampling the data to cancel out associations between site and target related to data imbalance. Similarly, in multi-subject prediction with repeated measurements from the same subject, subject-wise cross-validation can rule out that prediction is driven by subject identification [27, 22]. More generally, for a categorical confound \mathbf{z} , having distinct values for \mathbf{z} in the train and the test set ensures that the prediction cannot be driven by \mathbf{z} . We note that this procedure is different from the stratification strategy used in classical statistics, but it clearly avoids any bias due to imperfectly corrected association between \mathbf{z} and the other variables.

Continuous confound

When \mathbf{z} is a continuous variable, such as age, it is more challenging to generate test sets on which \mathbf{y}_S and \mathbf{z}_S are independent. We describe here an algorithm to generate such sampling, “confound-isolating cross-validation” subsampling. It is based on iterative sampling to match a desired distribution: the goal is to have a test set with independence between \mathbf{y} and \mathbf{z} , i.e. $p(\mathbf{y}, \mathbf{z}) = p(\mathbf{y})p(\mathbf{z})$, where $p(\mathbf{y}, \mathbf{z})$ is the joint probability function of \mathbf{y} and \mathbf{z} , and $p(\mathbf{y})$ and $p(\mathbf{z})$ are the marginal probability distribution.

A related quantity is mutual information, which characterizes the level of dependency between the two variables: $\eta(\mathbf{y}, \mathbf{z}) = \mathbb{E} \left[\log \left(\frac{p(\mathbf{y}, \mathbf{z})}{p(\mathbf{y})p(\mathbf{z})} \right) \right]$. In practice we estimate the probability density functions with a kernel-density estimator (KDE) using Gaussian kernels. We iteratively create the test S set by removing subjects; at each iteration, we consider the problem as a distribution matching problem, matching $p(\mathbf{y}_S, \mathbf{z}_S)$ and $p(\mathbf{y}_S)p(\mathbf{z}_S)$. For this, we use importance sampling: we draw randomly 4 subjects to discard with a probability $\frac{p(\mathbf{y}_S, \mathbf{z}_S)}{p(\mathbf{y}_S)p(\mathbf{z}_S)}$ using inverse sampling method [28, sec 2.2]. Algorithm 3

gives the details. The choice of 4 samples is tailored to the sample size considered here: it makes the algorithm faster than using sample, yet is low enough not to compromise mutual information minimization. A Python implementation is available on GitHub https://github.com/darya-chyzyk/confound_prediction and on PyPI repository <https://pypi.org/project/confound-prediction/> and can be installed with *pip install confound-prediction*.

Note that if \mathbf{z} and \mathbf{y} are too strongly related, the subsampling procedure above does not have enough degrees of freedom and may always chose the same subset: the test set would be deterministically defined by the sampling procedures, in which case there would effectively be only one fold of cross-validation. In practice, it is important to check that such a situation does not occur when analyzing a given dataset.

Empirical study methodology

We now describe the experimental materials underlying our empirical study of confound-controlling approaches in predictive models.

Two classic confounded predictions in population imaging**Motion confounding brain-age prediction**

As brain aging is a risk factor of many pathologies, the prediction of brain age from MRI is a promising biomarker [12]. In childhood also, markers of functional brain development can help to recognize neurodevelopmental disorders [29, 30]. Many recent studies report age prediction, *eg* from resting-state functional connectivity [7, 29, 31], from structural imaging [32], or combining multiple imaging modalities [8, 11]. However, older people and children move more in the scanner than young adults [see fig. 2, 33, 34, 10, 35]. Thus, age-related changes observed in brain images may be confounded by head motion [36] and image quality [37].

Indeed, in-scanner motion creates complex MRI artifacts that are difficult to remove [36]. In addition, they severely impact measurements of functional connectivity [38].

Here the confounding effect is that of head motion during the few hundreds of scans of individual acquisitions. To build a variable summarizing head motion for each subject, we use the movement time-series computed during preprocessing. As suggested in [38], we create the confound \mathbf{z} from the root mean squared displacements (position differences across consecutive time points) for each subject $z = \sqrt{\frac{1}{I-1} \sum_{i=2}^I \left((t_x^i - t_x^{i-1})^2 + (t_y^i - t_y^{i-1})^2 + (t_z^i - t_z^{i-1})^2 \right)}$, where t_x is left/right, t_y anterior/posterior, and t_z - superior/inferior

Algorithm 3: Confound-isolating cross-validation

Input: Target $\mathbf{y} \in \mathbb{R}^n$, confound $\mathbf{z} \in \mathbb{R}^n$, size $m < n$

- 1 $S \leftarrow \{1 \dots n\}$ /* Initialize */
- 2 **while** $\text{card}(S) > m$ **do**
- 3 $p_y \leftarrow \text{KDE}(\mathbf{y}_S)$ /* Density estimation */
- 4 $p_z \leftarrow \text{KDE}(\mathbf{z}_S)$
- 5 $p_{(y,z)} \leftarrow \text{KDE}((\mathbf{z}_S, \mathbf{y}_S))$
- 6 $\mathbf{m}_i \leftarrow \frac{p_{(y,z)}(\mathbf{z}_i, \mathbf{y}_i)}{p_y(\mathbf{y}_i)p_z(\mathbf{z}_i)}, \forall i \in S$
- 7 $S \leftarrow S - \{j\}$ Draw one index j to remove from S with probability \mathbf{m}_j using inversion sampling.
- 8 **end**

Output: Set of test indices S

translation and $i \in [I]$ is the time index. The prediction target y is the age in years.

Age confounding fluid-intelligence measures

Various studies have predicted individual cognitive abilities from brain functional connectivity [39, 40]. In particular, [40] used machine-learning to predict fluid intelligence from rest fMRI. Fluid Intelligence quantifies the ability to solve novel problems independently from accumulated knowledge, as opposed to crystallized intelligence that involves experience and previous knowledge [41]. It is well known that cognitive abilities change with age [42, 43, 44, 45], in particular that fluid intelligence progressively declines in middle age [46], while crystallized intelligence continues to grow with age. Indeed, in a cohort with a large age span, the data display a strong relation between fluid intelligence and age (Figure 2). When extracting biomarkers of fluid intelligence, the danger is therefore to simply measure age. We study how to control the impact of age when predicting a fluid-intelligence score from rest-fMRI functional connectivity.

