

S8 Text: Connecting SCRaPL error model to likelihoods currently employed by practitioners.

SCRaPL: A Bayesian hierarchical framework for detecting technical associates in single cell multiomics data.

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In the field of transcriptomics it is widely accepted that over-dispersion is an important feature of state of the art models [1]. Following the example of several papers (e.g, [1; 2; 3]) where they provide evidence that their model can handle excessive zeros by using over-dispersion, we demonstrate that zero-inflated Poisson is also a valid alternative as the over-dispersion property exists. For the sake of completeness we provide mean and variance formulas for the binomial posterior.

1 Zero-inflated Poisson

$$\mathbb{P}(y) = \begin{cases} \int_{-\infty}^{\infty} \pi + (1 - \pi)e^{-ae^x} \mathbb{N}(x, \mu, \sigma^2) dx, & \text{if } y = 0 \\ \int_{-\infty}^{\infty} \frac{1-\pi}{y!} a^y e^{yx-ae^x} \mathbb{N}(x, \mu, \sigma^2) dx, & \text{else} \end{cases} \quad (1)$$

$$\begin{aligned} \mathbb{E}(y) &= \sum_{y=0}^{\infty} y \mathbb{P}(y) = \sum_{y=1}^{\infty} y \mathbb{P}(y) = \sum_{y=1}^{\infty} y \frac{1-\pi}{y!} a^y \int_{-\infty}^{\infty} e^{yx-ae^x} \mathbb{N}(x, \mu, \sigma^2) dx \\ &= \sum_{y=1}^{\infty} \frac{1-\pi}{(y-1)!} a^y \int_{-\infty}^{\infty} e^{yx-ae^x} \mathbb{N}(x, \mu, \sigma^2) dx \\ &= a(1-\pi) \int_{-\infty}^{\infty} e^x e^{-ae^x} \mathbb{N}(x, \mu, \sigma^2) \sum_{y=1}^{\infty} a^{y-1} \frac{e^{(y-1)x}}{(y-1)!} dx \\ &= a(1-\pi) \int_{-\infty}^{\infty} e^x e^{-ae^x} \mathbb{N}(x, \mu, \sigma^2) \sum_{y=0}^{\infty} a^y \frac{e^{yx}}{y!} dx \\ &= a(1-\pi) \int_{-\infty}^{\infty} e^x e^{-ae^x} \mathbb{N}(x, \mu, \sigma^2) e^{ae^x} dx \\ &= a(1-\pi) \int_{-\infty}^{\infty} e^x \mathbb{N}(x, \mu, \sigma^2) dx = a(1-\pi) e^{\mu + \frac{\sigma^2}{2}} \end{aligned}$$

Hence,

$$\mathbb{E}(y) = a(1-\pi) e^{\mu + \frac{\sigma^2}{2}} = m \quad (2)$$

$$\begin{aligned}
\mathbb{E}(y^2) &= \sum_{y=0}^{\infty} y^2 \mathbb{P}(y) = \sum_{y=1}^{\infty} y^2 \mathbb{P}(y) \\
&= \sum_{y=1}^{\infty} y^2 \frac{1-\pi}{y!} \int_{-\infty}^{\infty} a^y e^{yx-ae^x} \mathbb{N}(x, \mu, \sigma^2) dx = \sum_{y=1}^{\infty} y \frac{1-\pi}{(y-1)!} \int_{-\infty}^{\infty} a^y e^{yx-ae^x} \mathbb{N}(x, \mu, \sigma^2) dx \\
&= a(1-\pi) \int_{-\infty}^{\infty} e^x e^{-ae^x} \mathbb{N}(x, \mu, \sigma^2) \sum_{y=0}^{\infty} \frac{(y+1)a^y e^{yx}}{y!} dx \\
&= a(1-\pi) \int_{-\infty}^{\infty} e^x e^{-ae^x} \mathbb{N}(x, \mu, \sigma^2) \sum_{y=0}^{\infty} \frac{y a^y e^{yx}}{y!} dx + a(1-\pi) \int_{-\infty}^{\infty} e^x e^{-ae^x} \mathbb{N}(x, \mu, \sigma^2) \sum_{y=0}^{\infty} \frac{a^y e^{yx}}{y!} dx \\
&= a(1-\pi) \int_{-\infty}^{\infty} e^x e^{-ae^x} \mathbb{N}(x, \mu, \sigma^2) \sum_{y=1}^{\infty} \frac{a^y e^{yx}}{(y-1)!} dx + a(1-\pi) \int_{-\infty}^{\infty} e^x \mathbb{N}(x, \mu, \sigma^2) dx \\
&= a^2(1-\pi) \int_{-\infty}^{\infty} e^{2x} e^{-ae^x} \mathbb{N}(x, \mu, \sigma^2) \sum_{y=0}^{\infty} \frac{a^y e^{yx}}{y!} dx + a(1-\pi) \int_{-\infty}^{\infty} e^x \mathbb{N}(x, \mu, \sigma^2) dx \\
&= a^2(1-\pi) \int_{-\infty}^{\infty} e^{2x} \mathbb{N}(x, \mu, \sigma^2) dx + a(1-\pi) \int_{-\infty}^{\infty} e^x \mathbb{N}(x, \mu, \sigma^2) dx = a(1-\pi) \left[a e^{2\mu+2\sigma^2} + e^{\mu+\frac{\sigma^2}{2}} \right]
\end{aligned}$$

