

Supplementary Online Content

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This supplementary material has been provided by the authors to give readers additional information about their work.

eMethods. Study Population, Propensity Score and Kernel Weighting, Sensitivity and Subgroup Analyses, and Data Linkage and Quality Assessment

1. The 100 Million Brazilian Cohort

The 100 Million Brazilian Cohort baseline is an open cohort using data linkage compiled by the Centre
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of Data and Knowledge Integration for Health (CIDACS/FIOCRUZ). The cohort is based on the notion of a “cohort baseline” with information on over 114 million individuals who were registered on the Brazilian National Registry for Social Programmes – Cadastro Único (CadÚnico) between 2001 and 2015. The eligibility criteria for CadÚnico registration includes a monthly family income per capita of half the minimum salary in Brazil (e.g. BRL 778.00 in 2015), or a total family income (monthly) of up to three minimum salaries¹. It contains information on housing conditions, income, and demographic characteristics, and socioeconomic information on all members of a registered family. For those selected to receive Bolsa Família (BFP), information on BFP receipt was obtained through linking the cohort baseline with the BFP payroll database (2004-2015), which contains information on sociodemographic characteristics (i.e., date of birth and family income), start of receipt, and length of time receiving BFP.

1.1 Bolsa Família Programme (BFP)

The Brazilian cash transfer programme, BFP, is the flagship and largest socioeconomic programme implemented by the Brazilian government in 2004. It forms part of the Brazilian initiative to eradicate extreme poverty and has three objectives: an income guarantee for the immediate relief of extreme poverty; access to public services (improving the education, health, and citizenship of families); and productive inclusion to increase capacity, job opportunities, and income generation among the poorest families¹. Implementation of BFP has resulted in 22.2 million Brazilians overcoming extreme poverty.

Every BFP participant is registered on CadÚnico. BFP is available throughout all the regions of Brazil and is estimated as having over 90% coverage among those eligible in the country. BFP beneficiaries receive basic, variable benefits, with values that vary over time, according to extreme poverty (families with a monthly per capita income of up to BRL 89.00), and poverty cut-off points (families with a monthly per capita income of up to BRL 178.00). The basic benefit is currently BRL 89.00 for extremely poor families. Variable benefits of BRL 41.00 are allocated to families with children, pregnant/puerperal women, or adolescents¹. The BFP uses educational and health-related conditionalities to promote behavioural change. The conditionalities are the requirement that all children must attend a minimum of 85% of school days, and women and children must attend health care appointments. These conditionalities are based on the idea that making benefits conditional upon ‘positive’ behaviours can further increase the chance of families breaking out of the cycle of poverty through increased education, or improved health. For instance, increasing school attendance, and consequently improving educational levels, can also lead to improving the quality of social networks, i.e. making friends at school rather than on the streets, and reducing opportunities for certain types of crime and risky behaviour.

In our study, we have considered those exposed to BFP as women with records of live births who started to receive BFP before or during pregnancy, and continued to receive the benefit until childbirth, or before death. Women who had not received the benefit at any time, or until childbirth, were considered not exposed. The women who stopped receiving the benefit at some point before childbirth were excluded from the analysis. The characteristics of the population excluded from the analysis due to the definition of BFP exposure (women who stopped receiving the benefit at some point before childbirth) are described in Supplementary eTable 1.

eTable 1. Description of Women Excluded From the Main Analysis Following Definition of *Bolsa Família* Program (BFP) Exposure

| Socioeconomic and demographic variables | Population following definition of BFP exposure ¹ | | | Women excluded following definition of BFP exposure ² |
|---|--|--------------------|-----------------------|--|
| | Non-BFP n= 1,264,037 | BFP n=5,413,236 | Total n=6,677, 273 | Total N=1,234,794 |
| Sociodemographic | | | | |
| Race/Colour | | | | |
| Asian | 0.47 | 0.35 | 0.37 | 0.30 |
| Black | 7.36 | 8.78 | 8.52 | 6.90 |
| Indigenous | 0.29 | 1.03 | 0.90 | 0.34 |
| Mixed-race/brown | 55.69 | 63.69 | 62.23 | 52.31 |
| White | 36.19 | 26.15 | 27.98 | 40.13 |
| Missing data* | 9.35 | 5.76 | 5.23 | 5.53 |
| Education | | | | |

| | | | | |
|--|-------|-------|-------|-------|
| High school/College (≥ 8 years) | 69.64 | 58.88 | 60.92 | 78.24 |
| Elementary/Middle school (4-7 years) | 24.21 | 32.86 | 31.22 | 19.23 |
| Elementary school or illiterate (< 3 years) | 6.15 | 8.26 | 7.86 | 2.53 |
| <i>Missing data*</i> | 1.98 | 2.26 | 2.21 | 1.57 |
| Age Groups | | | | |
| 10-19 years old | 14.09 | 23.24 | 70.57 | 22.19 |
| 20-34 years old | 75.89 | 69.32 | 21.51 | 70.53 |
| ≥ 35 years old | 10.02 | 7.44 | 7.92 | 7.28 |
| <i>Missing data*</i> | 0.0 | 0.0 | | 0.0 |
| Parity | | | | |
| 1 child in the cohort | 39.74 | 44.35 | 43.48 | 34.24 |
| 2-3 children in the cohort | 57.80 | 49.66 | 51.20 | 62.85 |
| > 3 children in the cohort | 2.46 | 5.99 | 5.32 | 2.91 |
| <i>Missing data*</i> | 0.0 | 0.0 | 0.0 | 0.0 |
| Household | | | | |
| Density | | | | |
| ≤ 2 per room | 79.30 | 53.22 | 57.88 | 58.29 |
| > 2 per room | 20.70 | 46.78 | 42.12 | 41.71 |
| <i>Missing data*</i> | 11.63 | 5.10 | 6.33 | 5.37 |
| Water Supply | | | | |
| Public network | 77.51 | 65.42 | 67.62 | 76.60 |
| Well/natural source/other | 22.49 | 34.58 | 32.38 | 23.40 |
| <i>Missing data*</i> | 7.53 | 2.79 | 3.69 | 3.28 |
| Waste disposal system | | | | |
| Public network | 77.51 | 38.72 | 40.97 | 51.34 |
| Septic tank/ditch/other | 22.49 | 61.28 | 59.02 | 48.65 |
| <i>Missing data*</i> | 8.88 | 3.51 | 4.52 | 3.46 |
| Garbage disposal | | | | |
| Public collection system | 51.21 | 68.78 | 71.19 | 81.97 |
| Burned/buried/other | 49.79 | 31.22 | 28.81 | 18.03 |
| <i>Missing data*</i> | 7.53 | 2.79 | 3.69 | 3.28 |
| Geographical | | | | |
| Region | | | | |
| South | 14.19 | 8.55 | 9.61 | 16.47 |
| North | 9.78 | 13.21 | 12.56 | 8.73 |
| Northeast | 30.69 | 42.87 | 40.57 | 22.09 |
| Southeast | 35.10 | 29.29 | 30.39 | 43.46 |
| Central-West | 10.24 | 6.08 | 6.86 | 9.23 |
| <i>Missing data*</i> | 0 | 0.0 | 0.0 | 0.0 |
| Household location | | | | |
| Urban | 81.56 | 71.74 | 73.55 | 82.69 |
| Rural | 18.44 | 28.26 | 26.45 | 17.31 |
| <i>Missing data*</i> | 5.56 | 2.06 | 2.72 | 2.75 |
| Year | | | | |
| 2004 | 7.85 | 10.26 | 9.80 | 10.26 |
| 2005 | 7.12 | 7.92 | 7.76 | 8.97 |
| 2006 | 26.84 | 45.62 | 42.06 | 53.97 |
| 2007 | 11.18 | 13.71 | 13.23 | 12.66 |
| 2008 | 5.37 | 5.53 | 5.50 | 4.10 |
| 2009 | 4.33 | 4.85 | 4.75 | 3.15 |