Population-imaging rest-fMRI datasets

Datasets

We run experiments on 626 participants from the CamCan data set and 9 302 participants from UKBB. All participants are healthy subjects with no neurological disorders.

- **CamCan** Cambridge Center for Ageing and Neuroscience data [47] studies age-related changes in cognition and brain anatomy and function. Characteristics of interest of this dataset are *i*) a population lifespan of 18–88 years, *ii*) a large pool (626 subjects) of multi-modal MRI data and neurocognitive phenotypes.
- **UKBB** The UK Biobank project [48] is a prospective epidemiological study to understand the development of diseases of UK population over the years. The data used here contain 9 302 subjects from the first release of UK Biobank ongoing cohort study with available fMRI scans and extensive health and lifestyle information [49, 50].

Table 1 presents detailed information about the number of subjects and scores scales for each data set.

We give detailed information on pre-processing steps for each dataset in appendix 8, following COBIDAS recommendations [51].

Prediction from functional connectivity

To build predictive models from resting-state fMRI, we follow the recommendations in [52]. We use the BASC functional atlas [53] with 64 regions, based on which we extract fMRI time series from the CamCAN dataset. Next, we normalize, de-trend and bandpass-filter between 0.01 and 0.1Hz the signal. We represent connectivity matrices with tangent parametrization [54]. Finally, we use a ridge regression with nested cross-validation to learn predictive biomarkers from the functional-connectivity matrices. We use Nilearn [55] for the whole predictive pipeline.

Table 1. Characteristics of the data used. The scores for Fluid Intelligence differ on the two datasets: CamCan uses the Cattell test, and UKBB a specifically-designed touch-screen questionnaire.

Dataset Information	CamCan	UKBB
Number of subjects	626	9 302
Age	18 – 88	40 – 70
Fluid Intelligence scale	Cattell (11 – 44 scores)	UKBB-designed (1 – 13 scores)

Tabular (non-imaging) data

The considerations on confounds in predictive models are not specific to imaging data. We also study a confounded prediction without brain signals: on the UKBB data, we consider predicting an individual's income from socio-demographics and mental-health assessments. We investigate education as a potential confound: it may be reflected both in mental-health and in income. There are 8 556 individuals with no missing values on the outcome and confound. We use random forests for prediction, as it is a popular learner that is well suited to non-Gaussian marginals of these tabular data, and the many categorical variables.

Simulation studies

To better understand the findings on real data, we also present experiments on simulated data. We simulate a data set $\mathbf{X}_0 \sim \mathcal{N}(0, 1)$ with confound $\mathbf{z}_0 \sim \mathcal{N}(0, 1)$ to predict continuous variable $\mathbf{y} \sim \mathcal{N}(0, 1)$. We evaluate two samples sizes: $n = 100$ and $n = 1 000$. We use $p = 100$ features in \mathbf{X}_0 . We study 3 scenarios:

- **No direct link between target and brain** where the brain signal does not provide any direct information to predict \mathbf{y} , but is observed with a confound linked to \mathbf{y} :

$$\begin{aligned} \text{observed confound } \mathbf{z} &= \mathbf{y} + \mathbf{z}_0, \\ \text{observed signal } \mathbf{X} &= \mathbf{X}_0 + \mathbf{z}. \end{aligned}$$
- **Direct link between target and brain** where the brain signal does indeed provide information to predicts \mathbf{y} and has an additional confound linked to \mathbf{y} :

$$\begin{aligned} \text{observed confound } \mathbf{z} &= \mathbf{y} + \mathbf{z}_0, \\ \text{observed signal } \mathbf{X} &= \mathbf{X}_0 + \mathbf{y} + \mathbf{z}. \end{aligned}$$
- **Weak confound & direct link between target and brain**

$$\begin{aligned} \text{observed confound } \mathbf{z} &= 0.5 \mathbf{y} + \mathbf{z}_0, \\ \text{observed signal } \mathbf{X} &= \mathbf{X}_0 + \mathbf{y} + \mathbf{z}_0. \end{aligned}$$

Note that one could consider instead a canonical scheme in which \mathbf{z} would cause \mathbf{x} and \mathbf{y} . Since our work is not on causal inference *per se*, we aim at a statistical procedure that does not require a prescribed causal relationship between the variables, which is often unknown.

Experimental paradigm: cross-validation measures

We use cross-validation to assess prediction accuracy. We consider five predictive frameworks: (1) Without deconfounding, (2) Deconfounding test and train jointly, (3) Out-of-sampling deconfounding, (4) Confound-isolating cross-validation, (5) Prediction from confounds only. The code for these various strategy to control for confounds can be found on GitHub https://github.com/darya-chyzyk/confound_prediction and on PyPI repository <https://pypi.org/project/confound-prediction/> and can be installed with `pip install confound-prediction`. We use 10 folds, with random splits of 20% of the data in the test set. For *confound-isolating cross-validation*, different seeds in the random number generator lead to different folds. We assess the null distribution of predictions with permutations (20 000 folds on permuted labels \mathbf{y}).

Figure 2. Joint distribution of target and confound. The first column presents the scatter plot of age and motion variable for CamCan (top) and UKBB (bottom). The second column shows the case of fluid intelligence prediction with age as confound for CamCan. In all cases, the target is clearly associated with the confound; the corresponding p -values are below 10^{-5} .

Figure 3. Evolution of the test set created by Confound-isolating cross-validation. The joint distribution of the target (Fluid intelligence) and the confound (Age) for the CamCan dataset is taken for demonstration. We show the process of selecting proper samples for the test set. We begin with the entire dataset, Pearson correlation is -0.67 with infinitesimal p -value (right subplot). After half of the iterations we have already reached a correlation -0.17 with p -value = 0.009 (middle subplot). The final test set is shown on the right subplot, correlation -0.007 with p -value = 0.02 . It presents negligible residual dependency between targets and confounds.

