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Author manuscript

Development and Validation of a Genomic Knowledge Scale to Advance Informed Decision Making Research in Genomic Sequencing

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Abstract

Background—This study evaluated the psychometric properties of a new, comprehensive measure of knowledge about genomic sequencing, the University of North Carolina Genomic Knowledge Scale (UNC-GKS).

Methods—The UNC-GKS assesses knowledge in four domains thought to be critical for informed decision making about genomic sequencing. The scale was validated using classical test theory and item response theory in 286 adult patients and 132 parents of pediatric patients undergoing diagnostic whole exome sequencing (WES) in the NCGENES study.

Results—The UNC-GKS assessed a single underlying construct (genomic knowledge) with good internal reliability (Cronbach's alpha $= 0.90$). Scores were most informative (able to discriminate between individuals with different levels of genomic knowledge) at one standard deviation above the scale mean or lower, a range that included most participants. Convergent validity was supported by associations with health literacy and numeracy ($rs=0.41-0.46$). The scale functioned well across subgroups differing in sex, race/ethnicity, education, and English proficiency.

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Conflict of interest statement

None of the authors have conflicts of interest to disclose. The study sponsor was not involved in study design, data collection, analysis or interpretation of data, writing this report, or the decision to submit the report for publication.

Discussion—Findings supported the promise of the UNC-GKS as a valid and reliable measure of genomic knowledge among people facing complex decisions about WES and comparable sequencing methods. It is neither disease- nor population-specific, and it functioned well across important subgroups, making it usable in diverse populations.

Keywords

genomic sequencing; knowledge; whole exome sequencing; informed decision making

Rapidly evolving genetic testing practices have begun to include panels of dozens of genes in some clinical scenarios, or even more comprehensive tests ranging from thousands of genes to the entire genome or exome (e.g., whole exome sequencing, or WES).¹ Genomescale tests ("genomic sequencing") are more complex than single gene tests in important ways. For instance, they yield a wide spectrum of potential results, many of which have uncertain meaning. Providing informed consent for genomic sequencing requires that people have knowledge that includes, but goes beyond, knowledge needed for informed consent for single gene testing.^{2,3} People with greater knowledge of the nature of genes and their effects on health, how genes are inherited in families, and the potential benefits, harms, and limitations of genomic sequencing are better equipped than their less knowledgeable peers to make informed decisions about undergoing sequencing, comprehend the meaning and limitations of their results, and take appropriate actions upon learning these results^{4,5}. Yet, many people currently offered genomic sequencing have inadequate knowledge and misconceptions about basic genetics^{2,6} and are unfamiliar with genomic sequencing. The unique and complex issues raised by genome-scale tests are not well-covered by existing knowledge measures. Having a validated, comprehensive measure of genomic knowledge could help identify knowledge gaps and reduce the chance that people's decisions and responses to genomic sequencing are based on false assumptions, unrealistic expectations, or misconceptions.

To meet the need for a valid way to assess genomic knowledge, we conducted the present study to evaluate a new measure, the University of North Carolina Genomic Knowledge Scale (UNC-GKS). We conceptualized "genomic knowledge" as encompassing four domains: the structure and function of genes, how they are inherited, their relation to health, and potential benefits, harms, and limitations of WES—a sequencing method that identifies variants in the subset of the genome that encodes the genes. These domains are based on our clinical experiences with patients offered sequencing and reflect a pragmatic approach that is highly relevant to typical research applications of genomic sequencing. The domains are also consistent with a framework discussed by Smerecnik and colleagues⁷ that includes awareness knowledge (knowing that there are genetic risk factors for disease), how-to knowledge (knowing how those risk factors influence risk for developing a disease), and principles knowledge (knowledge of pathways through which genes are theorized to influence health). Our comprehensive definition allows evaluation of the extent to which people have a basic working knowledge they can use to evaluate pros and cons, risks, uncertainties, and alternatives to genomic sequencing.

The UNC-GKS was also designed to address several limitations in existing measures. First, many existing knowledge measures focus on testing for mutations in a single genee.^{g., 8} or are specific to a particular disease^{e.g., 9,10}. However, genome-scale tests, in addition to becoming increasingly common, raise unique and complex issues not well-covered by these kinds of existing measures. Moreover, because the UNC-GKS is general rather than diseasespecific, scores can be compared across populations affected by different diseases. The UNC-GKS can also be used in populations affected by diseases for which disease-specific measures do not exist. A second issue addressed by the UNC-GKS relates to the sociodemographic diversity of most patient populations. Given recognized subgroup differences in knowledge and views of genomic sequencing, $11,12$ it is often useful to compare knowledge scores across subgroups in a study's sample. However, these comparisons are only informative if observed subgroup knowledge differences reflect real differences rather than measurement artifacts. It is rare to see formal analyses investigating the psychometric functioning of a knowledge measure across different subgroups. Finally, some existing knowledge measures assess agreement or disagreement with statements about genetics or genomics¹³—a common approach for measuring beliefs or attitudes that may, but does not necessarily, correspond to knowledge. Studies in educational testing and knowledge assessment more often use items with multiple choice or true/false response options that can be unambiguously scored as correct or incorrect^{14,15}.

