SUPPLEMENTAL MATERIAL

DATA S1. STATISTICAL ANALYSIS PLAN

The purpose of this analysis plan is to provide guide to our analyst when conducting the study. Most of the content will be included in the manuscript in order to guide researchers who want to replicate our findings or conduct similar studies. We also provided justifications for our methods and decisions so other researchers can make a choice or adjust their methods accordingly.

ABBREVIATIONS

Key Definition

First AF Date (variable name first AF date)

The date of the first AF diagnosis within the study period.

Index Date (variable name index_date)

The date 12 months after the first AF date and start of the follow up period.

Baseline Period (variable name baseline)

Time (≥12 months) before the first AF date, used to establish a patient's medical history, and to exclude prior AF diagnosis.

Study Period

The study population will be patients who were newly diagnosed with AF between 7/28/2011- 12/30/2016, which is the enrollment period of the EAST trial, but patients were followed up until 12/31/2019.

Early Rhythm Control Therapy

The study aimed to compare patients treated with early rhythm control therapy (AF ablation and/or AADs), here defined as within the first year of AF diagnosis, and those treated with usual care (rate control drugs). Some patients may be treated with both AF ablation and AADs.

BACKGROUND AND OBJECTIVES

Atrial fibrillation (AF) imposes an increased risk for cardiovascular complications such as death, stroke and myocardial infarction, particularly in the first year after diagnosis.^{1,2} Restoring and maintaining sinus rhythm is associated with reduced mortality.³ Despite improved efficacy and safety of rhythm control therapy, previous trials have failed to demonstrate superiority over rate control.⁴⁻⁶ However, rhythm control therapy appears to be more effective when applied early.^{7,8}

Recently, the Early Treatment of Atrial Fibrillation for Stroke Prevention Trial (EAST-AFNET 4) randomized patients with early-onset AF and increased cardiovascular risk $(CHA₂DS₂-VASC-Score \geq 2) to early rhythm control therapy or current guideline-based usual$ care.⁹ In this trial, stopped for efficacy, early rhythm control was associated with a lower risk of death from cardiovascular causes, stroke, or hospitalization with worsening of heart failure or acute coronary syndrome.

To further assess the generalizability of the EAST-AFNET 4 trial in routine practice in a large cohort of US patients with AF, we assessed the proportion of patients who would have met trial eligibility and examined the association between early rhythm control and clinical outcomes, stratified by trial eligibility.

STUDY DESIGN AND DATA SOURCE

We will conduct a retrospective cohort analysis using OptumLabs Data Warehouse, which contains over 160 million privately insured and Medicare Advantage enrollees of all ages and races from all 50 states.^{10,11} In 2014, this amounted to 19% of all commercially insured and Medicare Advantage beneficiaries in the U.S.

STUDY POPULATION

The study population will be adult patients $(\geq 18$ years) who were newly diagnosed with AF between 7/28/2011-12/30/2016, which is the enrollment period of the EAST trial.

The study population will include two treatment groups: early rhythm control therapy (EAST) group and usual group. The EAST group will include patients who underwent early rhythm control therapy, i.e. AF ablation and/or any AAD therapy, within the first year after AF diagnosis. Some patients may be treated with both AF ablation and AAD. The usual care group will include patients who did not undergo early rhythm control therapy within the first year after AF diagnosis.

We will then limit to those who were older than 75 years of age or had had a previous transient ischemic attack or stroke, or met two of the following criteria: age greater than 65 years, female sex, heart failure, hypertension, diabetes mellitus, severe coronary artery disease, chronic kidney disease (Modification of Diet in Renal Disease stage 3 or 4 [glomerular filtration rate, 15 to 59 ml per minute per 1.73 m2 of body-surface area]), and left ventricular hypertrophy (diastolic septal wall width, >15 mm).

		Generic Names
Anti-arrhythmic drugs		amiodarone, dofetilide, dronedarone, flecainide, propafenone, sotalol, quinidine, disopyramide, moricizine, procainamide, azimilide
Rate control drugs	Beta Blockers	atenolol, bisoprolol, carvedilol, metoprolol, nadolol, nebivolol, propranolol, labetalol
	Calcium Blockers	diltiazem, verapamil
	Cardiac glycosides	digoxin, digitoxin

Table 1. Generic Names of Anti-Arrhythmic Drug Therapy

Patients will be required to have at least 12 months of continuous enrollment in health insurance plans (both medical and prescription drug plans) before the index date, in order to capture an adequate prior medical history and to exclude AF diagnoses prior to the first AF date. Also, Patients were required to have AF diagnosis on at least two different days. Patients whose demographic or residence data are invalid will be excluded. We anticipate that few patients will be under 18 years, but if any patient is under 18 years, they will be excluded as well. We will need to fill out the flow diagram on the next page.

MEASUREMENTS

Baseline Characteristics

Baseline characteristics include socio-demographic characteristics, medical history, concurrent medication use, and previous treatment with rate control drugs. Socio-demographic characteristics include age, sex, race/ethnicity, and region, determined at the time of index date. Race/ethnicity is provided by OptumLabs, classified as non-Hispanic White (White), non-Hispanic Black (Black), Asian, Hispanic, or other/unknown. Self-report was the primary source, and when it was missing, imputation was made by the data provider based on other available administrative data.¹²

Medical history will be determined using patients' physician, facility and pharmacy claims before the index date. We will use all data available to us to establish patients' medical history, and the length of baseline period will be included in the propensity score model to avoid any potential bias. In our previous studies, the baseline period was on average 3-4 years, and there was no substantial difference in the length of the baseline period among different treatment groups, especially after propensity score matching or weighting.

