THE LANCET Haematology

Supplementary appendix 1

This appendix formed part of the original submission and has been peer reviewed. We post it as supplied by the authors.

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Supplementary Appendix

A sex-informed approach to improve personalized decision-making process in myelodysplastic syndromes

This appendix has been provided by the authors to give readers additional information about their work

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Supplementary Table_1A_SF1. Demographic, haematological and clinical features of 2,025 patients from EuroMDS cohort, collected at the time of diagnosis and information on treatment.

Treatments. 426 out of 1,904 patients (22·4%) received red blood cell transfusions; 304 patients (15%) were treated with erythroid stimulating agents; 316 patients (15·6%) were treated with hypomethylating agents; 300 patients (14·8%) were treated with AML-like chemotherapy; 492 patients (24·2%) received allogeneic stem cell transplantation; 131 patients (6·5%) were treated with other treatments (lenalidomide, immunosuppressive drugs). No significant difference was noticed in the prevalence of different treatment strategies between men and women (not shown)

Supplementary Table_1B_SF1. Demographic, haematological and clinical features of 2,387 patients from IWG-PM cohort, collected at the time of diagnosis and information on treatment.

*Treatments. 488 out of 2,359 evaluable patients (*20·7%) *received red blood cell transfusions; 459 patients (19*·*9%) were treated with hypomethylating agents; 45 patients (1*·*6%) were treated with AML-like chemotherapy; 232 patients (10%) received allogeneic stem cell transplantation; 161 patients (7%) were treated with other treatments (lenalidomide, immunosuppressive drugs). No significant difference was noticed in the prevalence of different treatment strategies between men and women (not shown)*

Supplementary Table_1C_SF1. Demographic, haematological and clinical features of 7,687 patients from Spanish MDS Group registry (GESMD), collected at the time of diagnosis and information on treatment.

Treatments: 2,047 out of 5,336 patients (38·4%) received red blood cell transfusions; 854 patients (16%) were treated with erythroid stimulating agents; 1,238 patients (16·1%) were treated with hypomethylating agents; 369 patients (4·8%) were treated with AML-like chemotherapy; 300 patients (3·9%) received allogeneic stem cell transplantation. A higher prevalence of transfusion dependency was noticed in men vs. women (P=0·023).

Supplementary Table_1D_SF1. Demographic, haematological and clinical features of 1,185 patients from Düsseldorf MDS registry, collected at the time of diagnosis and information on treatment.

Treatments. 432 out of 1,110 patients (38·9%) received red blood cell transfusions; 109 patients (9·8%) were treated with hypomethylating agents; 152 patients (12·8%) received allogeneic stem cell transplantation. No significant *difference was noticed in the prevalence of different treatment strategies between men and women (not shown).*

** The diagnosis of myeloid neoplasm was formulated according to the criteria of the 2016 revision of WHO classification of myeloid neoplasms. Peripheral blood and bone marrow dysplasia was performed using established consensus criteria. (MDS with 5q-, MDS with isolated deletion of long arm of chromosome 5; MDS-SLD, MDS with single lineage dysplasia; MDS-MLD, MDS with multilineage dysplasia; MDS-RS-SLD, MDS with ring sideroblasts and single lineage dysplasia; MDS-RS-MLD, MDS with ring sideroblasts and multilineage dysplasia; MDS-EB1, MDS with excess of blasts, type 1; MDS-EB1, MDS with excess of blasts, type 2; IPSS-R, Revised International Prognostic Scoring System).*

Supplementary Figure_1_SF1. Probability of survival (since the time of diagnosis) of MDS patients belonging to EuroMDS cohort, IWG-PM cohort, registry of Spanish MDS Group (GESMD) and Düsseldorf MDS registry according to 2016 WHO categories (A, B, C and D, respectively) and to IPSS-R risk groups (E, F, G and H, respectively). (MDS with 5q-, MDS with isolated deletion of long arm of chromosome 5; MDS-SLD, MDS with single lineage dysplasia; MDS-MLD, MDS with multilineage dysplasia; MDS-RS-SLD, MDS with ring sideroblasts and single lineage dysplasia; MDS-RS-MLD, MDS with ring sideroblasts and multilineage dysplasia; MDS-EB1, MDS with excess of blasts, type 1; MDS-EB1, MDS with excess of blasts, type 2; MDS-U, unclassified MDS; IPSS-R, Revised International Prognostic Scoring System).