Accordingly, our goal was to develop a measure that met the following criteria: (1) it covers domains of knowledge relevant to the complex decision contexts created by genome-scale sequencing; (2) it applies to a broad range of contexts and populations rather than being specific to a particular disease or population; (3) it has adequate validity and reliability across important sociodemographic subgroups; and (4) it uses a true/false response scale. No existing measures meet these criteria. We note that a recently introduced measure of basic knowledge about genetics and genetic causes of disease¹⁶ used true/false response options; however, it was evaluated in a sample with little diversity and analyses did not examine the psychometric functioning of its items across subgroups, so it is unclear whether the measure can assess knowledge with similar validity and reliability across important subgroups. Moreover, it does not include items about genomic sequencing and implications of potential sequencing results. The UNC-GKS includes those types of items, and we examined its item functioning across subgroups varying by sex, race/ethnicity, education, and English language proficiency. Consequently, the present study was expected to yield a new tool for research with diverse populations offered genomic sequencing.

Methods

Participants

Participants were enrolled in the North Carolina Clinical Genomic Evaluation by Nextgeneration Exome Sequencing (NCGENES) study, which is investigating the performance and best use of WES in the diagnosis and clinical care of patients with suspected genetic disorders³. Adult and pediatric patients evaluated at a UNC-affiliated hospital or at Vidant Medical Center (Greenville, NC) were eligible for NCGENES if they had symptoms or an illness with a possible genetic etiology (as determined by the referring physician) and if they

were in one of these diagnostic groups: hereditary cancers, cardiovascular disorders (mainly cardiomyopathies), neurodevelopmental disorders, congenital disorders, retinopathies, or selected other disorders (e.g., mitochondrial disorders). Adult patients and parents of pediatric patients sequenced in NCGENES completed study measures including the UNC-GKS. Participants also included a small sample of guardians of adult patients whose cognitive or physical functioning precluded completion of study procedures; however, none were included in the present sample. We recruited 418 participants (286 adult patients, 132 parents) for the present study's sample between August 2012 and December 2014. All completed study measures in English.

Procedures

Eligible individuals were contacted by study staff to schedule a study visit and then were mailed an appointment letter, consent and HIPAA forms, educational brochures designed for the study, and an intake questionnaire that included the UNC-GKS. Potential participants then met with a certified genetic counselor who obtained informed consent for sequencing. At this meeting, consenting participants returned their completed intake questionnaire, completed health literacy measures, and had their blood drawn. Data for the present study came from the intake questionnaire, the literacy measures, and UNC Hospitals chart abstraction. Prior to completing the UNC-GKS, participants had the opportunity to read the mailed educational brochures, which provided an overview of genomic sequencing and potential sequencing results. They had not yet received the more specific, personalized information provided during informed consent procedures. Their intake UNC-GKS scores therefore may not reflect the level of genomic knowledge in the general public; instead, these scores approximate knowledge likely to be found in candidates for sequencing. The institutional review boards of the University of North Carolina and Vidant Medical Center approved the study protocols.

Measures

Development of UNC-GKS—The UNC-GKS was developed in an iterative process that gathered feedback on measure domains and on item content and clarity from a team that included certified genetic counselors and medical geneticists with extensive clinical experience educating patients, behavioral scientists with formal training in communication and measure development, and others with and without genetics expertise. This team identified four key domains for the measure: the structure and function of genes, how they are inherited in families, their relation to health, and strengths and limitations of WES. We viewed the latter domain as a potentially separate module that future users could adapt or replace for other sequencing contexts (e.g., newborn or population screening). We reviewed existing measures and adapted or drafted knowledge items with the goal of ensuring good content validity across the four domains and cohesion across the underlying construct of genomic testing knowledge¹⁷. Some items within each domain specifically addressed misconceptions that could affect informed decision making (e.g., A mother and daughter who look alike are more genetically similar than a mother and daughter who do not look *alike*). We used genetic terms that participants were exposed to throughout the study (e.g., in consent procedures, brochures, and counseling). Because the term "gene variant" was especially important, the instructions included the following reminder: "We are using the

term 'gene variant' to mean a version of a gene. Sometimes two people have the same version of a gene (they have the same gene variant) and other times two people have different versions of a gene (they have different *gene variants*)."

The resulting measure includes 25 items framed as statements and uses the response categories: true, false, and not sure/don't know (with the latter provided to minimize guessing). The statements and correct answers appear in Table 1. We scored correct responses as 1 and incorrect responses and not sure/don't know responses as 0.

Sociodemographic and medical characteristics—Sociodemographic variables came from clinical records or self-report in the intake questionnaire. They included participant sex, race/ethnicity, educational attainment, and annual household income. Nominating clinicians reported patients' diagnosis or symptoms; this information was supplemented and confirmed during the informed consent session.

English proficiency—We used 3-item subscale of the Cultural Identity Scale¹⁸ to assess English proficiency in speaking, reading, and writing. Responses, ranging from 1 (Poor) to 4 (*Excellent*), were summed to create a single score for which higher scores indicated greater proficiency (Cronbach's alpha=0.90).

