Cooperation and heterogeneity of the autistic mind

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Abstract

Individuals with autism spectrum conditions (ASC) have a core difficulty in recursively inferring the intentions of others. The precise cognitive dysfunctions that determine the heterogeneity at the heart of this spectrum, however, remains unclear. Furthermore, it remains possible that impairment in social interaction is not a fundamental deficit but a reflection of deficits in distinct cognitive processes. To better understand heterogeneity within ASC, we employed a game-theoretic approach to characterise *unobservable* computational processes implicit in social interactions. Using a social hunting game with autistic adults, we found that a selective difficulty representing the level of strategic sophistication of others, namely inferring others' mindreading strategy, specifically predicts symptom severity. In contrast, a reduced ability in iterative planning was predicted by overall intellectual level. Our findings provide the first quantitative approach that can reveal the underlying computational dysfunctions that generate the autistic 'spectrum'.

Keywords

autism spectrum conditions; theory of mind; belief inference; executive function; social decision-making; computational model

Introduction

Despite recent progress in social and affective neuroscience, there are no simple models that explain and predict complex human social behaviour and its pathologies. Autism spectrum conditions (ASC) involve severe deficits in reciprocal social interaction (American Psychiatric Association, 1994) and thus represent a behavioural deficit model for social functioning. A lack of 'Theory of Mind' (ToM), *i.e.*, the representation of others' goals and intentions (Frith and Happe, 2005) is thought to be a key pathognomonic feature defining social dysfunctions of ASC. Although many tasks are sensitive to impairments in various facets of ToM (Tager-Flusberg et al., 2001), a systematic approach to the common underlying mechanism has yet to be realised. Furthermore, it is widely assumed that aberrant ToM is not a fundamental deficit in itself but rather a reflection of multiple underlying cognitive processes, each of which may be manifested in different forms, or subtypes, of ASC. Here, we employed a game-theoretic approach which has recently provided new

insights into the computations and neural correlates of social behaviours (Behrens et al., 2009; King-Casas et al., 2008).

We used a Stag-hunt game in which participants interacted with a computerized agent to hunt stags together (high value) or defect in order to hunt rabbits alone (low value) (Yoshida W. et al., 2008) (Figure 1). In the Stag-hunt, cooperation depends on recursive representations of another's intentions, since, if I decide to hunt the stag, I must believe that you believe that I will cooperate with you; thus interactions of highly sophisticated players allow cooperation to emerge (Yoshida W. et al., 2008). During the experiment, a computerized agent shifted its sophistication (by three degrees of recursion) without notice. Thus, to behave optimally, participants were required to (i) estimate the agent's sophistication level, (ii) update their own strategies continuously, and (iii) behave optimally on the basis of their inference. This calls on recursive belief inference, cognitive flexibility, and interactive planning, respectively, all of which have been reported impaired in ASC (Sigman et al., 2006). To tease apart tightly-entangled cognitive processes implicit in social interaction, we employed a previously developed theoretical model, in which participants behave optimally with respect to the goal of maximizing the payoff based on these three processes. This model-based behavioural analysis allowed us to capture the linkage of *unobservable* computational processes with *observable* intellectual and diagnosis scores that contain the heterogeneity within the autism spectrum.

Materials and Methods

Participants

Seventeen adults with ASC (12 males) and 17 healthy control adults (11 males) participated in the study. All participants were German speaking and age (ASC: 33.1±7.8; control: 31.0±5.7) was matched between two groups. For each participant in the ASC group, diagnoses were made according to DSM-IV criteria (American Psychiatric Association, 1994) using a videotaped semi-structured interview. All participants gave informed written consent and the research protocol was approved by the Max Planck Institute for Human Development, Berlin, Germany.

Diagnostic and intellectual measures

To compare the ASC and control subjects in autistic diagnosis, we used the Autism-spectrum quotient (ASQ); a screening questionnaire used in clinical practice. The ASQ score of the control group (mean ± std=10.8 ± 5.2) was significantly lower than that of the ASC group (38.7 ± 8.0, p<0.1e-12) (Figure 2A). We also confirmed that all control individuals scored lower than 26 and were ruled out as ASC. For the participants in the ASC group, autistic symptomatology was quantified using the Autism Diagnostic Interview-Revised (ADI-R) (Lord et al., 1994) in 16 individuals and the Asperger Syndrome and High Functioning Autism Diagnostic Interview (ASDI) (Gillberg et al., 2001) in 14 individuals. The ADI-R is a semi-structured interview administered to the parents or the caretakers, while the ASDI is designed to cover diagnostic criteria for Asperger and performed by the individuals with autism themselves.