Figure 4. Evolution of the link between confound and target with the number of subjects for different subsampling methods on the CamCan dataset with Age prediction case. Applying Algorithm 3 effectively reduces statistical dependencies between confound and target (red curve). In our experiments, we stop the sampling when the test set size is $1/5$ of the dataset.

Results of the empirical study

Experiments on resting-state fMRI data

Potential confounds

Figure 2 shows the relationships between target variable y and confounds z . Fluid Intelligence (target) is strongly negatively correlated with age (confound) on the CamCan dataset (second column of Figure 2). Also, on the CamCan data, Age and Motion are very correlated (first column of Figure 2). On the more homogeneous and larger UKBB sample (9 302 subjects), this link is weaker.

Confound-isolating cross-validation

Figure 3 displays the evolution of the association between confound and target during *Confound-isolating cross-validation* in the CamCan dataset, predicting Fluid Intelligence with Age as a confound. In the full dataset, comprising 608 subjects, the correlation between confound and target is $\rho = -0.67$. Iterating the algorithm to remove half of the subjects leads to $\rho = -0.17$. The final test set contains $1/5$ of the subjects and achieves $\rho = -0.07$, showing that it indeed cancels the dependency between aging and motion. The joint distribution between target and confound displayed in Figure 3 shows that the initial statistical dependency between these two variables vanishes after a few tens of iterations of the algorithm. Quantitative evaluation, measuring both Pearson correlation and mutual information (Figure 4) confirms that the confound-isolating procedure efficiently creates a subset of the data without the dependency as soon as it reduces the data to 300 subjects or less. Figure 8 shows similar success on the other prediction problems that we study.

In a cross-validation setting, the different test sets should probe different subjects to maximize testing power. A risk, when using *confound-isolating cross-validation*, is that it could repeatedly generate test sets with the same samples. To measure the diversity of the test sets, we compute the average fraction of common samples between two tests sets created with different seeds. The value is in the range from 0 to 1, where 1 means that all test sets contain the same samples and 0 that

test sets have no sample in common; the expected value is $\frac{1}{5}$. We find an average intersection of 0.30 for age prediction with CamCan and 0.27 with UKBB; for Fluid Intelligence prediction with CamCan, we find 0.36. This demonstrates that the test sets do not repeat much, hence that there is no hidden determinism in the cross-validation scheme of the proposed method.

Testing for confounded prediction

Figure 5 reports the mean absolute error² for the different approaches to control for confounds. The figure also reports the p -value of predictive accuracy, from permutations³. The first thing to note is that without controlling for confounding effects, all models lead to significant prediction. But are these driven by the confounds? Given that the various approaches measure predictions on different data, we compare how far these predictions are above chance, rather than their absolute value.

Deconfounding test and train sets jointly – removing the linear effect of the confounding variable on the full data– has little impact on the prediction performance on all datasets. On the other hand, *out-of-sample deconfounding* changes significantly prediction performance in a way that varies across tasks. Prediction accuracy of fluid intelligence on CamCan falls to chance level. Age prediction on CamCan is little impacted. However, Age prediction on UKBB gives results worse than chance, i.e. worse than a model that learns to predict age on data where this relationship has been shuffled by permutation (see Fig. 5). *Confound-isolating cross-validation* also gives varying results on different datasets. For fluid-intelligence prediction on CamCan, it also gives results at chance level. For age prediction on CamCan, it does alter significantly prediction accuracy, and on UKBB, it leads to a slightly worse prediction, but still above chance. Finally, *Prediction from confounds* leads to chance-level or good prediction of the target depending on the dataset. In particular, it does better than chance for Fluid intelligence prediction.

These results show that in all these datasets, the confounds z are associated with both the data X and the target y . For fluid intelligence prediction on CamCan, all the prediction of y from X is mediated by z . However, for age prediction in CamCan, there exists within X some signal that is unrelated to z but predicts y . Age prediction in UKBB is a more subtle situation: X contains signals from z and y with shared variance, but there is enough signal beyond the effect of z to achieve a good prediction, as demonstrated by *confound-isolating cross-validation*, where the prediction cannot be driven by z . Yet, *out-of-sample deconfounding* removes the shared variance and hence creates predictions that are worse than chance.

2 Mean absolute error is a good metric to compare across different test sets as it gives an absolute error measure in the unit of y , unlike explained variance, that depends on the variance of y .

3 Technically, there is one p -value per fold; to report only one number, we use p -value aggregation [56].

Figure 5. Comparisons on population-imaging data Each sub-figure shows one prediction setting: (a) CamCan Age prediction, (b) CamCan Fluid Intelligence prediction, (c) UKBB Age prediction. The left column of each sub-figure reports the prediction performance by the mean absolute error for the five approaches considered: Prediction from the data without deconfounding, prediction after deconfounding test and train jointly, prediction with out-of-sample deconfounding, prediction with confound-isolating cross-validation, and prediction from confounds. The left column displays the distribution across validation folds for the actual data (top, orange), and for permuted data distribution (bottom, gray). The right column displays the distribution of p -values across folds, obtained by permutation, and the text yields the aggregated p -value across folds (see the main text), testing whether prediction accuracy is better than chance. The figure clearly displays different behaviors across the three problems: without deconfounding, and Deconfounding test and train jointly yield significant, but probably spurious accuracy; Out-of-sample deconfounding can be overconservative (the prediction is worse than chance on UKBB) suggesting that the deconfounding model removes too much variance; confound-isolating cross validation yields more nuanced results, and prediction from confounds yields variable results.

