

**A Bias Correction Method in Meta-analysis of Randomized
Clinical Trials with no Adjustments for Zero-inflated Outcomes:
Supplementary Materials**

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Appendix A. Analogue between the ZIBC method and EM algorithm

We can obtain Equation (5) following the idea of the EM algorithm. Denote $\log L_{CV}(\boldsymbol{\beta}) \triangleq \sum_{i=1}^n \{-\exp(\mathbf{x}_i^t \boldsymbol{\beta}) + y_i \mathbf{x}_i^t \boldsymbol{\beta}\}$ as the log-likelihood function of $\boldsymbol{\beta}$ under conventional model minus a constant, the analogous expectation and maximization steps are shown as follows.

- (Analogous) expectation-step:

Taking the expectation of $\log L_{CV}(\boldsymbol{\beta})$, we have

$$\begin{aligned} \mathbb{E}[\log L_{CV}(\boldsymbol{\beta})] &= \sum_{i=1}^n \{-\exp(\mathbf{x}_i^t \boldsymbol{\beta}) + \mathbb{E}[y_i] \cdot \mathbf{x}_i^t \boldsymbol{\beta}\} \\ &= \sum_{i=1}^n \{-\exp(\mathbf{x}_i^t \boldsymbol{\beta}) + (1 - \pi_i) \exp(\mathbf{x}_i^t \boldsymbol{\beta}^0) \cdot \mathbf{x}_i^t \boldsymbol{\beta}\} \\ &\approx \sum_{i=1}^n \{-\exp(\mathbf{x}_i^t \boldsymbol{\beta}) + (1 - \bar{\pi}) \exp(\bar{\mathbf{x}}^t \boldsymbol{\beta}^0) \cdot \mathbf{x}_i^t \boldsymbol{\beta}\} \\ &\triangleq Q(\boldsymbol{\beta}). \end{aligned}$$

- (Analogous) maximization step:

To maximize $Q(\boldsymbol{\beta})$, we derive its first derivative with respect to $\boldsymbol{\beta}$, which is given by

$$\begin{aligned} \frac{\partial Q(\boldsymbol{\beta})}{\partial \boldsymbol{\beta}} &= \sum_{i=1}^n \{-\exp(\mathbf{x}_i^t \boldsymbol{\beta}) + (1 - \bar{\pi}) \exp(\bar{\mathbf{x}}^t \boldsymbol{\beta}^0)\} \mathbf{x}_i \\ &= [(1 - \bar{\pi}) \exp(\bar{\mathbf{x}}^t \{\boldsymbol{\beta}^0 - \boldsymbol{\beta}\}) - 1] \sum_{i=1}^n \exp(\mathbf{x}_i^t \boldsymbol{\beta}) \mathbf{x}_i. \end{aligned}$$

Note that $\boldsymbol{\beta}^*$ is the solution of $S_{CV}(\boldsymbol{\beta}) = 0$. By plugging $\boldsymbol{\beta}^*$ in $\frac{\partial Q(\boldsymbol{\beta})}{\partial \boldsymbol{\beta}}$, we obtain Equation (5).

Appendix B. Proof of Lemma 1

Proof. Consider three “average” subjects in the control group, intervention group, and the overall sample (denoted as $\{\text{average}, C\}$, $\{\text{average}, T\}$, and $\{\text{average}\}$), with $\mathbf{x}_{\text{average}, C, p-2} = \bar{\mathbf{x}}_{C, p-2}$, $\mathbf{x}_{\text{average}, T, p-2} = \bar{\mathbf{x}}_{T, p-2}$, and $\mathbf{x}_{\text{average}, p-2} = \bar{\mathbf{x}}_{p-2}$, respectively. Without loss of generality, assuming that observed covariates excluding the intervention assignment are grand mean centered before data analysis, we have $\mathbf{x}_{\text{average}, p-2} = \bar{\mathbf{x}}_{p-2} = \mathbf{0}$. Since $\bar{\mathbf{x}}_{C, p-2} = \bar{\mathbf{x}}_{T, p-2}$, we also have $\mathbf{x}_{\text{average}, C, p-2} = \mathbf{x}_{\text{average}, T, p-2} = \mathbf{x}_{\text{average}, p-2} = \mathbf{0}$. Therefore, we have