A. Probability of overall survival of EuroMDS cohort according to WHO categories (P<0·0001)

B. Probability of overall survival of IWG-PM cohort according to WHO categories (P<0·0001)

MDS-EB2

C. Probability of overall survival of Spanish MDS Group (GESMD) registry according to WHO categories (P<0·0001)

D. Probability of overall survival of Düsseldorf MDS registry according to WHO categories (P<0·0001)

- MDS-EB1
- MDS-EB2

E. Probability of overall survival of EuroMDS cohort according to IPSS-R categories (P<0·0001)

F. Probability of overall survival of IWG-PM cohort according to IPSS-R categories (P<0·0001)

G. Probability of overall survival of Spanish MDS Group (GESMD) registry according to IPSS-R categories (P<0·0001)

-
-
- H. Probability of overall survival of Düsseldorf MDS registry according to IPSS-R categories (P<0·0001)

IPSS-R Categories

- Very Low Low
- Intermediate
- High
- Very High

SUPPLEMENTARY_FILE_2 (SF2) – Cytogenetics and mutation screening

At diagnosis, cytogenetic analysis was performed using standard G-banding and karyotypes were classified using the International System for Cytogenetic Nomenclature Criteria.

In patients belonging to EuroMDS cohort, we analyzed in addition somatic mutations in 47 genes related to myeloid neoplasms, obtained by analyzing tumor DNA derived from bone marrow mononuclear cells (94% of cases) or peripheral blood granulocytes (6% of cases). Sample for DNA sequencing was collected within 30 days from diagnosis for 93% of cases (in all cases within 6 months since the date of diagnosis).

Sequencing strategy was performed using a targeted multiplexed amplicon-based approaches (Illumina, San Diego, CA, USA) starting from genomic DNA; the resulting libraries were sequenced on Illumina platforms (NextSeq500) in pairedend mode. Targeted regions are listed in *Supplementary Table_1_SF2* (see below).

Variants with a variant allele frequency (VAF) lower than 0.01 and/or variants with a coverage <200x were filtered out. Functionally annotated variants were then also excluded based on the information retrieved from public databases (dbSNP, gnomAD) and the expected germ line allele frequency. Single nucleotide polymorphisms (SNP) were annotated according to the NCBI dbSNP [\(http://www.ncbi.nlm.nih.gov/snp;](http://www.ncbi.nlm.nih.gov/snp) Build 137) and gnomAD

[\(http://gnomad.broadinstitute.org;](http://gnomad.broadinstitute.org/) gnomAD r2.0.1) databases.

The remaining variants were considered as possible somatic mutations and their pathogenic value was evaluated in order to differentiate known and putative pathogenic mutations from variants of unclear significance by using a multistep algorithm:

1) All variants (missense, in-frame insertions/deletions, frameshift, nonsense and splice site) were considered pathogenic if they were previously reported in the publicly accessible Catalogue Of Somatic Mutations In Cancer (COSMIC, version 69) [\(http://cancer.sanger.ac.uk/cancergenome/projects/cosmic\)](http://cancer.sanger.ac.uk/cancergenome/projects/cosmic) at least in two hematological sample. 2) Internal-tandem-duplication of FLT3 and in-frame insertions/deletions of CALR (exon 9) genes were included as pathogenic variants.

3) Loss of function mutations (Nonsense, frameshift and splice site) were considered pathogenic.

4) Missense variants and in-frame insertions/deletions not fulfilling these criteria were individually assessed based on the available data from COSMIC (the tissues they were found in, whether any other COSMIC variants were reported affecting the same amino-acid positions or were within 3 amino-acids) and their predicted functional consequences using the Mutation Taster algorithm [\(http://www.mutationtaster.org\)](http://www.mutationtaster.org/).

5) Nonsynonymous variants not fulfilling these criteria were then classified on the basis of their functional interpretation using in silico prediction effect by SIFT 1.03 [\(http://sift.jcvi.org\)](http://sift.jcvi.org/), PolyPhen 2.0

[\(http://genetics.bwh.harvard.edu/pph2\)](http://genetics.bwh.harvard.edu/pph2) and MutationTaster 1.0 algorithms [\(http://www.mutationtaster.org\)](http://www.mutationtaster.org/). Variants with less than 2/3 deduced damaging consequences on the amino acid level were discarded.