| | | | | | |
|----------------------|-------|------|------|--|------|
| 2010 | 6.12 | 4.36 | 4.70 | | 2.96 |
| 2011 | 5.18 | 2.51 | 3.02 | | 1.63 |
| 2012 | 10.41 | 2.79 | 4.23 | | 1.21 |
| 2013 | 7.44 | 1.61 | 2.71 | | 0.62 |
| 2014 | 6.09 | 0.72 | 1.73 | | 0.31 |
| 2015 | 2.07 | 0.12 | 0.49 | | 0.14 |
| <i>Missing data*</i> | 0 | 0.0 | 0.0 | | 0 |

¹Main analysis population: BFP - women who started to receive the BFP before or during pregnancy, and continued to receive the benefit until childbirth, or before death. Non-BFP: Women who had not received the benefit at any time, or until childbirth.

² Excluded women - women who stopped receiving the benefit at some point before childbirth

2. Propensity score and kernel weighting

2.1 Propensity score estimation

BFP is allocated according to the criteria of poverty and extreme poverty, which is classified by per capita family income.¹ However, a set of socioeconomic characteristics are related to program receipt. Since as the receipt of the program was not randomly assigned to families, in line with our research protocol² and previous quasi-experimental studies using the CadÚnico dataset and the 100 Million Brazilians Cohort, the association between BFP and maternal mortality was estimated based on the propensity score-based method (PS).

We estimated the probability of receiving the BFP benefit from the baseline covariates using multiple logistic regressions. The following covariates were considered to estimate the propensity score (PS): self-declared race/skin colour (white, mixed-race-brown, black, indigenous, or of Asian descent), level of education (≤ 3 years, 4-7 years, or ≥ 8 years), age (10-19 years old; 20-34 years old; and ≥ 35 years old), parity (number of childbirths in the cohort: 1; 2-3 and ≥ 3 children), urban/rural residency), region (North, Northeast, Southeast, South and Central-West), household density (≤ 2 inhabitants per room, or > 2 inhabitants per room), water supply (public network, or a well, natural source or other), waste disposal system (public network, septic tank, ditch, or other), garbage disposal (public collection system, burned, buried, or other) and year of registration on Cadastro Único. eTable 2 shows the estimates from the logit model for the probability of receiving BFP (propensity scores) based on women baseline characteristics and socioeconomic conditions.

eTable 2. Logistic Models to Estimate the Propensity Scores for *Bolsa Família* Participation

| Variables | Odds Ratio (95% CI) | p-value | Standard Error |
|----------------------------|------------------------|---------|----------------|
| Parity | | | |
| 1 child in the cohort | Ref. | | - |
| 2-3 children in the cohort | 0.57 (0.56 - 0.57) | <0.001 | 0.0014 |
| >3 children in the cohort | 2.20 (2.17 - 2.23) | <0.001 | 0.0016 |
| Education | | | |

| | | | |
|--|-----------------------|--------|--------|
| High school/College (≥8 years) | Ref | | |
| Elementary/Middle school (4-7 years) | 1.14 (1.13-1.14) | <0.001 | 0.0073 |
| Elementary school or illiterate (<3 years) | 0.91 (0.90-0.92) | <0.001 | 0.0038 |
| Age Groups | | | |
| 20-34 | Ref. | | - |
| ≤ 19 | 2.09 (2.08 - 2.11) | <0.001 | 0.0073 |
| ≥ 35 | 0.92* (0.92- 0.93) | <0.001 | 0.0038 |
| Maternal race/skin color | | | |
| White | Ref. | | - |
| Asian descendants | 1.15 (1.11 - 1.19) | <0.001 | 0.0218 |
| Black | 1.31 (1.30 - 1.33) | <0.001 | 0.0071 |
| Indigenous | 3.15 (3.02 - 3.29) | <0.001 | 0.0692 |
| Mixed/brown | 1.20 (1.19 - 1.21) | <0.001 | 0.0342 |
| Region | | | |
| South | Ref. | | - |
| North | 1.49 (1.48-1.51) | <0.001 | 0.0082 |
| Northeast | 1.94 (1.92 - 1.96) | <0.001 | 0.0082 |
| Southeast | 1.47 (1.46 - 1.49) | <0.001 | 0.0607 |
| Center-West | 0.90 (0.89 - 0.91) | <0.001 | 0.0048 |
| | | | |
| Local domiclio | | | |
| Urban | Ref. | | - |
| Rural | 0.99 (0.98 - 1.00) | <0.146 | 0.0043 |
| Household density | | | |

| | | | |
|--|-----------------------|--------|--------|
| ≤ 2 people per room | Ref. | | - |
| > 2 people per room | 2.44 (2.43-2.46) | <0.001 | 0.0684 |
| Water Supply | | | |
| Public network | Ref. | | - |
| Well/natural source/other | 1.15 (1.14 - 1.16) | <0.001 | 0.0141 |
| Waste disposal system | | | |
| Public network | Ref. | <0.001 | - |
| Septic tank/ditch/other | 1.88 (1.86 - 1.91) | <0.001 | 0.0406 |
| Garbage disposal | | | |
| Public collection system | | | |
| Burned/buried/other | 1.31 (1.29-1.32) | <0.001 | 0.0057 |
| Registration year in Cadastro Unico | | | |
| 2004 | Ref | | |
| 2005 | 0.99 (0.98-1.00) | 0.292 | 0.0058 |
| 2006 | 1.16 (1.15-1.17) | <0.001 | 0.0051 |
| 2007 | 0.78 (0.78-0.79) | <0.001 | 0.0039 |
| 2008 | 0.70 (0.69-0.71) | <0.001 | 0.0042 |
| 2009 | 0.81 (0.80-0.82) | <0.001 | 0.0051 |
| 2010 | 0.53 (0.53-0.54) | <0.001 | 0.0034 |
| 2011 | 0.37 (0.36-0.37) | <0.001 | 0.0027 |
| 2012 | 0.21 (0.21-0.22) | <0.001 | 0.0012 |
| 2013 | 0.18 (0.18-0.18) | <0.001 | 0.0013 |
| 2014 | 0.09 (0.09-0.09) | <0.001 | 0.0741 |

| | | | |
|-------------|---------------------|--------|--------|
| 2015 | 0.45 (0.43-0.47) | <0.001 | 0.0072 |
|-------------|---------------------|--------|--------|

2.2 Kernel matching/weighting

Kernel matching establishes a non-parametric relationship between the PS and outcome - maternal mortality, selecting observations of non-beneficiaries (non-BFP) who were similar to the set of treated groups (BFP beneficiaries), according to observable characteristics (PS).⁶ There are weighting schemes for all of the untreated groups (non-BFP), and the weights depend on the distance between each individual from the control group, and the participant observation for which the counterfactual is estimated⁷. The basic idea behind kernel estimation is giving different weight to observations with a different distance.