The intelligence level of the participants in the ASC and the control group was measured by a vocabulary test (MWT (Tewes, 1991)) and a nonverbal strategic thinking test (LPS (Horn, 1962)). There was no significant difference between two groups for both verbal (control: 117.1 ± 15.8, ASC: 105.9 ± 16.0) and nonverbal (strategic) intelligence scores (control: 124.2 ± 13.3, ASC: 124.2 ± 12.0) (Figure 2A). To assess fluid and crystallized intellectual functioning of the ASC, we build the mean of the two scores, and based on these, WAIS-R full scale IQ was estimated.

Experimental task and computational model

We used a Stag-hunt game (Skyrms, 2003) in which participants interacted with a computerized agent, thereby utilizing a Theory of Mind (ToM) model which is a generative model to infer participants' mental states in the Stag-hunt. As the game and the model have been described in detail elsewhere (Yoshida W. et al., 2008), we briefly summarize the key concepts here.

The Stag-hunt game has a pro-cooperative pay-off matrix, in which each of two players choose whether to hunt highly valued stags together (20 points each), and share the proceeds, or defect in order to hunt rabbits of smaller value (10 points). The task was embedded into a move-by-move maze-based hunting game (Figure 1), where each game finishes when either player has caught a prey. Each move incurs a small cost (1 point), such that the best strategy is to hunt bigger prey as swiftly as possible. In the experiment, the participants played two sessions each of which comprised 40 games. At the end of the experiment, the participants were paid money based on accrued total points.

The player's strategy is defined by the hierarchical value-functions using optimal control theory. In sequential games with multiple players, like the Stag-hunt game, the payoffs are defined over a joint-space for each player and the values for one player become a function of the other player's values. Thus the optimal value-functions (strategies) are specified by an order based on *recursive sophistication*. Namely, the first-order strategies discount the strategies of other players (*i.e.*, I will ignore your goals). Second-order strategies are optimised under the assumption that other players are using a first-order strategy (*i.e.*, you are ignoring my goals). Third-order strategies pertain when I assume that you assume I am using a first-order strategy and so on. In general, a player with the lower-order (competitive) strategy would try to catch a rabbit, provided both players were not close to the stag, and a player with a higher-order (cooperative) strategy would tend to chase the stag even if it was close to a rabbit. In the experiment, the computer agent changed strategy, which is defined with first-, third-, or fifth-order sophistication, without notice.

Using the hierarchical value-functions, we tested two computational models of player's behaviour. The ToM model assumes that players infer the other's strategy; the sophistication level of computerised agent in the Stag-hunt, based on Bayesian inference under flat priors bounded by K (*bounded rationality*) after each observation (moves in the maze at each trial):

$$\text{agent's level}\,(T) = \underset{\text{agent's level}\in\{1:K\}}{\text{argmax}} \prod_{t=1}^{T-1} \kappa^{T-1} p\,(\text{observation}\,(t+1)\,|\text{subject's level}\,(t)\,,\,\text{agent's level}\,(t))$$

where κ is a forgetting parameter that exponentially discounts previous evidence and allows the player to respond more quickly to changes in the other's strategy. The parameter value ranges from zero; the most flexible setting in which players update the strategy based only on the latest action of others, to one; when players take the history of all actions into account. Accordingly, under the definition of value-functions with order, the players optimise their own strategy, which is one level higher than the estimated other's sophistication level. As a reference model without belief inference, we assumed a 'fixed strategy' model in which the players use a fixed strategy at the level of k and do not change strategy based on other's behaviour.

Model selection and parameter estimation

We computed the evidence of the fixed strategy models, $p\left(M_k^{\text{FIX}}\right)$, with $k=1,\ldots,6$ and the ToM models, $p\left(M_K^{\text{TOM}}\right)$, with $K=1,\ldots,6$; *i.e.*, 12 models in total, given the actual behavioural data. To evaluate the expectation of belief inference, we calculated the marginalised posterior probabilities of the fixed strategy model and the ToM model by accumulating evidence for different models specified by different k-level and bound K; *i.e.*, 6 models each, as $\sum_k p\left(M_k^{\text{FIX}}\right)$ and $\sum_K p\left(M_K^{\text{TOM}}\right)$.