6) Additionally, TP53 variants were verified using the IARC repository [\(https://p53.iarc.fr/\)](https://p53.iarc.fr/).

Variants that did not satisfy any of the above criteria were not considered as pathogenic mutations in downstream analyses

Supplementary Table 1_SF2. Panel of sequenced genes in the EuroMDS and IWG-PM cohorts. The column description from left to right: Gene, name of gene; Pathway, main biological pathways in which the gene is involved or has a determinant function; NCBI gene ID, National Center for Biotechnology Information gene ID; Position, Chromosomal location.

*# Data available only in EuroMDS cohort; * Data available only in IWG-PM cohort*

SUPPLEMENTARY_FILE_4 (SF4) – Genomic landscape of MDS by sex

Supplementary Table_1_SF4. Prevalence of chromosomal abnormalities in patients from EuroMDS cohort stratified by sex. *For each class of comparisons, p-values were adjusted for multiple testing using the Benjamini-Hochberg procedure.*

**Only for men*

Supplementary Table_2_SF4. Prevalence of chromosomal abnormalities in patients from IWG-PM cohort stratified by sex. *For each class of comparisons, p-values were adjusted for multiple testing using the Benjamini-Hochberg procedure.*

**Only for men*

Supplementary Table_3_SF4. Prevalence of mutated genes in patients from EuroMDS cohort stratified by sex. For each class of comparisons, p-values were adjusted for multiple testing using the Benjamini-Hochberg procedure.

Supplementary Table_4_SF4. Prevalence of mutated genes in patients from IWG-PM cohort stratified by sex. For each class of comparisons, p-values were adjusted for multiple testing using the Benjamini-Hochberg procedure.

Supplementary Figure_1 _SF4. Cumulative counts and densities per age in sex-biased genes in 2,025 MDS patients from EuroMDS_cohort

Men

Supplementary Figure_2_SF4. Cumulative counts and densities per age in sex-biased genes in 2,387 MDS patients from IWG-PM_cohort

Supplementary Figure_3_SF4. Distribution of gene mutations in patients from EuroMDS cohort stratified sex across age categories (there are reported only the genes that showed at least 15 mutated patients).

Men

Supplementary Figure_4_SF4. Distribution of gene mutations in patients from IWG-PM cohort stratified sex across age categories (there are reported only the genes that showed at least 15 mutated patients).

Women

Men

SUPPLEMENTARY_FILE_5 (SF5) - Mutation acquisition order

In order to determine the relative order of mutation acquisition, comparisons were made for each pair of mutations in each patient (for additional details on the methodology, see: *<https://ascopubs.org/doi/suppl/10.1200/JCO.20.01659>*). Even without a time course experiment, it is possible to infer the relative order in which two events occurred. Such ordered pairings were used to determine the relative probabilities of a gene mutation occurring first or second for a given pairing with the use of Bradley–Terry (BT) modeling, which provided an estimate of the overall timing of mutation acquisition. For each patient the proportions of cells carrying each mutation, the variant allele fractions corrected for any copy number change at the site of the variant were considered. BT was applied to the set of genes in which genes mutations co-occurring with other gene mutations in at least 15 patients were considered, as it was done in previous works. The data used for the BT model inference were retrieved at site level for 1761 MDS patients from EuroMDS cohort. The R package BradleyTerry2 (version 1.0-8) was used to generate estimates of relative mutation timing. The results for the determination of mutation order in EuroMDS cohort are available in *Supplementary Figure_1_SF6.*

Supplementary Figure_1_SF5. Determination of MDS mutation order in EuroMDS cohort. The number of pairs n in which the event occurred is shown for each gene on the right of correspondent gene in the plot. Only genes mutations co-occurring with other gene mutations in at least 15 patients were considered. The horizontal axis shows the log odds of a gene occurring second in a gene pair. Any pair of genes can be assessed by calculating the exponential of the difference in log odds for gene A and gene B. Blue asterisks mark statistically significant man-biased genomic abnormalities, and yellow asterisks mark statistically significant woman-biased genomic abnormalities.