Tabular data

Figure 6 gives the results of analysis on the tabular data. There is a significant prediction of income from socio-demographic and mental-health information, without any deconfounding. However, prediction from confounds shows that qualifications also predict income well. To control for qualification, deconfounding removes the signal explained by these in X . However, it does not make the prediction worse but rather improves it in the case of out-of-sample deconfounding. Such an improvement can be explained if the deconfounding adds information about the confound to the signal rather than removing it, as can happen when the model of the confounds is mis-specified. To limit this problem, we used as a confounds model a random forest, with algorithm 2 for deconfounding. Finally, confound-isolating cross-validation shows very variable results, but overall that prediction does not work better than chance on balanced datasets, so that qualification is no longer specifically related to income.

Here, deconfounding leads to the conclusion that the prediction of income from social-demographic and mental-health information is not at all driven by qualifications while the other approaches suggest otherwise. The discrepancy is probably due to the complex non-linear interactions between these variables. The reality is probably that qualifications contribute to the prediction of income, as well as mental health and socio-demographics information, and that teasing out these contributions is hard.

Simulated data

We now turn to simulated data, for which there is a ground truth. Figure 7 shows the results of the different methods to control for confounds on 3 different simulated cases (Figure 9 gives results for the same simulations with 1000 samples).

- (a) In the case where there is no direct relationship between the data and the target, the performance of the prediction model should not be better than chance after controlling for the confound. Both joint deconfounding and *confound-isolating cross-validation* clearly reveal that all the prediction is mediated by the confound. Out-of-sample deconfounding displays a less clear signal, as there seems to be a slight prediction even after deconfounding, though it is not significant.
- (b) For a direct link between the data and the target, joint deconfounding yields a false negative, in the sense that it fully removes the prediction from the brain signal: it is too aggressive in removing signal. Other approaches correctly support a successful prediction.
- (c) For a weaker confounding signal, results are similar, however it is worth noting that the target can no longer be well predicted from the confound.

Overall, on the simulations, both *out-of-sample deconfounding* and *confound-isolating cross-validation* give reliable answers,

Figure 6. Comparisons on tabular data: predicting income from socio-demographics and mental-health, controlling for qualifications. The left column of the figure reports the prediction performance by the mean absolute error for the five approaches considered: Prediction from the data without deconfounding, prediction after deconfounding test and train jointly, prediction with out-of-sample deconfounding, prediction with confound-isolating cross-validation, and prediction from confounds. The left column displays the distribution across validation folds for the actual data (top, orange), and for permuted data distribution (bottom, gray). The right column displays the distribution of p-values across folds, obtained by permutation, and the text yields the aggregated p-value across folds (see the main text), testing whether prediction accuracy is better than chance.

while deconfounding the test and train jointly as well as measuring the prediction from confounds cannot be trusted.

Discussion and conclusion

Measuring the accuracy of predictive models, *eg* for biomarkers or brain decoding, must account for the presence of confounding effects that can contribute to the prediction. Indeed, an imaging biomarker that solely picks up head motion may detect pathologies with some success, but be overall a waste of scanner time. An accurate prediction of fluid intelligence from brain functional connectivity might simply be a consequence of indirectly capturing the subjects' age. Standard cross-validation procedures ignoring the confounds can overestimate prediction accuracy.

Addressing confounds in predictive modeling

Approaches must be adapted to out-of-sample settings

Deconfounding approaches used in standard GLM-based analysis must be adapted to out-of-sample data by separating estimation of the confounds' model from removal of the effect of confounds on the data, as detailed in section and algorithm 1. Importantly, applying deconfounding to the whole data without separating train and test set is not only wrong in theory—because it breaks independence of train and test data—but also leads to incorrect conclusions in practice, as clearly visible from the simulations.

Even done right, deconfounding in predictive settings can lead to pessimistic evaluations, as stressed by [26] and shown in our experiments. This is because the signal explained by the confound is removed from the brain signal before it is passed to the predictive model. The corresponding correction can remove too much information when there is plenty of shared signal between the confound and the target—*eg* aging and Alzheimer's disease. Such problem does not arise in a GLM-based standard analysis because the confounds and the effects of interest are modeled *simultaneously*, and the consequences of shared signal are easier to handle.

To give a measure of predictive accuracy that is not pessimistic, we also study a different approach: testing the predictive model on a subset of the data crafted such that the effect of interest is independent from the confound. When the confounding effect is represented as a categorical variable, for instance the effect of acquisition site, the approach can be simple as it amounts to splitting the data so as to ensure that generalization occurs for a category not observed in the training set. Creating an adequate test set for continuous confounds requires a dedicated method, as with *confound-isolating cross-validation* (Algorithm 3). It enables a test of the predictive power from brain imaging without discarding the potentially useful shared signal. In addition, it is non-parametric and does not rely on a linear confounding model. Empirical studies, on both brain-imaging data and simulations, show that both out-of-sample deconfounding and *confound-isolating cross-validation* can control correctly for confounds. Deconfounding before fitting a predictive model brings the benefit of building a predictor free of the confounding effect. However, it can remove shared variance and lead to pessimistic evaluations. *Confound-isolating cross-validation* brings the benefit of measuring the predictive power in the absence of the confounding effect. Such measure is of direct importance to gauge the practical value of a biomarker. As an attractive complementary approach, note that deep learning approaches for learning confound-free models have been proposed in [57].

To summarize, our main claim is that it is possible to learn a

Figure 7. Comparisons on simulated data. The figure is organized as Figure 5. There are three simulation settings: (a) No direct link between target and brain, (b) A direct link between target and brain in the presence of a confound and (c) A weak confound and a direct link between target and brain. Green ticks indicate correct conclusions, red crosses mark incorrect ones, and warning signs the weak results.

confounded model yet evaluate it in an unbiased fashion. What matters in this logic is that the predictive accuracy after CICV remains better than chance, which amounts to performing an omnibus test of the variables of the model. The case where CICV would yield a null result certainly means that one should be cautious in claiming a conditional association between X and y , as slight variations in the confounding model may render the association significant or not: indeed the apparent association between features and target is dominated by the confounder and thus, not a reliable one.

We think that even in a neuroscience context, practitioners should be made aware that the claimed association between covariates and target is dominated by a confounding effect, and in that sense, spurious. In other words, this has an impact on the practical significance of claimed associations.