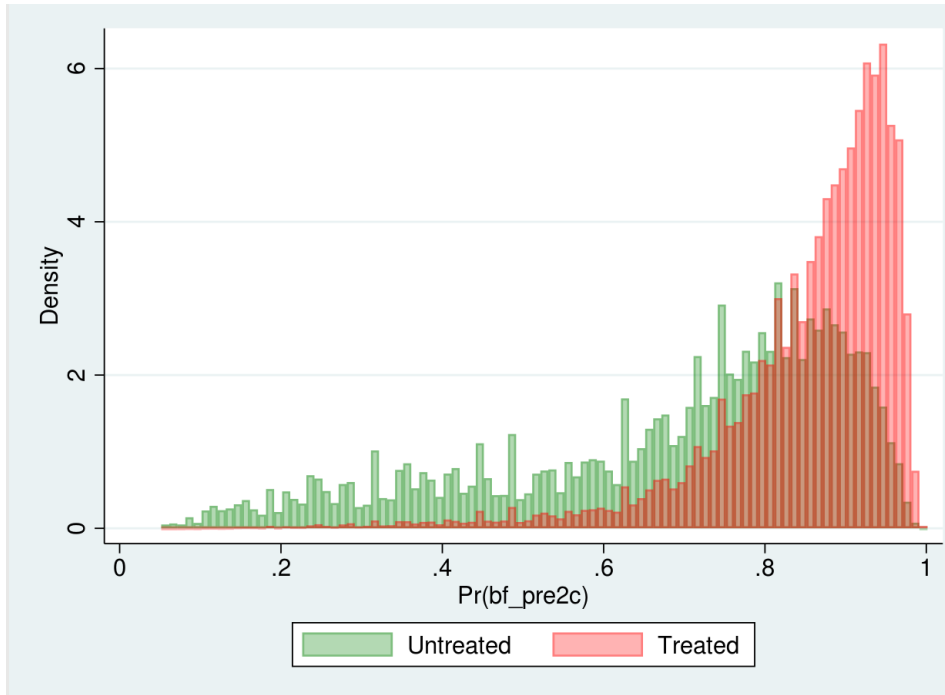
We estimated the average treatment effect on the treated (ATT) by kernel matching (eTable 3). eFig 1 shows the distribution of estimated propensity scores given the covariates. Overlapping histograms are displayed for the beneficiaries and non-beneficiaries. eFig2 shows the balance between variables before and after Kernel matching.

eTable 3. ATT of Maternal Mortality for BFP Receipt Between 2004 and 2015 Using Kernel Matching

| | Kernel Weighting |
|--|-------------------------------------|
| | ATT* (95%CI) |
| ATT | -0.0001096 (-0.0002052, -0.0000139) |
| N | 5,757,188 |
| *Average treatment effect on the treated (ATT) estimated using kernel matching (PS variables). | |

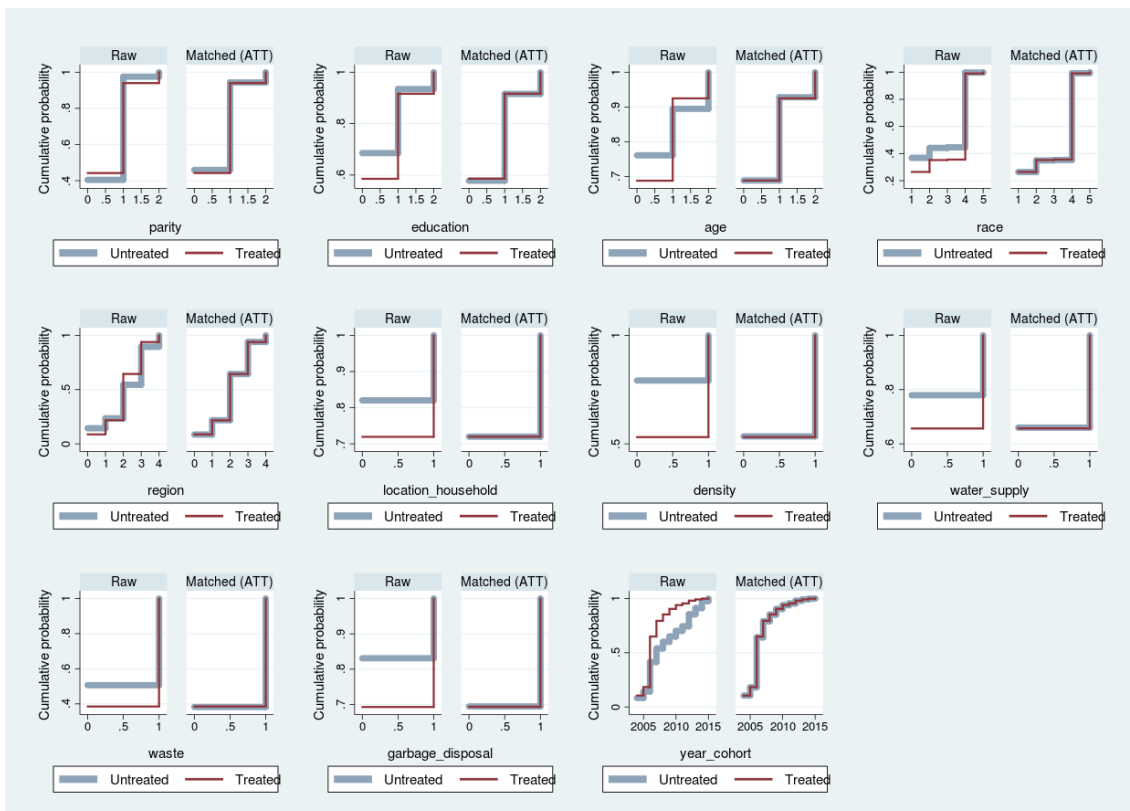
eFigure 1. Propensity Scores Common Support Area

Distribution of propensity scores across beneficiaries and non-beneficiaries



eFigure 2. Cumulative Distribution Balancing Plot

Balance between variables before and after Kernel matching]



3. Sensitivity analysis

3.1 Inverse probability of treatment weighting (IPTW)

We used the same framework of analysis for the kernel weighting, to estimate the effect of treatment on the treated (ATT) using weights. First, we estimated the propensity score (PS) for receiving BFP from the cohort baseline sociodemographic covariates. Second, we estimated the weights for BFP beneficiary (weight=1) and non-BFP beneficiary families (weight= $E(PS)/(1-E(PS))$). We estimated the logistic regression using inverse probability of treatment weighting (IPTW) and compared the differences in the distribution of PS covariates between beneficiaries and non-beneficiaries using proportions, to assess the balance of potential confounders before and after IPTW (eTable 4). This process yielded very similar results to the kernel approach, with similar rates with the IPTW balance (61.41 (59.35-63.53) and 70.67 (62.48-79.92)) for beneficiaries and non-beneficiaries (non-BFP), respectively (eTable 5). We also performed unweighted multivariate logistic regression and beneficiaries were slightly higher than non-beneficiaries (61.41(59.35-63.56) and 57.91 (53.86-62.26)). Adjusted IPTW results (weighted OR:0.84, 95%CI:0.74-0.96) and unweighted multivariate logistic regression (OR:0.89, 95%CI:0.81-0.98) analysis, consistently showed a similar magnitude of associations (eTable 5).