Health literacy—Study staff assessed general health literacy and genetics-related health literacy in person using the 66-item Rapid Estimate of Adult Literacy Measure¹⁹ and the 8item Rapid Estimate of Adult Literacy-Genetics⁶. For both scales, we created scores by summing the number of words a participant pronounced correctly; words that a participant pronounced incorrectly or skipped were not counted. Higher scores indicate greater general health literacy and genetics-related health literacy, respectively. REALM raw scores can be used to categorize people as having low health literacy (scores of $0-44$, sixth grade reading level), marginal literacy (scores of 45–60, seventh to eighth grade reading level), and functional health literacy (scores of $61–66$, ninth grade reading level)¹⁹. REAL-G scores of three or less have been interpreted as indicating low genetics-related health literacy (≤ sixth grade level)⁶.

Numeracy—We measured subjective numeracy with a validated 3-item version of the Subjective Numeracy Scale^{20–22}. Items assess perceived numerical aptitude and preference for numbers on a scale from 1 (Not at all good/helpful) to 6 (Extremely good/helpful). Summing responses yields a single score for which higher scores indicated stronger preference for numerical over textual information and greater perceived ability to perform mathematical tasks (Cronbach's alpha=0.89). We measured objective numeracy with a validated measure that presents three arithmetic problems testing the use of proportions, fractions, and percentages²³. Summing correct responses yields an objective numeracy score.

Data analysis

We examined the psychometric properties, factor structure (to evaluate the assumption that all items reflect a single underlying construct—in this case, genomic knowledge), and convergent validity (to evaluate the scale's association with conceptually related variables)

of the UNC-GKS. We also conducted item response theory (IRT) analyses that offer more in-depth information than classical test theory methods, including evaluation of variation in item performance (differential item functioning) across demographic subgroups.

Item-level descriptive statistics—First, we examined the proportion of participants correctly answering each UNC-GKS item to evaluate whether items were too easy (*ceiling* effects, indicated by >90% of participants answering them correctly) or too hard (floor effects, indicated by >90% of participants answering them incorrectly). Second, we computed inter-item tetrachoric correlations²⁴ to evaluate whether UNC-GKS items were positively associated with each other, as we would expect. Third, we evaluated whether responses to each item were consistent with the sum of the responses to the remaining items by examining whether item-total correlations were positive. Negative or low item-total correlations indicate items that may need to be reworded or discarded.

Factor analyses—Before completing IRT analyses, we checked the assumption that the measure was unidimensional by conducting a confirmatory factor analysis of the inter-item tetrachoric correlation matrix using the $M plus^{25}$ "weighted least squares with robust standard errors, mean- and variance-adjusted" algorithm. We evaluated model fit using the root mean square error of approximation (RMSEA, acceptable if $\langle 0.05 \rangle^{26}$, the Tucker-Lewis index (TLI, acceptable if >0.95)²⁷, the comparative fix index (CFI, acceptable if >0.95)²⁸, and residual correlations between items via modification indices. Large modification indices (>10) reveal possible local dependence for sets of items, indicating possible violation of the local dependence assumption of IRT. Local dependencies indicate content redundancy or similar wording between two or more items and may suggest additional factors exist in the scale.

IRT analyses—We evaluated performance of the UNC-GKS items by fitting one-, two-, and three-parameter logistic IRT models (1PL, 2PL, and 3PL, respectively) using the software program IRTPRO²⁹. The 1PL model³⁰ characterizes each item by a single parameter—the difficulty parameter, b, which indicates the level of genetic knowledge at which there is a 50% chance of answering the item correctly (that is, how *difficult* the item is). The 2PL model³¹ estimates both b and an additional parameter, the *discrimination* parameter, a, which reflects the degree to which item responses are associated with the latent construct being measured (how effectively an item *discriminates* between individuals with higher versus lower genomic knowledge). The 3PL model³¹ estimates the a and b parameters and an additional parameter, c, which accounts for guessing. We chose the best fitting model by examining chi-square tests of the likelihood ratio for each model pair, then examined this model's goodness of fit to the data using Orlando and Thissen's $S-X^2$ statistic $32,33$, for which a nonsignificant result indicates adequate model fit at the item-level (i.e., how well each item fits the model). We controlled for multiple comparisons using the Benjamini-Hochberg procedure^{34,35}.

In addition to items flagged for potential local dependence in the confirmatory factor analysis, we used the IRT-based local dependence statistic³⁶ to identify items that were excessively related after controlling for the underlying construct (genomic knowledge)—an undesirable characteristic. Values >10 indicate substantial local dependence. We then

conducted an additional check on the dimensionality of the data by estimating a bifactor IRT model in which each locally dependent set of items was specified as a second order factor. Violations of local dependence were deemed negligible if the variance accounted for by first-order or general factor (common variance) $37-40$ was at least 0.85.