$$\begin{aligned}\log(\mu_{\text{average}, C}) &= \beta_{0, C}^0 \\ \log(\mu_{\text{average}, T}) &= \beta_{0, T}^0 + \beta_{1, T}^0 \\ \log(\mu_{\text{average}}) &= \beta_0^0 + \beta_1^0 \mathbb{1}_{\{A_{\text{average}}=T\}}\end{aligned}$$

under the true method. If an average subject in the overall sample belongs to the control group, then

$$\begin{aligned}\log(\mu_{\text{average}}) &= \log(\mu_{\text{average}, C}) \\ \Rightarrow \beta_{0, C}^0 &= \beta_0^0 \\ \Rightarrow \hat{\beta}_{0, C, \text{MLE}} &\approx \hat{\beta}_{0, \text{MLE}}.\end{aligned}\tag{A.1}$$

Similarly, if an average subject belongs to the intervention group, then

$$\begin{aligned}\log(\mu_{\text{average}}) &= \log(\mu_{\text{average}, T}) \\ \Rightarrow \beta_{0, T}^0 + \beta_{1, T}^0 &= \beta_0^0 + \beta_1^0 \\ \Rightarrow (\widehat{\beta_0 + \beta_1})_{T, \text{MLE}} &\approx \hat{\beta}_{0, \text{MLE}} + \hat{\beta}_{1, \text{MLE}}.\end{aligned}\tag{A.2}$$

Under similar arguments, for the conventional method, we have

$$\hat{\beta}_{0, C, \text{CV}} \approx \hat{\beta}_{0, \text{CV}}\tag{A.3}$$

and

$$(\widehat{\beta_0 + \beta_1})_{T,CV} \approx \hat{\beta}_{0,CV} + \hat{\beta}_{1,CV}. \quad (\text{A.4})$$

Plug Equations (A.1) and (A.3) into Equation (8), and plug Equations (A.2) and (A.4) into Equation (9), we have

$$\begin{aligned} \hat{\beta}_{0,\text{MLE}} &\approx \hat{\beta}_{0,\text{CV}} - \log(1 - \bar{\pi}_C) \\ \hat{\beta}_{0,\text{MLE}} + \hat{\beta}_{1,\text{MLE}} &\approx \hat{\beta}_{0,\text{CV}} + \hat{\beta}_{1,\text{CV}} - \log(1 - \bar{\pi}_T), \end{aligned} \quad (\text{A.5})$$