Concurrent medication*,* such as anti-hypertensive and anti-diabetic medications, will be captured within 3 months of the index date. Previous treatment with rhythm or rate control drugs will be captured during the entire baseline, in the form of the number of previous AADs and the number of previous rate control drugs. Although patients with longer baseline period are more likely to have a larger number of previous drugs, the baseline period will not differ between treatment groups, and thus, this should not introduce any undue bias when comparing early rhythm control and usual care patients.

Follow up and Outcomes

OptumLabs Data Warehouse is continuously updated on a monthly basis and the data are complete within 6 months of the service being provided. To avoid potential interaction of the current COVID-19 pandemic with the outcomes, patients will be followed until December 31st, 2019, the end of enrollment in health insurance plans, or death, whichever happened first.

The primary outcomes will be a composite endpoint of all-cause mortality, stroke, or hospitalization with the diagnoses heart failure or acute coronary syndrome, and second, the number of nights spent in the hospital per year, i.e. the same primary endpoints as the EAST trial. The secondary outcomes will include each of these outcomes considered separately.

Mortality will be identified based on the Social Security Death Master File and discharge status. Before November 2011, the Social Security Death Master File has complete mortality data. However, effective on November $1st$, 2011, Section 205(r) of the Social Security Act prohibits the Social Security Administration (SSA) from disclosing state death records that SSA receives through its contracts with the states, except in limited circumstances. Thus, if the SSA knows of a death only from the states and not from any of its other sources of death information, which happens roughly one-third of the time, those death data will not appear on the Death Master File.¹³ Using discharge status (i.e. in-hospital death), we typically capture an additional 30% of deaths in addition to what has been captured by Death Master File. Therefore, most of the deaths missing from Death Master File should be captured by discharge status, particularly since most deaths occur in an institutional setting. We acknowledge that a small proportion of patients who died out of hospital and were not captured by Death Master File could be missing, however, this should be non-differential between treatment groups and should not influence our comparison. In fact, the mortality data is more reliable than most measures derived from administrative data, since its specificity is nearly perfect, and the sensitivity is also very high.

Missing Data

Studies using administrative claims data generally do not have the problem of missing data, *per se*. We will define the presence of a condition, outcome or drug use by the presence of a claim with eligible diagnosis or procedure codes or prescription fills. Patients will be considered to have a comorbidity, outcome or drug exposure if they have a claim, and will be considered not having a comorbidity, outcome or drug exposure if they do not have a claim. Therefore, we do not have missing data in comorbidities, drug use, or outcomes. However, misclassification may exist. This is a limitation of using claims data, but the algorithms used to define our outcomes of interest and important covariates are commonly used and have demonstrated good performance in previous studies.^{14–18} Our internal validation also suggested good performance of the algorithms. We anticipate that any existing residual misclassification will be non-differential between treatment groups and should not meaningfully impact our findings.

For the demographic data, we typically will delete a very small percentage (<1%) of patients with invalid demographic data during the cohort creation process (e.g., missing residence region or inconsistent birth year). For race/ethnicity, the categories in the database are non-Hispanic white, non-Hispanic black, Hispanic, Asian, other and unknown. The other and unknown will be used as a separate category in the propensity score model.

Internal Validation of Diagnosis Codes

The codes and algorithms used herein have been commonly used and validated in many previous studies.14–22

We also leveraged the ability to link to laboratory results and electronic health records to validate our diagnosis codes. For example, we compared the ejection fraction documented in electronic health records and the diagnosis codes for HF. Using a cutoff of LVEF ≤40% for heart failure with reduced ejection fraction (HFrEF) diagnosis codes and LVEF \geq 50% for heart failure with preserved ejection fraction (HFpEF) codes, we observed the specificity of 91% and 81%, respectively, and sensitivity of 81% and 91%, respectively.

We also compared eGFR with the presence of a diagnosis code of Stage 3-4 chronic kidney disease (CKD) in those who did not have renal failure. We found 88% of patients who had a diagnosis of Stage 3-4 CKD had eGFR <60 mL/min/1.73m², and 90% of those who did not have a diagnosis had eGFR ≥ 60 mL/min/1.73m², which indicates good performance of the diagnosis codes. Moreover, the discrepancy between the diagnosis codes and eGFR could be because some patients may have a temporary decline in eGFR, but later recovered and did not develop to CKD or some patients had serum creatinine tests in facilities that did not submit data to the OptumLabs Data Warehouse.