Which approach to use when: deconfounding versus confound-isolating cross-validation

Out-sample deconfounding and *confound-isolating cross-validation* give valid and complementary information. In the worst case, these approaches can be conservative, but they don't yield spurious associations. From a prediction perspective, when the training population reflects adequately the target population, changing the training data to remove the effect of the confounder may not improve prediction accuracy [24]. For instance, for many pathologies, patients move more in the scanner than healthy individuals. Should an imaging-biomarker of the pathology be developed, this effect will be most likely true in the population on which the biomarker is applied. Hence it is counter-productive to force the biomarker to discard this information. Rather, *confound-isolating cross-validation* should be used to check that the imaging biomarker does bring in value in addition to capturing motion.

On the other hand, *confound-isolating cross-validation* is not a universal remedy: removing a confounding effect from the training data may be useful if the confounder is incidentally correlated with X or y without any clear causal relationship. This is the case if the confounder is a feature of the measurement process. For instance, if the data are acquired across two imaging sites with different scanners, but one site recruited a much larger fraction of patients than the other, the risk is that the predictor learns to use information about the scanner rather than the pathology. In such a case, the training strategy must be adapted, for instance by removing the effect of the confound.

Finally, if the goal is to *interpret* successful prediction as evidence of a link between brain signals and the predicted outcome, modifying the training data is more likely to disentangle the biomarker pattern of interest from the confounding effect. In such a situation, deconfounding should be preferred, to give a model, with its parameters, that is not driven by the confounding signal.

Limitation: with many confounds the problem is harder

Here we have studied the case of one, clearly-identified, confound. The case of multiple confounds (*eg* age, education, gender, ethnicity), is more challenging. In such situations, deconfounding approaches may remove fully the signal of interest. For *confound-isolating cross-validation*, reliable estimation of mutual information will require larger sample sizes than with a single confound.

Elements to interpret analyses with confounds

Defining confounds calls for modeling choices

Whether a variable should be considered as a confounding effect or not is not dictated by the data, but by the question at hand. The actual notion of confound comes from causal modeling, to give a causal interpretation to model parameters [15, 58]. Confound variables are then chosen so as to model the difference between the measurements at hand and those obtained with a hypothetical intervention. Such choices are implicitly based on a model of which variables are causes or consequences of the fictional intervention and the outcome of interest [see 59, for guidelines in the case of UKBB].

In pure biomarker settings, the focus is not on potential interventions, but on detecting or predicting an outcome. The concern is then that the measured accuracy might not reflect the actual application settings [27, 22]. Here also, the choice of variables to control for must be governed by an understanding of how the data at hand may differ from ideal data to reflect the target application. More concretely, Confounds can indeed relate to any aspect of the setup, *e.g.* acquisition devices, data processing routines when these are not homogeneous across all the dataset, measurement-related covariate such as motion, individual conditions, such as age, sex or genetics, that is correlated with the imaging variable and with the outcome.

Deconfounding for causal interpretations: the collider-bias danger

Using deconfounding to cancel the impact of a putative confound z removes any bias incurred by the spurious association between the data X and the prediction target y , when z is associated with both X and y . However, z may be a *consequence* of both the target and the data. In such a situation conditioning on it can create a form of selection bias, sometimes known as "collider bias" [60, 61]. Conditioning on the third variable z can then reverse the correlation between two variables X and y , a phenomenon known as Berkson's or Simpson's statistical paradox [62, 63]. It can be understood from a simple example: when studying a population of hospital patients, individuals may have been admitted to the hospital because they have disease A or B . On this specific population, the two diseases are anti-correlated. However, concluding that disease A protects from disease B would be incorrect. Another example can be found in a cognitive experiment where both a visible-enough stimuli and a timely motor response are needed for a successful response. When learning a model decoding stimuli visibility from brain response, deconfounding on successful responses would lead this model to rely on motor-cortex activity, while the link between visual stimuli and motor cortex is not neuroscientifically relevant as such. Deconfounding by itself does not suffice to yield associations with clear interpretations.

A sampling view on confounds

Confound-isolating cross-validation strives to sample an ideal sub-population. This is also one of the best strategies to avoid the presence of confounds in experimental settings: targeting the recruitment of participants so that the design is balanced, for instance with matched controls or randomized controlled trials. But this can only be done at study design, and targeted acquisitions, with matching and restriction, can make it hard to collect large samples or tackle many covariates. At analysis time, researchers have to rely on statistical methods to adapt the analysis to the presence of confounds. For in-sample analy-

sis, propensity scores are a classic reweighting technique used to obtain reliable effect estimates from confounded datasets [64, 65]. The use of subsampling in *confound-isolating cross-validation* can be seen as an extension of these ideas for out-sample validation of predictive accuracy.

Conclusion: deconfounding and isolating confounds are complementary

Deconfounding strives to remove confounding effects from the data, after which successful prediction can be interpreted as a direct link from the remaining brain signals to the outcome of interest. However, in biomarkers settings, the primary focus may be on the quality of detection, rather than interpretation, for instance to improve diagnosis or prognosis. In such settings, an important question is: how much do the brain signals improve the prediction upon a simpler measure of the confounding effect? Answering this question calls for a cross-validation procedure isolating this confounding effect. The corresponding prediction accuracy can then safely be interpreted as not resulting in any way from the confounding effect.

Acknowledgment

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Appendices

Data preprocessing

CamCan data were preprocessed using [Pyprocess](#), a collection of Python scripts for preprocessing fMRI data, that is based on the SPM12 software and the nipy toolbox [66]. We preprocessed CamCan data only. For UKBB data the preprocessed and connectivity matrices are available from the data repository. We apply a commonly used protocol that includes the following steps: Motion correction, correction for subject's head motion during the acquisition. Estimated six motion parameters (three translational parameters and three rotational parameters) are used as confounds in the age prediction experiments. For each subject we expressed the head motion using translation across all three axes as a square root of the mean of the sum of square finite difference of each translation axes over the time: $\sqrt{\frac{\overline{\Delta translation_x^2} + \overline{\Delta translation_y^2} + \overline{\Delta translation_z^2}}{3}}$ The resting fMRI data are coregistered to the anatomical T1-MRI and then normalized to MNI template.