eTable 4. Baseline Characteristics of Bolsa Família Program (BFP) Nonbeneficiaries (non-BFP) and Beneficiaries From *Cadastro Único* Registration Between 2004 and 2015

| Socioeconomic and demographic variables | Before IPTW | | | After IPTW | | |
|--|----------------------------|------------------------|-------------------------------|---------------------|------------------------|---------------------------|
| | Non-BFP n= 1,264,037 | BFP n=5,413,23 6 | Diff (BFP- Non- BFP) | Non-BFP n=1,020,639 | BFP n=5,413,23 6 | Diff (BFP- Non-BFP) |
| Sociodemographic | | | | | | |
| Race/Colour | | | | | | |
| Asian | 0.47 | 0.35 | -0.12 | 0 | 0 | 0 |
| Black | 7.36 | 8.78 | 1.42 | 9.04 | 8.90 | -0.14 |
| Indigenous | 0.29 | 1.03 | 0.74 | 0 | 0 | 0 |
| Mixed-race | 55.69 | 63.69 | 8 | 64.63 | 64.58 | -0.05 |
| White | 36.19 | 26.15 | -10.04 | 26.33 | 26.52 | 0.19 |
| Education | | | | | | |
| High school/College (>=8 years) | 69.64 | 58.88 | -10.76 | 54.81 | 58.88 | 4.07 |
| Elementary/Middle school (4-7 years) | 24.21 | 32.86 | 8.65 | 35.59 | 32.86 | -2.73 |
| Elementary school or illiterate (<3 years) | 6.15 | 8.26 | 2.11 | 9.59 | 8.26 | -1.33 |
| Age Groups | | | | | | |
| 10-19 years old | 14.09 | 23.24 | 9.15 | 22.42 | 23.24 | 0.82 |
| 20-34 years old | 75.89 | 69.32 | -6.57 | 70.19 | 69.32 | -0.87 |
| >=35 years old | 10.02 | 7.44 | -2.58 | 7.39 | 7.44 | 0.05 |

| Sociodemographic | Non-BFP n= 1,264,037 | BFP n=5,413,23 6 | Diff (BFP- Non- BFP) | | Non-BFP n=1,020,639 | BFP n=5,413,23 6 | Diff (BFP- Non-BFP) |
|----------------------------|-----------------------------------|-------------------------------|---|--|----------------------------|-------------------------------|--|
| Parity | | | | | | | |
| 1 child in the cohort | 39.74 | 44.35 | 4.61 | | 44.35 | 44.35 | 0 |
| 2-3 children in the cohort | 57.80 | 49.66 | -8.14 | | 49.66 | 49.66 | 0 |
| >3 children in the cohort | 2.46 | 5.99 | 3.53 | | 5.99 | 5.99 | 0 |
| Household | | | | | | | |
| Density | | | | | | | |
| <=2 per room | 79.30 | 53.22 | -26.08 | | 53.56 | 53.22 | -0.34 |
| >2 per room | 20.70 | 46.78 | 26.08 | | 46.44 | 46.78 | 0.34 |
| Water Supply | | | | | | | |
| Public network | 77.51 | 65.42 | -12.09 | | 65.36 | 65.42 | 0.06 |
| Well/natural source/other | 22.49 | 34.58 | 12.09 | | 34.64 | 34.58 | -0.06 |
| Waste disposal system | | | | | | | |
| Public network | 77.51 | 38.72 | -38.79 | | 38.71 | 38.72 | 0.01 |
| Septic tank/ditch/other | 22.49 | 61.28 | 38.79 | | 61.29 | 61.28 | -0.01 |
| <i>Missing data</i> | 8.88 | 3.51 | | | | 3.51 | 3.51 |
| Garbage disposal | | | | | | | |
| Public collection system | 51.21 | 68.78 | 17.57 | | 69.39 | 68.78 | -0.61 |

| | | | | | | | | |
|---------------------|--|-------|-------|--------|--|-------|-------|-------|
| Burned/buried/other | | 49.79 | 31.22 | -18.57 | | 30.61 | 31.22 | 0.61 |
| Geographical | | | | | | | | |
| Region | | | | | | | | |
| South | | 14.19 | 8.55 | -5.64 | | 8.43 | 8.55 | 0.12 |
| North | | 9.78 | 13.21 | 3.43 | | 13.52 | 13.21 | -0.31 |
| Northwest | | 30.69 | 42.87 | 12.18 | | 41.19 | 42.87 | 1.68 |
| Southwest | | 35.10 | 29.29 | -5.81 | | 30.22 | 29.29 | -0.93 |
| Central-West | | 10.24 | 6.08 | -4.16 | | 6.64 | 6.08 | -0.56 |
| Household location | | | | | | | | |
| Urban | | 81.56 | 71.74 | -9.82 | | 72.07 | 71.74 | -0.33 |
| Rural | | 18.44 | 28.26 | 9.82 | | 27.93 | 28.26 | 0.33 |
| Year | | | | | | | | |
| 2004 | | 7.85 | 10.26 | 2.41 | | 11.00 | 10.26 | -0.74 |
| 2005 | | 7.12 | 7.92 | 0.8 | | 7.70 | 7.92 | 0.22 |
| 2006 | | 26.84 | 45.62 | 18.78 | | 44.19 | 45.62 | 1.43 |
| 2007 | | 11.18 | 13.71 | 2.53 | | 15.55 | 13.71 | -1.84 |
| 2008 | | 5.37 | 5.53 | 0.16 | | 6.41 | 5.53 | -0.88 |
| 2009 | | 4.33 | 4.85 | 0.52 | | 5.28 | 4.85 | -0.43 |
| 2010 | | 6.12 | 4.36 | -1.76 | | 3.65 | 4.36 | 0.71 |
| 2011 | | 5.18 | 2.51 | -2.67 | | 1.56 | 2.51 | 0.95 |
| 2012 | | 10.41 | 2.79 | -7.62 | | 2.61 | 2.79 | 0.18 |

| | | | | | | | | |
|------|--|------|------|-------|--|------|------|-------|
| 2013 | | 7.44 | 1.61 | -5.83 | | 1.21 | 1.61 | 0.4 |
| 2014 | | 6.09 | 0.72 | -5.37 | | 0.71 | 0.72 | 0.01 |
| 2015 | | 2.07 | 0.12 | -1.95 | | 0.13 | 0.12 | -0.01 |

eTable 5. Kernel-Weighted Regression and Other Strategies for Associations Between BFP Receipt and Maternal Death in 100 Million Brazilian Cohort, 2004-2015

| | Kernel Weighting | | | | IPTW | | | | Unweighted Logistic Regression | | | |
|--|------------------------|------------------------|---------------------|--------------------------|------------------------|------------------------|---------------------|--------------------------|--------------------------------|------------------------|---------------------|--------------------------|
| | Weighted Rates | | OR(95%IC) | | Weighted Rates | | OR(95%IC) | | Rates | | OR(95%IC) | |
| | Non-BFP _a | BFP _a | OR | Adjusted OR _b | Non-BFP _a | BFP _a | OR | Adjusted OR _b | Non-BFP _a | BFP _a | OR | Adjusted OR _c |
| Rates /OR | 71.59 (62.87-81.53) | 60.63 (58.46-62.89) | 0.84 (0.73-0.96) | 0.82 (0.71-0.93) | 70.67 (62.48-79.92) | 61.41 (59.35-63.53) | 0.87 (0.76-0.98) | 0.84 (0.74-0.96) | 57.91 (53.86-62.26) | 61.41 (59.35-63.56) | 0.91 (0.83-1.00) | 0.89 (0.81-0.98) |
| n | 1,017,154 | 4,731,624 | 5,748,917 | 5,542,230 | 1,020,639 | 5,413,236 | 6,433,875 | 6,175,248 | 1,264,037 | 5,413,236 | 5,879,995 | 5,665,567 |
| ^a Non-BFP and BFP correspond to maternal rates/100,000 parturients of SINASC to non-beneficiaries and beneficiaries, respectively | | | | | | | | | | | | |
| ^b Adjusted by prenatal care, gestational age, type of delivery, and multiple pregnancy | | | | | | | | | | | | |
| ^c Multivariate logistic regression adjusted by PS variables. | | | | | | | | | | | | |
| | | | | | | | | | | | | |
| | | | | | | | | | | | | |