We have also conducted validation of the major bleeding diagnosis codes based on the International Society on Thrombosis and Haemostasis (ISTH) criteria²³: (1) fatal bleeding, and/or, (2) symptomatic bleeding in a critical area or organ, such as intracranial, intraspinal, intraocular, retroperitoneal, intraarticular or pericardial, or intramuscular with compartment syndrome, and/or, (3) bleeding causing a fall in hemoglobin level of 2 g/dL or more, or leading to transfusion of two or more units of whole blood or red cells. We used ICD-9 and CPT procedure codes to identify transfusion, but we were not able to know the units of whole blood or red cells used in the transfusion. We also identified other procedures to control or manage bleeding, such as endoscopic procedures to address gastrointestinal bleeding, neurosurgical decompression for intracranial bleeding, evacuation of hematoma, or vascular embolization procedures to control bleeding. Among all bleeding events, one in four was bleeding in critical areas, and one third required transfusion. This is generally consistent with previous studies that adapted ISTH definition using administrative data.²⁴ Nearly 80% of patients had a procedure to control or manage bleeding. In patients with hemoglobin test results, we abstracted the most recent test performed within six months prior to the bleeding. The median time from the previous hemoglobin test to the date of bleeding is 29 (IQR 8-66) days. The median hemoglobin level during the bleeding was 8.2 (IQR 7.3-11.2) g/dL, with a median drop of 2.1 (IQR 1.1-3.6) g/dL. Among patients with transfusion, the median hemoglobin level was 7.3 (IQR 6.5-8.1) g/dL with a median drop of 2.7 (IQR 1.1-3.6) g/dL. In patients without transfusion, the median hemoglobin level was 10.4 (IQR 8.2-12.3) g/dL, with a median drop of 2.1 (IQR 1.2-3.6) g/dL. Overall, 95% of patients identified using diagnosis codes had bleeding in critical area, or a transfusion, or a procedure used to control bleeding, which suggests high specificity of our algorithm. Even in the remaining 5% patients, the hemoglobin level was low, a median of 10.5 (IQR 8.7-12.0), with a median drop of 2.1 (IQR 1.2-3.5) g/dL.

STATISTICAL METHODS

Main Analyses

We will calculate the proportion of patients who would be excluded from the trial based on the operational definition below (Table 2). We will divide patients to three subgroups: (1) patients who would be eligible for EAST; (2) patients who failed to meet the inclusion criterion, i.e. those under 75 years without any stroke risk factors; (3) patients who met at least one of the exclusion criteria. Some patients may have both failed to meet the inclusion criterion and met the exclusion criteria. In the stratified analyses for clinical outcomes, such patients will be classified as those who met the exclusion criteria.

Note: Two EAST enrollment criteria could not be considered due to lack of availability in our dataset: medical conditions limiting expected survival to <1 year and contraindications for rhythm control therapy

AAD denotes anti-arrhythmic drug, AF atrial fibrillation, CABG coronary artery bypass graft, MI myocardial infarction, PCI percutaneous coronary intervention.

We will use propensity score overlap weighting to account for the differences in baseline characteristics between patients who underwent early rhythm control therapy and those who were treated with usual care (See the next section 5.2). Standardized mean difference will used to assess the balance of covariates after weighting and a difference less than 0.1 will be considered acceptable.25

Cox proportional hazards regression will be used to compare patients treated with early rhythm control therapy and patients treated with usual care in the propensity-score weighted cohort, with a robust sandwich estimator for variance estimation. The regression will be

performed in the overall cohort as well as in each of the three subgroups. The Fine and Gray method will be used to consider death as a competing risk when assessing non-fatal outcomes (i.e., stroke, bleeding, or cardiac arrest when considered separately).26 The proportional hazards assumption will be tested on the basis of Schoenfeld residuals.²⁷ If the proportional hazards assumption does not hold, the hazard ratios will be interpreted as average effects over the observed times, and we will provide the cumulative risks and hazard ratios at different time points to facilitate the interpretation of the effects over time. $28,29$

A *P* value less than 0.05 will considered statistically significant for all tests. All tests will be 2-sided. All analyses will be conducted using SAS Enterprise Guide 7.1 (SAS Institute Inc.) and Stata 16.0 (Stata Corp).

Propensity Score Methods

A propensity score, the probability of undergoing early rhythm control therapy, will be estimated using logistic regression based on socio-demographics, medical history, concurrent medication use, the year of the index date, and the length of baseline period. We will use the overlap weight method to balance treatment groups. The overlap weight will be calculated as 1 minus propensity score for the early rhythm control therapy patients, and the propensity score for the usual care-treated patients. The propensity score and weight will be calculated in each of the three subgroups (patients who were eligible for EAST, patients who fail to meet the inclusion criteria, and patients who meet one of the exclusion criteria) in order to ensure optimal balance in each of the subgroups.

Other commonly used propensity score methods include propensity score matching and inverse probability treatment weighting (IPTW). We will not use propensity score matching as the main method because a large amount of patients may be dropped during matching, however, we will perform a sensitivity analysis using propensity score matching. We will not use IPTW, since IPTW gave imprecise estimates of treatment effect and undue influence to a small number of observations when substantial confounding was present.³⁰ The performance of IPTW often gets worse when the prevalence of treatment is low.31

We chose the overlap weight because this approach minimizes the asymptotic variance of the treatment effect, while also possessing a desirable exact balance property.32 Unlike IPTW, the overlap weights are bounded between 0 and 1 and thus are less sensitive to extreme weights. Compared to the common practice of truncating weights or discarding patients with extreme weights, the overlap weights avoid this arbitrary choice of a cutoff point for inclusion. The overlap weight also possesses an attractive exact balance property, i.e., the means of all variables (including the proportions of a binary or categorical variable) will be exactly the same between treatment and control groups after weighting.

The results using the overlap weight should be interpreted as the average treatment effect for the overlap population. The overlap population typically represents a target population of intrinsic substantive interest, i.e. patients who could appear in either treatment groups. In such patients, clinical consensus regarding the treatment choice is often ambiguous and thus research is most needed to guide decision making.

Sensitivity Analyses

We will conduct a few sensitivity analyses to assess the robustness of the findings. First, propensity score matching will be used instead of propensity score weighting for the primary outcome. One-to-one nearest neighborhood caliper matching will be used to match patients based on the logit of the propensity score using a caliper equal to 0.2 of the standard deviation of the logit of the propensity score.³³ Patients will be exact matched on whether they were eligible for the trial, failed to meet the inclusion criterion, or met at least one of the exclusion criteria.