Next, we examined differential item functioning (DIF), which enables evaluation of whether items behave differently across subgroups after holding the underlying construct (genomic knowledge) constant⁴¹. It detects a form of measurement bias that occurs when people in different groups with the same level of the underlying construct have a different probability of getting a particular score on a scale. DIF may indicate that attributes other than those the scale is intended to measure are affecting responses. In the present study, we examined DIF across sex, race/ethnicity, education, and English proficiency groups. For each item, we used a logistic regression model to evaluate whether item responses were associated with group membership after controlling for participants' IRT score on the UNC-GKS. Uniform DIF (of a similar magnitude across the range of the underlying construct) was evaluated with a likelihood ratio test comparing a logistic regression model with one predictor (IRT score) to a model with both IRT score and an additional predictor (group membership); this approach allowed us to evaluate whether, after controlling for overall level of genomic knowledge, one group was more likely than the other to answer the item correctly. Non-uniform DIF (for which magnitude may differ across the range of the underlying construct) was evaluated with a likelihood ratio test comparing a model with both predictors (IRT score and group membership) to a model that also included their interaction term. This model allowed us to evaluate whether an item provided better measurement of genomic knowledge for one group versus another. We used the Benjamini-Hochberg procedure to make inferential decisions in multiple comparisons.

According to common practice, we planned to drop items if they did not fit well, substantially violated local dependence, or functioned differently for key groups. The remaining items would then be used to calibrate a final IRT model to use in subsequent analyses.

IRT scoring and reliability—We computed IRT scores for the UNC-GKS based on the parameters from the final IRT model. These scores are relative to the population of this sample, assuming a normal distribution with a mean of 0 and standard deviation of 1. To be more easily interpretable, we scaled the IRT scores to the T-score metric with a mean of 50 and a standard deviation of 10. Analysts typically compute IRT scaled scores based on response patterns, essentially weighting item responses by their IRT a parameters so that items more strongly related to the underlying construct have a greater impact on the score. However, analysts also often use summed scores because they do not require special software. To enable practical use of scaled scores, we computed a scoring table to convert summed scores to expected scaled scores. We also computed a scoring table for a 19-item version of the UNC-GKS that omitted the WES items, for use when those items are not needed. The 19-item version was scored using the 19 IRT parameters from the 25-item calibration so that scores would be on the same scale and comparable.

Next, we used the IRT test information function (TIF) to examine the precision of scale scores—the extent to which an estimate of genomic knowledge at a given scale score is reasonably close to the true value. Given that these scores estimate individuals' genomic knowledge, greater precision improves the scale's ability to distinguish between individuals with different levels of genomic knowledge in addition to providing other useful information. The TIF sums information functions for each individual item into a single function. Greater test "information" indicates greater precision⁴². TIF is depicted in a graphical format where the amount of information is plotted against the latent construct (here, genomic knowledge) to show how well the test estimates the construct over the full range of individuals' ability or knowledge. The areas of greatest measurement precision are indicated by the highest points of the curve.

Classical test theory reliability—We evaluated internal consistency reliability of the 25 UNC-GKS items by computing Cronbach's coefficient α^{43} . Ideal α values are at least 0.70^{44} , indicating a set of items that are strongly related to one another.

Convergent validity—We calculated Pearson correlations between the UNC-GKS scale score and the REAL-G, REALM, and subjective and objective numeracy scales to evaluate convergent validity, or the extent to which measures that should be associated with each other are in fact associated. We predicted that the UNC-GKS would correlate positively with genetics-related literacy and general health literacy, because individuals with greater ability to obtain, process, and understand health information should be more able to learn the domains of information assessed by the UNC-GKS. We also predicted positive correlations between the UNC-GKS and both measures of numeracy because the ability to reason and apply numerical concepts influences ability to learn these domains of information⁴⁵.

Results

Sample

The final sample included 286 adult patients and 132 parents (Table 2). Three quarters were women and 17% were racial/ethnic minorities. Participants' mean age was 47 years (range 17–84 years). Nearly 20% of participants had not attended college; just over half had a college degree. The median annual household income category was \$60,000–74,999. Nearly 13% of the sample had marginal or worse general health literacy and 6% had low genetic literacy. About 18% reported less than "excellent" proficiency in speaking, writing, and/or reading English.

Scale and item descriptive statistics

The number of missing responses was minor, ranging from 2–8 of 418 participants per item. The mean proportion of participants correctly answering each item was 0.73 (SD=0.12) with a range of 0.48 to 0.89 across the items. Figure 1 shows the distribution of responses for each item, revealing items for which knowledge and uncertainty were highest and lowest. Items from the four domains assessed by the measure (the structure and function of genes, how they are inherited, their relation to health, and the strengths and limitations of WES) were well distributed across the range of items answered correctly, incorrectly, or for which

there was uncertainty. There were no floor or ceiling effects; thus, no items were too easy or too hard for this sample. All correlations among items were positive and statistically significant, with a mean of $r=0.46$ ($SD=0.15$) and range of 0.10 to 0.90. Similarly, all itemtotal correlations were positive and of medium to large magnitude, with a mean of $r=0.49$ $(SD=0.11)$ and range of 0.29 to 0.68. Thus, item-level statistics did not identify any items as candidates for revision or removal and subsequent analyses considered all 25 items. Distributions of the summed score for the 19 and 25 item versions of the UNC-GKS appear in Figures 2 and 3. Scores are skewed to the left, indicating that participants correctly answered most items.