Disease evolution

SUPPLEMENTARY_FILE_6 (SF6) - Identification of co-mutational patterns and mutually exclusive mutations in MDS patients stratified by sex

Pairwise associations among genes and cytogenetic abnormalities in MDS patients stratified by sex

In order to assess pairwise association among genes and/or cytogenetics abnormalities, we calculated the cooccurrence of genomics abnormalities across patients. In more details, for each couple of genomic abnormalities, the number of patients showing mutation co-occurrence were quantified. 2x2 contingency tables were generated by each pair present and the significance was evaluated with Fisher's exact test. Furthermore, for each possible pairing of genes and/or cytogenetic abnormalities the odds ratio was calculated. Odds ratios less than 1 indicates that the pairs of mutation were mutually exclusion, while odds ratios greater than 1 implies mutation co-occurrence.

Supplementary Figure_1_SF6. Pairwise associations among genes that happen to be mutated in at least 40 patients and cytogenetic abnormalities in 2025 MDS patients from EuroMDS cohort, stratified by sex. In the upper triangle, for each couple of genomic abnormalities, the number of patients showing mutation co-occurrences are illustrated using a blue color scale. In the lower triangle the gene-gene co- occurrence and mutual exclusivity is assessed using odds ratio and significance is evaluated using Fisher test. Multiple hypothesis testing was performed using the Benjamini-Hochberg adjustment in order to control the false discovery rate, meaning the expected proportion of false discoveries amongst the rejected hypotheses. Such a correction is necessary when dealing with such a high number of comparisons and help to identify the most significant associations.

A. Men from EuroMDS cohort

 $\begin{array}{ccc}\n & 9 & 4 & 0.1 \\
& 9 & 6 & 0.05 \\
& 9 & 6 & 0.01\n\end{array}$

Supplementary Figure_2_SF6. Pairwise associations among genes that happen to be mutated in at least 40 patients and cytogenetic abnormalities in 2,387 MDS patients from IWG-PM cohort, stratified by sex. In the upper triangle, for each couple of genomic abnormalities, the number of patients showing mutation co-occurrences are illustrated using a blue color scale. In the lower triangle the gene-gene co- occurrence and mutual exclusivity is assessed using odds ratio and significance is evaluated using Fisher test. Multiple hypothesis testing was performed using the Benjamini-Hochberg adjustment in order to control the false discovery rate, meaning the expected proportion of false discoveries amongst the rejected hypotheses. Such a correction is necessary when dealing with such a high number of comparisons and help to identify the most significant associations.

A. Men from IWG-PM cohort

Bayesian networks

We used Bayesian Networks (BN) to define in a more comprehensive way the relationships between genomic abnormalities in MDS stratified by sex. We included gene mutations and cytogenetic abnormalities as random variables in the model and we investigated conditional dependency among them.

Given a set of variables (in our study the set of cytogenetic and genetic mutations), a BN is a graphical way to highlight conditional dependencies among variables, i.e. how the values taken by a given variable influences the probability of the others. They main hypothesis underlying BN is that joint probability distribution (JPD) over the set of variables could be represented as a Directed Acyclic Graph (DAG), i.e. a directed graph with no loops.

DAG nodes represent random variables; each node i is associated with the probability distribution P_i, the probability of observing a mutation at the i-th position; while a link represents a dependence among two variables (i.e. how the presence of a given mutation influences the presence of the other). For instance, an arrow from node A to node B is a probabilistic direct dependence between A and B. Directed dependence means that the value taken by B is influenced by the value taken by A while the vice versa in not true, i.e. we have a causal connection between the variables. More formally, given the set of variables $x = (x_1,...,x_n)$, BNs are a way to represent a specific factorization of their JPF. Given the DAG structure (S), the joint probability distribution is given by:
 $\frac{n}{2}$

$$
P(x_1, ..., x_n) = \prod_{i=1}^n P_i(x_i | Parents_i^S)
$$

where the factorized probabilities are conditioned on the parents of the node i in the directed acyclic graph (S). This is equivalent to say that each variable is independent of its non-child nodes in the graph given the state of its parents. Given the training data we estimated the network structure (S) and the parameters of the JPD in the BN (i.e. P_i for $i =$ 1,...,n). We inferred the network structure from data using the GOBNILP software:⁴ given a set of random variables, GOBNILP assigns a score (based on data) to each Directed Acyclic Graph and choose the structure which maximizes the score (according to previous literature^{1,3} we set the maximum number of parents to 3). For each variable in which conditional dependency was found (i.e. a link in the inferred structure is present), the definition of mutually exclusivity was used to define a significant negative dependency, while the definition of co-occurrence was used to define a positive dependency.