3.2 Varying definitions of exposure

In order to verify whether different definitions of BFP exposure could affect the results, we fitted the same kernel weighting procedure with varying exposure definitions, obtaining similar findings (eTables 6 and 7). In the main strategy (main document), we considered as exposed the BFP beneficiary women with records of children born alive who started to receive BFP before or during pregnancy, and continued to receive the benefit before the outcome, or until childbirth. Unlike our main exposure, we only consider as unexposed those who did not receive BFP at any time in the entire cohort (eTable 6). The results are similar to the main definition of non-exposed (women who did not receive the benefit at any time, or until childbirth, were considered not exposed) (eTable 6).

We also verified other definitions considering the period of pregnancy covered by the exposure. Thus, we considered as exposed the BFP beneficiary women with records of children born alive who started to receive PBF, 22 weeks (foetal viability proxy), 6, or 9 months before childbirth. The results were similar to those found in the main analysis (eTable 7).

eTable 6. Kernel-Weighted Regression for Associations Between BFP Receipt and Maternal Death in 100 Million Brazilian Cohort, 2004-2015¹

| Kernel weighting | | | | Unweighted estimates | | | |
|---|------------------------|-------------------------------------|----------------------------------|-------------------------------------|------------------------|--------------------------------|----------------------------------|
| Kernel weighted rates by 100,000 SINASC parturients | | Kernel weighted logistic regression | | Rates by 100,000 SINASC parturients | | Unweighted logistic regression | |
| Non-BFP | BFP | OR (95%CI) | Adjusted OR (95%CI) ² | Non-BFP (95%CI) | BFP (95%CI) | OR (95%CI) ³ | Adjusted OR (95%CI) ² |
| 62.56 (56.58-69.10) | 61.41 (59.35-63.53) | 0.56 (0.42, 0.75) | 0.58 (0.44, 0.78) | 107.19 (81.04-141.75) | 60.47 (58.29-62.72) | 0.79 (0.69, 0.90) | 0.82 (0.71, 0.95) |
| n=609,049 | n=4,720,036 | n=5,185,459 | n=4,992,866 | n=465423 | n=4,720,036 | n=5,204,589 | n=5,011,102 |

¹ Exposed: BFP beneficiary women with records of children born alive who started to receive PBF before or during pregnancy, and continued to receive the benefit before the outcome or until childbirth; Not Exposed: Did not receive BFP at any time in the entire cohort.

² Adjusted by prenatal care, gestational age, type of delivery and multiple pregnancy.

³ Multivariate logistic regression adjusted by PS variables.

eTable 7. Kernel-Weighted Regression for Associations Between Duration of BFP Receipt and Maternal Death per Year in 100 Million Brazilian Cohort, 2004-2015

| | 22 weeks before childbirth | 6 months before childbirth | 9 months before childbirth |
|----|----------------------------|----------------------------|----------------------------|
| | OR (95%CI) | OR (95%CI) | OR (95%CI) |
| OR | 0.83(0.72-0.95) | 0.82(0.71-0.94) | 0.83(0.72-0.96) |
| N | 5,498,605 | 5,459,798 | 5,325,764 |

¹ Exposed: BFP beneficiary women with records of children born alive who started to receive BFP 22 weeks (foetal viability proxy), 6, or 9 months before childbirth; Not Exposed: women who did not receive the benefit at any time, or until childbirth.

3.3 Length of exposure to receipt of Bolsa Família

As described in the main text, in order to evaluate BFP duration effects, we fitted the same kernel weighting procedure with varying exposure definitions, comparing women exposed to different ranges of years between the BFP receipt until delivery (1-4 years, 5-8 years, or ≥ 9 years), with women not exposed to the programme. Additionally, we calculated an indicator considering the years of exposure to the BFP, divided by the woman's age on delivery, multiplied by 100, resulting in a proxy of the percentage of the life of women exposed to PBF until delivery (mean=30.44%), classified as <30%, 30 to 70%, and $\geq 70\%$. The results for the analysis with varying exposure definitions can be found on eTables 8 and 9.

eTable 8. Kernel-Weighted Regression for Associations Between Length of Exposure to Receiving BFP and Maternal Death per Year in 100 Million Brazilian Cohort, 2004-2015¹

| | <u>Length of exposure to receipt of Bolsa Família</u> | | |
|-----------------------------|--|-------------------|-----------------------------------|
| | 1-4 years* | 5-8 years* | ≥ 9 years* |
| Weighted Odds Ratio* | 0.85 | 0.70 | 0.68 |
| (95% CI) | 0.75-0.97 | 0.60-0.82 | 0.53-0.88 |
| | | | |
| N | 5,220,204 | 2,189,711 | 1,071,053 |

¹All the analytical steps (PS estimation, kernel matching and weighted logistic regressions) were conducted separately considering varying exposure definitions: Exposed- BFP beneficiary women who started to receive PBF for 1-4, 5-8, or ≥ 9 years before childbirth; Not exposed: women who did not receive the benefit at any time, or until childbirth.

eTable 9. Kernel-Weighted Regression for Associations Between Length of Exposure to Receiving BFP/Age and Maternal Death per Year in 100 Million Brazilian Cohort, 2004-2015¹

| | <u>Length of exposure to receipt of Bolsa Família/age</u> | | |
|--------------|--|---------------|-------------------------------|
| | <30% | 30-70% | $\geq 70\%$ |
| | | | |
| OR | 0.92 | 0.52 | 0.39 |
| CI95% | 0.77-1.09 | 0.40-0.69 | 0.18-0.82 |
| | | | |
| N | 3,256,101 | 4,650,341 | 1,298,905 |

All the analytical steps (PS estimation, kernel matching and weighted logistic regressions) were conducted separately considering varying exposure definitions: Exposed: years of exposure of the BFP, divided by the woman's age on delivery, multiplied by 100, resulting in a proxy of the percentage of the life of women exposed to PBF until delivery (<30, 30-70, ≥ 70) ; Not exposed: women who did not receive the benefit at any time, or until childbirth.