To estimate the cognitive flexibility, we used the average forgetting effect, which is calculated as the weighted sum of the parameter κ with the posterior probability of each model. The forgetting parameter was optimised for each ToM model with different K, and set at one for the fixed strategy models.

$$\sum_k p\left(M_k^{\text{FIX}}\right)+\sum_K \kappa_K p\left(M_K^{\text{TOM}}\right)$$

The average sophistication for the ASC participants was estimated as the weighted sum of parameter k with the model probability of each fixed strategy model.

Results

To first confirm a basic cognitive understanding of the task demands, we examined behavioural strategies in the ASC group. Overall behavioural measurements including the average reaction times, total earnings, and cooperation rates of the ASC group did not differ from the control group providing evidence that an understanding of the social task was intact (Chiu et al., 2008; Uddin et al., 2008). The participants in the control group attempted to catch a stag, i.e., they behaved more cooperatively, when the computer agent was more sophisticated, and this tendency was also observed for the ASC group on average (Figure 2B). The behavioral profiles, however, differed greatly in individuals with ASC. Thus, while all control participants dealt with both hunting a stag and hunting a rabbit according to the situation (e.g., initial positions), there are extreme participants in the ASC group, who never cooperate or compete (Figure 2C). This suggests a behavioural diversity or even multiple phenotypes among the ASC participants

To identify functional abnormalities in the computational processes involved in the task, we employed the ToM model and the fixed strategy model. The ToM model includes two model parameters characterizing the cognitive processing; one is the upper bound of sophistication (K), which defines the capacity of strategic planning, and the other is a forgetting effect, which controls how quickly player responds to changes in the other's sophistication, thereby representing a measure of cognitive flexibility. For the fixed strategy model, as it is assumed that players do *not* change their strategy, only the sophistication level (k) is estimated.

First, to compare the behavioural fit of these two models, we calculated the log likelihoods of the models for the control and the ASC participants (Control: ToM −5776, fixed −6060; ASC: ToM −5330, fixed −4937; greater values indicate better fit). Bayesian model selection based on these log likelihoods showed that, for the control participants, the ToM model *with* belief inference accounted for the behaviour significantly better than the fixed strategy model *without* belief inference. Conversely, the fixed strategy model explained individual behaviour better for 12 out of 17 participants with ASC (Figure 3A). We also evaluated the quality of model fitting with a pseudo-r^2 statistics; defined as $(m_0- m)/m_0$ where m and m_0

are the log likelihood of the data under the model and under purely random choices, respectively. The results showed that the ToM model (r^2=0.230) provided a better fit than the fixed strategy model (0.194) for the controls, while the fixed model (0.262) provided a better than the ToM model (0.203) for the ASC participants.

We divided the ASC participants into two sub-groups based on the better explained models, and compared their diagnosis and intellectual scores (Figure 2B). The diagnosis scores (the ADI-R and ASDI) were significantly higher for the ASC participants whose behaviour fit better with the fixed strategy model (n=12) than others with the ToM model (n=5), while there was no difference on the intellectual (IQ) scores. Furthermore, the probability of the ToM model, how likely recursive belief inference is used, correlated negatively with individual autistic symptomatology as measured by the sum of scores on the ADI-R and the ASDI (Figure 3C). The correlation was not significant (p=0.055) with the rather small number of sample size (n=14), however, this negative correlation was also observed with the ADI-R (n=14, r=−0.45, p=0.105) and the ASDI score (n=16, r=−0.44, p=0.090) individually. We thereby characterized unobservable computational processes implicitly involved in ToM quantifying the individual ability for recursive belief inference.