Second, we will conduct a stratified analysis based on whether the early rhythm controltreated patients were treated with AF ablation or without AF ablation. To conduct the stratified analysis, we will first recalculate the propensity score weights to balance patients treated with early rhythm control and patients treated with usual care, and perform regression analyses to compare early rhythm control to usual care; we will then recalculate the weights to balance patients treated with AF ablation and patients treated with usual care, and perform regression analyses to compare AF ablation to usual care. Some of the early rhythm control-treated patients may have been treated with both AADs and AF ablation, and such patients will be classified to the ablation group.

Third, a similar stratified analysis will be conducted based on the adherence to AADs in the early rhythm control-treated patients, i.e., patients with proportion of days covered (PDC)<80% and those with PDC≥80%, since the adherence to AAD therapy in practice is often lower than that in clinical trials. The adherence will consider all rhythm control drugs that a patient used during follow up, even if they were different from the initial treatment. To conduct the stratified analysis, we will first recalculate the propensity score weights to balance patients who were treated with AADs and adherent and patients who were treated with usual care, and perform regression analyses to compare usual care-treated patients to adherent AAD-treated patients; we will then recalculate the weights to balance patients who were treated with usual care and patients who were treated with AADs and not adherent, and perform regression analyses to compare usual care-treated patients to non-adherent AAD-treated patients.

Subgroup Analyses

We will perform subgroup analyses for the primary outcome stratified by age, sex, race, CHA2DS2-VASc, hypertension with left ventricular hypertrophy, heart failure, cardiomyopathy, sleep apnea, and prior thromboembolism. The subgroup analyses will be performed separately in patients who were eligible for the trial, patients who failed to meet the trial inclusion criterion, and patients who met at least one of the trial exclusion criteria. Patients who failed the trial inclusion criterion are those under 75 years without stroke risk factors, therefore, we will perform subgroup analyses only by sex and race.

Since an increasing number of subgroup analyses could increase the chance of false positive results, we pre-specified the above subgroups since they are either key demographic characteristics or risk factors strongly associated with the primary outcome. The subgroup analyses will not only explore whether there is any heterogeneity in treatment effects, but also help understand whether there is any subgroup of patients who may benefit from ablation but were not adequately represented in the trial.

For all analyses performed in this study, we will not perform any adjustment for multiple testing. The sample size will be large and thus even with the conservative Bonferroni adjustment, many tests will still be statistically significant. We will consider all the analyses except those related to the primary outcome exploratory. However, if the exploratory results, e.g., treatment heterogeneity in certain subgroups, are consistent with the EAST trial or are confirmed by future studies, the results will more likely to be a true finding.

Residual Confounding

We will assess falsification endpoints to test for residual confounding. Treatment effects estimated in observational studies are prone to unmeasured confounding. In recent years, falsification end point, also called control outcome, has become a popular method to assess for unmeasured confounding.^{34–36} A falsification endpoint is a health outcome that researchers believe is highly unlikely to be casually related to the treatment in question. If a significant relationship is found between the treatment and a falsification endpoint, it may indicate the treatment groups are different in some unmeasured ways, i.e. the existence of unmeasured confounding. This method is similar to a negative control, a routine precaution taken in the

design of biologic laboratory experiments, and is recommended to be used to detect confounding and bias in observational studies.^{35,37,38} We selected three endpoints that that are unlikely to be a result of undergoing early-rhythm control therapy – emergency room visit or hospitalization related to chronic obstructive pulmonary disease (COPD), pneumonia, and fracture.

LIMITATIONS

Our study relies on administrative data to ascertain baseline characteristics and outcomes, which could be subject to misclassification. However, it is unlikely there is any systematic difference in the ascertainment of comorbidities and outcomes between different treatment groups, and thus, the misclassification should not meaningfully impact our comparisons between drugs. The diagnosis and procedure codes used in this study have been commonly used in previous studies, and demonstrated good performance in our internal validation using linked laboratory results and electronic health records (described in Section 4.4) as well as other validation studies with positive predictive value around 90%.^{14,39-42}

Second, our study will only include privately insured and Medicare Advantage patients. The patient characteristics and outcomes could be different in the Medicaid, Medicare Fee-for-Service, and uninsured populations. However, the insurance coverage rates are high in older Americans. Over 90% of Americans aged 50-64 have health insurance and over 75% had private health insurance.⁴³ One in three Medicare patients is enrolled in Medicare Advantage.⁴⁴ Although traditionally Medicare Advantage attracted healthier people, after the risk adjustment system was phased in from 2004-2007, the favorable risk selection has been largely reduced.⁴⁵

In fact, the results from this study will be more generalizable than most observational studies using other data sources. Observational studies largely use either administrative data or registries. Some cardiovascular registries focused on cardiology practices for recruitment and patients have to sign informed consent and agree to participate and to be actively followed, and thus the patients in these registries were more selective. Some administrative data are limited within a health system, within a region, or within an age range (e.g., Medicare, Kaiser, etc.). The OptumLabs Data Warehouse contains patients of all ages and races managed at heterogeneous practice settings from all 50 states.^{10,11} The distribution of patient characteristics (e.g., age, sex and race/ethnicity) in the database is similar to those of the general U.S.

population.11 The data are updated monthly and are generally believed to be timely, accurate, and reflective of contemporary practice patterns. The concordance between OptumLabs and everyday practice is a major strength of the data source.