For additional details on the methodology, please see[: https://ascopubs.org/doi/suppl/10.1200/JCO.20.01659](https://ascopubs.org/doi/suppl/10.1200/JCO.20.01659)

Supplementary Figure_3_SF6. Genomic Landscape of MDS through Bayesian Networks in 2,025 patients from EuroMDS cohort, stratified by sex. Given a set of variables (in our study the set of cytogenetic and genetic mutations), a Bayesian Networks is a graphical way to highlight conditional dependencies among variables, i.e. how the values taken by a given variable influences the probability of the others. The size of each node accounts for the number of correspondent genomic or cytogenetic alterations. The color of each link reflects odds ratio of co- occurrence or mutually exclusivity as calculated previously in Figure_1_SF6. The thickness of edges grows with increasing significance of mutual exclusivity / co-occurrence between alterations.

A. Men from EuroMDS cohort

B. Women from EuroMDS cohort

Supplementary Figure_4_SF6. Genomic Landscape of MDS through Bayesian Networks in 2,387 patients from IWG-PM cohort, stratified by sex. Given a set of variables (in our study the set of cytogenetic and genetic mutations), a Bayesian Networks is a graphical way to highlight conditional dependencies among variables, i.e. how the values taken by a given variable influences the probability of the others. The size of each node accounts for the number of correspondent genomic or cytogenetic alterations. The color of each link reflects odds ratio of co- occurrence or mutually exclusivity as calculated previously in Figure_2_SF6. The thickness of edges grows with increasing significance of mutual exclusivity / co-occurrence between alterations.

A. Men from IWG-PM cohort

B. Women from IWG-PM cohort

Supplementary Table_1_SF6. Description of statistically significant (P value <0.05) co-occurring and exclusive mutations/chromosomal abnormalities (representative genes and chromosomal abnormalities are showed) in men vs women from EuroMDS and IWG-PM cohorts (only significant relationships in both populations are reported).

Supplementary Figure_5_SF6 Genomic Landscape of MDS through Bayesian Networks. Comparison between EuroMDS and IWG-PM cohorts inferred relations. All the represented relations are significantly determined in EuroMDS cohort and are confirmed in IWG-PM cohort.

Dirichelet Processes Multinomial Mixture Model

In order to identify MDS molecular subtypes we carried out Dirichelet Process Clustering (DP). The DP infinite multinomial mixture model allows to capture broad dependencies among all gene mutations assuming them to be extracted from a mixture of multinomials. The rationale underlying the model is that we expect mutations to be clustered together according to the specific molecular mechanism at work in a given tumor. Using an infinite mixture with DP prior, instead of finite mixture, allows not to specifying a priori the number of mutations categories, which, instead, is inferred from the data. Importantly, the usage of advanced clustering methods such as DP for patient clustering allows to avoid overfitting issues. To carry out the analysis we used the R package HDP available online [https://github.com/nicolaroberts/hdp\)](https://github.com/nicolaroberts/hdp).

The input data consists of a patient by genes binary matrix. The genotype of a patient is a row of the matrix: G= $(G_1,...,G_n)$; where n is the number of features per patient (in our case: 12 cytogenetic and 47 genomic variables). G_{ii} is a binary variable which denotes the presence or absence of i-th alteration. Missing data where imputed with R package copynumber. The analysis was performed using different kinds of imputation with comparable results. Patients with no alterations were excluded from the DP clustering and classified as a class on their own.

More formally, DP mixture model assumes data to be generate according to the following process:

- $\theta \sim \text{DP(Dirichlet}(\alpha), \alpha_0)$
- \bullet X | θ, N ~ Multinomial(θ, N_i)

where: θ are the parameters of the multinomials, α_0 is the concentration parameter of the DP process and α are the parameters of the base distribution with parameter $\alpha = (1/n,...,1/n)$.

We carried out Monte Carlo Markov Chain (MCMC) sampling of DP posterior for 4 different initial conditions (n. of different chains). For each chain we discarded the first 3000 iterations and we sampled 4000 realizations at intervals of 20 iterations. Components are built by grouping raw clusters of DP posterior samples according to the following conditions: 1) clusters are merged if their cosine similarity is above a give threshold (0.95 in our case) and 2) clusters are assigned to component 0 if they have no significant data categories or sample exposure. Components 1-5 account for the 97% of the data while component 0 accounts for data that cannot be explained by the model.