$$\begin{aligned}
\mathbb{V}(y) &= \mathbb{E}(y^2) - \mathbb{E}(y)^2 = a(1-\pi) e^{\mu+\frac{\sigma^2}{2}} \left(1 + a e^{\mu+\frac{3\sigma^2}{2}} \right) - a^2(1-\pi)^2 e^{2\mu+\sigma^2} \\
&= a(1-\pi) e^{\mu+\frac{\sigma^2}{2}} \left[1 + a e^{\mu+\frac{3\sigma^2}{2}} - a(1-\pi) e^{\mu+\frac{\sigma^2}{2}} \right] \\
&= a(1-\pi) e^{\mu+\frac{\sigma^2}{2}} \left[1 + a e^{\mu+\frac{\sigma^2}{2}} \left(e^{\sigma^2} - 1 + \pi \right) \right] = m \left(1 + m \frac{e^{\sigma^2} - 1 + \pi}{1 - \pi} \right)
\end{aligned}$$

Hence,

$$\mathbb{V}(y) = m \left(1 + m \frac{e^{\sigma^2} - 1 + \pi}{1 - \pi} \right) \quad (3)$$

2 Binomial

$$\begin{aligned}
\mathbb{P}(k) &= \int_{-\infty}^{\infty} \binom{n}{k} \Phi(x)^k (1 - \Phi(x))^{n-k} \mathbb{N}(x, \mu, \sigma^2) dx \quad (4) \\
\mathbb{E}(k) &= \sum_{k=0}^n k \mathbb{P}(k) = \int_{-\infty}^{\infty} \mathbb{N}(x, \mu, \sigma^2) \sum_{k=0}^n k \binom{n}{k} \Phi(x)^k (1 - \Phi(x))^{n-k} dx \\
&= n \int_{-\infty}^{\infty} \Phi(x) \mathbb{N}(x, \mu, \sigma^2) dx = n \Phi \left(\frac{\mu}{\sqrt{1 + \sigma^2}} \right)
\end{aligned}$$

Hence,

$$\mathbb{E}(k) = n \Phi \left(\frac{\mu}{\sqrt{1 + \sigma^2}} \right) \quad (5)$$

$$\begin{aligned}
\mathbb{E}(k(k-1)) &= \sum_{k=0}^n k(k-1)\mathbb{P}(k) = \int_{-\infty}^{\infty} \mathbb{N}(x, \mu, \sigma^2) \sum_{k=0}^n k(k-1) \binom{n}{k} \Phi(x)^k (1-\Phi(x))^{n-k} dx \\
&= n(n-1) \int_{-\infty}^{\infty} \Phi(x)^2 \mathbb{N}(x, \mu, \sigma^2) = n(n-1) \left[\Phi\left(\frac{\mu}{\sqrt{1+\sigma^2}}\right) - 2T\left(\frac{\mu}{\sqrt{1+\sigma^2}}, \frac{1}{\sqrt{1+2\sigma^2}}\right) \right], \\
&\qquad\qquad\qquad T(h, a) = \mathbb{N}(h, 0, 1) \int_0^a \frac{\mathbb{N}(hx, 0, 1)}{1+x^2} dx
\end{aligned}$$

$$\begin{aligned}
\mathbb{V}(k) &= \mathbb{E}(k^2) - \mathbb{E}(k)^2 = \mathbb{E}(k(k-1)) + \mathbb{E}(k) - \mathbb{E}(k)^2 \\
&= n(n-1) \left[\Phi\left(\frac{\mu}{\sqrt{1+\sigma^2}}\right) - 2T\left(\frac{\mu}{\sqrt{1+\sigma^2}}, \frac{1}{\sqrt{1+2\sigma^2}}\right) \right] + n\Phi\left(\frac{\mu}{\sqrt{1+\sigma^2}}\right) - n^2\Phi\left(\frac{\mu}{\sqrt{1+\sigma^2}}\right)^2 \\
&\qquad\qquad\qquad = n^2\Phi\left(\frac{\mu}{\sqrt{1+\sigma^2}}\right) \left[1 - \Phi\left(\frac{\mu}{\sqrt{1+\sigma^2}}\right) \right] - 2n(n-1)T\left(\frac{\mu}{\sqrt{1+\sigma^2}}, \frac{1}{\sqrt{1+2\sigma^2}}\right)
\end{aligned}$$

hence,

$$\mathbb{V}(k) = n^2\Phi\left(\frac{\mu}{\sqrt{1+\sigma^2}}\right) \left[1 - \Phi\left(\frac{\mu}{\sqrt{1+\sigma^2}}\right) \right] - 2n(n-1)T\left(\frac{\mu}{\sqrt{1+\sigma^2}}, \frac{1}{\sqrt{1+2\sigma^2}}\right)$$

References

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