References

- 1. Benjamin EJ, Wolf PA, D'Agostino RB, Silbershatz H, Kannel WB, Levy D. Impact of atrial fibrillation on the risk of death: The Framingham Heart Study. *Circulation*. 1998;98(10):946-952. doi:10.1161/01.CIR.98.10.946
- 2. Friberg L, Hammar N, Pettersson H, Rosenqvist M. Increased mortality in paroxysmal atrial fibrillation: Report from the Stockholm Cohort-Study of Atrial Fibrillation (SCAF). *Eur Heart J*. 2007;28(19):2346-2353. doi:10.1093/eurheartj/ehm308
- 3. Van Gelder IC, Hagens VE, Bosker HA, et al. A Comparison of Rate Control and Rhythm Control in Patients with Recurrent Persistent Atrial Fibrillation. *N Engl J Med*. 2002;347(23):1834-1840. doi:10.1056/NEJMoa021375
- 4. Wyse DG, Waldo AL, DiMarco JP, et al. A Comparison of Rate Control and Rhythm Control in Patients with Atrial Fibrillation. *N Engl J Med*. 2002;347(23):1825-1833. doi:10.1056/NEJMoa021328
- 5. Carlsson J, Miketic S, Windeler J, et al. Randomized trial of rate-control versus rhythm-control in persistent atrial fibrillation: The strategies of treatment of atrial fibrillation (STAF) study. *J Am Coll Cardiol*. 2003;41(10):1690-1696. doi:10.1016/S0735-1097(03)00332-2
- 6. Roy D, Talajic M, Nattel S, et al. Rhythm Control versus Rate Control for Atrial Fibrillation and Heart Failure. *N Engl J Med*. 2008;358(25):2667-2677. doi:10.1056/NEJMoa0708789
- 7. Kirchhof P, Bax J, Blomstrom-Lundquist C, et al. Early and comprehensive management of atrial fibrillation: Executive summary of the proceedings from the 2nd AFNET-EHRA consensus conference "research perspectives in AF." *Eur Heart J*.

2009;30(24):2969-2980. doi:10.1093/eurheartj/ehp235

- 8. Kirchhof P. Can we improve outcomes in AF patients by early therapy? *BMC Med*. 2009;7(1):72. doi:10.1186/1741-7015-7-72
- 9. Kirchhof P, Camm AJ, Goette A, et al. Early Rhythm-Control Therapy in Patients with Atrial Fibrillation. *N Engl J Med*. 2020;383(14):1305-1316. doi:10.1056/nejmoa2019422
- 10. Wallace PJ, Shah ND, Dennen T, Bleicher PA, Crown WH. Optum labs: Building a novel node in the learning health care system. *Health Aff*. 2014;33(7):1187-1194. doi:10.1377/hlthaff.2014.0038
- 11. *Optum Research Data Assets*. https://www.optum.com/content/dam/optum/resources/productSheets/5302_Data_Asse ts Chart Sheet ISPOR.pdf. Accessed October 17, 2020.
- 12. Hershman DL, Tsui J, Wright JD, Coromilas EJ, Tsai WY, Neugut AI. Household net worth, racial disparities, and hormonal therapy adherence among women with earlystage breast cancer. *J Clin Oncol*. 2015;33(9):1053-1059. doi:10.1200/JCO.2014.58.3062
- 13. Da Graca B, Filardo G, Nicewander D. Consequences for healthcare quality and research of the exclusion of records from the death master file. *Circ Cardiovasc Qual Outcomes*. 2013;6(1):124-128. doi:10.1161/CIRCOUTCOMES.112.968826
- 14. Tirschwell DL, Longstreth WT. Validating administrative data in stroke research. *Stroke*. 2002;33(10):2465-2470. doi:10.1161/01.STR.0000032240.28636.BD
- 15. Cunningham A, Stein CM, Chung CP, Daugherty JR, Smalley WE, Ray WA. An automated database case definition for serious bleeding related to oral anticoagulant

use. *Pharmacoepidemiol Drug Saf*. 2011;20(6):560-566. doi:10.1002/pds.2109

- 16. Arnason T, Wells PS, van Walraven C, Forster AJ. Accuracy of coding for possible warfarin complications in hospital discharge abstracts. *Thromb Res*. 2006;118(2):253- 262. doi:10.1016/j.thromres.2005.06.015
- 17. Hwang YJ, Shariff SZ, Gandhi S, et al. Validity of the International Classification of Diseases, Tenth Revision code for acute kidney injury in elderly patients at presentation to the emergency department and at hospital admission. *BMJ Open*. 2012;2(6). doi:10.1136/bmjopen-2012-001821
- 18. Wilchesky M, Tamblyn RM, Huang A. Validation of diagnostic codes within medical services claims. *J Clin Epidemiol*. 2004;57(2):131-141. doi:10.1016/S0895- 4356(03)00246-4
- 19. Yao X, Shah ND, Sangaralingham LR, Gersh BJ, Noseworthy PA. Non–Vitamin K Antagonist Oral Anticoagulant Dosing in Patients With Atrial Fibrillation and Renal Dysfunction. *J Am Coll Cardiol*. 2017;69(23):2779-2790. doi:10.1016/j.jacc.2017.03.600
- 20. Yao X, Tangri N, Gersh BJ, et al. Renal Outcomes in Anticoagulated Patients With Atrial Fibrillation. *J Am Coll Cardiol*. 2017;70(21):2621-2632. doi:10.1016/j.jacc.2017.09.1087
- 21. Noseworthy PA, Gersh BJ, Kent DM, et al. Atrial fibrillation ablation in practice: Assessing CABANA generalizability. *Eur Heart J*. 2019;40(16):1257-1264. doi:10.1093/eurheartj/ehz085
- 22. Noseworthy PA, Van Houten HK, Gersh BJ, et al. Generalizability of the CASTLE-AF trial: Catheter ablation for patients with atrial fibrillation and heart failure in routine