Factor analyses

The one factor confirmatory factor analysis model fit the data well ($X^2=476.4$, $d\hat{t}=275$, $p<0.001$; RMSEA=0.04; CFI=0.96; TLI=0.95), indicating that the items represented a single underlying construct. The standardized factor loadings were positive, statistically significant, and of moderate to large magnitude (0.40–0.95); thus, all items were associated with the underlying construct of genomic knowledge (Table 4). Two modification indices indicated possible local dependence between items 5 and 6, and between items 9 and 12. Fitting a model that allowed the errors for these two item pairs to correlate led to good fit (X^2 =456.0, $d=273$, $p<0.001$; RMSEA=0.04; CFI=0.96; TLI=0.96) with a statistically significant improvement in fit over the simpler model ($\chi^2_{diff} = 20.4$, $df=2$, $p<0.001$). However, the residual correlations were not large ($r=0.33$ for items 5 and 6, $r=0.42$ for items 9 and 12) and the items' text did not suggest content redundancy so we retained them for further evaluation.

Item response theory analyses

The UNC-GKS items varied in their ability to discriminate between participants with differing amounts of genomic knowledge, as indicated by the better fit of the 2PL rather than the 1PL model (χ^2_{diff} = 221.37, *df*=24, *p*<0.001). The 2PL model fit the items well; no items exhibited misfit or local dependence ($p<0.05$). Participants were not likely to have been guessing when they answered "true" or "false," given that the 3PL model did not fit better than the 2PL model (χ^2_{diff} = 2.92, df=24, p=1.00); participants who were unsure of their responses were likely selecting "not sure/don't know." The $S-X^2$ and LD statistics appear in Table 3.

To further evaluate the residual correlations found in the confirmatory factor analysis model between items 5 and 6 and items 9 and 12, we estimated a bifactor 2PL model with each of these item pairs loading on a second-order factor in addition to the overall genomic knowledge factor. Explained common variance for this model was 0.91; thus, most of the variance in the items was attributable to the overall factor. This finding supported our decision to retain these items and led us to consider the data to be "essentially unidimensional," meaning that there were no other meaningful underlying dimensions. Thus, these four items could be considered part of the overall factor.

DIF detection analyses, estimated with logistic regression models, did not reveal any statistically significant parameters ($p<0.05$) for uniform or non-uniform DIF. That is, after controlling for level of genomic knowledge, the scale items did not function differently in terms of their difficulty or how related they were to genomic knowledge when answered by individuals who differed on each of the key demographic groups of sex, race/ethnicity, education, and English proficiency. The measure performed comparably across these groups.

Our analytic plan called for items to be dropped prior to calibration of a final IRT model if they did not fit well, substantially violated local dependence, or functioned differently for key groups. Because no items were flagged for removal using these criteria, the original 2PL calibration was the final IRT model used for scoring. As shown in Table 3, the difficulty (b) parameters for the items ranged from −2.56 to 0.09; higher b parameters indicate more difficult items. The discrimination (a) parameters indicated that all items were highly related to the underlying construct of genomic knowledge and that they were able to discriminate between individuals with different levels of genomic knowledge. Items specific to WES (items $20-25$) had *a* parameters >2.0 and thus were particularly good at discriminating between these individuals, perhaps because the items covered information that was relatively unfamiliar to participants, many of whom had not known about WES prior to the study.

Item response theory scoring and reliability

We examined score precision (TIF) and reliability for both forms of the scale (see Figure 4). For the 25-item UNC-GKS, precision and reliability was especially high for T-scores between 32 and 49, and reliability was good (0.70) for all scores below 61 (approximately one standard deviation above the mean). For the 19-item UNC-GKS, precision and reliability was highest for T-scores between 36 and 40, and reliability was good (≥ 0.70) for all T-scores below 60. Thus, reliability was above accepted cut-offs for individuals who correctly answered 23 items on the 25-item UNC-GKS or 17 items on the 19-item UNC-GKS. Table 4 provides a scoring translation table for converting summed scale scores to Tscores based on the final (2PL) IRT model using the overall sample as a reference group. The correlation between T-scores resulting from the two versions of the scale was 0.97.

Classical test theory reliability

Internal consistency reliability was high (Cronbach's α =0.90) for the 25-item UNC-GKS and for the 19-item UNC-GKS $(a=0.86)$.

Convergent validity

As expected, the UNC-GKS correlated moderately and positively with general health literacy, genetics-related health literacy, and objective and subjective numeracy (Table 5), providing evidence of convergent validity.

Discussion

In order to make informed decisions about accepting or declining sequencing, understand their sequencing results, and decide on appropriate ways to apply these results, people need knowledge about the structure and function of genes, how genes are inherited in families,

effects of genes on health, and potential benefits, harms, and limitations of sequencing^{4,5}. Our study used rigorous methods to evaluate the psychometric properties of the UNC-GKS, which assesses these domains of knowledge. Findings indicate that all of the scale's items measured genomic knowledge well and, taken together, they cover a broad range of critical genomic knowledge. Thus, the UNC-GKS can generate a single score representing a person's genomic knowledge.

Importantly, findings indicated that the UNC-GKS measures genomic knowledge in the same way regardless of people's sociodemographic background. That is, different subgroup scores indicate different levels of genomic knowledge rather than differences in the way people in certain subgroups understood scale items or differences in their response patterns that are unrelated to their level of genomic knowledge. As such, the UNC-GKS is valid for a variety of subgroups and researchers can use it with diverse populations. For instance, group differences in UNC-GKS scores could be used to advance understanding of disparities in genomic sequencing.