In terms of the model parameters, the mean estimated forgetting effect in the ASC group was significantly higher than that of the control group (Figure 3D). The higher value of the forgetting effect means that participants are tied to their past strategies and this result indicates that the ASC participants had an additional impairment in cognitive flexibility. For the sophistication level, we inferred that the upper bound under the ToM model (K) is equal to 5 for 13 out of 17 control participants, which means the participants used strategies with up to a sixth sophistication level. This is reasonable given the computer agent's policies never exceeded level five. In contrast, we found that the average sophistication under the fixed model (k) varied among the ASC participants (mean ± std = 4.3 ± 1.2) and correlated positively with individual IQ scores (Figure 3E). This correlation was not observed for control participants' IQ scores and the sophistication level which is calculated as the expected upper bound of the ToM models (Figure. 3E). Note that, from our model-based behavioural analysis, we concluded that the ASC participants have a difficulty to change their sophistication level (k-level) while the controls flexibly change it but under the fixed upper bound (K). Thus, these parameters of sophistication levels (K and k) are functionally different for the two groups. However, we obtained almost identical results for both the ASC and control group using the expected sophistication *and* its bound using Bayesian model averaging over both fixed and ToM models.

Discussion

In summary, our model-based approach isolates specific factors that contribute to ASC social impairments, which parametrically covary with symptom severity and intelligence level. Thereby our findings offer a partial account of the heterogeneity of autism spectrum conditions.

While we found evidence that the general understanding of our social stag-hunt task was intact in ASC, the observed game behavior of individuals with ASC was guided to a significantly less degree of belief inference than that of the control participants. Moreover, severity of autism symptomatology predicted the extent to which ASC participants behaved according to a fixed strategy, i.e., disregarded the other player's beliefs. This is line with research showing that ToM impairments in autism are core deficits and underlie individual's social dysfunction (Baron-Cohen and Belmonte, 2005; Tager-Flusberg et al., 2001).

Individuals with autism also showed deficits in cognitive flexibility as indicated by a tendency to be tied to their past strategies during the social game rather than a policy of flexibly changing strategies by taking account the other's new actions. In addition, the level of sophistication (you think that I think that you think *etc*) for participants in the ASC group was related to IQ scores. A previous study using a 'Beauty Contest' game has indicated an association between higher-level reasoning in healthy individuals and higher intelligence scores (Coricelli and Nagel, 2009). Our result showed that highly intelligent individuals with ASC behave cooperatively as if they make predictions over a longer time-horizon. This suggests that the level of sophistication, a key component of higher-level reasoning, can be inferred in more complex dynamic social exchanges.

Impairments in cognitive flexibility and planning are known characteristics of neuropsychological profiles in ASC and have fostered the development of theories of executive dysfunction in autism (Hill, 2004; Russo et al., 2007). However, the precise contribution and nature of the relationship between ToM and executive functions remain a matter of debate (Beauchamp and Anderson, 2010). Using our Stag-hunt game we show that the domains of both belief inference and executive function contribute to determining social function. Future studies with our model-based approach could potentially to disentagle ToM from executive functions and shed light on the question how these congitive dysfunctions influence one another and generate autism social symptomatology.

Recent studies using game-theoretic frameworks have shown that frontal brain regions such as the medial prefrontal cortex, orbitofrontal cortex, and cingulate gyrus, as well as the amygdala and superior temporal sulcus are engaged by social exchanges (Lee, 2008; Sanfey, 2007; Hampton et al., 2008). Function in these regions has been repeatedly reported as abnormal in ASC (Frith, 2001; Amaral et al., 2008). However, none of the studies to date throws light on how different subtypes or symptomatic manifestations in ASC are determined by dysfunction in distinct regions, and how abnormal connectivity between these regions results in deficient social reciprocal behaviour. Our findings highlight the power of simple assessment tools for psychopathology predicated on the idea that core psychiatric deficits can be understood in terms of computational dysfunction. Thus, future studies might fruitfully explore how the specific mechanistic deficit that we observe here relates to function or dysfunction in the social brain.

Acknowledgments

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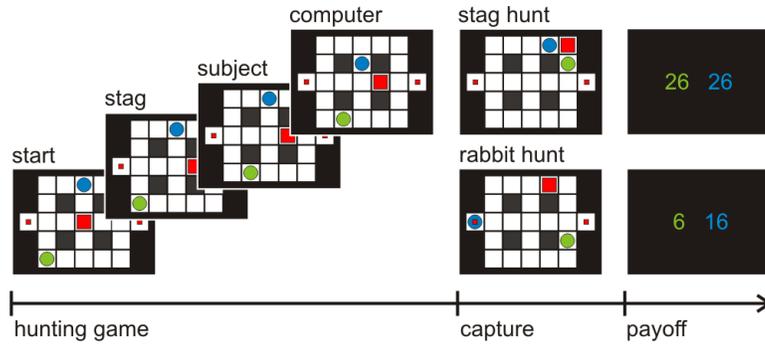
Figure 1.
Stag-hunt game.