practice. *Hear Rhythm*. 2020;17(7):1057-1065. doi:10.1016/j.hrthm.2020.02.030

- 23. Schulman S, Kearon C. Definition of major bleeding in clinical investigations of antihemostatic medicinal products in non-surgical patients. *J Thromb Haemost*. 2005;3(4):692-694. doi:10.1111/j.1538-7836.2005.01204.x
- 24. Jasuja GK, Reisman JI, Miller DR, et al. Identifying major hemorrhage with automated data: Results of the Veterans Affairs Study to Improve Anticoagulation (VARIA). *Thromb Res*. 2013;131(1):31-36. doi:10.1016/j.thromres.2012.10.010
- 25. Austin PC. Balance diagnostics for comparing the distribution of baseline covariates between treatment groups in propensity-score matched samples. *Stat Med*. 2009;28(25):3083-3107. doi:10.1002/sim.3697
- 26. Fine JP, Gray RJ. A Proportional Hazards Model for the Subdistribution of a Competing Risk. *J Am Stat Assoc*. 1999;94(446):496. doi:10.2307/2670170
- 27. Grambsch PM, Therneau TM. Proportional hazards tests and diagnostics based on weighted residuals. *Biometrika*. 1994;81(3):515-526. doi:10.1093/biomet/81.3.515
- 28. Therneau TM, Grambsch PM. *Modeling Survival Data: Extending the Cox Model*. New York, NY: Springer New York; 2000. doi:10.1007/978-1-4757-3294-8
- 29. Weintraub WS, Grau-Sepulveda M V., Weiss JM, et al. Comparative effectiveness of revascularization strategies. *N Engl J Med*. 2012;366(16):1467-1476. doi:10.1056/NEJMoa1110717
- 30. Elze MC, Gregson J, Baber U, et al. Comparison of Propensity Score Methods and Covariate Adjustment: Evaluation in 4 Cardiovascular Studies. *J Am Coll Cardiol*. 2017;69(3):345-357. doi:10.1016/j.jacc.2016.10.060
- 31. Austin PC, Schuster T. The performance of different propensity score methods for

estimating absolute effects of treatments on survival outcomes: A simulation study. *Stat Methods Med Res*. 2016;25(5):2214-2237. doi:10.1177/0962280213519716

- 32. Li F, Morgan KL, Zaslavsky AM. Balancing Covariates via Propensity Score Weighting. *J Am Stat Assoc*. 2018;113(521):390-400. doi:10.1080/01621459.2016.1260466
- 33. Austin PC. Optimal caliper widths for propensity-score matching when estimating differences in means and differences in proportions in observational studies. *Pharm Stat*. 2011;10(2):150-161. doi:10.1002/pst.433
- 34. Prasad V, Jena AB. Prespecified falsification end points: Can they validate true observational associations? *JAMA - J Am Med Assoc*. 2013;309(3):241-242. doi:10.1001/jama.2012.96867
- 35. Dusetzina SB, Brookhart MA, MacIejewski ML. Control outcomes and exposures for improving internal validity of nonrandomized studies. *Health Serv Res*. 2015;50(5):1432-1451. doi:10.1111/1475-6773.12279
- 36. Ioannidis JPA. Are mortality differences detected by administrative data reliable and actionable? *JAMA - J Am Med Assoc*. 2013;309(13):1410-1411. doi:10.1001/jama.2013.3150
- 37. Lipsitch M, Tchetgen Tchetgen E, Cohen T. Negative Controls: A tool for detecting confounding and bias in observational studies. *Epidemiology*. 2010;21(3):383-388. doi:10.1097/EDE.0b013e3181d61eeb
- 38. Wimmer NJ, Resnic FS, Mauri L, Matheny ME, Yeh RW. Comparison of transradial versus transfemoral percutaneous coronary intervention in routine practice: Evidence for the importance of "falsification hypotheses" in observational studies of comparative

effectiveness. *J Am Coll Cardiol*. 2013;62(22):2147-2148. doi:10.1016/j.jacc.2013.07.036

- 39. Kumamaru H, Judd SE, Curtis JR, et al. Validity of claims-based stroke algorithms in contemporary medicare data: Reasons for geographic and racial differences in stroke (REGARDS) study linked with medicare claims. *Circ Cardiovasc Qual Outcomes*. 2014;7(4):611-619. doi:10.1161/CIRCOUTCOMES.113.000743
- 40. Kokotailo RA, Hill MD. Coding of stroke and stroke risk factors using International Classification of Diseases, revisions 9 and 10. *Stroke*. 2005;36(8):1776-1781. doi:10.1161/01.STR.0000174293.17959.a1
- 41. Jensen PN, Johnson K, Floyd J, Heckbert SR, Carnahan R, Dublin S. A systematic review of validated methods for identifying atrial fibrillation using administrative data. *Pharmacoepidemiol Drug Saf*. 2012;21(SUPPL. 1):141-147. doi:10.1002/pds.2317
- 42. Fan J, Arruda-Olson AM, Leibson CL, et al. Billing code algorithms to identify cases of peripheral artery disease from administrative data. *J Am Med Informatics Assoc*. 2013;20(E2). doi:10.1136/amiajnl-2013-001827
- 43. Barnett J, Vornovitsky M. Health Insurance Coverage in the United States: 2015 Current Population Reports. *28th Annu Data Users Conf Ref Doc*. September 2016. https://digitalcommons.unomaha.edu/datausers_reference_2017/4. Accessed October 17, 2020.
- 44. Medicare Advantage 2016 Spotlight: Enrollment Market Update | KFF. https://www.kff.org/medicare/issue-brief/medicare-advantage-2016-spotlightenrollment-market-update/. Accessed October 17, 2020.
- 45. McWilliams JM, Hsu J, Newhouse JP. New risk-adjustment system was associated