In order to test the robustness of the analysis by length of exposure, we estimated kernel matching and weighted logistic regressions for each year of duration of BFP receipt in the study period. There was a dose-response association with BFP and maternal mortality until the 8th year of exposure, confirming the increase in association over time found in the main analysis (eTable 10).

| eTable 10. Kernel-Weighted Regression for Associations Between Length of Exposure to Receiving BFP and Maternal Death per Year in 100 Million Brazilian Cohort, 2004-2015 ¹ | | | | | | | | | | | |
|---|-----------|-----------|-----------|-----------|-----------|-----------|-----------|-----------|-----------|-----------|-----------|
| | 1 year | 2 years | 3 years | 4 years | 5 years | 6 years | 7 years | 8 years | 9 years | 10 years | 11 years |
| OR | 0.80 | 0.78 | 0.74 | 0.70 | 0.67 | 0.62 | 0.60 | 0.56 | 0.57 | 0.57 | 0.34 |
| CI95% | 0.70-0.93 | 0.68-0.91 | 0.64-0.87 | 0.59-0.82 | 0.56-0.80 | 0.52-0.76 | 0.49-0.73 | 0.45-0.70 | 0.43-0.76 | 0.33-0.98 | 0.10-1.13 |
| N | 5,186,550 | 4,056,807 | 3,466,353 | 3,458,245 | 2,914,401 | 2,334,259 | 1,778,948 | 1,216,999 | 646,838 | 196,083 | 93,616 |
| ¹ Exposed: BFP beneficiary women with records of children born alive who started to receive PBF for 1, 2, 3, 4, 5, 6, 7, 8, 9, 10, or 11 years before childbirth; Not exposed: women who did not receive the benefit at any time, or until childbirth. | | | | | | | | | | | |

3.4 Subgroup analysis

To justify our hypothesis that the association between BFP and maternal mortality varies across subgroups, we conducted an analysis stratified by subgroups of some sociodemographic indicators, such as race (white, black/mixed black), area of residence (rural, urban) and Municipal Human Development Index – MHDI (high or very high, medium, low, very low). We estimated the propensity score for each subgroup of these sociodemographic indicators, with the same variables as the previous steps, as well as conducted kernel-weighted logistic models separately within each subgroup (eTables 11-14).

eTable 11. Regression Results: Coefficients on Unadjusted Kernel-Weighted Logistic Regressions Within Subgroups of Municipal Human Development Index (MHDI)¹

| Municipal Human Development Index – MHDI | Weighted Odds Ratio* (95% CI) | Robust Standard Error | p-value | N |
|--|----------------------------------|-----------------------|---------|-----------|
| Model c1 | | | | |
| High or very high | 1.01 (0.85 to 1.20) | 0.004 | 0.510 | 3,149,053 |
| Constant | 0.002 (0.001 to 0.002) | 0.0001 | < 0.001 | |
| Model c2 | | | | |
| Medium | 0.94 (0.76 to 1.17) | 0.055 | 0.053 | 2,062,570 |
| Constant | 0.002 (0.002 to 0.003) | 0.0001 | < 0.001 | |
| Model c3 | | | | |
| Low, very low | 0.62 (0.47 to 0.84) | 0.061 | 0.010 | 1,465,433 |
| Constant | 0.003 (0.003 to 0.004) | 0.0002 | < 0.001 | |

* Beneficiary status (BFP=1)

¹ All the analytical steps (PS estimation, kernel matching and weighted logistic regressions) were conducted separately for each category of this subgroup.

eTable 12. Regression Results: Coefficients on Unadjusted Kernel-Weighted Logistic Regressions Within Subgroups of Maternal Race¹

| Race | Weighted Odds Ratio* (95% CI) | Robust Standard Error | p-value | N |
|--------------------|--|------------------------------|----------------|-----------|
| Model d1 | | | | |
| White | 0.98 (0.68 to 1.42) | 0.183 | 0.942 | 1,645,603 |
| Constant | 0.005 (0.003 to 0.007) | 0.0001 | < 0.001 | |
| Model d2 | | | | |
| Black/ Mixed Black | 0.79 (0.69 to 0.93) | 0.064 | < 0.001 | 4,126,296 |
| Constant | 0.007 (0.007 to 0.009) | 0.0001 | < 0.001 | |

* Beneficiary status (BFP=1)

¹ All the analytical steps (PS estimation, kernel matching and weighted logistic regressions) were conducted separately for each category of this subgroup.

eTable 13. Regression Results: Coefficients on Unadjusted Kernel-Weighted Logistic Regressions Within Subgroups of Area of Residence (Rural and Urban)¹

| Area of residence | Weighted Odds Ratio* (95% CI) | Robust Standard Error | p-value | N |
|-------------------|----------------------------------|-----------------------|---------|-----------|
| Model e1 | | | | |
| Urban | 0.89 (0.77 to 1.04) | 0,183 | 0.160 | 4,262,349 |
| Constant | 0.006 (0.006 to 0.008) | 0.005 | < 0.001 | |
| Model e2 | | | | |
| Rural | 0.69 (0.53 to 0.92) | 0,099 | 0.008 | 1,486,571 |
| Constant | 0.004 (0.003 to 0.005) | 0.001 | < 0.001 | |

* Beneficiary status (BFP=1)

¹All the analytical steps (PS estimation, kernel matching and weighted logistic regressions) were conducted separately for each category of this subgroup.

eTable 14. Regression Results: Coefficients on Unadjusted Kernel-Weighted Logistic Regressions Within Subgroups of Municipal Quintiles of Family Health Program Coverage¹

| Family Health Programme Coverage | Weighted Odds Ratio* (95% CI) | Robust Standard Error | p-value | N |
|----------------------------------|----------------------------------|-----------------------|---------|-----------|
| Model b1 | | | | |
| >=70% | 0.76 (0.57 to 0.97) | 0.205 | 0.001 | 3,294,584 |
| Constant | 0.007 (0.006 to 0.009) | 0.0007 | < 0.001 | |
| Model b2 | | | | |
| 30-70% | 0.98 (0.78 to 1.22) | 0.215 | 0.677 | 2,233,112 |
| Constant | 0.006 (0.005 to 0.008) | 0.0006 | < 0.001 | |
| Model b3 | | | | |
| <=30% | 1.04 (0.78 to 1.39) | 0.147 | 0.757 | 1,134,887 |
| Constant | 0.007 (0.005 to 0.009) | 0.0001 | < 0.001 | |

* Beneficiary status (BFP=1)

¹ All the analytical steps (PS estimation, kernel matching and weighted logistic regressions) were conducted separately for each category of this subgroup.

3.5 Variables for matching

The selection of variables was guided by the literature and the availability of such variables in our dataset, considering factors related to both treatment and outcome, as suggested by Sturmes et al, 2014⁸; Austin, 2011⁹, Brookhart et al, 2010¹⁰, Ali, et al, 2019. The 100 Million Brazilian Cohort comprises the poorest half of the Brazilian population. Therefore, socioeconomic factors must be considered when estimating the PS, as they are important confounders of this relationship. However, household conditions variables, such as type of water supply, waste disposal system (sewer system) and location of household, could influence the BFP receipt, as well be impacted by the recipient of the benefit. This may violate the conditions necessary for these covariates to be considered exogenous. To reduce possible biases, We performed robustness analyses excluding water type, location of household (rural and urban), and waste type from the set of covariates used to match on for and we obtained similar results (OR:0,81 CI: 0.70-0.93) (eTable 15-17).