Two players: a subject (green circle) and a computer agent (blue circle), try to catch prey: a mobile stag (big square, big payoff) by cooperation or two stationary rabbits (small squares, small payoff), by moving in sequential manner. At the end of each game, both players receive points equal to the sum of prey and points relating to the remaining time.
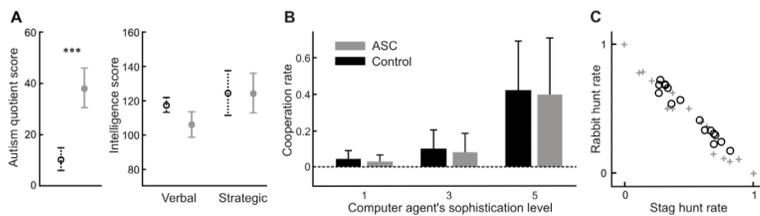
Figure 2.
Subjects and behavioural results

A. Mean scores of diagnostic (the left panel) and intellectual measurements (the right panel) of the control group (black dot line) and the ASC group (grey solid line). The errorbars show the standard deviations. Autism-spectrum quotient (ASQ) score was significantly higher for the ASC group than for the control group ($p<0.1e-12$), while there was no significant difference between the groups for both verbal and strategic intelligence scores.

B. Both the control and the ASC group attempted to catch a stag, in effect when they behaved more cooperatively, when the computer agent was more sophisticated.

C. The rate of stag hunt and rabbit hunt in the games with sophisticated computer agent with the fifth-order strategy. The participants in the ASC group (grey crosses) showed larger variety of behaviour than the control participants (black circles).
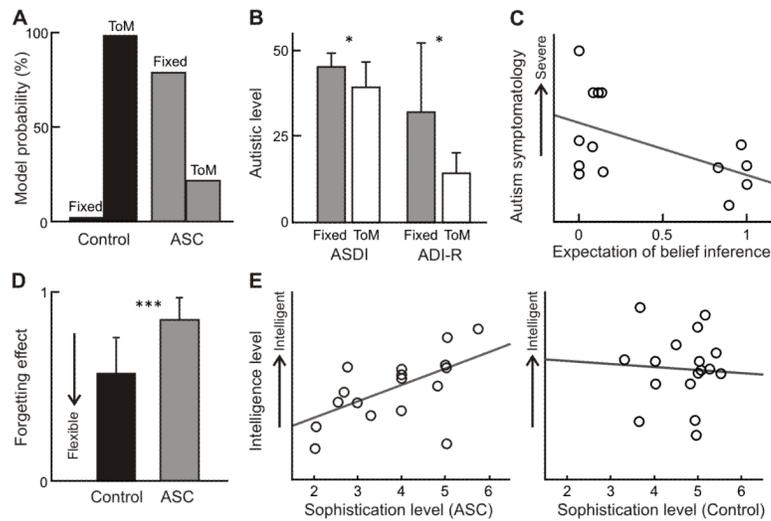
Figure 3. Model-based behavioural results
A. The probability for ToM model (98.2%) was higher than for the fixed strategy model for controls, while the fixed strategy model was dominant (78.6%) in individuals with ASC.
B. Diagnosis measurement scores; Autism Diagnostic Interview-Revised (ADI-R) and Asperger Syndrome and High-functioning Autism Diagnostic Interview (ASDI), were significantly higher for the ASC participants whose behaviour fit better with the fixed strategy model (n=12) than those showing a better fit with the ToM model (n=5).
C. In the ASC group, the greater expectation of recursive belief inference, the severer autism symptomatology (n=14, r=−0.52, p=0.055), as measured by the sum of scores on the ADI-R and the ASDI.
D. The estimated forgetting parameter for the ASC group (mean ± std=0.57 ± 0.19) was significantly higher than that for the control group (0.93 ± 0.13) (p<0.1e-6).
E. The estimated sophistication for the individuals with autism showed significant positive correlation with individual IQ scores (the left panel: n=17, r=0.54, p=0.026), while there was no correlation for the control participants (the right panel: n=17, r=0.02).