with reduced favorable selection in medicare advantage. *Health Aff*. 2012;31(12):2630-

2640. doi:10.1377/hlthaff.2011.1344

ICD-9-CM denotes International Classification of Diseases, 9th Revision, Clinical Modification, ICD-10-CM International Classification of Diseases, 10th Revision, Clinical Modification, and CPT current procedural terminology.

Table S3. EAST-AFNET 4 Trial Eligibility Criteria

Table S3. EAST-AFNET 4 Trial Eligibility Criteria

Note: Two EAST enrollment criteria could not be considered due to lack of availability in our dataset: medical conditions limiting expected survival to <1 year and contraindications for rhythm control therapy

AAD denotes anti-arrhythmic drug, AF atrial fibrillation, CABG coronary artery bypass graft, MI myocardial infarction, PCI percutaneous coronary intervention.

AAD denotes anti-arrhythmic drug, ACE angiotensin-converting enzyme, AF atrial fibrillation, ARB angiotensin II receptor blockers, CABG coronary artery bypass grafting, CAD coronary artery disease, CKD chronic kidney disease, COPD chronic obstructive pulmonary disease, CRT cardiac resynchronization therapy, HCM hypertrophic cardiomyopathy, ICD implantable cardioverter defibrillators, ILR implantable loop recorder, NSAID nonsteroidal anti-inflammatory drug, PAD peripheral artery disease, PCI percutaneous coronary intervention, PS propensity score, TIA transient ischemic attack.The CHA₂DS₂-VASc score is a 0- to 9-point stroke risk score where a higher point score indicates higher risk of stroke. The point score is calculated as follows: 1 point each for heart failure, hypertension, diabetes, vascular disease, age 65 to 74 years, and female sex and 2 points for age 75 years or older and prior thromboembolism (including ischemic stroke, TIA or systemic embolism).

Concurrent medication use was defined as prescriptions within three months prior to the index date.

* To maintain de-identification, OptumLabs does not allow researchers to disclose the number of events when the number is 10 or fewer.

AAD denotes anti-arrhythmic drug, ACE angiotensin-converting enzyme, AF atrial fibrillation, ARB angiotensin II receptor blockers, CABG coronary artery bypass grafting, CAD coronary artery disease, CKD chronic kidney disease, COPD chronic obstructive pulmonary disease, CRT cardiac resynchronization therapy, HCM hypertrophic cardiomyopathy, ICD implantable cardioverter defibrillators, ILR implantable loop recorder, NSAID nonsteroidal anti-inflammatory drug, PAD peripheral artery disease, PCI percutaneous coronary intervention, PS propensity score, TIA transient ischemic attack. The CHA₂DS₂-VASc score is a 0- to 9-point stroke risk score where a higher point score indicates higher risk of stroke. The point score is calculated as follows: 1 point each for heart failure, hypertension, diabetes, vascular disease, age 65 to 74 years, and female sex and 2 points for age 75 years or older and prior thromboembolism (including ischemic stroke, TIA or systemic embolism). Concurrent medication use was defined as prescriptions within three months prior to the index date. * To maintain de-identification, OptumLabs does not allow researchers to disclose the number of events when the number is 10 or fewer.

AAD denotes anti-arrhythmic drug, ACE angiotensin-converting enzyme, AF atrial fibrillation, ARB angiotensin II receptor blockers, CABG coronary artery bypass grafting, CAD coronary artery disease, CKD chronic kidney disease, COPD chronic obstructive pulmonary disease, CRT cardiac resynchronization therapy, HCM hypertrophic cardiomyopathy, ICD implantable cardioverter defibrillators, ILR implantable loop recorder, NSAID nonsteroidal anti-inflammatory drug, PAD peripheral artery disease, PCI percutaneous coronary intervention, PS propensity score, TIA transient ischemic attack.The CHA₂DS₂-VASc score is a 0- to 9-point stroke risk score where a higher point score indicates higher risk of stroke. The point score is calculated as follows: 1 point each for heart failure, hypertension, diabetes, vascular disease, age 65 to 74 years, and female sex and 2 points for age 75 years or older and prior thromboembolism (including ischemic stroke, TIA or systemic embolism). Concurrent medication use was defined as prescriptions within three months prior to the index date. * To maintain de-identification, OptumLabs does not allow researchers to disclose the number of events when the number is 10 or fewer.

CI, confidence interval; LVH, left ventricular hypertrophy; SHF, systolic heart failure; CM, cardiomyopathy; OSA, obstructive sleep apnea; TE, thromboembolism.