which directly gives Equation (10). □

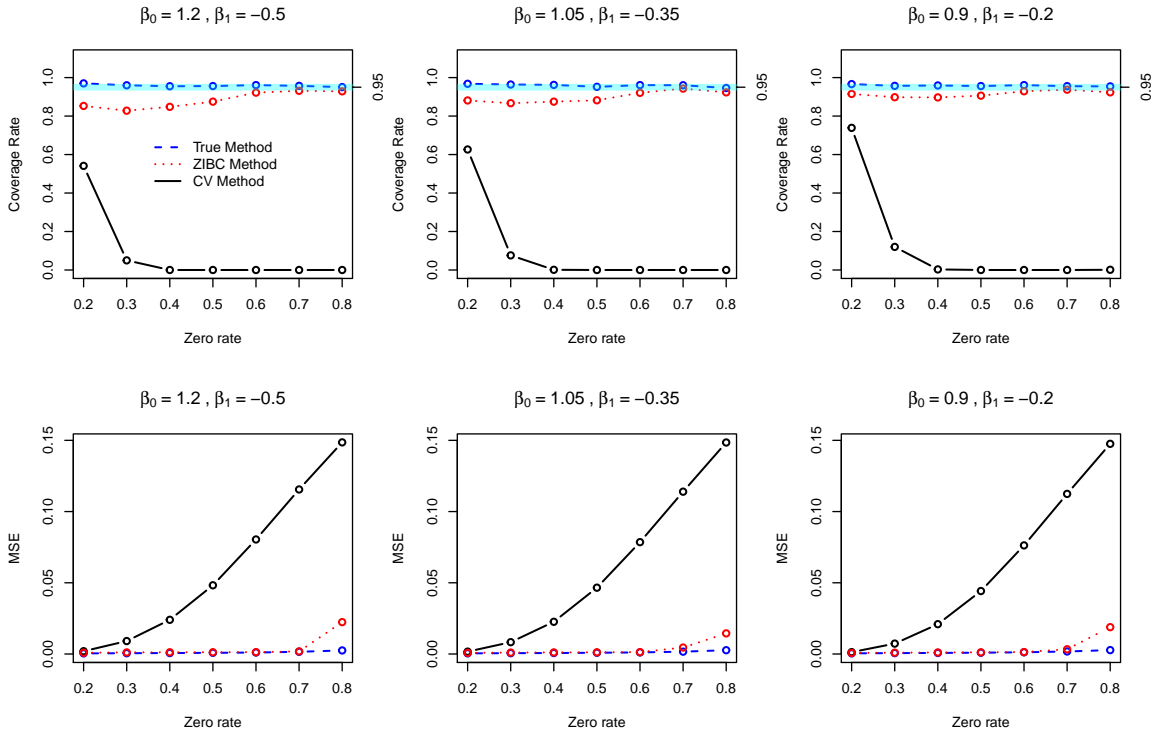


Figure S1: Coverage rates and MSE values of the true (blue dashed line), ZIBC (red dotted line) and conventional (black solid line) methods from 1000 replications ($K=16$)

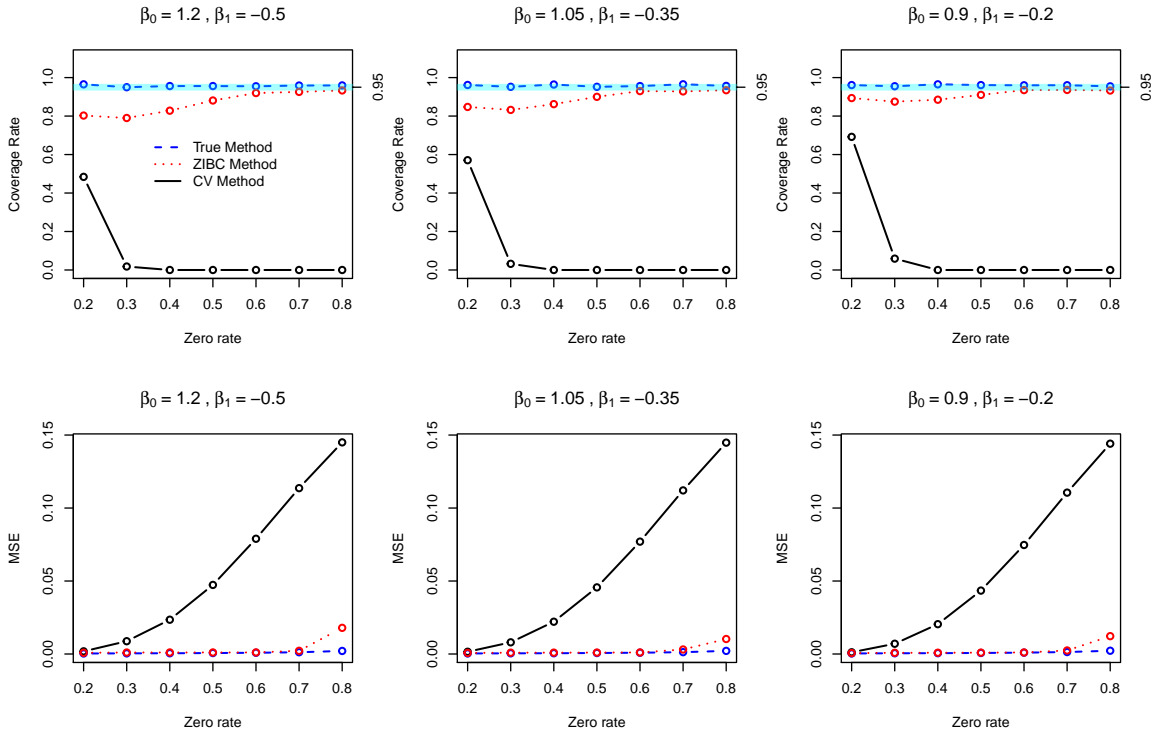


Figure S2: Coverage rates and MSE values of the true (blue dashed line), ZIBC (red dotted line) and conventional (black solid line) methods from 1000 replications ($K=20$)