The model found a mixture of multinomials with 5 components, plus an additional one of unexplained data.

For additional details on the methodology, please see[: https://ascopubs.org/doi/suppl/10.1200/JCO.20.01659](https://ascopubs.org/doi/suppl/10.1200/JCO.20.01659)

Supplementary Figure_6_SF6 Distribution of MDS genomic-based groups in patients from EuroMDS (A) and IWG-PM cohorts (B), stratified by sex. (MDS genomic-based groups were defined according to Bersanelli M, et al. Classification and Personalized Prognostic Assessment on the Basis of Clinical and Genomic Features in Myelodysplastic Syndromes. J Clin Oncol 39: 1223-1233, 2021)

A)

Percent count (%)

B)

SUPPLEMENTARY_FILE_7 (SF7) - Sex effect on MDS clinical outcome

Supplementary Figure_1_SF7. Probability of overall survival of MDS patients belonging to retrospective EuroMDS cohort (plot A), retrospective IWG-PM cohort (B), prospective registry of Spanish MDS Group (GESMD) (C) and prospective Düsseldorf MDS registry (D).

Men

A. Probability of overall survival of EuroMDS cohort stratified by sex (men vs women HR 1·40, CI 1·26-1·52, P<0·0001)

B. Probability of overall survival of IWG-PM cohort stratified by sex (men vs women HR 1·33, CI 1·13-1·57, P<0·0001)

C. Probability of overall survival of Spanish MDS Group (GESMD) cohort stratified by sex (men vs women HR 1·30, CI 1·24-1·35, P<0·0001)

D. Probability of overall survival of Düsseldorf MDS registry cohort stratified by sex (men vs women HR 1·23, CI 1·07-1·36, P=0·0061)

*Supplementary Figure_2_SF7. Competing risk analysis of leukemic death (LD) vs non leukemic death (NLD) in MDS patients from EuroMDS cohort with early disease stage (defined by IPSS-R score ≤3·5), stratified by sex. When estimating the occurrence of non-leukemic death, only deaths for all causes except leukemic evolution were considered as events. * The 5-year risk of non-leukemic death was 32·1% in men vs 18.4% in women (P<0·0001), while no difference was found regarding the risk of leukemic death*

- Probability of non-leukemic death in men
- Probability of non-leukemic death in women
- Probability of leukemic death in men **COL**
- **PROBABILITY OF LEADER IS A LIGAN COMPOOL**

Supplementary Figure_3_SF7. Probability of overall survival of patients stratified by sex, according to different haemoglobin values. This analysis was conducted on retrospective EuroMDS cohort (A), retrospective IWG-PM cohort (B), prospective registry of Spanish MDS Group (GESMD) (C) and prospective Düsseldorf MDS registry (D).

- A. Probability of survival of EuroMDS cohort according to haemoglobin values stratified by sex **Women**
	-

B. Probability of survival of IWG-PM cohort according to haemoglobin values stratified by sex

Women

C. Probability of survival of Spanish MDS Group (GESMD) cohort according to haemoglobin values stratified by sex Women

D. Probability of overall survival of Düsseldorf MDS registry cohort according to haemoglobin values stratified by sex Women

*Supplementary Table_1_SF7***.** *Prognostic impact of different haemoglobin (Hb) value in men and women from EuroMDS and IWG-PM cohorts; HR for probability of overall survival were calculated by using Hb normal value as reference (12-14 g/dl for women and 13-15 g/dl for men). Analyses were adjusted for age. Our analysis showed that anaemia start to have* significantly negative prognostic impact below 11 g/dl in men and below 10 g/dl in women. This effect was maintained in multivariable analysis including age, neutrophil and platelet count, % of bone marrow blast and cytogenetics stratified according to IPSS-R criteria (EuroMDS cohort: men HR 2·17[1·23-4·27], P<0·001; women HR 2·51 [1·32-4·42], P<0·0001; IWG-PM cohort: *men HR 2·04[1·47-3·66], P<0·0001; women HR 2·29 [1·39-3·84], P<0·0001).*