Furthermore, UNC-GKS scores were most informative below one standard deviation above the scale's mean. Thus, the scores are precise for most individuals, with slightly less reliability, although still well within an acceptable range for those who answer all items correctly or those who give one incorrect answer (i.e., scores at one standard deviation above the scale's mean or higher). These high-scoring individuals may have less need for educational interventions, so this psychometric property of the UNC-GKS does not detract from its usefulness. In sum, the measure is reliable and especially sensitive for identifying people with varying degrees of low to moderate genomic knowledge—the group in greatest need of education and decision support when offered genomic sequencing.

We created two versions of the UNC-GKS: a 25-item version that includes items assessing knowledge about strengths and limitations of WES (the context for which we developed the scale) and a 19-item version that excludes these WES-specific items. T-scores were comparable for these two forms—an advantage for researchers and clinicians who use only the first 19 items because they are not offering WES, or they wish to develop a new set of items specific to the type of testing they are using. Our team has developed and will validate versions of the scale that substitute the WES items with new items relevant to the application of genomic sequencing to newborn screening and to screening adults for highly penetrant mutations that confer risk for treatable or preventable diseases.

Moreover, findings indicated that participants were not guessing, likely as a positive result of including the "not sure/don't know" response option. This response option also allows evaluation of the extent to which individuals are unsure of responses versus having been forced to indicate that each item is true or false, without being able to indicate uncertainty. Being able to evaluate uncertainty may have practical utility in some applications. For instance, items with high unsure/don't know responses may represent specific areas that should be targeted with educational efforts.

Limited literacy and numeracy—basic abilities necessary for seeking, comprehending, and using health information^{46,47}—is a prevalent problem and one that is highly relevant for

researchers and clinicians who need to educate individuals about genomic sequencing. Both were associated with lower scores on the UNC-GKS. These associations have been identified as important in the context of genetic testing and genomic sequencing 48 , and they have contributed to calls to increase educational efforts necessary for making scientific knowledge about genetics and genomics more broadly accessible⁴.

Additional research will be needed to evaluate additional aspects of validity, including other correlates of genomic knowledge measured by the UNC-GKS, group differences in genomic knowledge that may contribute to disparities in decision making and decision outcomes, and whether UNC-GKS scores identify people at risk for experiencing poor psychosocial or medical outcomes after testing. Research using this measure may also help guide development of psychoeducational interventions to address knowledge deficits, targeting clinicians (e.g., training in communicating effectively about genomic sequencing), patients and research participants (e.g., more useful and accessible educational resources), or both.

Future studies should also address limitations of the present study. Some of our subgroups were smaller than a suggested subgroup size of 100 for DIF analyses⁴⁹. This rule of thumb is an estimate and, given the distribution of our item responses, we believe our results to be reliable. Yet, a more diverse sample would enable more in-depth examination of differential item functioning across subgroups, including the potential for more difficult or complex items to function poorly in some subgroups. We note that we did not find evidence for this type of problem, perhaps in part because our items were iteratively reviewed by measurement and clinical experts. However, it would be valuable to recruit a larger proportion of underserved minority groups, individuals with less than a high school education, and those with low health literacy. In addition, the research was conducted primarily at an academic medical center in a single geographic area. Future research should evaluate the measure in community-based samples and other geographic areas. Finally, the scale is likely to be too long for some purposes, and validation of a briefer form could increase its usefulness.

Conclusion

This study used a rigorous approach to demonstrate that the UNC-GKS is a promising tool for advancing research on people's decisions about and responses to genomic sequencing. The scale has promising psychometric properties across different sociodemographic subgroups, making it appropriate for research in diverse populations. In addition, it covers domains of knowledge considered by a multidisciplinary team of experts to be critical for informed decision making about genomic sequencing, which is rapidly replacing single gene testing in some populations. Moreover, the scale is not disease or population specific, making it usable across a wide range of populations facing complex decisions involving next generation genomic sequencing.