Table S8. Subgroup Analysis for the Secondary Outcome Hospitalization with the Diagnosis Heart Failure in Propensity Score Weighted Patients (Overall Cohort)

Table S8. Subgroup Analysis for the Secondary Outcome Hospitalization with the Diagnosis Heart Failure in Propensity Score Weighted Patients (Overall Cohort)

CI, confidence interval; LVH, left ventricular hypertrophy; SHF, systolic heart failure; CM, cardiomyopathy; OSA, obstructive sleep apnea; TE, thromboembolism.

Table S9. Subgroup Analysis for the Secondary Outcome Hospitalization with the Diagnosis Myocardial Infarction in Propensity Score Weighted Patients (Overall Cohort)

Table S9. Subgroup Analysis for the Secondary Outcome Hospitalization with the Diagnosis Myocardial Infarction in Propensity Score Weighted Patients (Overall Cohort)

CI, confidence interval; LVH, left ventricular hypertrophy; SHF, systolic heart failure; CM, cardiomyopathy; OSA, obstructive sleep apnea; TE, thromboembolism.

Table S10. Subgroup Analysis for the Secondary Outcome All-Cause Mortality in Propensity Score Weighted Patients (Overall Cohort)

Table S10. Subgroup Analysis for the Secondary Outcome All-Cause Mortality in Propensity Score Weighted Patients (Overall Cohort)

CI, confidence interval; LVH, left ventricular hypertrophy; SHF, systolic heart failure; CM, cardiomyopathy; OSA, obstructive sleep apnea; TE, thromboembolism.

Table S11. Subgroup Analysis for the Primary Outcome in Propensity Score Weighted Patients (Trial Eligible Patients)

CI, confidence interval; LVH, left ventricular hypertrophy; SHF, systolic heart failure; CM, cardiomyopathy; OSA, obstructive sleep apnea; TE, thromboembolism.

Table S12. Subgroup Analysis for the Primary Outcome in Propensity Score Weighted Patients (Trial Ineligible Patients)

CI, confidence interval; LVH, left ventricular hypertrophy; SHF, systolic heart failure; CM, cardiomyopathy; OSA, obstructive sleep apnea; TE, thromboembolism.

Table S13. Sensitivity Analyses Stratified by Treatment with AF Ablation or without AF Ablation in the Early Rhythm-Control Therapy Cohort

Table S13. Sensitivity Analyses Stratified by Treatment with AF Ablation or without AF Ablation in the Early Rhythm-Control Therapy Cohort

Table S13. Sensitivity Analyses Stratified by Treatment with AF Ablation or without AF Ablation in the Early Rhythm-Control Therapy Cohort

First, we recalculated the propensity score weights to balance patients treated with early rhythm-control and patients treated without early rhythm-control and performed regression analyses to compare early rhythm-control to the control group; we then recalculated the weights to balance patients treated with AF ablation and patients treated without early rhythm-control and performed regression analyses to compare AF ablation to the control group. Patients treated with both AAD therapy and AF ablation were classified to the ablation group. AAD, anti-arrhythmic drug; AF, atrial fibrillation; CI, confidence interval; HF, hospitalization with the diagnosis heart failure; MI, hospitalization with the diagnosis myocardial infarction.

Adherence was defined as proportion of days covered (PDC) ≥80% in the timeframe between first AF date to index date. The adherence considered all rhythmcontrol drugs that patients used, even if they were different from the initial treatment. We first recalculated the propensity score weights to balance patients who were treated with AADs who were adherent and patients who were treated without early rhythm-control, and performed regression analyses to compare patients treated without early rhythm-control to adherent AAD-treated patients; we then recalculated the weights to balance patients who were treated without early rhythm-control and patients who were treated with AADs who were not adherent, and performed regression analyses to compare patients treated without early rhythm-control to non-adherent AAD-treated patients. AAD, anti-arrhythmic drug; CI, confidence interval; HF, hospitalization with the diagnosis heart failure; MI, hospitalization with the diagnosis myocardial infarction.

Adherence was defined as proportion of days covered (PDC) ≥80% in the timeframe between first AF date to index date. The adherence considered all rhythmcontrol drugs that patients used, even if they were different from the initial treatment. We first recalculated the propensity score weights to balance patients who were treated with AADs who were adherent and patients who were treated without early rhythm-control, and performed regression analyses to compare patients treated without early rhythm-control to adherent AAD-treated patients; we then recalculated the weights to balance patients who were treated without early rhythm-control and patients who were treated with AADs who were not adherent, and performed regression analyses to compare patients treated without early rhythm-control to non-adherent AAD-treated patients. AAD, anti-arrhythmic drug; CI, confidence interval; HF, hospitalization with the diagnosis heart failure; MI, hospitalization with the diagnosis myocardial infarction.

Adherence was defined as proportion of days covered (PDC) ≥80% in the timeframe between first AF date to index date. The adherence considered all rhythmcontrol drugs that patients used, even if they were different from the initial treatment. We first recalculated the propensity score weights to balance patients who were treated with AADs who were adherent and patients who were treated without early rhythm-control, and performed regression analyses to compare patients treated without early rhythm-control to adherent AAD-treated patients; we then recalculated the weights to balance patients who were treated without early rhythm-control and patients who were treated with AADs who were not adherent, and performed regression analyses to compare patients treated without early rhythm-control to non-adherent AAD-treated patients. AAD, anti-arrhythmic drug; CI, confidence interval; HF, hospitalization with the diagnosis heart failure; MI, hospitalization with the diagnosis myocardial infarction.

Table S17. Falsification Endpoint Test in Propensity Score Weighted Cohort

Outcomes were captured by primary diagnosis during an emergency room visit or an inpatient stay.