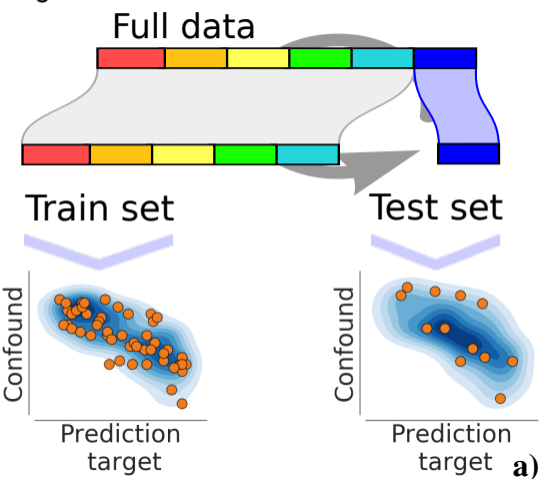
Supplementary results on the resting state data sets

Figure 8. Evolution of mutual information and correlation with number of subjects for different subsampling methods on the CamCan dataset with Fluid Intelligence prediction and UKBB Age prediction. This figure shows that the proposed method effectively reduces statistical dependencies between confound and target (red curve) for both data sets and both predictors.

Supplementary results on simulated data, 1000 samples

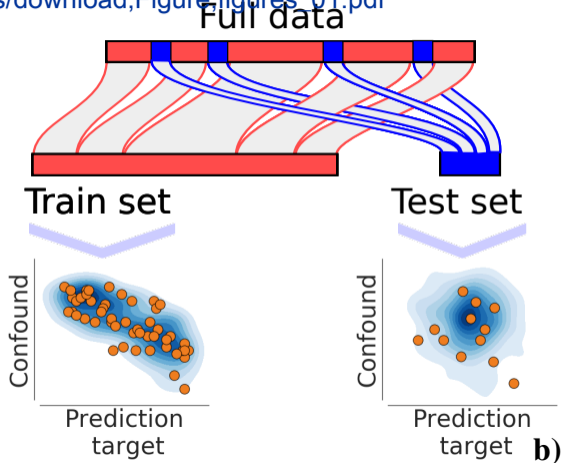
Figure 9. Benchmarking approaches to control confounded predictions on simulated data with many samples. The left column of each sub-figure assesses the prediction performance through the mean absolute error (in signal units). We display the error distribution across validation folds for the data (top, orange), and for permuted data distribution (bottom, gray). The right column displays the distribution of p-values across folds, obtained by permutation, and the text reports the aggregated p-value across folds (see the main text). Five approaches are benchmarked: Without deconfounding, Deconfounding test and train jointly, Out-of-sampling deconfounding, Confound-isolating cross-validation, and Prediction from confounds. There are three simulation settings: (a) No direct link between target and brain, (b) A direct link between target and brain and (c) A weak confound and a direct link between target and brain. Green ticks indicate correct conclusions, red crosses mark incorrect ones, and warning signs the weak results.

K-fold cross-validation strategy



Confound-isolating cross-validation

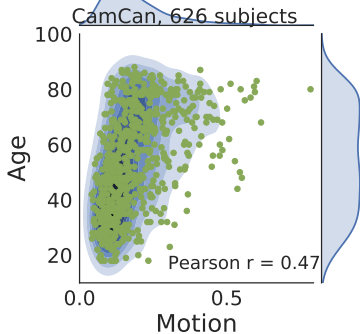
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Figure

Age prediction

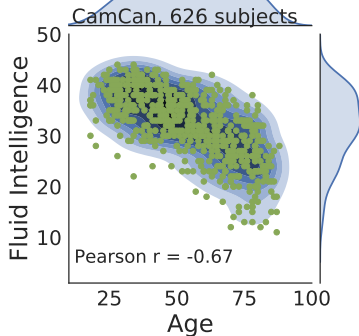
CamCan



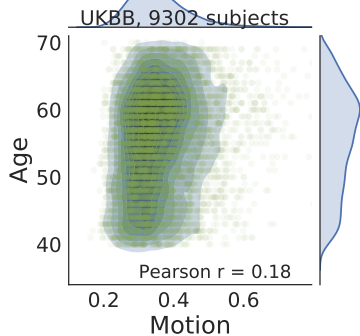
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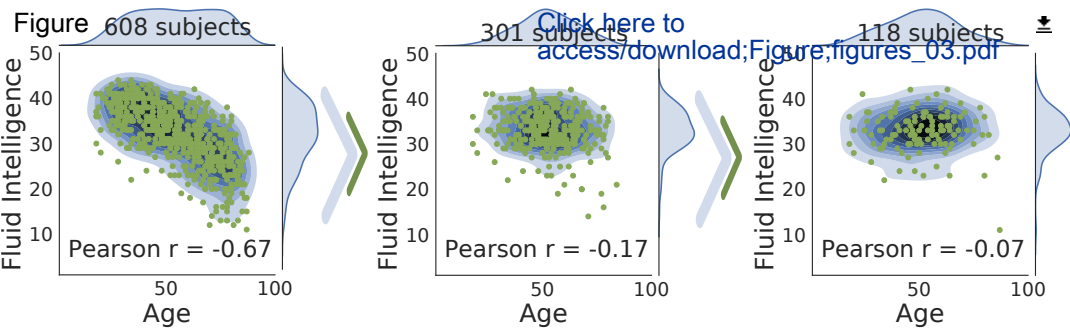
Fluid Intelligence prediction