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References

- 1. Warman CJ, Beaulieu C, Hartley T, et al. 2015; Axons to Exons: the Molecular Diagnosis of Rare Neurological Diseases by Next-Generation Sequencing. Curr Neurol Neurosci Rep. 15(9):64. [PubMed: 26289954]
- 2. Bernhardt BA, Roche MI, Perry DL, et al. 2015; Experiences with obtaining informed consent for genomic sequencing. Am J Med Genet A. 167A(11):2635–2646. [PubMed: 26198374]
- 3. Roche MI, Berg JS. 2015; Incidental findings with genomic testing: implications for genetic counseling practice. Curr Genet Med Rep. 3(4):166–176. [PubMed: 26566463]
- 4. Secretary's Advisory Committee on Genetics, Health, and Society. Genetics Education and Training. Washington, DC: Department of Health & Human Services; 2011.
- 5. Lautenbach DM, Christensen KD, Sparks JA, et al. 2013; Communicating genetic risk information for common disorders in the era of genomic medicine. Annu Rev Genomics Hum Genet. 14:491– 513. [PubMed: 24003856]
- 6. Erby LH, Roter D, Larson S, et al. 2008; The rapid estimate of adult literacy in genetics (REAL-G): a means to assess literacy deficits in the context of genetics. Am J Med Genet. 146A(2):174–181. [PubMed: 18076116]
- 7. Smerecnik CM, Mesters I, de Vries NK, et al. 2008; Educating the general public about multifactorial genetic disease: applying a theory-based framework to understand current public knowledge. Genet Med. 10(4):251–258. [PubMed: 18414207]
- 8. Green MJ, Biesecker BB, McInerney AM, et al. 2001; An interactive computer program can effectively educate patients about genetic testing for breast cancer susceptibility. Am J Med Genet. 103(1):16–23. [PubMed: 11562929]
- 9. Richman AR, Tzeng JP, Carey LA, et al. 2011; Knowledge of genomic testing among early-stage breast cancer patients. Psychooncology. 20(1):28–35. [PubMed: 20200857]
- 10. Scuffham TM, McInerny-Leo A, Ng SK, et al. 2014; Knowledge and attitudes towards genetic testing in those affected with Parkinson's disease. J Community Genet. 5(2):167–177. [PubMed: 24018619]
- 11. Singer E, Antonucci T, Van Hoewyk J. 2004; Racial and ethnic variations in knowledge and attitudes about genetic testing. Genet Test. 8(1):31–43. [PubMed: 15140372]
- 12. Haga SB, Barry WT, Mills R, et al. 2013; Public knowledge of and attitudes toward genetics and genetic testing. Genet Test Mol Biomarkers. 17(4):327–335. [PubMed: 23406207]
- 13. Kaphingst KA, Facio FM, Cheng MR, et al. 2012; Effects of informed consent for individual genome sequencing on relevant knowledge. Clin Genet. 82(5):408–415. [PubMed: 22694298]
- 14. Frisbie DA. 1973; Multiple Choice versus true-false comparison of reliabilities and concurrent validities. J Educ Meas. 10(4):297–304.
- 15. Ebel, RL, Frisbie, DA. Essentials of Educational Measurement. 5. Englewood Cliffs, NJ: Prentice Hall; 1991.
- 16. Fitzgerald-Butt SM, Bodine A, Fry KM, et al. 2016; Measuring genetic knowledge: a brief survey instrument for adolescents and adults. Clin Genet. 89(2):235–243. [PubMed: 26032340]
- 17. DeVellis, RF. Scale development: Theory and applications. 2. Vol. 26. Newbury Park: Sage Publications; 2003.
- 18. Félix-Ortiz M, Newcomb MD, Myers H. 1994; A multidimensional measure of cultural identity for Latino and Latina adolescents. Hisp J Behav Sci. 16(2):99–115.
- 19. Davis TC, Crouch MA, Long SW, et al. 1991; Rapid assessment of literacy levels of adult primary care patients. Fam Med. 23(6):433–435. [PubMed: 1936717]
- 20. Fagerlin A, Zikmund-Fisher BJ, Ubel PA, et al. 2007; Measuring numeracy without a math test: development of the Subjective Numeracy Scale. Med Decis Making. 27(5):672–680. [PubMed: 17641137]
- 21. McNaughton CD, Wallston KA, Rothman RL, et al. 2011; Short, subjective measures of numeracy and general health literacy in an adult emergency department. Acad Emerg Med. 18(11):1148– 1155. [PubMed: 22092896]