SUPPLEMENTARY_FILE_8 (SF8) - Personalized prognostic assessment in myelodysplastic syndromes based on demographics, clinical and genomic features

Multistate Cox's proportional-hazards model (coxph)

The association between the overall survival time and possible predictor variables was investigated fitting multistate Cox proportional-hazards models using the survival and *mstate* R packages. Specifically, we considered 3 possible states (Diagnosis, Acute Myeloid Leukemia, AML, and Death) and 3 possible transitions (Diagnosis to AML, Diagnosis to Death, AML to Death). The model was fitted without any proportionality assumption on the baseline hazards, meaning that separate baseline hazards were allowed for the different transitions, and considering transition specific effects. The analysis of the transition probabilities showed that the transition from AML to Death was particularly fast and highly probable. For this reason, in the following analyses we considered a simplified version of the multistate model in which only two transitions are considered: Diagnosis to AML and Diagnosis to Death, were the second transition also includes subjects that died after passing through the AML state.

Random effects Cox proportional-hazard multistate model (CoxHD)

Further innovative prognostic multistate models were developed fitting the random effects Cox proportional-hazards model developed by Gerstung et al (Nat Genet 49: 332–3340, 2017) and implemented in the R package CoxHD available at (*<http://github.com/mg14/CoxHD>*). Here, we considered 3 possible states (Diagnosis, AML, Death) and 2 possible transitions: Diagnosis to AML and Diagnosis to Death. In the last case, we also included subjects who died after being affected by AML. As for the previous Cox's model, we did not introduce any proportionality assumption on the baseline hazards, meaning that separate baseline hazards were allowed for the different transitions, and we considered transition specific effects.

Here, the covariates included in the design matrix Z are categorized in groups and the model parameters *u^j* for each group of variables *g* are assumed to be i.i.d. normally distributed. Letting the hazard be

$$
\lambda = \lambda_0(t)e^{(u^T Z)}
$$

where $\lambda_0(t)$ is the baseline hazard of the coxph model; this means that

$$
\forall j \in g: \qquad u_i \sim N(\mu_g; \sigma_g^2), \qquad i.i.d.
$$

The shared means are motivated by the assumption that on average the effect of variables belonging to the same category is comparable.

This model can be interpreted as a hierarchical model in which we assume that variables belonging to the same group have the same prior (gaussian) distribution and is equivalent to a ridge penalized model in which the parameters are penalized group by group. The log-likelihood of the model is

$$
l(u, \mu, \sigma^2; Z) = l_0(u; Z) - \sum_g \frac{\sum_{j \in g} (u_j - \mu_g)^2}{\sigma_g^2} = l_0(u; Z) + l_2(u, \mu, \sigma^2; Z)
$$

where $u = \{u_j : j = 1, \cdots, p\}$, $u_g = \{u_j : j \in g\}$.

The term $l_0(u; Z)$ is the likelihood of the coxph, while the second term is a sum of ridge penalties resulting from the assumption of normal prior distributions for each group of variables, which penalizes large values of u_i – μ_q (encourages the model parameters to be close to the mean of the corresponding Gaussian distributions) with strength $1/\sigma_q$.

Goodness of fit and model comparison based on the concordance statistic

The goodness of fit of both the coxph and the CoxHD models was evaluated computing the concordance. Concordance is defined as the probability for any two randomly chosen observations that the one with the shorter survival time of the two also has the larger predicted risk score (i.e. is concordant). The concordance C of each model was estimated using the survival R package as

$$
\mathcal{C}=\left(A+\tfrac{Tp}{2}\right)/(A+D+T_P),
$$

where A , D and T_P indicate the number of pairs of observations that are concordant, discordant, and tied on the predictor p but not on the observed data.

Supplementary Figure_1_SF8. Fraction of explained variation that was attributable to different prognostic factors for non-leukemic death and leukemic death by using Sex-informed Genomic Scoring System on merged EuroMDS and IWG-PM cohorts

Supplementary Table_1_SF8. Comparison of concordance (with standard deviation, sd) between IPSS-R categories, IPSS-R and age, and IPSS-R, age and sex on EuroMDS cohort

In order to test the improvement of the model due to the introduction of the Sex variable, we used the R function anova.coxph to compute an analysis of deviance for these Cox models considering IPSS-R and age vs. IPSS-R, age and sex. We obtained a p-value of 0.0033, confirming the importance of Sex in the model.