eTable 15. Logistic Models to Estimate the Propensity Scores for *Bolsa Família* Participation

| Variables | Odds Ratio (95% CI) | p-value | Standard Error |
|--|--------------------------------|----------------|---------------------------|
| Parity | | | |
| 1 child in the cohort | Ref. | | - |
| 2-3 children in the cohort | 0.57 (0.57- 0.58) | <0.001 | 0.0014 |
| >3 children in the cohort | 2.18 (2.15 - 2.11) | <0.001 | 0.0015 |
| Education | | | |
| High school/College (≥8 years) | Ref | | |
| Elementary/Middle school (4-7 years) | 1.16 (1.15-1.16) | <0.001 | 0.0031 |
| Elementary school or illiterate (<3 years) | 0.91 (0.90-0.92) | <0.001 | 0.0045 |
| Age groups | | | |
| 20-34 | Ref. | <0.001 | - |
| ≤ 19 | 2.06 (2.05 - 2.08) | <0.001 | 0.0071 |
| ≥ 35 | 0.90 (0.90- 0.92) | <0.001 | 0.0037 |
| Race | | | |
| White | Ref. | | - |
| Asian descent | 1.16 (1.12 - 1.21) | <0.001 | 0.0216 |
| Black | 1.32 (1.31 - 1.34) | <0.001 | 0.0063 |
| Indigenous | 3.41 (3.27 - 3.55) | <0.001 | 0.0703 |
| Mixed/Black | 1.22 (1.21 - 1.23) | <0.001 | 0.0344 |
| Region | | | |
| South | Ref. | | - |

| | | | |
|--|-----------------------|--------|--------|
| North | 1.66 (1.65-1.68) | <0.001 | 0.0089 |
| Northeast | 2.00 (1.98 - 2.01) | <0.001 | 0.0084 |
| Southeast | 1.42 (1.41 - 1.43) | <0.001 | 0.0057 |
| Center-West | 0.93 (0.92 - 0.94) | <0.001 | 0.0049 |
| Garbage disposal | | | |
| Public collection system | Ref. | | - |
| Burned/buried/other | 1.36 (1.35 - 1.37) | <0.001 | 0.0866 |
| Registration year in Cadastro Unico | | | |
| 2004 | Ref | | |
| 2005 | 1.00 (0.99-1.01) | 0.647 | 0.0059 |
| 2006 | 1.74 (1.16-1.18) | <0.001 | 0.0052 |
| 2007 | 0.79 (0.78-0.80) | <0.001 | 0.0040 |
| 2008 | 0.70 (0.69-0.71) | <0.001 | 0.0041 |
| 2009 | 0.80 (0.80-0.81) | <0.001 | 0.0051 |
| 2010 | 0.53 (0.52-0.53) | <0.001 | 0.0034 |
| 2011 | 0.37 (0.37-0.38) | <0.001 | 0.0027 |
| 2012 | 0.21 (0.21-0.22) | <0.001 | 0.0012 |
| 2013 | 0.18 (0.18-0.18) | <0.001 | 0.0013 |
| 2014 | 0.09 (0.09-0.09) | <0.001 | 0.0070 |
| 2015 | 0.42 (0.42-0.44) | <0.001 | 0.0067 |

eTable 16. Average Treatment Effect on the Treated (ATT) Estimated Using Kernel-Matching (Excluding Type of Water Supply, Waste Disposal System [Sewer System], and Location of Household of the PS)

| | Kernel Weighting |
|--|-------------------------------------|
| | ATT* (95%CI) |
| ATT | -0.0001275 (-0.0002252, -0.0000297) |
| N | 5,879,643 |
| *Average treatment effect on the treated (ATT) estimated using kernel matching (PS variables). | |

eTable 17. Kernel-Weighted Regression for Associations Between BFP Receipt and Maternal Death (Excluding Type of Water Supply, Waste Disposal System [Sewer System], and Location of Household of the PS) in 100 Million Brazilian Cohort, 2004-2015

| | Weighted Odds Ratio* | |
|---|-----------------------------|--|
| | (95% CI) | |
| | 0.81(0.70-0.93) | |
| N | 5,662,191 | |

4. Data linkage and quality assessment

4.1 100 Million Brazilian Cohort baseline versus Mortality Information System linkage

The women who died by maternal causes in the cohort were identified by linking data from the Brazilian Mortality Information System (SIM) with the 100 Million Cohort dataset for a 15 year period (2000 to 2015). The 100 Million Cohort and SIM databases were linked by the Centre for Data and Knowledge Integration for Health (Barreto et al., 2019), using record linkage software developed by the centre. This section provides a detailed description of how this linkage was carried out.

The linkage algorithm used five variables to identify matching records from the two databases (SIM and 100 Million Cohort), and each one was recorded in both of the linked datasets: beneficiary's name, mother's name, municipality of residence code, and date of birth.

The record linkage software performed two main steps:

First, the record corresponding to data for each 100 Million Cohort individual (the larger dataset) was indexed in the Lucene Apache library¹¹. Lucene Apache has a method of indexing files, and performed a data search in these files. After this initial structuring, the data was saved in an indexed document with an identification number. At the end of the process, a folder was generated with the indexed database.

Second, for each death in the SIM dataset, the algorithm searched the 100 Million cohort indexed database for a potential match. The variables used to match the individual records were: the beneficiary's name, mother's name, date of birth, municipality of residence, and sex. All of the records must include the beneficiary's name, in order to perform the linkage. The records which did not include the beneficiary's name were excluded. The software performs this search at three levels:

Level 1: A first search is made through all of the five selected variables. The mortality data (SIM) found in the 100 Million Cohort baseline is saved, generating a list of linked records. A similarity calculation is performed for the linked records, generating a score which indicates the similarity between the two linked records. The similarity calculation is performed by comparing names, dates of birth, states, municipalities, and sex for the two records (one in each dataset) of the linked individuals. This process generates a "weighted average," which is the value of the similarity score. If the score is equal to, or greater than 0.95, the link is accepted (as being a correct link).

Level 2: If the similarity calculation is less than 0.95, a new search is performed (this time only including those matched under the 0.95 score), using a group of four (of the five previous variables) per time, in five different combinations. Subsequently, all search results with each group of variables are pooled, and the similarity test is performed. Again, if the score is equal to, or greater than 0.95, the linked data is accepted.

Level 3: Lastly, if a SIM individual is not linked to a 100 Million Cohort record in the two previous searches, a final search is undertaken. For this search, the beneficiary's name, mother's name, municipality of residence code, and date of birth are divided into smaller pieces. For example, the date of birth is divided into three smaller categories: day, month, and year. Several attempts are made, to try to match the person in both datasets. In order to obtain confirmation on whether the record linked is actually valid, a distance-editing method is applied, based on Jaro-Winkler.

In order to check the entire linked dataset, robust accuracy tests are performed, to assess the overall linkage quality.

4.2 Statistical accuracy of the linkage

The quality of the linkage was tested manually and using a receiver operating characteristic (ROC) curve by a team of statisticians working at CIDACS. The entire process is described below.

For the manual test, the team took a sample of 10,000 linked pairs (i.e. each pair represents a maternal death linked to a 100 Million Cohort record). Sampling was stratified by similarity scores, to enable assessment of the linkage quality for a range of similarity scores. The sampling process was developed in five stages:

Similarity scores generated in the linkage process were divided into five strata;

The number of observations in each was used to decide the stratum range;

- A variable was created for each linked record pair, to indicate the stratum of similarity to which the record pair belonged.
- The proportion of records in each similarity stratum was calculated.
- A random sample proportional to the size of the stratum was taken, within each similarity stratum.