- 22. McNaughton CD, Cavanaugh KL, Kripalani S, et al. 2015; Validation of a short, 3-item version of the subjective numeracy scale. Med Decis Making. 35(8):932–936. [PubMed: 25878195]
- 23. Schwartz LM, Woloshin S, Black WC, et al. 1997; The role of numeracy in understanding the benefit of screening mammography. Ann Intern Med. 127(11):966–972. [PubMed: 9412301]
- 24. Wirth RJ, Edwards MC. 2007; Item factor analysis: current approaches and future directions. Psycho Methods. 12(1):58–79.
- 25. Muthén, LK, Muthén, BO. Mplus User's Guide. Seventh. Los Angeles, CA: Muthén & Muthén; 1998–2012.
- 26. Steiger JH. 1990; Structural model evaluation and modification an interval estimation approach. Multivar Behav Res. 25(2):173–180.
- 27. Tucker LR, Lewis C. 1973; Reliability coefficient for maximum likelihood factor-analysis. Psychometrika. 38(1):1–10.
- 28. Bentler PM. 1990; Comparative fit indexes in structural models. Psychol Bull. 107(2):238–246. [PubMed: 2320703]
- 29. Cai, L, Thissen, D, du Toit, SHC. IRTPRO for Windows. Lincolnwood, IL: Scientific Software International; 2011.
- 30. Thissen D. 1982; Marginal maximum likelihood estimation for the one-parameter logistic model. Psychometrika. 47:201–214.
- 31. Birnbaum, AS. Some latent trait models and their use in inferring an examinee's ability. In: Lord, FM, Novick, MR, editors. Statistical Theories lf Mental Test Scores. Reading, MA: Addison-Wesley; 1968. 392–479.
- 32. Orlando M, Thissen D. 2003; Further investigation of the performance of S-X-2: An item fit index for use with dichotomous item response theory models. Appl Psychol Meas. 27(4):289–298.
- 33. Orlando M, Thissen D. 2000; Likelihood-based item-fit indices for dichotomous item response theory models. Applied Psychological Measurement. 24(1):48–62.
- 34. Benjamini Y, Hochberg Y. 1995; Controlling the false discovery rate a practical and powerful approach to multiple testing. J Roy Stat Soc B Met. 57(1):289–300.
- 35. Williams VSL, Jones LV, Tukey JW. 1999; Controlling error in multiple comparisons, with examples from state-to-state differences in educational achievement. J Educ Behav Stat. 24(1):42– 69.
- 36. Chen WH, Thissen D. 1997; Local dependence indexes for item pairs: using item response theory. J Educ Behav Stat. 22(3):265–289.
- 37. Reise SP, Moore TM, Haviland MG. 2010; Bifactor models and rotations: exploring the extent to which multidimensional data yield univocal scale scores. J Pers Assess. 92(6):544–559. [PubMed: 20954056]
- 38. ten Berge JMF, Socan G. 2004; The greatest lower bound to the reliability of a test and the hypothesis of unidimensionality. Psychometrika. 69(4):613–625.
- 39. Bentler PM. 2009; Alpha, Dimension-free, and model-based internal consistency reliability. Psychometrika. 74(1):137–143. [PubMed: 20161430]
- 40. Sijtsma K. 2009; On the use, the misuse, and the very limited usefulness of cronbach's alpha. Psychometrika. 74(1):107–120. [PubMed: 20037639]
- 41. Thissen, D, Steinberg, L, Wainer, H. Detection of differential item functioning using the parameters of item response models. In: Holland, PW, Wainer, H, editors. Differential Item Functioning. Hillsdale, NJ: Erlbaum; 1993. 67–113.
- 42. Thissen D, Wainer H. 2002; Test Scoring. J Educ Meas. 39(3):265–268.
- 43. Cronbach LJ. 1951; Coefficient alpha and the internal structure of tests. Psychometrika. 16(3):297– 334.
- 44. Streiner, DL, Norman, GR. A Practical Guide to Their Development and Use. Second. Oxford: Oxford University Press; 1995. Health Measurement Scales.
- 45. Lea DH, Kaphingst KA, Bowen D, et al. 2011; Communicating genetic and genomic information: health literacy and numeracy considerations. Public Health Genomics. 14(4–5):279–289. [PubMed: 20407217]

- 46. Medicine Io. Health literacy: A prescription to end confusion. Washington, DC: Institute of Medicine of the National Academies, Committee on Health Literacy, Board on Neuroscience and Behavioral Health; 2004.
- 47. Kutner, M, Greenberg, E, Jin, Y. , et al. The health literacy of America's adults: results From the 2003 National Assessment of Adult Literacy (NCES 2006–483). Washington, DC: U.S. Department of Education; 2006.
- 48. Lea DH, Kaphingst KA, Bowen D, et al. 2011; Communicating genetic and genomic information: health literacy and numeracy considerations. Public Health Genomics. 14(4–5):279–289. [PubMed: 20407217]
- 49. Lai JS, Teresi J, Gershon R. 2005; Procedures for the analysis of differential item functioning (DIF) for small sample sizes. Eval Health Prof. 28(3):283–294. [PubMed: 16123258]

Figure 1.

Item response distributions for the University of North Carolina Genomic Knowledge (UNC-GKS)

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Figure 4. Item Response Theory test information and reliability for the University of North Carolina Genomic Knowledge (UNC-GKS) T-scores

Table 1

UNC Genomic Knowledge Scale Items

Note: Correct answer to the items is true unless followed by an asterisk (*).

Table 2

Participant Characteristics (N=418)

Ages for participating parents of pediatric patients were not collected early in the study; therefore, descriptive statistics for participant age are based on all adult patients and 27 of the 132 participating parents.

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Results of the one-factor confirmatory factor analysis model and two-parameter logistic item response theory model for University of North Carolina Results of the one-factor confirmatory factor analysis model and two-parameter logistic item response theory model for University of North Carolina Genomic Knowledge Scale items Genomic Knowledge Scale items

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Note. CFA = confirmatory factor analysis; 2PL IRT = two parameter logistic item response theory; s.e. = standard error; $d\vec{t}$ = degrees of freedom. Note. CFA = confirmatory factor analysis; 2PL IRT = two parameter logistic item response theory; s.e. = standard error; df = degrees of freedom. Author Manuscript **Author Manuscript**

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Table 4

Table for converting summed scores to item response theory-scaled T-scores for the University of North Carolina Genomic Knowledge Scale (UNC-GKS)

Note. This table can be used to convert the summed count of a participant's correct responses (summed scores) to item response theory scaled scores in the T score metric, allowing easy comparison across studies and populations. These T scores were calculated using this study's sample, a population mean of 50, and standard deviation of 10. To convert a score, find a participant's summed score in column 1 and follow the row to determine the IRT-scaled T score and standard error for the 25- and 19-item versions of the UNC-GKS.

Table 5

Pearson correlations between University of North Carolina Genomic Knowledge (UNC-GKS) scale score and other measures

* $p<0.001$