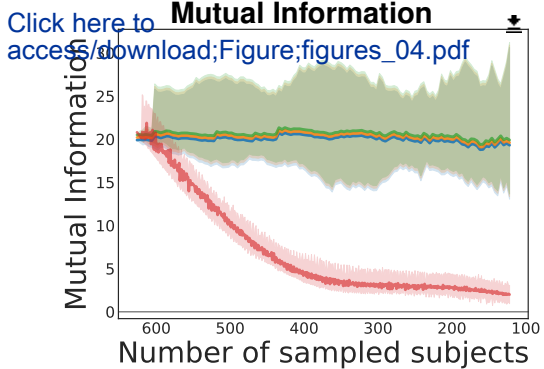
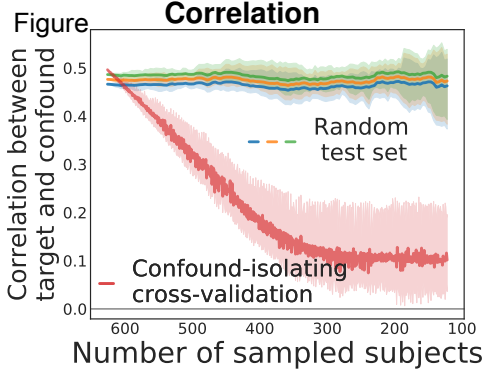
CamCan

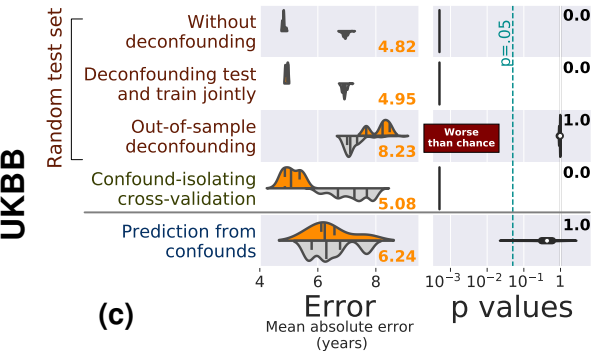
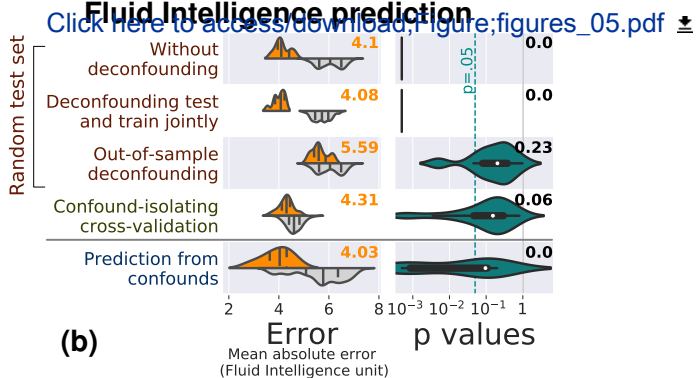
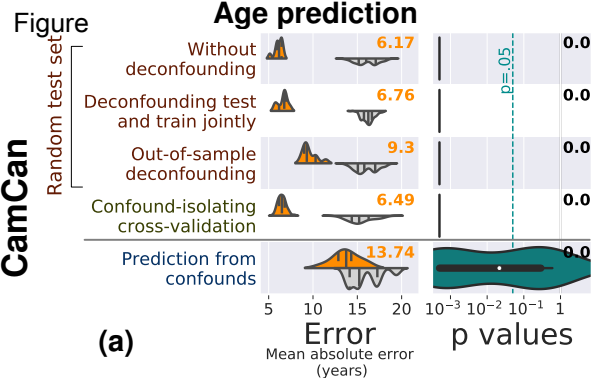


UKBB









Not permuted

Permuted

Figure
Random test set

Without deconfounding



Deconfounding test and train jointly



Out-of-sample deconfounding



Confound-isolating cross-validation



Prediction from confounds

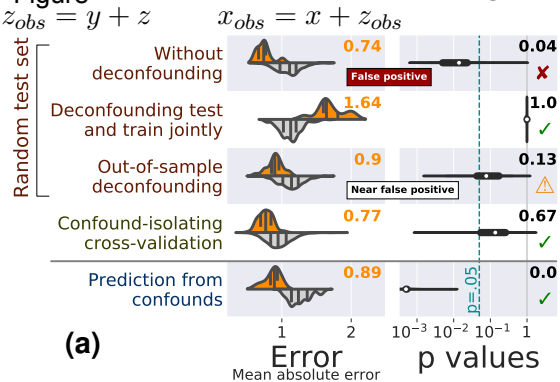


0.7 1.0 10^{-3} 10^{-2} 10^{-1} 1

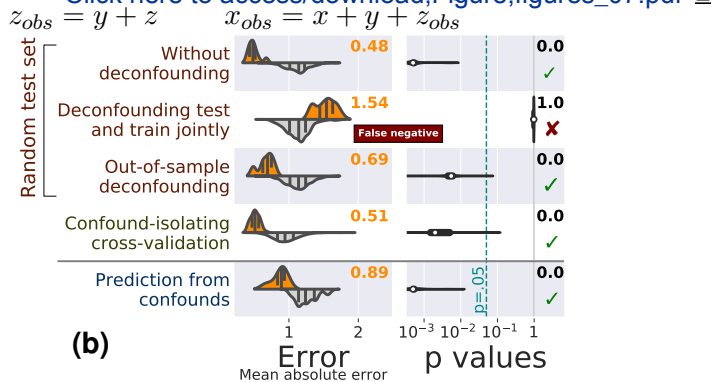
Error p values

Mean absolute error
(unit: # of income bins)