In the random sample of 10,000 linked pairs, three of the five items of information used for the linkage (beneficiary's name, mother's name, and date of birth) were checked for each pair. Three new variables were created, indicating whether the information on the two databases (100 Million Cohort and SIM) agreed for each of the three variables separately, assigning a value for each variable of 1 for agreement, and 0 for disagreement. Records with agreement in all three variables were declared a true match and those with some disagreement were investigated further. For disagreements in the name fields, these were considered to agree when the names in the two databases contained different letters with similar phonetics. For foreign or uncommon names, disagreements of up to three digits were accepted (and the agreement variable was reset to 1, indicating agreement). For disagreements in the date of birth field, differences of only one number were accepted.

For remaining disagreements in the name fields, if any of the following occurred:

- The mother's name was completely different in the two databases.
- Three or more different letters in the name in the two databases.
- Completely different surnames in the two databases
- One of the pairs of records did not contain a surname in the two databases.
- The surname contained two or more abbreviations in one of the pairs of records, or different surnames.

The records were then declared to be a false match for remaining disagreements in the date of birth field, if any of the following occurred:

- At least one of the digits in the year was different, resulting in an age difference of at least 7 (seven) years; or
- At least one digit in the month was different, resulting in a difference of at least 10 (ten) months; or,
- Both the month and year were different.

The records were then declared to be a false match.

In addition to these cases, all other records with some disagreement were declared true matches following manual inspection. At this stage, if any doubts remained, they were reanalysed using two extra matching variables: sex and municipality.

ROC curve

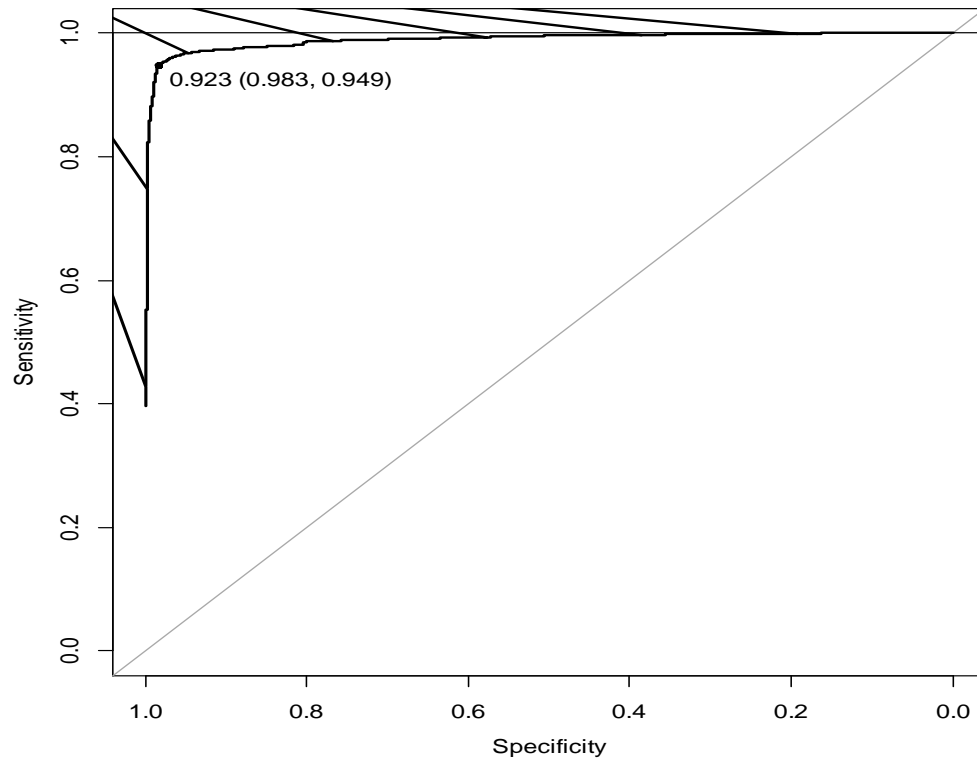
Following the manual verification process, the record linkage algorithm sensitivity and specificity were estimated for a range of cut-off values (the criteria for declaring a true match using the record linkage algorithm), viewing the result of the manual verification as the gold standard classification of the links (eTable 18). Using these estimated specificities and sensitivities, ROC curves were constructed to identify the global accuracy (as measured by the area under the curve) of the results obtained by the similarity score (eFigure 3).

From the ROC curve, the optimal cut-off point of 0.92 (ROC curve area [Sensitivity/Specificity]: 0.923 [0.983/0.949]) was identified. Using this optimal cut-off to declare matches, 97.8% of the linked pairs were estimated to be true matches, and 2.2% were estimated to be false matches. An estimated 5% of the true matches were not linked (eFigure 3).

eTable 18. Accuracy Analysis of the Linkage Between *Cadastro Único* and the Mortality Information System in a Sample of 10 000 Record Pairs

| Cut-off point | Specificity; sensitivity | Total matches (%) | True matches (% of linked cases) | False matches (% of linked cases) | Lost matches (% of true matches) |
|---------------|--------------------------|-------------------|----------------------------------|-----------------------------------|----------------------------------|
| ≥0.83 | SP=0.459; S=0.996 | 7551 (75.5) | 4686 (62.1) | 2865 (37.9) | 17 (0.4) |
| ≥0.84 | SP=0.576; S=0.993 | 6916 (69.2) | 4668 (67.5) | 2248 (32.5) | 35 (0.7) |
| ≥0.85 | SP=0.692; S=0.990 | 6290 (62.9) | 4657 (74.0) | 1633 (26.0) | 46 (1.1) |
| ≥0.86 | SP=0.789; S=0.987 | 5759 (57.6) | 4641 (80.6) | 1118 (19.4) | 62 (1.3) |
| ≥0.87 | SP=0.863; S=0.977 | 5322 (53.2) | 4595 (86.3) | 727 (13.7) | 108 (2.3) |
| ≥0.88 | SP=0.908; S=0.973 | 5060 (50.6) | 4575 (90.4) | 485 (9.6) | 128 (2.7) |
| ≥0.89 | SP=0.941; S=0.969 | 4872 (48.7) | 4557 (93.5) | 315 (6.5) | 146 (3.1) |
| ≥0.90 | SP=0.961; S=0.964 | 4741 (47.4) | 4534 (95.6) | 207 (4.4) | 169 (3.6) |
| ≥0.91 | SP=0.974; S=0.956 | 4638 (46.4) | 4498 (97.0) | 140 (3.0) | 205 (4.4) |
| ≥0.92 | SP=0.981; S=0.950 | 4570 (45.7) | 4470 (97.8) | 100 (2.2) | 233 (5.0) |
| ≥0.93 | SP=0.986; S=0.940 | 4495 (45.0) | 4423 (98.4) | 72 (1.6) | 280 (6.0) |
| ≥0.94 | SP=0.989; S=0.919 | 4381 (43.8) | 4323 (98.7) | 58 (1.3) | 380 (8.0) |
| ≥0.95 | SP=0.991; S=0.895 | 4258 (42.6) | 4211 (98.9) | 47 (1.1) | 492 (10.5) |
| ≥0.96 | SP=0.995; S=0.855 | 4049 (40.5) | 4022 (99.3) | 27 (0.7) | 681 (14.5) |
| ≥0.97 | SP=0.998; S=0.750 | 3540 (35.4) | 3527 (99.6) | 13 (0.4) | 1176 (25.0) |
| ≥0.98 | SP=0.998; S=0.603 | 2844 (28.4) | 2835 (99.7) | 9 (0.3) | 1868 (39.7) |

eFigure 3. ROC Curve of the Linkage Between Mortality and *Cadastro Único* Between 2001 and 2015



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