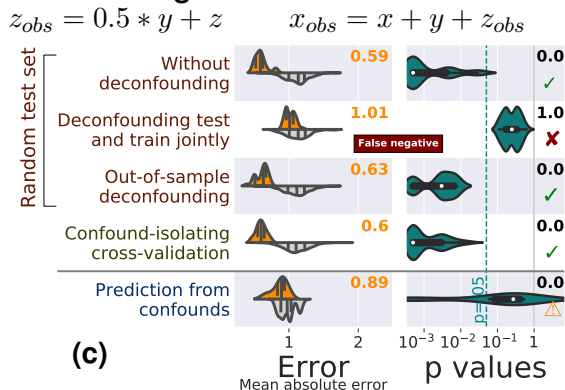
No direct link between data and target



Direct link between data and target



Weak confound & direct link between data and target



Notations

Data: $x \sim \mathcal{N}(0, 1)$

Target: $y \sim \mathcal{N}(0, 1)$

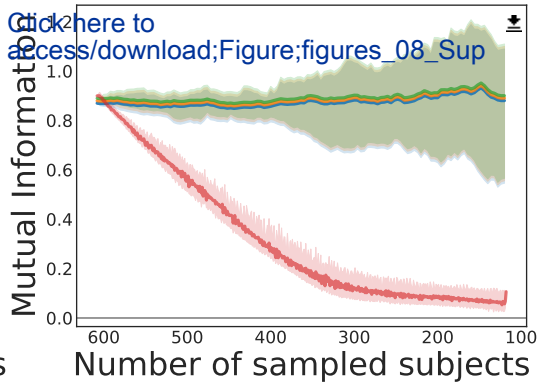
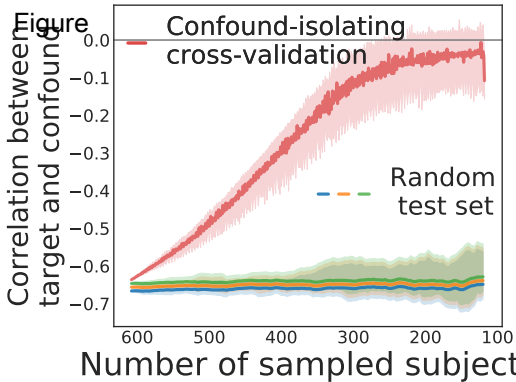
Confound: $z \sim \mathcal{N}(0, 1)$

Observed data: x_{obs}

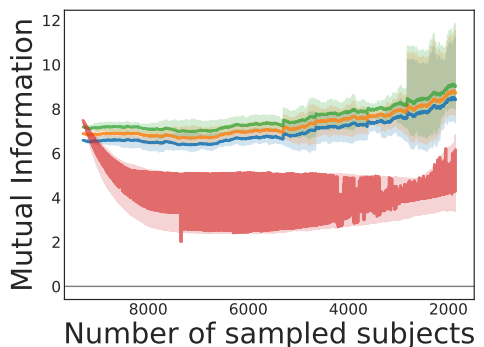
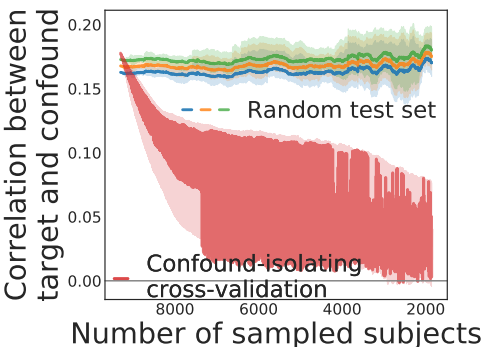
Observed confound: z_{obs}

Not permutated

Permutated



CamCan, Fluid Intelligence prediction

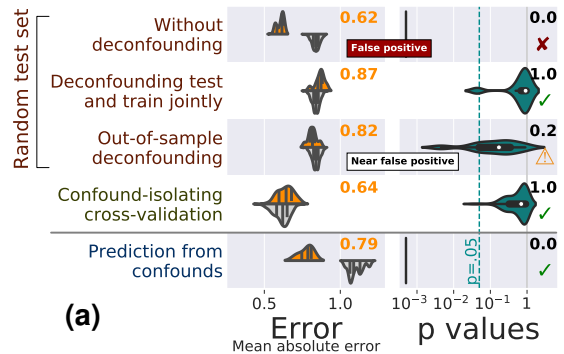


UKBB, Age prediction

No direct link between data and target

$$z_{obs} = y + z$$

$$x_{obs} = x + z_{obs}$$



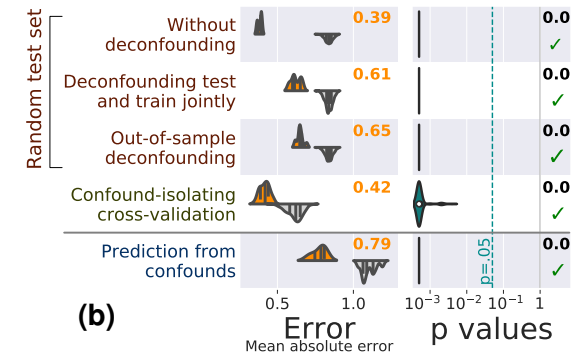
(a)

Direct link between data and target

Click here to access/download;Figure;figures_09_Supplementary

$$z_{obs} = y + z$$

$$x_{obs} = x + y + z_{obs}$$



(b)

Weak confound & direct link between data and target

$$z_{obs} = 0.5 * y + z$$

$$x_{obs} = x + y + z_{obs}$$

Notations

Data: $x \sim \mathcal{N}(0, 1)$

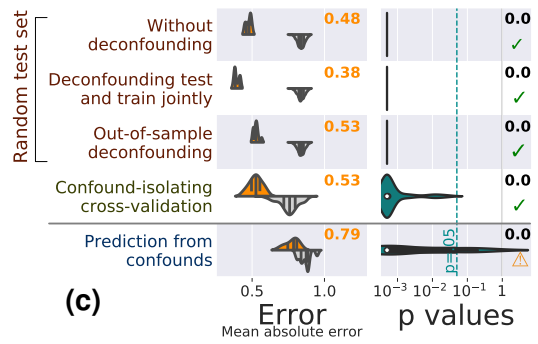
Target: $y \sim \mathcal{N}(0, 1)$

Confound: $z \sim \mathcal{N}(0, 1)$

Observed data: x_{obs}

Observed confound: z_{obs}

Not permuted
 Permuted



(c)